

# Interpreting Models for Categorical and Count Outcomes

Rose Medeiros  
StataCorp LLC

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## 1 Introduction

### 1.1 Goals

#### Goals

- Learn how to fit models that include categorical variables and/or interactions using factor variable syntax
  - Get an overview of tools available for investigating models
  - Learn a bit about how Stata partitions model fitting and model testing tasks
- 

## 2 Estimation

### 2.1 Factor Variables

#### A Logistic Regression Model

- We'll use data from the National Health and Nutrition Examination Survey (NHANES) for our examples
  - `. webuse nhanes2`
- We'll start with a model for high blood pressure (`highbp`) using age, body mass index (`bmi`) and sex (`female`)

- Before we fit the model, let's investigate the variables

```
. codebook highbp age bmi female
```

```
-----
highbp                                1 if bpsystol >= 140|bpdiast >= 90, 0 otherwise
-----
```

```

      type: numeric (byte)
      range: [0,1]                units: 1
unique values: 2                  missing .: 0/10,351

      tabulation: Freq. Value
                   5,975  0
                   4,376  1

```

```
-----
age                                    age in years
-----
```

```

      type: numeric (byte)
      range: [20,74]              units: 1
unique values: 55                 missing .: 0/10,351

      mean: 47.5797
      std. dev: 17.2148

      percentiles:    10%    25%    50%    75%    90%
                       24     31     49     63     69

```

```
-----
bmi                                    Body Mass Index (BMI)
-----
```

```

      type: numeric (float)
      range: [12.385596,61.129696] units: 1.000e-07
unique values: 9,941              missing .: 0/10,351

      mean: 25.5376
      std. dev: 4.91497

      percentiles:    10%    25%    50%    75%    90%
                       20.1037  22.142  24.8181  28.0267  31.7259

```

```
-----
female                                1=female, 0=male
-----
```

```

      type: numeric (byte)
      range: [0,1]                units: 1
unique values: 2                  missing .: 0/10,351

      tabulation: Freq. Value
                   4,915  0
                   5,436  1

```

- Now we can fit the model

```
. logit highbp age bmi female
```

```
Iteration 0:  log likelihood = -7050.7655
Iteration 1:  log likelihood = -5859.5273
Iteration 2:  log likelihood = -5845.5355
Iteration 3:  log likelihood = -5845.4948
Iteration 4:  log likelihood = -5845.4948
```

```
Logistic regression               Number of obs   =    10,351
                                LR chi2(3)       =    2410.54
                                Prob > chi2        =    0.0000
                                Pseudo R2         =    0.1709
```

```
Log likelihood = -5845.4948
```

highbp	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
age	.0459846	.0013974	32.91	0.000	.0432457 .0487235
bmi	.1371553	.0050802	27.00	0.000	.1271983 .1471123
female	-.4824464	.0451382	-10.69	0.000	-.5709156 -.3939771
_cons	-5.84101	.1523939	-38.33	0.000	-6.139697 -5.542324

## Working with Categorical Variables

- Now we would like to include region in the model, let's take a look at this variable

```
. codebook region
```

```
-----
```

```
region                                     1=NE, 2=MW, 3=S, 4=W
```

```
-----
```

```

      type: numeric (byte)
      label: region

      range: [1,4]                units: 1
unique values: 4                  missing .: 0/10,351

      tabulation:  Freq.  Numeric  Label
                   2,096      1  NE
                   2,774      2  MW
                   2,853      3  S
                   2,628      4  W
```

- region cannot simply be added to the list of covariates because it has 4 categories

- To include a categorical variable, put an `i.` in front of its name—this declares the variable to be a categorical variable, or in Stataese, a *factor variable*
- For example

```
. logit highbp age bmi i.female i.region
```

```
Iteration 0: log likelihood = -7050.7655
Iteration 1: log likelihood = -5857.277
Iteration 2: log likelihood = -5843.2102
Iteration 3: log likelihood = -5843.169
Iteration 4: log likelihood = -5843.169
```

```
Logistic regression                Number of obs    =    10,351
                                   LR chi2(6)        =    2415.19
                                   Prob > chi2        =    0.0000
Log likelihood = -5843.169         Pseudo R2       =    0.1713
```

highbp	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
age	.0459318	.0013982	32.85	0.000	.0431914	.0486722
bmi	.1372797	.0050825	27.01	0.000	.1273182	.1472411
female						
0	0	(base)				
1	-.4811765	.0451517	-10.66	0.000	-.5696723	-.3926807
region						
NE	0	(base)				
MW	-.1324591	.0662441	-2.00	0.046	-.2622952	-.002623
S	-.0887067	.0653787	-1.36	0.175	-.2168466	.0394331
W	-.0403994	.0667441	-0.61	0.545	-.1712154	.0904166
_cons	-5.772271	.1584937	-36.42	0.000	-6.082913	-5.461629

## Niceities

- Starting in Stata 13, value labels associated with factor variables are displayed in the regression table
- We can tell Stata to show the base categories for our factor variables

```
. set showbaselevels on
```

- ◊ This means the base category will always be clearly documented in the output

## Factor Notation as Operators

- The `i.` operator can be applied to many variables at once:

```
. logit highbp age bmi i.(female region)
```

```
Iteration 0: log likelihood = -7050.7655
Iteration 1: log likelihood = -5857.277
Iteration 2: log likelihood = -5843.2102
Iteration 3: log likelihood = -5843.169
Iteration 4: log likelihood = -5843.169
```

```
Logistic regression                Number of obs   =    10,351
                                   LR chi2(6)         =    2415.19
                                   Prob > chi2        =     0.0000
Log likelihood = -5843.169         Pseudo R2      =     0.1713
```

---

highbp	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
age	.0459318	.0013982	32.85	0.000	.0431914	.0486722
bmi	.1372797	.0050825	27.01	0.000	.1273182	.1472411
female						
0	0 (base)					
1	-.4811765	.0451517	-10.66	0.000	-.5696723	-.3926807
region						
NE	0 (base)					
MW	-.1324591	.0662441	-2.00	0.046	-.2622952	-.002623
S	-.0887067	.0653787	-1.36	0.175	-.2168466	.0394331
W	-.0403994	.0667441	-0.61	0.545	-.1712154	.0904166
_cons	-5.772271	.1584937	-36.42	0.000	-6.082913	-5.461629

---

- In other words, it understands the distributive property
  - ◊ This is useful when using variable ranges, for example
- For the curious, factor variable notation works with wildcards
  - ◊ If there were many variables starting with `u`, then `i.u*` would include them all as factor variables

---

## Using Different Base Categories

- By default, the smallest-valued category is the base category
- This can be overridden within commands
  - ◊ `b#`. specifies the value `#` as the base
  - ◊ `b(##)`. specifies the `#`'th largest value as the base
  - ◊ `b(first)`. specifies the smallest value as the base
  - ◊ `b(last)`. specifies the largest value as the base
  - ◊ `b(freq)`. specifies the most prevalent value as the base
  - ◊ `bn`. specifies there should be no base
- The base can also be permanently changed using `fvset`; see `help fvset` for more information

## Playing with the Base

- We can use `region=3` as the base class on the fly:  

```
. logit highbp age bmi i.female b3.region
```
  - We can use the most prevalent category as the base  

```
. logit highbp age bmi i.female b(freq).region
```
  - Factor variables can be distributed across many variables  

```
. logit highbp age bmi b(freq).(female region)
```
  - The base category can be omitted (with some care here)  

```
. logit highbp age bmi i.female bn.region, noconstant
```
  - We can also include a term for `region=4` only  

```
. logit highbp age bmi i.female 4.region
```
- 

## Specifying Interactions

- Factor variables are also used for specifying interactions
  - ◊ This is where they really shine
- To include both main effects and interaction terms in a model, put `##` between the variables
- To include only the interaction terms, put `#` between the terms
- Variables involved in interactions are treated as categorical by default
  - ◊ Prefix a variable with `c.` to specify that a variable is continuous

- Here is our model with an interaction between age and female

```
. logit highbp bmi c.age##female i.region
```

```
Iteration 0: log likelihood = -7050.7655
Iteration 1: log likelihood = -5824.3249
Iteration 2: log likelihood = -5795.4621
Iteration 3: log likelihood = -5795.4025
Iteration 4: log likelihood = -5795.4025
```

```
Logistic regression          Number of obs   =    10,351
                             LR chi2(7)           =    2510.73
                             Prob > chi2          =     0.0000
Log likelihood = -5795.4025  Pseudo R2       =     0.1780
```

highbp	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
bmi	.1378163	.005139	26.82	0.000	.1277441 .1478886
age	.0334439	.0018514	18.06	0.000	.0298151 .0370727
female					
0	0 (base)				
1	-1.883645	.1530275	-12.31	0.000	-2.183574 -1.583717
female#c.age					
1	.0276653	.0028606	9.67	0.000	.0220585 .033272
region					
NE	0 (base)				
MW	-.1359488	.0664206	-2.05	0.041	-.2661308 -.0057668
S	-.0902012	.0655982	-1.38	0.169	-.2187713 .0383689
W	-.0379412	.066944	-0.57	0.571	-.169149 .0932666
_cons	-5.176679	.1687139	-30.68	0.000	-5.507352 -4.846006

### Some Factor Variable Notes

- If you plan to look at marginal effects of any kind, it is best to
  - ◊ Explicitly mark all categorical variables with *i*.
  - ◊ Specify all interactions using *#* or *##*
  - ◊ Specify powers of a variable as interactions of the variable with itself
- There can be up to 8 categorical and 8 continuous interactions in one expression
  - ◊ Have fun with the interpretation

## 3 Postestimation

### 3.1 Tests of Coefficients

#### Introduction to Postestimation

- In Stata jargon, postestimation commands are commands that can be run after a model is fit, for example

