

Linear Regression Models with Interaction/Moderation

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1 Introduction

1.1 Goals

Goals

- Learn how to use factor variable notation when fitting models involving
 - ◇ Categorical variables
 - ◇ Interactions
 - ◇ Polynomial terms
 - Learn how to use postestimation tools to interpret interactions
 - ◇ Tests for group differences
 - ◇ Tests of slopes
 - ◇ Graphs
-

A Linear Model

- We'll use data from the National Health and Nutrition Examination Survey (NHANES) for our examples

```
. webuse nhanes2
```

- We'll start with a basic a model for bmi using age and sex (female).

- Before we fit the model, let's investigate the variables using codebook

```
. codebook bmi age female
```

```
-----
bmi                                     Body Mass Index (BMI)
-----
      type:  numeric (float)

      range:  [12.385596,61.129696]      units:  1.000e-07
unique values: 9,941                    missing .:  0/10,351

      mean:   25.5376
      std. dev: 4.91497

percentiles:      10%      25%      50%      75%      90%
                  20.1037  22.142  24.8181  28.0267  31.7259

-----
age                                     age in years
-----
      type:  numeric (byte)

      range:  [20,74]                    units:  1
unique values: 55                        missing .:  0/10,351

      mean:   47.5797
      std. dev: 17.2148

percentiles:      10%      25%      50%      75%      90%
                  24       31       49       63       69

-----
female                                1=female, 0=male
-----
      type:  numeric (byte)

      range:  [0,1]                      units:  1
unique values: 2                          missing .:  0/10,351

      tabulation:  Freq.  Value
                  4,915  0
                  5,436  1
```

- Now we can fit the model

```
. regress bmi age female
```

```
-----+-----
Source |      SS      df      MS      Number of obs   =    10,351
-----+-----
Model | 7330.98402      2  3665.49201  F(2, 10348)      =    156.29
-----+-----
Prob > F                =    0.0000
```

Residual		242693.178	10,348	23.4531483	R-squared	=	0.0293
-----+					Adj R-squared	=	0.0291
Total		250024.162	10,350	24.1569239	Root MSE	=	4.8428

bmi		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
-----+							
age		.0488667	.0027653	17.67	0.000	.0434462	.0542872
female		.0380616	.0953249	0.40	0.690	-.1487936	.2249168
_cons		23.19255	.1482223	156.47	0.000	22.90201	23.48309

2 Estimation

2.1 Including Categorical Variables

Working with Categorical Variables

- We would now like to include region in the model, let's take a look at this variable

```
. codebook region
```

```
-----
region                                     1=NE, 2=MW, 3=S, 4=W
-----
```

```

      type: numeric (byte)
      label: region

      range: [1,4]                units: 1
unique values: 4                  missing .: 0/10,351

      tabulation: Freq.  Numeric  Label
                  2,096      1  NE
                  2,774      2  MW
                  2,853      3   S
                  2,628      4   W

```

◇ It cannot simply be added to the list of covariates because it has 4 categories

- To include a categorical variable, put an *i.* in front of its name—this declares the variable to be a categorical variable, or in Stataese, a *factor variable*
- For example, to add region to our model we use

```
. regress bmi age i.female i.region
```

Source		SS	df	MS	Number of obs	=	10,351
-----+					F(5, 10345)	=	63.02
Model		7390.19781	5	1478.03956	Prob > F	=	0.0000
Residual		242633.964	10,345	23.4542256	R-squared	=	0.0296
-----+					Adj R-squared	=	0.0291
Total		250024.162	10,350	24.1569239	Root MSE	=	4.843

bmi		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
-----+							
age		.0488851	.0027674	17.66	0.000	.0434605	.0543097

female							
0		0	(base)				
1		.0372717	.0953357	0.39	0.696	-.1496047	.2241481
region							
NE		0	(base)				
MW		.0064779	.1402121	0.05	0.963	-.268365	.2813207
S		.0387957	.1393383	0.28	0.781	-.2343342	.3119256
W		-.1537648	.1418286	-1.08	0.278	-.4317762	.1242466
_cons		23.2187	.1760452	131.89	0.000	22.87362	23.56378

Niceities

- Value labels associated with factor variables are displayed in the regression table
- We can tell Stata to show the base categories for our factor variables

```
. set showbaselevels on
```

Factor Notation as Operators

- The `i.` operator can be applied to many variables at once:

```
. regress bmi age i.(female region)
```

Source		SS	df	MS	Number of obs	=	10,351
					F(5, 10345)	=	63.02
Model		7390.19781	5	1478.03956	Prob > F	=	0.0000
Residual		242633.964	10,345	23.4542256	R-squared	=	0.0296
					Adj R-squared	=	0.0291
Total		250024.162	10,350	24.1569239	Root MSE	=	4.843

bmi		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
age		.0488851	.0027674	17.66	0.000	.0434605 .0543097
female						
0		0	(base)			
1		.0372717	.0953357	0.39	0.696	-.1496047 .2241481
region						
NE		0	(base)			
MW		.0064779	.1402121	0.05	0.963	-.268365 .2813207
S		.0387957	.1393383	0.28	0.781	-.2343342 .3119256
W		-.1537648	.1418286	-1.08	0.278	-.4317762 .1242466
_cons		23.2187	.1760452	131.89	0.000	22.87362 23.56378

- In other words, it understands the distributive property
 - ◇ This is useful when using variable ranges, for example
 - For the curious, factor variable notation works with wildcards
 - ◇ If there were many variables starting with `u`, then `i.u*` would include them all as factor variables
-

Using Different Base Categories

- By default, the smallest-valued category is the base category
 - This can be overridden within commands
 - ◊ `b#`. specifies the value `#` as the base
 - ◊ `b(##)`. specifies the `#`'th largest value as the base
 - ◊ `b(first)`. specifies the smallest value as the base
 - ◊ `b(last)`. specifies the largest value as the base
 - ◊ `b(freq)`. specifies the most prevalent value as the base
 - ◊ `bn`. specifies there should be no base
-

Playing with the Base

- We can use `region=3` as the base class on the fly:

```
. regress bmi age i.female b3.region
```
 - We can use the most prevalent category as the base

```
. regress bmi age i.female b(freq).region
```
 - Factor variables can be distributed across many variables

```
. regress bmi age b(freq).(female region)
```
 - The base category can be omitted (with some care here)

```
. regress bmi age i.female bn.region, noconstant
```
 - We can also include a term for `region=4` only

```
. regress bmi age i.female 4.region
```
-

2.2 Including Interactions

Specifying Interactions

- Factor variables are also used for specifying interactions
 - ◊ This is where they really shine
 - To include both main effects and interaction terms in a model, put `##` between the variables
 - To include only the interaction terms, put `#` between the terms
-

Categorical by Categorical Interactions

- For example, to fit a model that includes main effects for age, female, and region, as well as the interaction of female, and region

```
. regress bmi age female##region
```