- ◇ Predictions
- ◇ Additional hypothesis tests
- ◇ Checks of assumptions
- We'll explore postestimation tools that can be used to help interpret model results
  - ◇ The main example here is after `logit` models, but these tools can be used with most estimation commands
- The usefulness of specific tools will depend on the types of hypotheses you wish to examine

### Finding the Coefficient Names

- Some postestimation commands require that you know the names used to store the coefficients
- To see these names we can replay the model showing the *coefficient legend*

```
. logit, coeflegend

Logistic regression                Number of obs   =    10,351
                                LR chi2(7)       =    2510.73
                                Prob > chi2          =     0.0000
Log likelihood = -5795.4025        Pseudo R2       =     0.1780

-----+-----
      highbp |      Coef.  Legend
-----+-----
      bmi |   .1378163  _b[bmi]
      age |   .0334439  _b[age]
      |
    female |
      0 |           0  _b[0b.female]
      1 |  -1.883645  _b[1.female]
      |
  female#c.age |
      1 |   .0276653  _b[1.female#c.age]
      |
     region |
      NE |           0  _b[1b.region]
      MW |  -.1359488  _b[2.region]
      S  |  -.0902012  _b[3.region]
      W  |  -.0379412  _b[4.region]
      |
      _cons | -5.176679  _b[_cons]
-----+-----
```

- From here, we can see the full specification of the factor levels:
  - ◇ `_b[2.region]` corresponds to `region=2` which is "MW" or midwest
  - ◇ `_b[3.region]` corresponds to `region=3` which is "S" or south
- The coefficient for the female by age interaction is stored as `_b[1.female#c.age]`



## Joint Tests

- The test command performs a Wald test of the specified null hypothesis
  - ◊ The default test is that the listed terms are equal to 0
- test takes a list of terms, which may be variable names, but can also be terms associated with factor variables
- To specify a joint test of the null hypothesis that the coefficients for the levels of region are all equal to 0

```
. test 2.region 3.region 4.region

( 1) [highbp]2.region = 0
( 2) [highbp]3.region = 0
( 3) [highbp]4.region = 0

      chi2( 3) =    4.96
Prob > chi2 =    0.1744
```

---

## Testing Sets of Coefficients

- If you are testing a large number of terms, typing them all out can be laborious
- testparm also performs Wald tests, but it accepts lists of variables, rather than coefficients in the model
- For example, to test all coefficients associated with i.region

```
. testparm i.region

( 1) [highbp]2.region = 0
( 2) [highbp]3.region = 0
( 3) [highbp]4.region = 0

      chi2( 3) =    4.96
Prob > chi2 =    0.1744
```

---

## Likelihood Ratio Tests

- Likelihood ratio tests provide an alternative method of testing sets of coefficients
- To test the coefficients associated with region we need to store our model results. The name is arbitrary, we'll call them m1

```
. estimates store m1
```

- Now we can rerun our model without region

```
. logit highbp bmi c.age##female if e(sample)
```

```
Iteration 0: log likelihood = -7050.7655
Iteration 1: log likelihood = -5826.855
Iteration 2: log likelihood = -5797.9206
Iteration 3: log likelihood = -5797.8856
Iteration 4: log likelihood = -5797.8856
```

```
Logistic regression          Number of obs   =    10,351
                             LR chi2(4)           =    2505.76
                             Prob > chi2          =    0.0000
Log likelihood = -5797.8856  Pseudo R2       =    0.1777
```

highbp	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
bmi	.1376855	.0051366	26.80	0.000	.127618 .147753
age	.0335286	.0018501	18.12	0.000	.0299025 .0371548
female					
0	0	(base)			
1	-1.882479	.1530115	-12.30	0.000	-2.182376 -1.582582
female#c.age					
1	.027615	.0028601	9.66	0.000	.0220092 .0332207
_cons	-5.247536	.1628488	-32.22	0.000	-5.566713 -4.928358

- Adding if e(sample) makes sure the same sample, what Stata calls the *estimation sample*, is used for both models

### Likelihood Ratio Tests (Continued)

- Now we store the second set of estimates

```
. estimates store m2
```

- And use the lrtest command to perform the likelihood ratio test

```
. lrtest m1 m2
```

```
Likelihood-ratio test          LR chi2(3) =    4.97
(Assumption: m2 nested in m1)  Prob > chi2 =    0.1743
```

- We'll restore the results from m1 which includes region even though the terms are not collectively significant

```
. estimates restore m1
```

```
(results m1 are active now)
```

- Now it's as though we just ran the model stored as m1

## Tests of Differences

- test can also be used to the equality of coefficients

```
. test 3.region = 4.region
```

```
( 1) [highbp]3.region - [highbp]4.region = 0
```

```
      chi2( 1) =    0.71
      Prob > chi2 =    0.3978
```

- A likelihood ratio test can also be used; see `help constraint` for information on setting the necessary constraints
- The `lincom` command calculates linear combinations of coefficients, along with standard errors, hypothesis tests, and confidence intervals
- For example, to obtain the difference in coefficients

```
. lincom 3.region - 4.region
```

```
( 1) [highbp]3.region - [highbp]4.region = 0
```

```
-----+-----
      highbp |      Coef.   Std. Err.      z    P>|z|     [95% Conf. Interval]
-----+-----
      (1) |   -0.05226   .0618078    -0.85   0.398    -0.1734012    .0688811
-----+-----
```

---

## 3.2 Predictions

### What are margins?

- Stata defines margins as “statistics calculated from predictions of a previously fit model at fixed values of some covariates and averaging or otherwise integrating over the remaining covariates.”
  - ◇ Also known as counterfactuals, or when we fix a categorical variable, potential outcomes
- What sorts of predictions does margins work with?
  - ◇ Predicted means, probabilities, and counts
  - ◇ Derivatives
  - ◇ Elasticities
- We'll also see contrasts and pairwise comparisons of the above

## Average Predictions

- Let's start with margins in its most basic form

```
. margins

Predictive margins                    Number of obs   =    10,351
Model VCE      : OIM

Expression   : Pr(highbp), predict()

-----+-----
           |               Delta-method
           |      Margin   Std. Err.      z    P>|z|     [95% Conf. Interval]
-----+-----
      _cons |   .4227611   .0042898   98.55   0.000   .4143533   .4311689
-----+-----
```

- What happened here?
  1. The predicted probability of `highbp=1` was calculated for each case, using each case's observed values of `bmi`, `age`, `female`, and `region`
  2. The average of those predictions was calculated and displayed
- Unless we tell it to do otherwise, `margins` works with the estimation sample

---

## Predictions at the Average

- An alternative is to calculate the predicted probability fixing all the covariates at some value, often the mean

```
. margins, atmeans

Adjusted predictions                    Number of obs   =    10,351
Model VCE      : OIM

Expression   : Pr(highbp), predict()
at           : bmi           =    25.5376 (mean)
              age           =    47.57965 (mean)
              0.female      =    .4748333 (mean)
              1.female      =    .5251667 (mean)
              1.region      =    .2024925 (mean)
              2.region      =    .2679934 (mean)
              3.region      =    .2756255 (mean)
              4.region      =    .2538885 (mean)

-----+-----
           |               Delta-method
           |      Margin   Std. Err.      z    P>|z|     [95% Conf. Interval]
-----+-----
      _cons |   .3929783   .0056167   69.97   0.000   .3819697   .4039869
-----+-----
```

- What happened here?
  1. The mean of each independent variable was calculated
  2. The predicted probability of `highbp=1` was calculated using the means from step 1

## Predictions at Each Level of a Factor Variable

- Adding a factor variable specifies that the predictions be repeated at each level of the variable, for example

```
. margins region

Predictive margins                                Number of obs   =    10,351
Model VCE      : OIM

Expression   : Pr(highbp), predict()

-----+-----
           |               Delta-method
           |      Margin   Std. Err.      z    P>|z|     [95% Conf. Interval]
-----+-----
    region |
      NE |   .4362592   .0095422   45.72   0.000   .4175568   .4549616
      MW |   .4103455   .0083278   49.27   0.000   .3940234   .4266677
      S  |   .4190352   .0081188   51.61   0.000   .4031226   .4349477
      W  |   .4290013   .0085434   50.21   0.000   .4122565   .4457461
-----+-----
```

- What happened here?
  - The predicted probability is calculated treating all cases as if `region=1` and using each case's observed values of `bmi`, `age`, and `female`
  - The mean of the predictions from step 1 is calculated
  - Repeat steps 1 and 2 for each value of `region`

## Multiple Factor Variables

- We can obtain margins for multiple variables

```
. margins region female

Predictive margins                                Number of obs   =    10,351
Model VCE      : OIM

Expression   : Pr(highbp), predict()

-----+-----
           |               Delta-method
           |      Margin   Std. Err.      z    P>|z|     [95% Conf. Interval]
-----+-----
    region |
      NE |   .4362592   .0095422   45.72   0.000   .4175568   .4549616
      MW |   .4103455   .0083278   49.27   0.000   .3940234   .4266677
      S  |   .4190352   .0081188   51.61   0.000   .4031226   .4349477
      W  |   .4290013   .0085434   50.21   0.000   .4122565   .4457461
           |
    female |
      0  |   .4692315   .006393   73.40   0.000   .4567014   .4817616
      1  |   .3766361   .0057397   65.62   0.000   .3653866   .3878857
-----+-----
```

- Or combinations of values, for example each combination of region and female

```
. margins region#female
```

```
Predictive margins          Number of obs   =   10,351
Model VCE      : OIM
```

```
Expression      : Pr(highbp), predict()
```

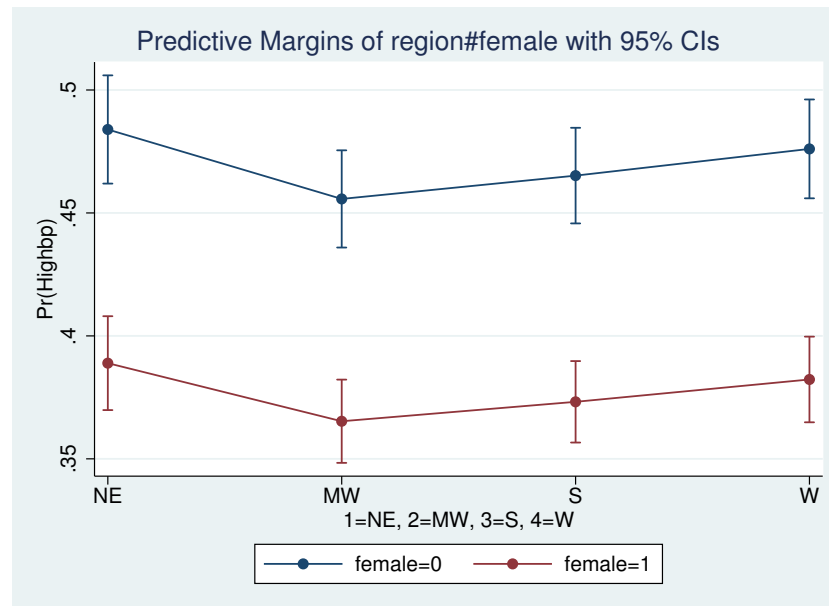
	Delta-method				
	Margin	Std. Err.	z	P> z	[95% Conf. Interval]
region#female					
NE#0	.4839466	.0112276	43.10	0.000	.461941 .5059522
NE#1	.3889131	.0097392	39.93	0.000	.3698246 .4080015
MW#0	.4556986	.0100844	45.19	0.000	.4359337 .4754636
MW#1	.3652888	.0086372	42.29	0.000	.3483602 .3822173
S#0	.4651826	.0099214	46.89	0.000	.4457369 .4846282
S#1	.3731942	.0084524	44.15	0.000	.3566278 .3897605
W#0	.4760455	.0102535	46.43	0.000	.455949 .496142
W#1	.3822812	.0088891	43.01	0.000	.3648589 .3997034

- We can graph the resulting predictions using the marginsplot command

### Graphing Predicted Probabilities

- For example to graph the last set of margins

```
. marginsplot
```



## Predictions at Specified Values of Covariates

- The `at()` option is used to specify values at which margins should be calculated
- To obtain the average predicted probability setting `age=40` specify

```
. margins, at(age=40)
```

```
Predictive margins          Number of obs   =   10,351
Model VCE      : OIM
```

```
Expression   : Pr(highbp), predict()
at           : age           =       40
```

	Margin	Std. Err.	z	P> z	[95% Conf. Interval]	
_cons	.3287856	.0053346	61.63	0.000	.31833	.3392413

- `at()` accepts number lists, so we can obtain predictions setting `age` to 20, 30, ..., 70

```
. margins, at(age=(20(10)70)) vsquish
```

```
Predictive margins          Number of obs   =   10,351
Model VCE      : OIM
```

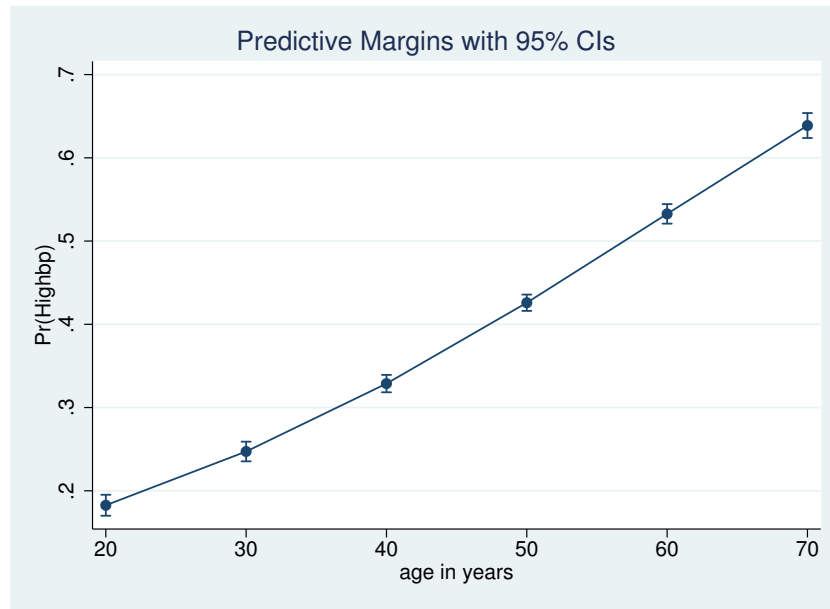
```
Expression   : Pr(highbp), predict()
1._at       : age           =       20
2._at       : age           =       30
3._at       : age           =       40
4._at       : age           =       50
5._at       : age           =       60
6._at       : age           =       70
```

	Margin	Std. Err.	z	P> z	[95% Conf. Interval]	
_at						
1	.1826464	.006436	28.38	0.000	.170032	.1952608
2	.247219	.0060245	41.04	0.000	.2354113	.2590268
3	.3287856	.0053346	61.63	0.000	.31833	.3392413
4	.425936	.0050064	85.08	0.000	.4161236	.4357485
5	.5326646	.0059775	89.11	0.000	.5209488	.5443804
6	.6387994	.0076524	83.48	0.000	.6238009	.6537979

- The `vsquish` option reduces the amount of vertical space the header for `margins` takes up

## Graphing Across Values of Continuous Variables

```
. marginsplot
```



## Specifying Values of Multiple Variables

- We can specify values of multiple variables using `at()`
- If we set values of all the independent variables in our model, we can ask very specific questions
- For example, what is the predicted probability of high blood pressure for an male who is age 40, with a bmi of 25 and living in the midwest (`region=2`)? What is the predicted probability if the person is female?

```
. margins female, at(age=40 bmi=25 region=2)
```

```
Adjusted predictions          Number of obs   =    10,351
Model VCE      : OIM

Expression   : Pr(highbp), predict()
at           : bmi           =    25
              age           =    40
              region        =     2
```

-----						
		Delta-method				
	Margin	Std. Err.	z	P> z	[95% Conf. Interval]	
-----						
female						
0	.3706418	.0118974	31.15	0.000	.3473232	.3939603
1	.2130731	.0096757	22.02	0.000	.194109	.2320372
-----						



- We can use the contrast operator `r.` to compare the predicted probabilities for males and females

```
. margins r.female, at(age=40 bmi=25 region=2)
```

```
Contrasts of adjusted predictions
Model VCE      : OIM
```

```
Expression    : Pr(highbp), predict()
at            : bmi          =          25
              : age          =          40
              : region       =           2
```

```
-----+-----
              |          df          chi2          P>chi2
-----+-----
      female |             1         200.44         0.0000
-----+-----
```

```
-----+-----
              |          Contrast      Std. Err.      [95% Conf. Interval]
-----+-----
      female |
(1 vs 0) |  -.1575687   .0111296   - .1793822   -.1357551
-----+-----
```

- We'll see more on contrasts below

### Specifying Ranges of Multiple Variables

- We can also specify ranges of values for multiple variables, for example multiple values of age and bmi

```
. margins, at(age=(20(10)70) bmi=(20(10)40))
```

- We can also combine the use of factor and continuous variables, for example

```
. margins female, at(age=(20(10)70)) vsquish
```

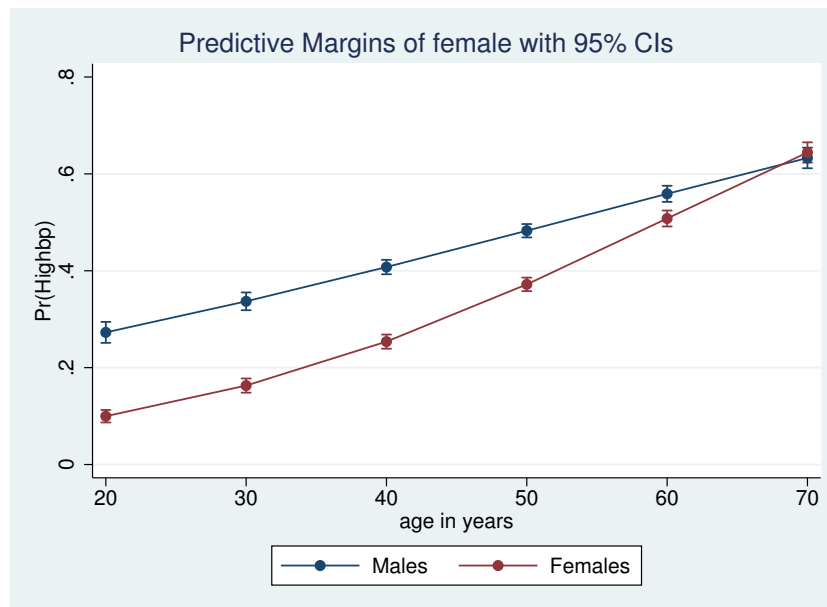
```
Predictive margins          Number of obs   =   10,351
Model VCE      : OIM
```

```
Expression   : Pr(highbp), predict()
1._at       : age           =       20
2._at       : age           =       30
3._at       : age           =       40
4._at       : age           =       50
5._at       : age           =       60
6._at       : age           =       70
```