Source	SS	df	MS	Number of obs	=	10,351
Model	7559.19099	8	944.898874	F(8, 10342)	=	40.30
Residual	242464.971	10,342	23.4446888	Prob > F	=	0.0000
				R-squared	=	0.0302
				Adj R-squared	=	0.0295
Total	250024.162	10,350	24.1569239	Root MSE	=	4.842

bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
age	.0488087	.0027671	17.64	0.000	.0433846	.0542328
female						
0	0	(base)				
1	-.2939562	.2116093	-1.39	0.165	-.7087514	.1208389
region						
NE	0	(base)				
MW	-.1420836	.2023593	-0.70	0.483	-.538747	.2545798
S	-.3347762	.2015721	-1.66	0.097	-.7298965	.0603441
W	-.2694841	.204234	-1.32	0.187	-.6698222	.1308541
female#region						
1#MW	.2897474	.280525	1.03	0.302	-.2601358	.8396306
1#S	.7124639	.2789251	2.55	0.011	.1657169	1.259211
1#W	.2266557	.2837887	0.80	0.424	-.3296251	.7829365
_cons	23.39271	.2013939	116.15	0.000	22.99793	23.78748

- Variables involved in interactions are assumed to be categorical, so no `i.` is needed
- To see all the omitted terms we can add the `allbaselevels` option

```
. regress bmi age female##region, allbaselevels
```

Source	SS	df	MS	Number of obs	=	10,351
Model	7559.19099	8	944.898874	F(8, 10342)	=	40.30
Residual	242464.971	10,342	23.4446888	Prob > F	=	0.0000
				R-squared	=	0.0302
				Adj R-squared	=	0.0295
Total	250024.162	10,350	24.1569239	Root MSE	=	4.842

bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
age	.0488087	.0027671	17.64	0.000	.0433846	.0542328
female						
0	0	(base)				
1	-.2939562	.2116093	-1.39	0.165	-.7087514	.1208389
region						
NE	0	(base)				
MW	-.1420836	.2023593	-0.70	0.483	-.538747	.2545798

S		-.3347762	.2015721	-1.66	0.097	-.7298965	.0603441
W		-.2694841	.204234	-1.32	0.187	-.6698222	.1308541
female#region							
0#NE		0	(base)				
0#MW		0	(base)				
0#S		0	(base)				
0#W		0	(base)				
1#NE		0	(base)				
1#MW		.2897474	.280525	1.03	0.302	-.2601358	.8396306
1#S		.7124639	.2789251	2.55	0.011	.1657169	1.259211
1#W		.2266557	.2837887	0.80	0.424	-.3296251	.7829365
_cons		23.39271	.2013939	116.15	0.000	22.99793	23.78748

Categorical by Continuous Interactions

- To include continuous variables in interactions use `c.` to specify that a variable is continuous
 - ◊ Otherwise it will be assumed to be categorical
- Here is our model with an interaction between age and region

```
. regress bmi c.age##region i.female
```

Source		SS	df	MS	Number of obs	=	10,351
Model		7568.54189	8	946.067737	F(8, 10342)	=	40.35
Residual		242455.62	10,342	23.4437846	Prob > F	=	0.0000
Total		250024.162	10,350	24.1569239	R-squared	=	0.0303
					Adj R-squared	=	0.0295
					Root MSE	=	4.8419

bmi		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
age		.0607829	.0062164	9.78	0.000	.0485975 .0729683
region						
NE		0	(base)			
MW		.3951518	.4106204	0.96	0.336	-.4097436 1.200047
S		1.051668	.4181868	2.51	0.012	.2319407 1.871395
W		.5921285	.4181932	1.42	0.157	-.227611 1.411868
region#c.age						
MW		-.0080245	.0081638	-0.98	0.326	-.0240272 .0079782
S		-.0211109	.008219	-2.57	0.010	-.0372217 -.0050002
W		-.0155977	.0082261	-1.90	0.058	-.0317225 .000527
female						
0		0	(base)			
1		.038259	.0953259	0.40	0.688	-.1485982 .2251161
_cons		22.64929	.3193208	70.93	0.000	22.02336 23.27522

Continuous by Continuous Interactions

- Prefix both variables in the interaction with `c.` to fit models with continuous by continuous variable interactions
- For example, we can interact age with serum vitamin c levels (`vitaminc`)

```
. regress bmi c.age##c.vitaminc i.female i.region
```

Source	SS	df	MS	Number of obs	=	9,973
Model	10298.9223	7	1471.27461	F(7, 9965)	=	63.61
Residual	230479.207	9,965	23.1288718	Prob > F	=	0.0000
				R-squared	=	0.0428
				Adj R-squared	=	0.0421
Total	240778.13	9,972	24.1454201	Root MSE	=	4.8092

bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
age	.0220407	.0059366	3.71	0.000	.0104038	.0336777
vitaminc	-2.331426	.2717928	-8.58	0.000	-2.864194	-1.798657
c.age#c.vitaminc	.029107	.0050017	5.82	0.000	.0193026	.0389115
female						
0	0	(base)				
1	.1858965	.0982311	1.89	0.058	-.0066564	.3784494
region						
NE	0	(base)				
MW	-.0936871	.1412331	-0.66	0.507	-.3705326	.1831584
S	-.2137082	.1431247	-1.49	0.135	-.4942615	.0668451
W	-.1626738	.1430181	-1.14	0.255	-.4430182	.1176706
_cons	25.45695	.3293507	77.29	0.000	24.81136	26.10255

- To include polynomial terms, interact a variable with itself
- For example, a model that includes both age and age²

```
. regress bmi c.age##c.age i.female i.region
```

Source	SS	df	MS	Number of obs	=	10,351
Model	10269.3919	6	1711.56532	F(6, 10344)	=	73.84
Residual	239754.77	10,344	23.1781487	Prob > F	=	0.0000
				R-squared	=	0.0411
				Adj R-squared	=	0.0405
Total	250024.162	10,350	24.1569239	Root MSE	=	4.8144

bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
age	.2731368	.0203077	13.45	0.000	.2333297	.3129439
c.age#c.age	-.0024099	.0002162	-11.15	0.000	-.0028337	-.0019861
female						
0	0	(base)				
1	.0462855	.0947764	0.49	0.625	-.1394945	.2320656
region						
NE	0	(base)				

MW		.0322091	.1394036	0.23	0.817	-.2410489	.3054671
S		.0289346	.1385186	0.21	0.835	-.2425886	.3004579
W		-.1105093	.1410448	-0.78	0.433	-.3869844	.1659657
_cons		18.6987	.4416971	42.33	0.000	17.83289	19.56451

◇ The coefficient for age-squared is next to c.age#c.age

Higher Order Interactions

- Factor variable syntax can be used to specify higher order interactions
- If the interactions are specified using ## all lower order terms are included
- For example, here we fit a model for bmi using a model that includes the three-way interaction of continuous variables age and vitaminc and categorical variable female

```
. regress bmi c.age##c.vitaminc##female
```