		Delta-method				
		Margin	Std. Err.	z	P> z	[95% Conf. Interval]
-----						
_at#female						
1	0	.2728133	.0110381	24.72	0.000	.2511789 .2944477
1	1	.0997683	.0065837	15.15	0.000	.0868644 .1126722
2	0	.3369363	.0093499	36.04	0.000	.3186108 .3552617
2	1	.1629921	.0074737	21.81	0.000	.148344 .1776402
3	0	.4076871	.0075866	53.74	0.000	.3928176 .4225566
3	1	.2537634	.0074437	34.09	0.000	.2391741 .2683527
4	0	.4826887	.0070403	68.56	0.000	.46889 .4964874
4	1	.3718821	.0071293	52.16	0.000	.357909 .3858552
5	0	.5588757	.0084852	65.87	0.000	.5422451 .5755063
5	1	.5079403	.0083938	60.51	0.000	.4914886 .5243919
6	0	.6329264	.0108508	58.33	0.000	.6116592 .6541935
6	1	.6442392	.0106744	60.35	0.000	.6233177 .6651607

### More Plots

```
. marginsplot, legend(order(3 "Males" 4 "Females"))
```



- ◇ The standard errors are drawn before the lines for the predictions, so we want the legend to show the third and fourth plots

## More Predictions

- We can use `at()` with the `generate()` suboption to answer different sorts of questions
- For example, what would the averaged predicted probability be if everyone aged 5 years, while their values `female` and `region` remained the same?
- The `generate(age+5)` requests margins calculated at each observations value of `age` plus 5

```
. margins, at(age=generate(age+5))
```

```
Predictive margins                                Number of obs    =    10,351
Model VCE      : OIM
```

```
Expression   : Pr(highbp), predict()
at           : age              = age+5
```

```
-----+-----
```

	Margin	Delta-method Std. Err.	z	P> z	[95% Conf. Interval]
_cons	.4672688	.004476	104.39	0.000	.458496 .4760416

```
-----+-----
```

- We can specify `at()` multiple times, to obtain predictions under different scenarios

```
. margins, at(age=generate(age)) ///
          at(age=generate(age+5)) at(age=generate(age+10))
```

```
Predictive margins                                Number of obs    =    10,351
Model VCE      : OIM
```

```
Expression   : Pr(highbp), predict()
1._at        : age              = age
2._at        : age              = age+5
3._at        : age              = age+10
```

```
-----+-----
```

	Margin	Delta-method Std. Err.	z	P> z	[95% Conf. Interval]
1	.4227611	.0042898	98.55	0.000	.4143533 .4311689
2	.4672688	.004476	104.39	0.000	.458496 .4760416
3	.512185	.0048335	105.97	0.000	.5027115 .5216585

```
-----+-----
```

## Predictions Over Groups

- The `over()` option produces predictions averaging within groups defined by the factor variable, for example, `female`

```
. margins, over(female)
```

```
Predictive margins                Number of obs    =    10,351
Model VCE      : OIM
```

```
Expression   : Pr(highbp), predict()
over         : female
```

		Delta-method				
	Margin	Std. Err.	z	P> z	[95% Conf. Interval]	
female						
0	.4687691	.0066113	70.90	0.000	.4558112	.4817269
1	.3811626	.0055567	68.47	0.000	.3702516	.3920737

- What happened here?
  1. The predicted probability for each case is calculated, using the case's observed values on all variables
  2. The average predicted probability is calculated using only cases where `female=0`
  3. Repeat step 2 using only cases where `female=1`

## Pairwise Comparisons of Predictions

- Earlier we obtained average predicted probabilities at each level of `region` using

```
. margins region
```

- For pairwise comparisons of these margins we can add the `pwcompare` option

```
. margins region, pwcompare
```

```
Pairwise comparisons of predictive margins
Model VCE      : OIM
```

```
Expression   : Pr(highbp), predict()
```

		Delta-method		Unadjusted	
	Contrast	Std. Err.	[95% Conf. Interval]		
region					
MW vs NE	-.0259137	.0126665	-.0507396	-.0010878	
S vs NE	-.017224	.0125288	-.0417801	.007332	
W vs NE	-.0072579	.0128075	-.0323601	.0178443	
S vs MW	.0086896	.0116321	-.0141089	.0314882	
W vs MW	.0186558	.0119339	-.0047343	.0420459	
W vs S	.0099661	.0117862	-.0131345	.0330667	

- Adding the groups option will allow us to see which levels are statistically distinguishable

```
. margins region, pwcompare(groups)

Pairwise comparisons of predictive margins
Model VCE      : OIM

Expression     : Pr(highbp), predict()

-----+-----
          |           Delta-method Unadjusted
          |           Margin  Std. Err.    Groups
-----+-----
region |
  NE   | .4362592  .0095422          B
  MW   | .4103455  .0083278          A
  S    | .4190352  .0081188         AB
  W    | .4290013  .0085434         AB
-----+-----
```

Note: Margins sharing a letter in the group label are not significantly different at the 5% level.

- The pwcompare() option can be used to specify other suboptions; see help margins pwcompare for more information

## Contrasts of Predictions

- The margins command allows *contrast operators* which are used to request comparisons of the margins
  - ◊ In this case the margins are predicted probabilities
- For example, to compare average predicted probabilities setting female=0 versus female=1 add the r. prefix

```
. margins r.female

Contrasts of predictive margins
Model VCE      : OIM

Expression     : Pr(highbp), predict()

-----+-----
          |           df           chi2     P>chi2
-----+-----
female |           1     116.16     0.0000
-----+-----

          |           Delta-method
          |           Contrast  Std. Err.    [95% Conf. Interval]
-----+-----
female |
(1 vs 0) | -.0925953  .0085912    -.1094338   -.0757569
-----+-----
```

- We can use the @ operator to contrast female at each level of region

```
. margins r.female@region
```

```
Contrasts of predictive margins
```

```
Model VCE      : OIM
```

```
Expression    : Pr(highbp), predict()
```

```
-----+-----
```

	df	chi2	P>chi2
female@region			
(1 vs 0) NE	1	117.89	0.0000
(1 vs 0) MW	1	109.28	0.0000
(1 vs 0) S	1	112.04	0.0000
(1 vs 0) W	1	115.96	0.0000
Joint	4	119.65	0.0000

```
-----+-----
```

```
-----+-----
```

	Delta-method		
	Contrast	Std. Err.	[95% Conf. Interval]
female@region			
(1 vs 0) NE	-.0950335	.0087525	-.1121881 -.0778789
(1 vs 0) MW	-.0904099	.0086485	-.1073606 -.0734592
(1 vs 0) S	-.0919884	.0086906	-.1090216 -.0749552
(1 vs 0) W	-.0937643	.0087074	-.1108305 -.0766982

```
-----+-----
```

- This reports the differences in predicted probabilities when female=1 versus female=0 at each level of region

## Contrasts of Predictions (Continued)

- To perform contrasts at different values of a continuous variable use the `at()` option

```
. margins r.female, at(age=(20(10)70)) vsquish
```

```
Contrasts of predictive margins
Model VCE      : OIM
```

```
Expression      : Pr(highbp), predict()
1._at           : age           =      20
2._at           : age           =      30
3._at           : age           =      40
4._at           : age           =      50
5._at           : age           =      60
6._at           : age           =      70
```

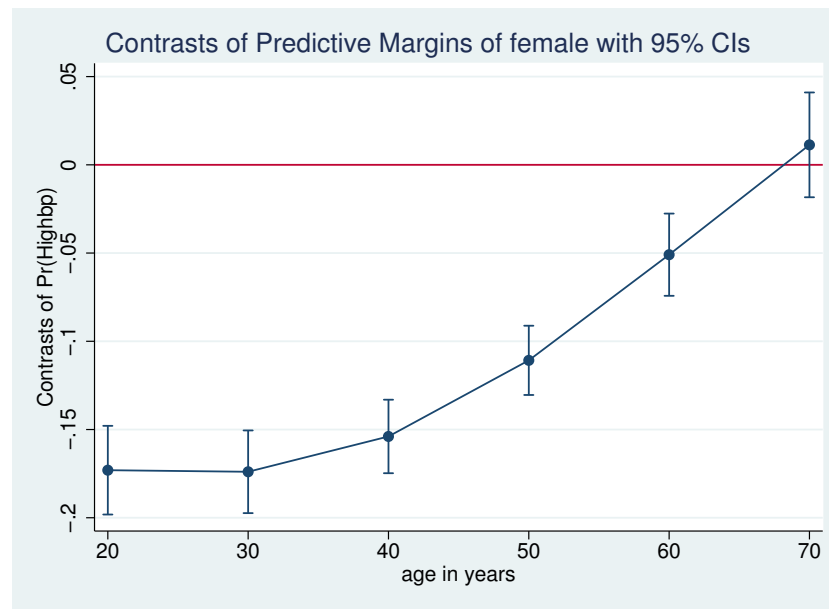
	df	chi2	P>chi2
-----			
female@_at			
(1 vs 0) 1	1	182.15	0.0000
(1 vs 0) 2	1	211.82	0.0000
(1 vs 0) 3	1	209.80	0.0000
(1 vs 0) 4	1	122.51	0.0000
(1 vs 0) 5	1	18.36	0.0000
(1 vs 0) 6	1	0.56	0.4552
Joint	6	123716.83	0.0000

	Contrast	Delta-method Std. Err.	[95% Conf. Interval]	
-----				
female@_at				
(1 vs 0) 1	-.173045	.0128218	-.1981752	-.1479147
(1 vs 0) 2	-.1739442	.0119516	-.1973689	-.1505195
(1 vs 0) 3	-.1539237	.0106268	-.1747518	-.1330956
(1 vs 0) 4	-.1108066	.0100111	-.130428	-.0911851
(1 vs 0) 5	-.0509354	.0118889	-.0742372	-.0276335
(1 vs 0) 6	.0113128	.0151483	-.0183773	.041003

- The output gives tests of the differences in predicted probabilities for `female=1` versus `female=0` at each of the specified values of age
  - The joint test is statistically significant
  - The differences get smaller in absolute value as age increases