Source	SS	df	MS	Number of obs	=	9,973
Model	12294.4386	7	1756.34837	F(7, 9965)	=	76.60
Residual	228483.691	9,965	22.9286193	Prob > F	=	0.0000
				R-squared	=	0.0511
				Adj R-squared	=	0.0504
Total	240778.13	9,972	24.1454201	Root MSE	=	4.7884

	bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
age		-.0038595	.0084263	-0.46	0.647	-.0203767 .0126578
vitaminc		-2.008713	.4231851	-4.75	0.000	-2.838241 -1.179185
c.age#c.vitaminc		.0313728	.0078481	4.00	0.000	.0159889 .0467566
female						
0		0	(base)			
1		-2.098183	.6208318	-3.38	0.001	-3.315138 -.8812268
female#c.age						
1		.0646392	.0119517	5.41	0.000	.0412115 .0880668
female#c.vitaminc						
1		.0314475	.5539279	0.06	0.955	-1.054363 1.117258
female#c.age#c.vitaminc						
1		-.0166002	.0102645	-1.62	0.106	-.0367206 .0035203
_cons		26.16464	.4416624	59.24	0.000	25.29889 27.03039

Some Factor Variable Notes

- If you plan to look at marginal effects of any kind, it is best to
 - Explicitly mark all categorical variables with i.
 - Specify all interactions using # or ##

- ◇ Specify powers of a variable as interactions of the variable with itself
 - There can be up to 8 categorical and 8 continuous interactions in one expression
 - ◇ Have fun with the interpretation
-

3 Postestimation

3.1 About Postestimation

Introduction to Postestimation

- In Stata jargon, postestimation commands are commands that can be run after a model is fit, for example
 - ◇ Predictions
 - ◇ Additional hypothesis tests
 - ◇ Checks of assumptions
 - We'll explore postestimation tools that can be used to help interpret the results of models that include interactions
 - The usefulness of specific tools will depend on the types of hypotheses you wish to examine
-

3.2 Investigating Categorical by Categorical Interactions

Estimating a Model

- Lets begin by running a model with main effects for age, female and region, and the interaction of female and region

```
. regress bmi age female##region
```

Source	SS	df	MS	Number of obs	=	10,351
Model	7559.19099	8	944.898874	F(8, 10342)	=	40.30
Residual	242464.971	10,342	23.4446888	Prob > F	=	0.0000
				R-squared	=	0.0302
				Adj R-squared	=	0.0295
Total	250024.162	10,350	24.1569239	Root MSE	=	4.842

bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
age	.0488087	.0027671	17.64	0.000	.0433846 .0542328
female					
0	0	(base)			
1	-.2939562	.2116093	-1.39	0.165	-.7087514 .1208389
region					
NE	0	(base)			
MW	-.1420836	.2023593	-0.70	0.483	-.538747 .2545798
S	-.3347762	.2015721	-1.66	0.097	-.7298965 .0603441
W	-.2694841	.204234	-1.32	0.187	-.6698222 .1308541
female#region					
1#MW	.2897474	.280525	1.03	0.302	-.2601358 .8396306
1#S	.7124639	.2789251	2.55	0.011	.1657169 1.259211

1#W		.2266557	.2837887	0.80	0.424	-.3296251	.7829365
_cons		23.39271	.2013939	116.15	0.000	22.99793	23.78748

- How might we begin?
 - ◇ Perform joint tests of coefficients
 - ◇ Estimate and test hypotheses about group differences
-

Finding the Coefficient Names

- Some postestimation commands require that you know the names used to store the coefficients
- To see these names we can replay the model and showing the *coefficient legend*

`. regress, coeflegend`

Source		SS	df	MS	Number of obs	=	10,351
					F(8, 10342)	=	40.30
Model		7559.19099	8	944.898874	Prob > F	=	0.0000
Residual		242464.971	10,342	23.4446888	R-squared	=	0.0302
					Adj R-squared	=	0.0295
Total		250024.162	10,350	24.1569239	Root MSE	=	4.842

bmi		Coef.	Legend
age		.0488087	_b[age]
female			
0		0	_b[0b.female]
1		-.2939562	_b[1.female]
region			
NE		0	_b[1b.region]
MW		-.1420836	_b[2.region]
S		-.3347762	_b[3.region]
W		-.2694841	_b[4.region]
female#region			
1#MW		.2897474	_b[1.female#2.region]
1#S		.7124639	_b[1.female#3.region]
1#W		.2266557	_b[1.female#4.region]
_cons		23.39271	_b[_cons]

- From here, we can see the full specification of the factor levels:
 - _b[2.region] corresponds to region=2 which is "MW" or midwest
 - _b[3.region] corresponds to region=3 which is "S" or south
 - We can also see the terms for the interaction:
 - _b[1.female#2.region] corresponds to the term for the interaction of region=2 and female=1
 - _b[1.female#3.region] corresponds to the term for the interaction of region=3 and female=1
-

Joint Tests

- The test command performs a Wald test of the specified null hypothesis
 - ◊ The default test is that the listed terms are equal to 0
- test takes a list of terms, which may be variable names, but can also be terms associated with factor variables
- To perform a joint test of the null hypothesis that the coefficients for the levels of `region` are all equal to 0

```
. test 2.region 3.region 4.region
```

```
( 1) 2.region = 0
( 2) 3.region = 0
( 3) 4.region = 0
```

```
F( 3, 10342) = 1.07
Prob > F = 0.3600
```

- ◊ Since the model contains an interaction, this is a test of the effect of `region` when `female=0`

Testing Sets of Coefficients

- To test that all of the coefficients associated with the interaction of `female` and `region` we would need to give the full name of all the coefficients

```
. test 1.female#2.region 1.female#3.region 1.female#4.region
```

- testparm also performs Wald tests, but it accepts lists of variables, rather than coefficients in the model
- So we can perform joint tests with less typing, for example

```
. testparm i.region#i.female
```

```
( 1) 1.female#2.region = 0
( 2) 1.female#3.region = 0
( 3) 1.female#4.region = 0
```

```
F( 3, 10342) = 2.40
Prob > F = 0.0656
```

An Alternative Test

- Likelihood ratio tests provide an alternative method of testing sets of coefficients
- To test the coefficients associated with the interaction of `female` and `region` we need to store our model results. The name is arbitrary, we'll call them `m1`

```
. estimates store m1
```

- Now we can rerun our model without `region`

```
. regress bmi age i.female i.region
```

Source	SS	df	MS	Number of obs	=	10,351
				F(5, 10345)	=	63.02
Model	7390.19781	5	1478.03956	Prob > F	=	0.0000
Residual	242633.964	10,345	23.4542256	R-squared	=	0.0296
				Adj R-squared	=	0.0291
Total	250024.162	10,350	24.1569239	Root MSE	=	4.843

bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
age	.0488851	.0027674	17.66	0.000	.0434605	.0543097
female						
0	0 (base)					
1	.0372717	.0953357	0.39	0.696	-.1496047	.2241481
region						
NE	0 (base)					
MW	.0064779	.1402121	0.05	0.963	-.268365	.2813207
S	.0387957	.1393383	0.28	0.781	-.2343342	.3119256
W	-.1537648	.1418286	-1.08	0.278	-.4317762	.1242466
_cons	23.2187	.1760452	131.89	0.000	22.87362	23.56378

- If we were removing one of these variables entirely, we would want to add `if e(sample)` to make sure the same sample, what Stata calls the *estimation sample*, is used for both models

Likelihood Ratio Tests (Continued)

- Now we store the second set of estimates

```
. estimates store m2
```

- And use the `lrtest` command to perform the likelihood ratio test

```
. lrtest m1 m2
```

```

Likelihood-ratio test                               LR chi2(3) =      7.21
(Assumption: m2 nested in m1)                       Prob > chi2 =    0.0654

```

- We'll restore the results from `m1`

```
. estimates restore m1
```

```
(results m1 are active now)
```

- Now it's as if we just ran the model stored in `m1`

Tests of Differences

- `test` can also be used to test the equality of coefficients

```
. test 3.region#1.female = 4.region#1.female
```

```
( 1) 1.female#3.region - 1.female#4.region = 0
```

```

F( 1, 10342) =    3.43
Prob > F =    0.0640

```

- A likelihood ratio test can also be used; see `help constraint` for information on setting the necessary constraints
- The `lincom` command can be used to calculate linear combinations of coefficients, along with standard errors, hypothesis tests, and confidence intervals
- For example, to obtain the difference in coefficients

```
. lincom 3.region#1.female - 4.region#1.female
```

```
( 1) 1.female#3.region - 1.female#4.region = 0
```

	bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
(1)		.4858082	.2622654	1.85	0.064	-.0282827	.9998991

Contrasts

- The `contrast` command allows us to test a wide variety of comparisons across groups
- For example comparing regions separately for men and women

```
. contrast region@female, effects
```

Contrasts of marginal linear predictions

Margins : asbalanced

	df	F	P>F
region@female			
0	3	1.07	0.3600
1	3	2.17	0.0890
Joint	6	1.62	0.1364
Denominator	10342		

	Contrast	Std. Err.	t	P> t	[95% Conf. Interval]	
region@female						
(MW vs base) 0	-.1420836	.2023593	-0.70	0.483	-.538747	.2545798
(MW vs base) 1	.1476637	.1943419	0.76	0.447	-.2332839	.5286114
(S vs base) 0	-.3347762	.2015721	-1.66	0.097	-.7298965	.0603441
(S vs base) 1	.3776878	.1927872	1.96	0.050	-.0002125	.755588
(W vs base) 0	-.2694841	.204234	-1.32	0.187	-.6698222	.1308541
(W vs base) 1	-.0428284	.1970381	-0.22	0.828	-.4290612	.3434044

- ◊ The `@` symbol requests comparisons of the levels of `region` at each value of `female`
- ◊ The `effects` option requests that individual contrasts be displayed along with their standard errors, hypothesis tests, and confidence intervals

Adjusting for Multiple Comparisons

- Use of contrast can result in a large number of hypothesis tests
- The `mcompare()` option can be used to adjust p-values and confidence intervals for multiple comparisons within factor variable terms
- The available methods are
 - ◊ `noadjust`
 - ◊ `bonferroni`
 - ◊ `sidak`
 - ◊ `scheffe`
- To apply Bonferroni's adjustment to our previous contrast

```
. contrast region@female, effects mcompare(bonferroni)
```

Contrasts of marginal linear predictions

Margins : asbalanced

	df	F	P>F
region@female			
0	3	1.07	0.3600
1	3	2.17	0.0890
Joint	6	1.62	0.1364
Denominator	10342		

Note: Bonferroni-adjusted p-values are reported for tests on individual contrasts only.

	Number of Comparisons
region@female	6

	Contrast	Std. Err.	Bonferroni t	Bonferroni P> t	Bonferroni [95% Conf. Interval]
region@female					
(MW vs base) 0	-.1420836	.2023593	-0.70	1.000	-.6760623 .391895
(MW vs base) 1	.1476637	.1943419	0.76	1.000	-.3651588 .6604863
(S vs base) 0	-.3347762	.2015721	-1.66	0.581	-.8666776 .1971252
(S vs base) 1	.3776878	.1927872	1.96	0.301	-.1310325 .886408
(W vs base) 0	-.2694841	.204234	-1.32	1.000	-.8084096 .2694415
(W vs base) 1	-.0428284	.1970381	-0.22	1.000	-.5627657 .4771089

Average Predicted Values

- We might want to explore predictions based on our model and data
- Predictions for individual observations can be made using the `predict` command, see `help predict`
- To find out about our model more generally, we may be more interested in average predicted values
 - ◊ Also known as predictive margins or recycled predictions
- To obtain the average predicted value of `bmi`

```
. margins
```

```
Predictive margins                                Number of obs    =    10,351
Model VCE      : OLS
```

```
Expression    : Linear prediction, predict()
```

		Delta-method				[95% Conf. Interval]	
		Margin	Std. Err.	t	P> t		
<hr/>							
	_cons	25.5376	.0475917	536.60	0.000	25.44431	25.63089

Predictions at Specified Values of Factor Variables

- Stata calls the list of variables that follow the `margins` command the *marginslist*
 - ◊ To appear in the *marginslist* a variable must have been specified as factor variable in the model
- To obtain the average predicted value of `bmi` at different values of `region`

```
. margins region
```

```
Predictive margins                                Number of obs    =    10,351
Model VCE      : OLS
```

```
Expression    : Linear prediction, predict()
```

		Delta-method				[95% Conf. Interval]	
		Margin	Std. Err.	t	P> t		
<hr/>							
region							
	NE	25.56063	.1057882	241.62	0.000	25.35327	25.768
	MW	25.57071	.09198	278.00	0.000	25.39042	25.75101
	S	25.60002	.0906777	282.32	0.000	25.42227	25.77776
	W	25.41018	.0944557	269.02	0.000	25.22503	25.59533

- How were these values generated?
 1. Calculate the predicted value of `bmi` setting `region=1` and using each case's observed values of `female` and `age`
 2. Find the mean of the predicted values
 3. Repeat steps 1 and 2 for each value of `region`

Predicted Values with Multiple Factor Variables

- We can obtain margins for multiple variables

```
. margins region female
```

```
Predictive margins                                Number of obs    =    10,351
Model VCE      : OLS
```

```
Expression    : Linear prediction, predict()
```

		Delta-method					
		Margin	Std. Err.	t	P> t	[95% Conf. Interval]	
region							
NE		25.56063	.1057882	241.62	0.000	25.35327	25.768
MW		25.57071	.09198	278.00	0.000	25.39042	25.75101
S		25.60002	.0906777	282.32	0.000	25.42227	25.77776
W		25.41018	.0944557	269.02	0.000	25.22503	25.59533
female							
0		25.51624	.0690736	369.41	0.000	25.38084	25.65164
1		25.55385	.0656788	389.07	0.000	25.42511	25.68259

- Or we can obtain predicted values of bmi at each combination of region and female

```
. margins region#female
```

```
Predictive margins                                Number of obs    =    10,351
Model VCE      : OLS
```