## Plotting Contrasts

```
. marginsplot, yline(0)
```



---

## Contrast Operators

- A few common contrast operators are
  - ◇ r. differences from the base (a.k.a. reference) level
  - ◇ a. differences from the next (adjacent) level
  - ◇ ar. differences from the previous level (reverse adjacent)
  - ◇ g. differences from the balanced grand mean
  - ◇ gw. differences from the observation-weighted grand mean
  - ◇ There are also operators for Helmert contrasts and contrasts using orthogonal polynomials for balanced and unbalanced cases

---

## contrast suboptions

- So far we've obtained contrasts using *contrast operators*, but *margins* also allows a `contrast()` option
  - The `contrast()` option is particularly useful for specifying options to contrast
  - For example, to obtain contrasts for continuous variables the `atcontrast()` suboption is used
    - ◇ The `effects` suboption requests a table showing the contrasts along with confidence intervals and p-values
    - ◇ In `atcontrast(a)` the `a` contrast operator requests comparisons of adjacent categories
- ```
. margins, at(age=(20(10)70)) contrast(atcontrast(a) effects) vsquish
```



Contrasts of predictive margins

Model VCE : OIM

Expression : Pr(highbp), predict()

```
1._at : age = 20
2._at : age = 30
3._at : age = 40
4._at : age = 50
5._at : age = 60
6._at : age = 70
```

|          | df | chi2      | P>chi2 |
|----------|----|-----------|--------|
| _at      |    |           |        |
| (1 vs 2) | 1  | 1947.70   | 0.0000 |
| (2 vs 3) | 1  | 1565.44   | 0.0000 |
| (3 vs 4) | 1  | 1191.96   | 0.0000 |
| (4 vs 5) | 1  | 1064.35   | 0.0000 |
| (5 vs 6) | 1  | 1301.80   | 0.0000 |
| Joint    | 5  | 278027.30 | 0.0000 |

|          | Contrast  | Std. Err. | z      | P> z  | [95% Conf. Interval] |           |
|----------|-----------|-----------|--------|-------|----------------------|-----------|
| _at      |           |           |        |       |                      |           |
| (1 vs 2) | -.0645726 | .0014631  | -44.13 | 0.000 | -.0674403            | -.0617049 |
| (2 vs 3) | -.0815666 | .0020616  | -39.57 | 0.000 | -.0856072            | -.077526  |
| (3 vs 4) | -.0971504 | .0028139  | -34.52 | 0.000 | -.1026656            | -.0916352 |
| (4 vs 5) | -.1067286 | .0032714  | -32.62 | 0.000 | -.1131405            | -.1003167 |
| (5 vs 6) | -.1061348 | .0029416  | -36.08 | 0.000 | -.1119002            | -.1003693 |

**Contrasts with generate()**

- Earlier we used the generate() suboption to obtain predicted probabilities modifying the observed values

- Specifically, we obtained predicted probabilities using each case's observed value of age and each case's observed value +5 years

```
. margins, at(age=generate(age)) at(age=generate(age+5))
```

Predictive margins  
Model VCE : OIM

Number of obs = 10,351

Expression : Pr(highbp), predict()

1.\_at : age = age

2.\_at : age = age+5

|     | Margin   | Std. Err. | z      | P> z  | [95% Conf. Interval] |          |
|-----|----------|-----------|--------|-------|----------------------|----------|
| _at |          |           |        |       |                      |          |
| 1   | .4227611 | .0042898  | 98.55  | 0.000 | .4143533             | .4311689 |
| 2   | .4672688 | .004476   | 104.39 | 0.000 | .458496              | .4760416 |

- Using the contrast option, we can compare the two

```
. margins, at(age=generate(age)) ///
    at(age=generate(age+5)) contrast(atcontrast(r))
```

Contrasts of predictive margins  
Model VCE : OIM

Expression : Pr(highbp), predict()

1.\_at : age = age

2.\_at : age = age+5

|     | df | chi2    | P>chi2 |
|-----|----|---------|--------|
| _at | 1  | 1728.47 | 0.0000 |

|          | Contrast | Std. Err. | [95% Conf. Interval] |          |
|----------|----------|-----------|----------------------|----------|
| _at      |          |           |                      |          |
| (2 vs 1) | .0445077 | .0010705  | .0424095             | .0466059 |

## Contrasts of Differences

- We can also request contrasts of contrasts by combining contrast operators
- For example, to compare the differences between males and females across levels of region use

```
. margins r.female#r.region
```

```
Contrasts of predictive margins
```

```
Model VCE      : OIM
```

```
Expression    : Pr(highbp), predict()
```

---

|                     | df | chi2 | P>chi2 |
|---------------------|----|------|--------|
| -----+-----         |    |      |        |
| female#region       |    |      |        |
| (1 vs 0) (MW vs NE) | 1  | 4.11 | 0.0426 |
| (1 vs 0) (S vs NE)  | 1  | 1.88 | 0.1703 |
| (1 vs 0) (W vs NE)  | 1  | 0.32 | 0.5709 |
| Joint               | 3  | 4.83 | 0.1851 |

---

---

|                     | Contrast | Delta-method<br>Std. Err. | [95% Conf. Interval] |          |
|---------------------|----------|---------------------------|----------------------|----------|
| -----+-----         |          |                           |                      |          |
| female#region       |          |                           |                      |          |
| (1 vs 0) (MW vs NE) | .0046236 | .0022806                  | .0001537             | .0090935 |
| (1 vs 0) (S vs NE)  | .0030451 | .0022208                  | -.0013077            | .0073979 |
| (1 vs 0) (W vs NE)  | .0012692 | .0022396                  | -.0031203            | .0056586 |

---

---

## Adjusting for Multiple Comparisons

- Use of contrast and pwcompare can result in a large number of hypothesis tests
- The mcompare() option can be used to adjust p-values and confidence intervals for multiple comparisons within factor variable terms
- The available methods are
  - ◇ noadjust
  - ◇ bonferroni
  - ◇ sidak
  - ◇ scheffe

---

## Using mcompare()

- To apply Bonferroni's adjustment to an earlier contrast

```
. margins r.female@region, mcompare(bonferroni)
```

Contrasts of predictive margins

Model VCE : OIM

Expression : Pr(highbp), predict()

|               | df | chi2   | P>chi2 | Bonferroni<br>P>chi2 |
|---------------|----|--------|--------|----------------------|
| female@region |    |        |        |                      |
| (1 vs 0) NE   | 1  | 117.89 | 0.0000 | 0.0000               |
| (1 vs 0) MW   | 1  | 109.28 | 0.0000 | 0.0000               |
| (1 vs 0) S    | 1  | 112.04 | 0.0000 | 0.0000               |
| (1 vs 0) W    | 1  | 115.96 | 0.0000 | 0.0000               |
| Joint         | 4  | 119.65 | 0.0000 |                      |

Note: Bonferroni-adjusted p-values are reported for tests on individual contrasts only.

|               | Number of<br>Comparisons |
|---------------|--------------------------|
| female@region | 4                        |

|               | Contrast  | Delta-method<br>Std. Err. | Bonferroni<br>[95% Conf. Interval] |
|---------------|-----------|---------------------------|------------------------------------|
| female@region |           |                           |                                    |
| (1 vs 0) NE   | -.0950335 | .0087525                  | -.1168946 -.0731723                |
| (1 vs 0) MW   | -.0904099 | .0086485                  | -.1120112 -.0688085                |
| (1 vs 0) S    | -.0919884 | .0086906                  | -.1136949 -.0702819                |
| (1 vs 0) W    | -.0937643 | .0087074                  | -.1155128 -.0720159                |

- Specifying adjusted p-values with the `pwcompare` option

```
. margins region, mcompare(sidak) pwcompare
```

Pairwise comparisons of predictive margins

Model VCE : OIM

Expression : Pr(highbp), predict()

```
-----+-----
```

|        | Number of<br>Comparisons |
|--------|--------------------------|
| region | 6                        |

```
-----+-----
```

|          | Delta-method<br>Contrast | Std. Err. | Sidak<br>[95% Conf. Interval] |
|----------|--------------------------|-----------|-------------------------------|
| region   |                          |           |                               |
| MW vs NE | -.0259137                | .0126665  | -.0592398 .0074124            |
| S vs NE  | -.017224                 | .0125288  | -.0501878 .0157398            |
| W vs NE  | -.0072579                | .0128075  | -.0409548 .026439             |
| S vs MW  | .0086896                 | .0116321  | -.021915 .0392943             |
| W vs MW  | .0186558                 | .0119339  | -.0127429 .0500544            |
| W vs S   | .0099661                 | .0117862  | -.0210439 .0409762            |

```
-----+-----
```

### 3.3 Marginal Effects

#### Marginal Effects

- In a straightforward linear model, the marginal effect of a variable is the coefficient  $b$

$$y = b_0 + b_1x_1 + b_2x_2 + e$$

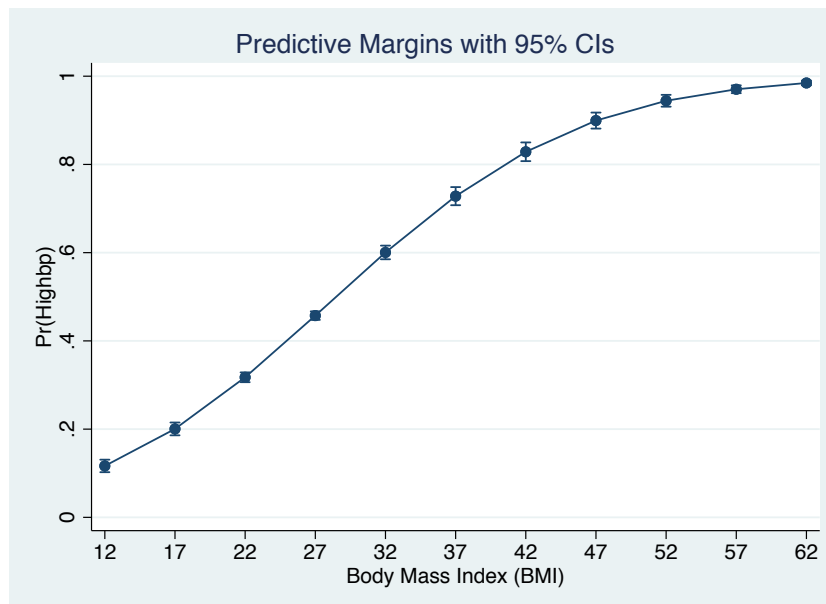
- In more complex models, this is no longer true
  - ◇ models with interactions
  - ◇ models with polynomial terms
  - ◇ generalized linear models when the margin is not on the linear scale
- For example, in a logistic regression model, the marginal effect of covariates is not constant on the probability scale
- `margins` can be used to estimate the margins of the derivative of a response