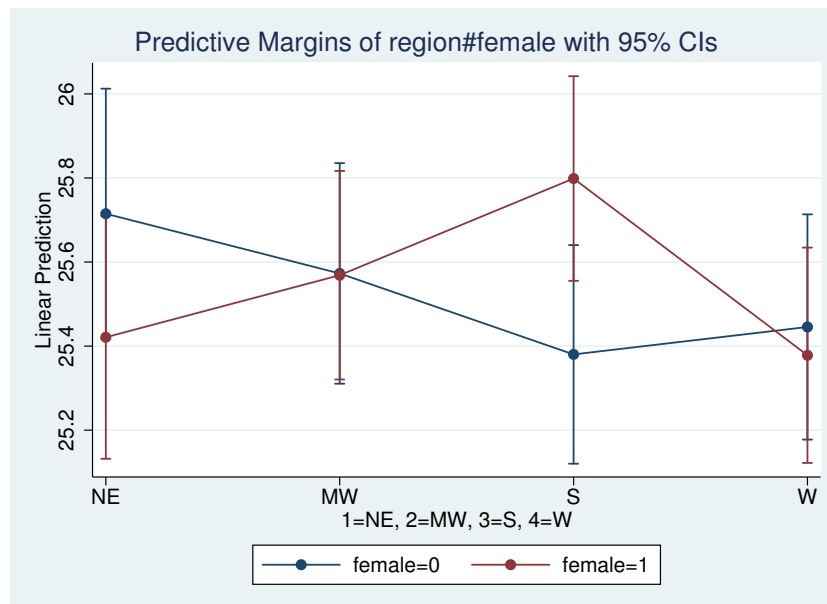
```
Expression    : Linear prediction, predict()
```

		Delta-method					
		Margin	Std. Err.	t	P> t	[95% Conf. Interval]	
region#female							
NE#0		25.71501	.1517587	169.45	0.000	25.41753	26.01248
NE#1		25.42105	.1474742	172.38	0.000	25.13197	25.71013
MW#0		25.57292	.1338383	191.07	0.000	25.31058	25.83527
MW#1		25.56872	.1265618	202.03	0.000	25.32063	25.8168
S#0		25.38023	.1326702	191.30	0.000	25.12017	25.64029
S#1		25.79874	.1241829	207.75	0.000	25.55532	26.04216
W#0		25.44552	.1366851	186.16	0.000	25.17759	25.71345
W#1		25.37822	.1306734	194.21	0.000	25.12208	25.63437

- We might prefer to graph these results, we can do so using the marginsplot command

Graphing Predicted Values

```
. marginsplot
```



- ◇ If our model did not include a region by female interaction, the lines would be parallel

Predicted Values for Specific Groups

- When we specify the variables in the *marginslist* Stata calculates predicted values treating each case as though it belonged to each group
- The `over()` option allows us to obtain predictions separately for each group, for example

```
. margins, over(female)
```

```
Predictive margins                                Number of obs    =    10,351
Model VCE      : OLS
```

```
Expression    : Linear prediction, predict()
over          : female
```

		Delta-method				
		Margin	Std. Err.	t	P> t	[95% Conf. Interval]
female						
0		25.50999	.0690654	369.36	0.000	25.37461 25.64538
1		25.56256	.0656723	389.24	0.000	25.43383 25.69129

- This time the table shows
 - ◇ The average predicted value of `bmi` for cases where `female=0` using each case's observed values of `age` and `region`
 - ◇ The average predicted value of `bmi` for cases where `female=1` using each case's observed values of `age` and `region`
- This can be useful when we want to compare groups

3.3 Investigating Categorical by Continuous Interactions

A Categorical by Continuous Interaction

- For this set of examples, we'll fit a model that includes an interaction between the continuous variable age and the categorical variable region

```
. regress bmi c.age##region i.female
```

Source	SS	df	MS	Number of obs	=	10,351
Model	7568.54189	8	946.067737	F(8, 10342)	=	40.35
Residual	242455.62	10,342	23.4437846	Prob > F	=	0.0000
				R-squared	=	0.0303
				Adj R-squared	=	0.0295
Total	250024.162	10,350	24.1569239	Root MSE	=	4.8419

bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
age	.0607829	.0062164	9.78	0.000	.0485975	.0729683
region						
NE	0 (base)					
MW	.3951518	.4106204	0.96	0.336	-.4097436	1.200047
S	1.051668	.4181868	2.51	0.012	.2319407	1.871395
W	.5921285	.4181932	1.42	0.157	-.227611	1.411868
region#c.age						
MW	-.0080245	.0081638	-0.98	0.326	-.0240272	.0079782
S	-.0211109	.008219	-2.57	0.010	-.0372217	-.0050002
W	-.0155977	.0082261	-1.90	0.058	-.0317225	.000527
female						
0	0 (base)					
1	.038259	.0953259	0.40	0.688	-.1485982	.2251161
_cons	22.64929	.3193208	70.93	0.000	22.02336	23.27522

- Let's take a look at how the coefficients are stored

```
. regress, coeflegend
```

Source	SS	df	MS	Number of obs	=	10,351
Model	7568.54189	8	946.067737	F(8, 10342)	=	40.35
Residual	242455.62	10,342	23.4437846	Prob > F	=	0.0000
				R-squared	=	0.0303
				Adj R-squared	=	0.0295
Total	250024.162	10,350	24.1569239	Root MSE	=	4.8419

bmi	Coef.	Legend
age	.0607829	_b[age]
region		
NE	0	_b[1b.region]
MW	.3951518	_b[2.region]
S	1.051668	_b[3.region]
W	.5921285	_b[4.region]
region#c.age		

```

      MW | -.0080245 _b[2.region#c.age]
      S | -.0211109 _b[3.region#c.age]
      W | -.0155977 _b[4.region#c.age]
      |
female |
      0 |          0 _b[0b.female]
      1 | .038259 _b[1.female]
      |
    _cons | 22.64929 _b[_cons]
-----

```

test and testparm

- As before, we can test the null hypothesis that all of the coefficients associated with the interaction of age and region are equal to 0 using testparm

```

. testparm c.age#i.region

( 1) 2.region#c.age = 0
( 2) 3.region#c.age = 0
( 3) 4.region#c.age = 0

F( 3, 10342) = 2.54
Prob > F = 0.0549

```

- We could also use lrtest
- We can test specific hypotheses about the slopes
- For example we might want to test whether the slope of age is significantly different in the south (region=3) versus the west (region=4)

```

. test 3.region#c.age = 4.region#c.age

( 1) 3.region#c.age - 4.region#c.age = 0

F( 1, 10342) = 0.52
Prob > F = 0.4689

```

Estimated Slopes

- We can use lincom to estimate the slope of age for the south (region=3)

```

. lincom c.age + 3.region#c.age

( 1) age + 3.region#c.age = 0

```

	bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
(1)		.0396719	.0053765	7.38	0.000	.0291329 .0502109

- We can also use margins with the dydx() option to calculate the slope of age for each region

```

. margins region, dydx(age)

```

```

Average marginal effects          Number of obs    =    10,351
Model VCE      : OLS

Expression   : Linear prediction, predict()
dy/dx w.r.t. : age