#### A Closer Look at Slopes

- Here is a graph of predicted probabilities across values of `bmi`

```
. margins, at(bmi=(12(5)62))
```

```
. marginsplot
```



### Average Marginal Effects

- The slope of bmi is not constant, but we might want to know what it is on average
- We can obtain the average marginal effect of bmi

```
. margins, dydx(bmi)
```

```
Average marginal effects          Number of obs   =   10,351
Model VCE      : OIM
```

```
Expression      : Pr(highbp), predict()
dy/dx w.r.t.    : bmi
```

|     | Delta-method |           |       |       |                      |
|-----|--------------|-----------|-------|-------|----------------------|
|     | dy/dx        | Std. Err. | z     | P> z  | [95% Conf. Interval] |
| bmi | .0262514     | .000852   | 30.81 | 0.000 | .0245816 .0279212    |

- What happened here?
  1. Calculate the derivative of the predicted probability with respect to bmi for each observaton
  2. Calculate the average of derivatives from step 1

- We can do the same for all variables in our model

```
. margins, dydx(*)
```

```
Average marginal effects          Number of obs   =    10,351
Model VCE      : OIM
```

```
Expression   : Pr(highbp), predict()
dy/dx w.r.t. : bmi age 1.female 2.region 3.region 4.region
```

```
-----
```

|        |           | Delta-method |        |       | [95% Conf. Interval] |           |
|--------|-----------|--------------|--------|-------|----------------------|-----------|
|        | dy/dx     | Std. Err.    | z      | P> z  |                      |           |
| bmi    | .0262514  | .000852      | 30.81  | 0.000 | .0245816             | .0279212  |
| age    | .0088181  | .0002145     | 41.11  | 0.000 | .0083976             | .0092385  |
| female |           |              |        |       |                      |           |
| 0      | 0         | (base)       |        |       |                      |           |
| 1      | -.0925953 | .0085912     | -10.78 | 0.000 | -.1094338            | -.0757569 |
| region |           |              |        |       |                      |           |
| NE     | 0         | (base)       |        |       |                      |           |
| MW     | -.0259137 | .0126665     | -2.05  | 0.041 | -.0507396            | -.0010878 |
| S      | -.017224  | .0125288     | -1.37  | 0.169 | -.0417801            | .007332   |
| W      | -.0072579 | .0128075     | -0.57  | 0.571 | -.0323601            | .0178443  |

```
-----
```

Note: dy/dx for factor levels is the discrete change from the base level.

## Marginal Effects Over the Response Surface

- It can also be informative to estimate the marginal effect of  $x$  at different values of  $x$
- For example, we can obtain the derivative with respect to age at age=20, 30, ..., 70

```
. margins, dydx(age) at(age=(20(10)70)) vsquish
```

```
Average marginal effects          Number of obs   =    10,351
Model VCE      : OIM
```

```
Expression      : Pr(highbp), predict()
dy/dx w.r.t.    : age
1._at           : age           =    20
2._at           : age           =    30
3._at           : age           =    40
4._at           : age           =    50
5._at           : age           =    60
6._at           : age           =    70
```

|     |     | Delta-method |           | z     | P> z  | [95% Conf. Interval] |          |
|-----|-----|--------------|-----------|-------|-------|----------------------|----------|
|     |     | dy/dx        | Std. Err. |       |       |                      |          |
| age |     |              |           |       |       |                      |          |
|     | _at |              |           |       |       |                      |          |
|     | 1   | .0056454     | .0001263  | 44.70 | 0.000 | .0053978             | .0058929 |
|     | 2   | .0072988     | .0001734  | 42.09 | 0.000 | .0069589             | .0076387 |
|     | 3   | .0089942     | .000245   | 36.71 | 0.000 | .008514              | .0094744 |
|     | 4   | .0103355     | .0003148  | 32.83 | 0.000 | .0097184             | .0109526 |
|     | 5   | .0108342     | .0003262  | 33.21 | 0.000 | .0101949             | .0114736 |
|     | 6   | .0102041     | .0002508  | 40.69 | 0.000 | .0097125             | .0106957 |



- Here we do something similar, setting female=0 and then female=1

```
. margins female, dydx(age) at(age=(20(10)70)) vsquish
```

```
Average marginal effects          Number of obs   =    10,351
Model VCE      : OIM
```

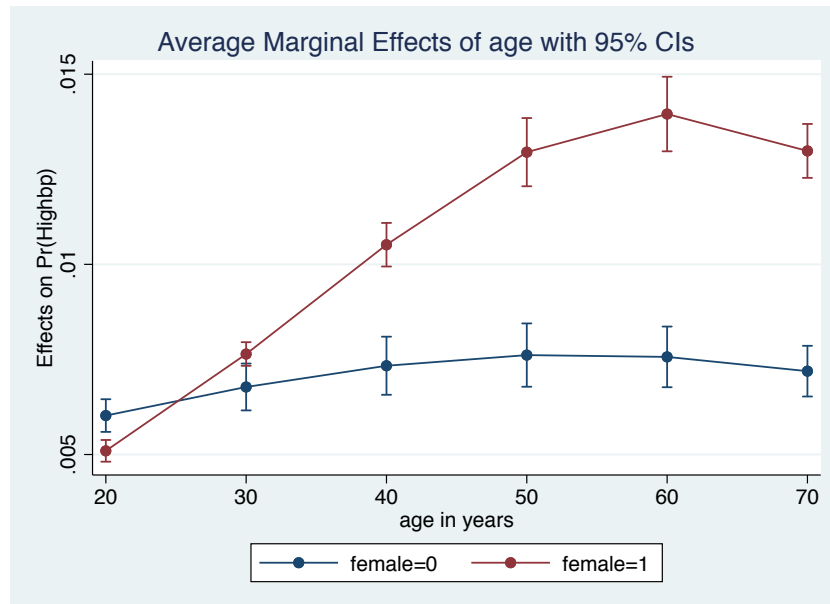
```
Expression   : Pr(highbp), predict()
dy/dx w.r.t. : age
1._at       : age           =    20
2._at       : age           =    30
3._at       : age           =    40
4._at       : age           =    50
5._at       : age           =    60
6._at       : age           =    70
```

|       |            | Delta-method |           |       |       |                      |
|-------|------------|--------------|-----------|-------|-------|----------------------|
|       |            | dy/dx        | Std. Err. | z     | P> z  | [95% Conf. Interval] |
| ----- |            |              |           |       |       |                      |
| age   |            |              |           |       |       |                      |
|       | _at#female |              |           |       |       |                      |
|       | 1 0        | .0060242     | .0002192  | 27.48 | 0.000 | .0055945 .0064538    |
|       | 1 1        | .0050964     | .0001457  | 34.98 | 0.000 | .0048108 .005382     |
|       | 2 0        | .0067761     | .0003143  | 21.56 | 0.000 | .0061601 .007392     |
|       | 2 1        | .0076423     | .0001587  | 48.17 | 0.000 | .0073313 .0079532    |
|       | 3 0        | .0073341     | .0003896  | 18.82 | 0.000 | .0065704 .0080978    |
|       | 3 1        | .0105163     | .0002922  | 35.99 | 0.000 | .0099436 .011089     |
|       | 4 0        | .0076144     | .0004244  | 17.94 | 0.000 | .0067825 .0084463    |
|       | 4 1        | .0129499     | .0004576  | 28.30 | 0.000 | .0120531 .0138467    |
|       | 5 0        | .0075668     | .000407   | 18.59 | 0.000 | .006769 .0083645     |
|       | 5 1        | .0139526     | .0005002  | 27.89 | 0.000 | .0129722 .0149331    |
|       | 6 0        | .0071918     | .00034    | 21.15 | 0.000 | .0065255 .0078581    |
|       | 6 1        | .0129829     | .0003617  | 35.90 | 0.000 | .012274 .0136917     |

### Plots of Marginal Effects

- We can, of course, plot these marginal effects, to see how they change with different values of female and age

```
. marginsplot
```



---

### 3.4 Other Models

#### `margins` with Other Estimation Commands

- `margins` works after most estimation commands
- The default prediction for `margins` is the same as the default prediction for `predict` after a given command
- See help `command` `postestimation` for information on postestimation commands and their defaults after a given command
- You can specify different predictions from `margins` using the `predict()` option

## Modeling Household Size

- For the next set of examples we will model the number of individuals in a household (`houssiz`) using a Poisson model
- Our model will include covariates `age`, `age2`, `region`, `rural`, and a `region` by `rural` interaction
- We've been working with `age` and `region` but we'll take a look at the new variables

```
. codebook houssiz rural
```

```
-----  
houssiz                                     # persons in household, 1-14  
-----
```

```
      type: numeric (byte)  
      range: [1,14]                units: 1  
unique values: 14                  missing .: 0/10,351  
  
      mean: 2.94377  
      std. dev: 1.69516  
  
percentiles:      10%      25%      50%      75%      90%  
                  1        2        2        4        5
```

```
-----  
rural   1=rural, 0=urban  
-----
```

```
      type: numeric (byte)  
      range: [0,1]                 units: 1  
unique values: 2                   missing .: 0/10,351  
  
tabulation: Freq. Value  
             6,548 0  
             3,803 1
```

- Now we can fit our model

```
. poisson houssiz i.region##i.rural age c.age#c.age
```

```
Iteration 0: log likelihood = -18385.275
Iteration 1: log likelihood = -18385.272
Iteration 2: log likelihood = -18385.272
```

```
Poisson regression                Number of obs   =    10,351
                                LR chi2(9)         =    1780.26
                                Prob > chi2        =    0.0000
Log likelihood = -18385.272      Pseudo R2       =    0.0462
```

| houssiz      | Coef.     | Std. Err. | z      | P> z  | [95% Conf. Interval] |           |
|--------------|-----------|-----------|--------|-------|----------------------|-----------|
| -----        |           |           |        |       |                      |           |
| region       |           |           |        |       |                      |           |
| NE           | 0         | (base)    |        |       |                      |           |
| MW           | -.0586473 | .0204129  | -2.87  | 0.004 | -.0986558            | -.0186387 |
| S            | .0021845  | .021345   | 0.10   | 0.918 | -.0396509            | .04402    |
| W            | -.0305816 | .0208232  | -1.47  | 0.142 | -.0713943            | .0102311  |
|              |           |           |        |       |                      |           |
| rural        |           |           |        |       |                      |           |
| 0            | 0         | (base)    |        |       |                      |           |
| 1            | .0441422  | .0278741  | 1.58   | 0.113 | -.0104901            | .0987745  |
|              |           |           |        |       |                      |           |
| region#rural |           |           |        |       |                      |           |
| MW#1         | .0474625  | .036487   | 1.30   | 0.193 | -.0240508            | .1189758  |
| S#1          | -.0013947 | .0352449  | -0.04  | 0.968 | -.0704734            | .0676839  |
| W#1          | .0300379  | .0366293  | 0.82   | 0.412 | -.0417541            | .10183    |
|              |           |           |        |       |                      |           |
| age          | .0561718  | .0025069  | 22.41  | 0.000 | .0512584             | .0610852  |
|              |           |           |        |       |                      |           |
| c.age#c.age  | -.0007312 | .0000272  | -26.87 | 0.000 | -.0007845            | -.0006779 |
|              |           |           |        |       |                      |           |
| _cons        | .2472973  | .0539633  | 4.58   | 0.000 | .1415311             | .3530634  |
| -----        |           |           |        |       |                      |           |

### margins after poisson

- predict's default after poisson is the predicted count
- To obtain the average predicted count, using the observed values of all covarites use

```
. margins
```

```
Predictive margins                Number of obs   =    10,351
Model VCE      : OIM
```

```
Expression   : Predicted number of events, predict()
```

|       | Margin   | Delta-method<br>Std. Err. | z      | P> z  | [95% Conf. Interval] |          |
|-------|----------|---------------------------|--------|-------|----------------------|----------|
| ----- |          |                           |        |       |                      |          |
| _cons | 2.943774 | .016864                   | 174.56 | 0.000 | 2.910721             | 2.976826 |
| ----- |          |                           |        |       |                      |          |

- As before, we can request predicted counts at specified values of factor variables

```
. margins region#rural
```

```
Predictive margins      Number of obs   =   10,351
Model VCE      : OIM
```

```
Expression      : Predicted number of events, predict()
```

|              |          | Delta-method |       |       | [95% Conf. Interval] |          |
|--------------|----------|--------------|-------|-------|----------------------|----------|
|              | Margin   | Std. Err.    | z     | P> z  |                      |          |
| region#rural |          |              |       |       |                      |          |
| NE#0         | 2.942144 | .0441807     | 66.59 | 0.000 | 2.855552             | 3.028737 |
| NE#1         | 3.074926 | .0722057     | 42.59 | 0.000 | 2.933405             | 3.216447 |
| MW#0         | 2.774558 | .0383527     | 72.34 | 0.000 | 2.699388             | 2.849728 |
| MW#1         | 3.040725 | .0579537     | 52.47 | 0.000 | 2.927138             | 3.154312 |
| S#0          | 2.948578 | .0447353     | 65.91 | 0.000 | 2.860899             | 3.036258 |
| S#1          | 3.077355 | .0472768     | 65.09 | 0.000 | 2.984695             | 3.170016 |
| W#0          | 2.853531 | .0411629     | 69.32 | 0.000 | 2.772853             | 2.934209 |
| W#1          | 3.073255 | .0580446     | 52.95 | 0.000 | 2.959489             | 3.18702  |

- And continuous variables

```
. margins, at(age=(20(10)70)) vsquish
```

```
Predictive margins      Number of obs   =   10,351
Model VCE      : OIM
```

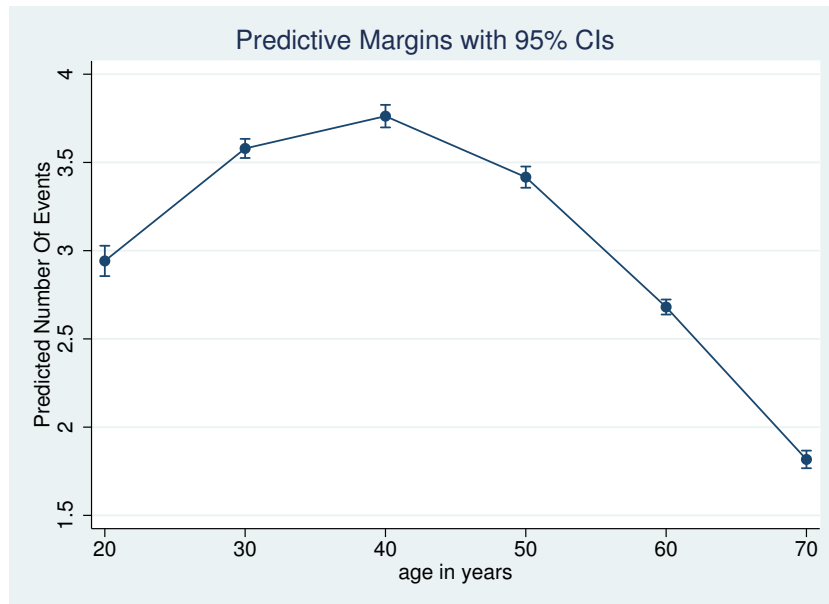
```
Expression      : Predicted number of events, predict()
```

```
1._at      : age      =      20
2._at      : age      =      30
3._at      : age      =      40
4._at      : age      =      50
5._at      : age      =      60
6._at      : age      =      70
```

|     |          | Delta-method |        |       | [95% Conf. Interval] |          |
|-----|----------|--------------|--------|-------|----------------------|----------|
|     | Margin   | Std. Err.    | z      | P> z  |                      |          |
| _at |          |              |        |       |                      |          |
| 1   | 2.94187  | .0438937     | 67.02  | 0.000 | 2.85584              | 3.0279   |
| 2   | 3.579277 | .0276575     | 129.41 | 0.000 | 3.525069             | 3.633484 |
| 3   | 3.762318 | .0326109     | 115.37 | 0.000 | 3.698402             | 3.826234 |
| 4   | 3.416678 | .0306675     | 111.41 | 0.000 | 3.356571             | 3.476785 |
| 5   | 2.680655 | .0216814     | 123.64 | 0.000 | 2.63816              | 2.72315  |
| 6   | 1.817047 | .0254912     | 71.28  | 0.000 | 1.767085             | 1.867009 |

## Plotting Predicted Counts

```
. marginsplot
```



### Other Margins

- After poisson, margins can be used to predict the following
  - ◊ n number of events; the default
  - ◊ ir incidence rate,  $\exp(xb)$ , n when the exposure variable = 1
  - ◊ pr(n) probability that  $y=n$
  - ◊ pr(a,b) probability that  $a \leq y \leq b$
  - ◊ xb the linear prediction

- Predicted probability that houssiz=1

```
. margins rural, predict(pr(1))
```

```
Predictive margins                                Number of obs   =   10,351
Model VCE    : OIM
```

```
Expression   : Pr(houssiz=1), predict(pr(1))
```

| ----- |          |              |       |       |                      |          |
|-------|----------|--------------|-------|-------|----------------------|----------|
|       |          | Delta-method |       |       |                      |          |
|       | Margin   | Std. Err.    | z     | P> z  | [95% Conf. Interval] |          |
| ----- |          |              |       |       |                      |          |
| rural |          |              |       |       |                      |          |
| 0     | .1714666 | .0020282     | 84.54 | 0.000 | .1674915             | .1754417 |
| 1     | .1541716 | .0025566     | 60.30 | 0.000 | .1491608             | .1591823 |
| ----- |          |              |       |       |                      |          |

- Predicted probability that  $3 \leq \text{houssiz} \leq 5$

```
. margins region#rural, predict(pr(3,5))
```

```
Predictive margins                                Number of obs    =    10,351
Model VCE      : OIM
```

```
Expression   : Pr(3<=houssiz<=5), predict(pr(3,5))
```

```
-----
```

|              |          | Delta-method |        |       | [95% Conf. Interval] |          |
|--------------|----------|--------------|--------|-------|----------------------|----------|
|              | Margin   | Std. Err.    | z      | P> z  |                      |          |
| region#rural |          |              |        |       |                      |          |
| NE#0         | .4557062 | .0047091     | 96.77  | 0.000 | .4464765             | .464936  |
| NE#1         | .4682528 | .0063677     | 73.54  | 0.000 | .4557723             | .4807332 |
| MW#0         | .4365671 | .0049383     | 88.41  | 0.000 | .4268883             | .4462459 |
| MW#1         | .4652407 | .005386      | 86.38  | 0.000 | .4546843             | .4757971 |
| S#0          | .4563673 | .0047189     | 96.71  | 0.000 | .4471185             | .4656162 |
| S#1          | .468461  | .004296      | 109.05 | 0.000 | .460041              | .4768809 |
| W#0          | .4460472 | .004858      | 91.82  | 0.000 | .4365256             | .4555688 |
| W#1          | .4681091 | .0051371     | 91.12  | 0.000 | .4580405             | .4781777 |

```
-----
```