```

		Delta-method				
		dy/dx	Std. Err.	t	P> t	[95% Conf. Interval]
age						
region						
	NE	.0607829	.0062164	9.78	0.000	.0485975 .0729683
	MW	.0527584	.0052919	9.97	0.000	.0423853 .0631315
	S	.0396719	.0053765	7.38	0.000	.0291329 .0502109
	W	.0451852	.0053875	8.39	0.000	.0346246 .0557457

- The `dydx()` option calculates derivative of the predicted values with respect to the specified variable, also known as the marginal effect

Predictions at Specified Values

- To obtain margins at set values of continuous variables use the `at()` option
- For example, the predicted value of `bmi` at each level of `region` setting `age=20`

```
. margins region, at(age=20) vsquish
```

```

Predictive margins          Number of obs    =    10,351
Model VCE      : OLS

Expression   : Linear prediction, predict()
at           : age                =          20

```

		Delta-method				
		Margin	Std. Err.	t	P> t	[95% Conf. Interval]
region						
	NE	23.88504	.2026955	117.84	0.000	23.48772 24.28236
	MW	24.1197	.1678019	143.74	0.000	23.79078 24.44862
	S	24.51449	.1766004	138.81	0.000	24.16832 24.86066
	W	24.16521	.1772397	136.34	0.000	23.81779 24.51264

◊ The `vsquish` option reduces the vertical space in the output

- The `at()` option accepts *numlists* so we aren't restricted to a single value of `age`

```
. margins region, at(age=(20(25)70)) vsquish
```

```

Predictive margins          Number of obs    =    10,351
Model VCE      : OLS

Expression   : Linear prediction, predict()
1._at       : age                =          20
2._at       : age                =          45
3._at       : age                =          70

```

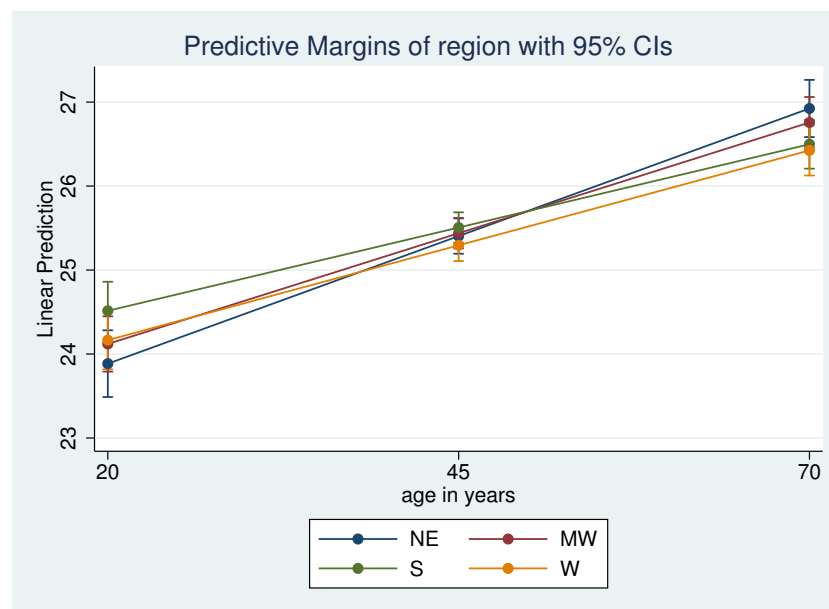
		Delta-method				[95% Conf. Interval]	
		Margin	Std. Err.	t	P> t		
<hr/>							
_at#region							
1#NE		23.88504	.2026955	117.84	0.000	23.48772	24.28236
1#MW		24.1197	.1678019	143.74	0.000	23.79078	24.44862
1#S		24.51449	.1766004	138.81	0.000	24.16832	24.86066
1#W		24.16521	.1772397	136.34	0.000	23.81779	24.51264
2#NE		25.40461	.1072029	236.98	0.000	25.19447	25.61475
2#MW		25.43866	.0922856	275.65	0.000	25.25776	25.61956
2#S		25.50629	.0922593	276.46	0.000	25.32544	25.68713
2#W		25.29484	.0956797	264.37	0.000	25.10729	25.48239
3#NE		26.92418	.1737943	154.92	0.000	26.58351	27.26485
3#MW		26.75762	.1545335	173.15	0.000	26.4547	27.06054
3#S		26.49809	.148221	178.77	0.000	26.20754	26.78863
3#W		26.42447	.1522388	173.57	0.000	26.12605	26.72289

◇ The observed values of age are from 20 to 74

Graphing Predicted Values

- And we can plot the results

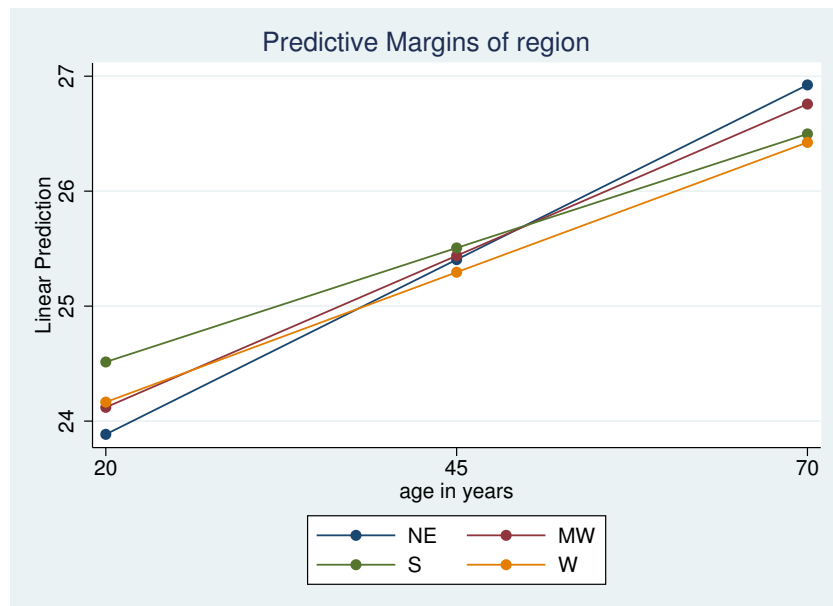
```
. marginsplot
```



Suppressing Confidence Intervals

- The confidence intervals can make the graph appear messy; we can suppress them

```
. marginsplot, noci
```



◇ This is dangerous because it makes the predictions look more precise than they are

Testing for Differences

- We might want to perform tests of differences at different levels of the continuous variable
- To obtain tests of differences between levels of region at each level of age

```
. margins region, at(age=(20(10)70)) vsquish contrast
```

```
Contrasts of predictive margins
Model VCE      : OLS
```

```
Expression      : Linear prediction, predict()
1._at           : age                =      20
2._at           : age                =      30
3._at           : age                =      40
4._at           : age                =      50
5._at           : age                =      60
6._at           : age                =      70
```

		df	F	P>F

region@_at				
1		3	1.94	0.1200
2		3	1.59	0.1884
3		3	1.06	0.3642
4		3	0.93	0.4251
5		3	1.56	0.1974
6		3	2.05	0.1041
Joint		6	1.69	0.1193

Denominator		10342		

Predicted Values Over Groups

- As with *marginslist*, when we specify `at()` Stata calculates predicted values treating each case as though they belong to each group or combination of values
- As before, we can use the `over()` option after models with categorical by continuous interactions
- For example, to obtain predicted values for each region using the observed values of female and age in that region

```
. margins, over(region)
```

```
Predictive margins                                Number of obs      =      10,351
Model VCE      : OLS
```

```
Expression    : Linear prediction, predict()
over          : region
```

		Delta-method				
		Margin	Std. Err.	t	P> t	[95% Conf. Interval]
region						
NE		25.57535	.1057592	241.83	0.000	25.36804 25.78266
MW		25.51936	.0919307	277.59	0.000	25.33916 25.69956
S		25.63317	.090649	282.77	0.000	25.45548 25.81086
W		25.42299	.0944498	269.17	0.000	25.23785 25.60813

3.4 Investigating Continuous by Continuous Interactions

A Continuous by Continuous Interaction

- For this example we'll use a similar model for *bmi* but we'll add a main effect of serum vitamin c (*vitaminc*), and an interaction between age and *vitaminc*
- Before we fit the model, let's take a closer look at *vitaminc*

```
. summ vitaminc, detail
```

serum vitamin C (mg/dL)				
Percentiles		Smallest		
1%	.2	.1		
5%	.3	.1		
10%	.3	.1	Obs	9,973
25%	.6	.1	Sum of Wgt.	9,973
50%	1		Mean	1.034814
		Largest	Std. Dev.	.5813791
75%	1.4	8.3		
90%	1.7	9.4	Variance	.3380017
95%	1.9	13.9	Skewness	4.539869
99%	2.4	18.1	Kurtosis	108.2617

◇ The distribution has a long tail, but most observations are between .2 and 2.

- Now let's fit the model

```
. regress bmi c.age##c.vitaminc i.female i.region
```


Source	SS	df	MS	Number of obs	=	9,973
				F(7, 9965)	=	63.61
Model	10298.9223	7	1471.27461	Prob > F	=	0.0000
Residual	230479.207	9,965	23.1288718	R-squared	=	0.0428
				Adj R-squared	=	0.0421
Total	240778.13	9,972	24.1454201	Root MSE	=	4.8092

bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
age	.0220407	.0059366	3.71	0.000	.0104038	.0336777
vitaminc	-2.331426	.2717928	-8.58	0.000	-2.864194	-1.798657
c.age#c.vitaminc	.029107	.0050017	5.82	0.000	.0193026	.0389115
female						
0	0	(base)				
1	.1858965	.0982311	1.89	0.058	-.0066564	.3784494
region						
NE	0	(base)				
MW	-.0936871	.1412331	-0.66	0.507	-.3705326	.1831584
S	-.2137082	.1431247	-1.49	0.135	-.4942615	.0668451
W	-.1626738	.1430181	-1.14	0.255	-.4430182	.1176706
_cons	25.45695	.3293507	77.29	0.000	24.81136	26.10255

- We can replay the model using `coeflegend`

```
. regress, coeflegend
```

Estimating Slopes

- We can use `lincom` to calculate the slope for `vitaminc` when `age=49` (it's median)

```
. lincom vitaminc + c.vitaminc#c.age*49
```

```
( 1) vitaminc + 49*c.age#c.vitaminc = 0
```

bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
(1)	-.9051806	.0870624	-10.40	0.000	-1.075841	-.7345206

- We could also calculate the slope of `age` when `vitaminc=1` (it's median)

```
. lincom age + c.vitaminc#c.age*1
```

```
( 1) age + c.age#c.vitaminc = 0
```

bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
(1)	.0511478	.0028212	18.13	0.000	.0456176	.0566779

- `margins` can produce estimates of the slopes for a range of values

Predicted Values

- Specifying multiple variables in the `at()` option results in predictions at each combination of values

```
. margins , at(age=(20(25)70) vitaminc=(.2(.6)2)) vsquish
```

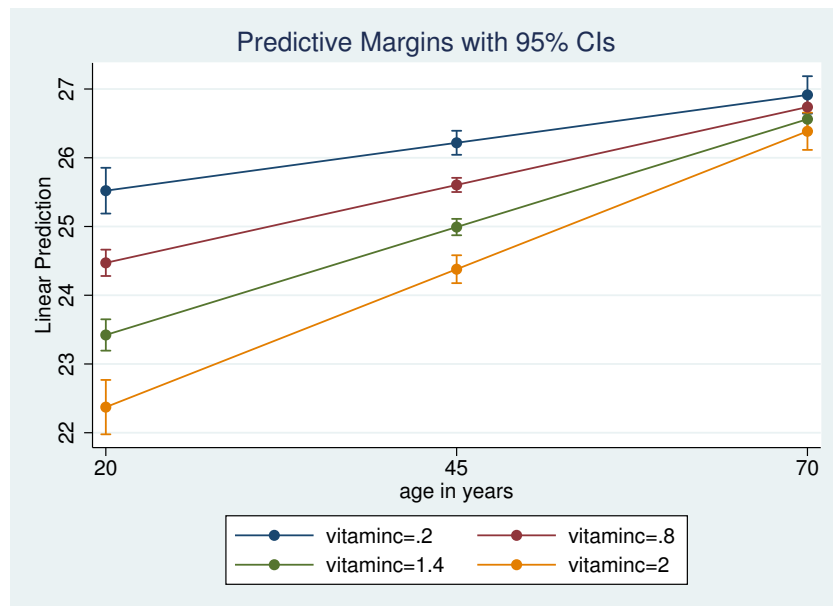
```
Predictive margins                                Number of obs    =      9,973
Model VCE      : OLS
```

```
Expression   : Linear prediction, predict()
1._at        : age           =      20
               vitaminc       =      .2
2._at        : age           =      20
               vitaminc       =      .8
3._at        : age           =      20
               vitaminc       =     1.4
4._at        : age           =      20
               vitaminc       =       2
5._at        : age           =     45
               vitaminc       =      .2
6._at        : age           =     45
               vitaminc       =      .8
7._at        : age           =     45
               vitaminc       =     1.4
8._at        : age           =     45
               vitaminc       =       2
9._at        : age           =     70
               vitaminc       =      .2
10._at       : age           =     70
               vitaminc       =      .8
11._at       : age           =     70
               vitaminc       =     1.4
12._at       : age           =     70
               vitaminc       =       2
```

		Delta-method					
		Margin	Std. Err.	t	P> t	[95% Conf. Interval]	

	_at						
1		25.52113	.1698638	150.24	0.000	25.18816	25.8541
2		24.47156	.097744	250.36	0.000	24.27996	24.66316
3		23.42199	.1162436	201.49	0.000	23.19413	23.64985
4		22.37242	.2018162	110.86	0.000	21.97682	22.76802
5		26.21768	.0891689	294.02	0.000	26.04289	26.39247
6		25.60472	.0525344	487.39	0.000	25.50174	25.70769
7		24.99175	.0606647	411.97	0.000	24.87284	25.11067
8		24.37879	.1034993	235.55	0.000	24.17591	24.58167
9		26.91423	.1388456	193.84	0.000	26.64207	27.1864
10		26.73788	.0879343	304.07	0.000	26.56551	26.91024
11		26.56152	.0875619	303.35	0.000	26.38988	26.73315
12		26.38516	.1381377	191.01	0.000	26.11438	26.65593