## Multiple Responses

- Starting in Stata 14, margins can compute margins for multiple responses at the same time
  - ◊ After, for example, ologit, mlogit, mvreg
- To demonstrate this, we'll model self-rated health in a different version of the NHANES dataset

```
. webuse nhanes2f
. codebook health
```

```
-----
health                                1=poor,..., 5=excellent
-----
```

```

      type:  numeric (byte)
      label:  hlthgrp

      range:  [1,5]
unique values: 5
      units:  1
missing   .:  2/10,337

      tabulation:  Freq.  Numeric  Label
                   729      1  poor
                   1,670    2  fair
                   2,938    3  average
                   2,591    4  good
                   2,407    5  excellent
                   2      .
```

- Our model is

```
. ologit health i.female age c.age#c.age
```

```
Iteration 0: log likelihood = -15764.397
Iteration 1: log likelihood = -15042.53
Iteration 2: log likelihood = -15036.362
Iteration 3: log likelihood = -15036.355
Iteration 4: log likelihood = -15036.355
```

```
Ordered logistic regression          Number of obs   =   10,335
                                      LR chi2(3)       =   1456.09
                                      Prob > chi2      =    0.0000
Log likelihood = -15036.355          Pseudo R2      =    0.0462
```

| health      | Coef.     | Std. Err. | z     | P> z  | [95% Conf. Interval] |           |
|-------------|-----------|-----------|-------|-------|----------------------|-----------|
| -----       |           |           |       |       |                      |           |
| female      |           |           |       |       |                      |           |
| 0           | 0         | (base)    |       |       |                      |           |
| 1           | -.1223788 | .0355107  | -3.45 | 0.001 | -.1919786            | -.052779  |
|             |           |           |       |       |                      |           |
| age         | -.0251916 | .0076063  | -3.31 | 0.001 | -.0400997            | -.0102834 |
|             |           |           |       |       |                      |           |
| c.age#c.age | -.00016   | .0000812  | -1.97 | 0.049 | -.0003191            | -9.73e-07 |
|             | -----     |           |       |       |                      |           |
| /cut1       | -4.442363 | .1659171  |       |       | -4.767554            | -4.117171 |
| /cut2       | -2.975821 | .1632372  |       |       | -3.29576             | -2.655882 |
| /cut3       | -1.573015 | .1618158  |       |       | -1.890168            | -1.255862 |
| /cut4       | -.3384551 | .1606298  |       |       | -.6532838            | -.0236264 |

## Specifying the Response

- By default margins will produce the average predicted probability of each value of health

```
. margins
```

```
Predictive margins          Number of obs   =   10,335
Model VCE      : OIM
```

```
1._predict : Pr(health==1), predict(pr outcome(1))
2._predict : Pr(health==2), predict(pr outcome(2))
3._predict : Pr(health==3), predict(pr outcome(3))
4._predict : Pr(health==4), predict(pr outcome(4))
5._predict : Pr(health==5), predict(pr outcome(5))
```

| _predict | Delta-method |           | z     | P> z  | [95% Conf. Interval] |          |
|----------|--------------|-----------|-------|-------|----------------------|----------|
|          | Margin       | Std. Err. |       |       |                      |          |
| 1        | .0709472     | .0024959  | 28.43 | 0.000 | .0660554             | .075839  |
| 2        | .1643302     | .0035781  | 45.93 | 0.000 | .1573172             | .1713432 |
| 3        | .2868785     | .0044083  | 65.08 | 0.000 | .2782384             | .2955187 |
| 4        | .2474815     | .004184   | 59.15 | 0.000 | .239281              | .255682  |
| 5        | .2303626     | .0039468  | 58.37 | 0.000 | .222627              | .2380981 |



- To request a single outcome we can use `predict(outcome(#))`

```
. margins, predict(outcome(2))
```

```
Predictive margins          Number of obs   =    10,335
Model VCE      : OIM
```

```
Expression   : Pr(health==2), predict(outcome(2))
```

```
-----+-----
```

|             | Delta-method |           |       |       |          | [95% Conf. Interval] |  |
|-------------|--------------|-----------|-------|-------|----------|----------------------|--|
|             | Margin       | Std. Err. | z     | P> z  |          |                      |  |
| -----+----- |              |           |       |       |          |                      |  |
| _cons       | .1643302     | .0035781  | 45.93 | 0.000 | .1573172 | .1713432             |  |
| -----+----- |              |           |       |       |          |                      |  |

- For multiple responses from a single command, repeat the `predict()` option

```
. margins, predict(outcome(1)) predict(outcome(2))
```

```
Predictive margins          Number of obs   =    10,335
Model VCE      : OIM
```

```
1. _predict   : Pr(health==1), predict(outcome(1))
```

```
2. _predict   : Pr(health==2), predict(outcome(2))
```

```
-----+-----
```

|             | Delta-method |           |       |       |          | [95% Conf. Interval] |  |
|-------------|--------------|-----------|-------|-------|----------|----------------------|--|
|             | Margin       | Std. Err. | z     | P> z  |          |                      |  |
| -----+----- |              |           |       |       |          |                      |  |
| _predict    |              |           |       |       |          |                      |  |
| 1           | .0709472     | .0024959  | 28.43 | 0.000 | .0660554 | .075839              |  |
| 2           | .1643302     | .0035781  | 45.93 | 0.000 | .1573172 | .1713432             |  |
| -----+----- |              |           |       |       |          |                      |  |

- To obtain predictions across values of age

```
. margins, at(age=(20(10)70)) pr(out(1)) pr(out(2)) vsquish
```

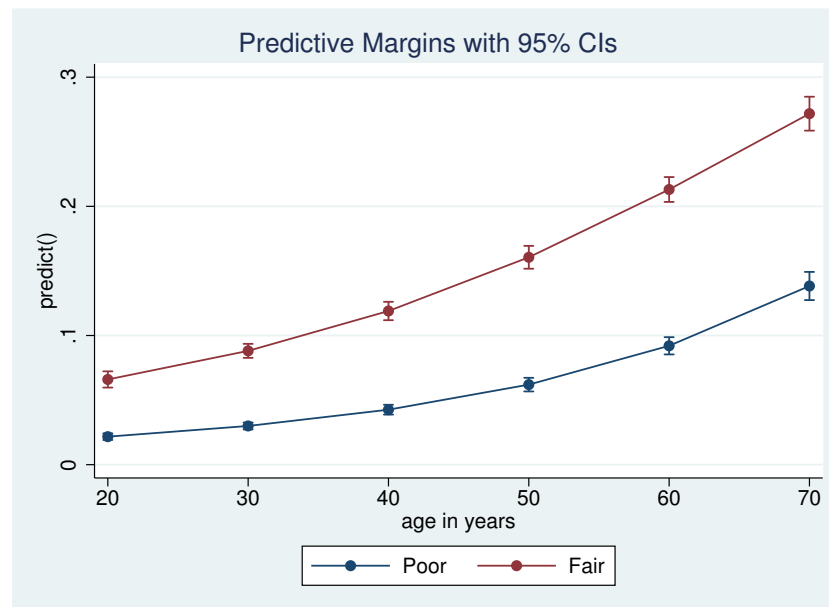
```
Predictive margins          Number of obs    =    10,335
Model VCE      : OIM
```

```
1._predict   : Pr(health==1), predict(out(1))
2._predict   : Pr(health==2), predict(out(2))
1._at        : age           =           20
2._at        : age           =           30
3._at        : age           =           40
4._at        : age           =           50
5._at        : age           =           60
6._at        : age           =           70
```

|               |   | Delta-method |           |       |       |                      |
|---------------|---|--------------|-----------|-------|-------|----------------------|
|               |   | Margin       | Std. Err. | z     | P> z  | [95% Conf. Interval] |
| ._predict#_at |   |              |           |       |       |                      |
| 1             | 1 | .0217005     | .0013107  | 16.56 | 0.000 | .0191315 .0242695    |
| 1             | 2 | .0299861     | .001366   | 21.95 | 0.000 | .0273087 .0326635    |
| 1             | 3 | .0425874     | .0019332  | 22.03 | 0.000 | .0387984 .0463763    |
| 1             | 4 | .0619896     | .0026898  | 23.05 | 0.000 | .0567177 .0672616    |
| 1             | 5 | .0920429     | .0034083  | 27.01 | 0.000 | .0853627 .0987231    |
| 1             | 6 | .1383404     | .0055654  | 24.86 | 0.000 | .1274324 .1492485    |
| 2             | 1 | .0659885     | .0032038  | 20.60 | 0.000 | .0597092 .0722678    |
| 2             | 2 | .0881333     | .0027672  | 31.85 | 0.000 | .0827097 .0935568    |
| 2             | 3 | .1189848     | .0036317  | 32.76 | 0.000 | .1118668 .1261029    |
| 2             | 4 | .1605636     | .0045152  | 35.56 | 0.000 | .151714 .1694132     |
| 2             | 5 | .2130434     | .0049117  | 43.37 | 0.000 | .2034167 .2226701    |
| 2             | 6 | .2717448     | .0066991  | 40.56 | 0.000 | .2586149 .2848748    |

### Plots with Multiple Responses

```
. marginsplot, legend(order(3 "Poor" 4 "Fair"))
```



---

## 4 Conclusion

### 4.1 Conclusion

#### Conclusion

- We've seen how to obtain a variety of predictions and marginal effects after regression models
  - We now know how to perform contrasts of predictions and marginal effects
  - We've also seen how to graph these results
-

# Index