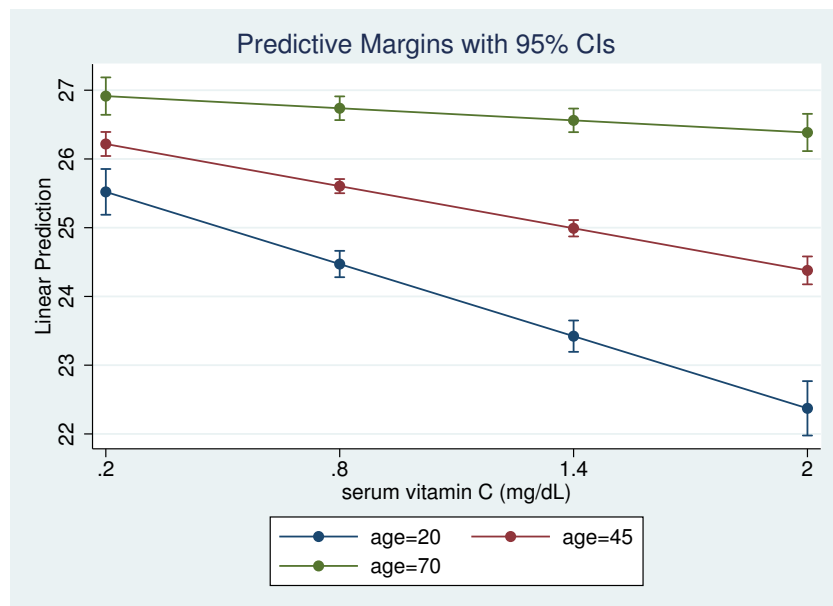
```
. marginsplot
```



Changing the X-axis Variable

- We can select which variable appears on the x-axis using the `xdimension()` option

```
. marginsplot, xdimension(vitaminc)
```



Models with Polynomial Terms

- We'll start by fitting a model that includes age and age^2

```
. regress bmi c.age##c.age i.female i.region
```

Source	SS	df	MS	Number of obs	=	10,351
				F(6, 10344)	=	73.84
Model	10269.3919	6	1711.56532	Prob > F	=	0.0000
Residual	239754.77	10,344	23.1781487	R-squared	=	0.0411
				Adj R-squared	=	0.0405
Total	250024.162	10,350	24.1569239	Root MSE	=	4.8144

bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
age	.2731368	.0203077	13.45	0.000	.2333297	.3129439
c.age#c.age	-.0024099	.0002162	-11.15	0.000	-.0028337	-.0019861
female						
0	0 (base)					
1	.0462855	.0947764	0.49	0.625	-.1394945	.2320656
region						
NE	0 (base)					
MW	.0322091	.1394036	0.23	0.817	-.2410489	.3054671
S	.0289346	.1385186	0.21	0.835	-.2425886	.3004579
W	-.1105093	.1410448	-0.78	0.433	-.3869844	.1659657
_cons	18.6987	.4416971	42.33	0.000	17.83289	19.56451

- Graphs can be particularly useful in understanding models with polynomial terms
- Here we predict values of bmi at different values of age

```
. margins, at(age=(20(10)70)) vsquish
```

```
Predictive margins                                Number of obs    =    10,351
Model VCE      : OLS
```

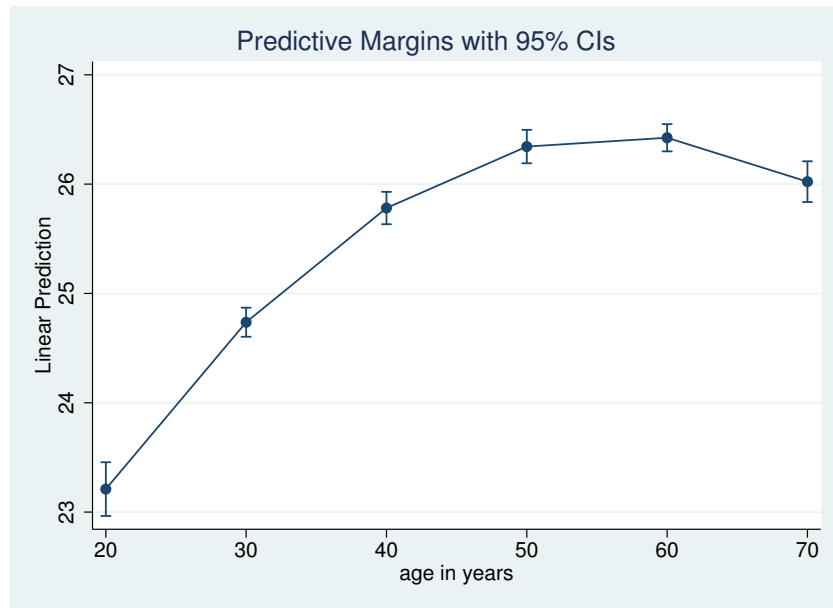
```
Expression   : Linear prediction, predict()
1._at        : age                =        20
2._at        : age                =        30
3._at        : age                =        40
4._at        : age                =        50
5._at        : age                =        60
6._at        : age                =        70
```

		Delta-method				[95% Conf. Interval]	
	Margin	Std. Err.	t	P> t			
_at							
1	23.21033	.1253478	185.17	0.000	22.96462	23.45604	
2	24.73675	.0678653	364.50	0.000	24.60372	24.86977	
3	25.78118	.0755647	341.18	0.000	25.63306	25.9293	
4	26.34363	.0780441	337.55	0.000	26.19065	26.49661	
5	26.4241	.0635204	415.99	0.000	26.29959	26.54861	
6	26.02259	.0951272	273.56	0.000	25.83612	26.20905	

Graphing Predicted Values

- And graph the predictions

```
. marginsplot
```



Slopes

- We can also obtain estimates of the slope of age across its range
- To do so we'll include age in both the dyed() and at() options

```
. margins, dydx(age) at(age=(20(10)70)) vsquish
```

Average marginal effects Number of obs = 10,351
Model VCE : OLS

Expression : Linear prediction, predict()

dy/dx w.r.t. : age

1._at	: age	=	20
2._at	: age	=	30
3._at	: age	=	40
4._at	: age	=	50
5._at	: age	=	60
6._at	: age	=	70

		Delta-method		t	P> t	[95% Conf. Interval]	
	dy/dx	Std. Err.					
age							
_at							
1	.1767405	.0117968	14.98	0.000	.1536164	.1998646	
2	.1285424	.0076583	16.78	0.000	.1135307	.143554	
3	.0803442	.0039415	20.38	0.000	.0726181	.0880703	
4	.032146	.0031343	10.26	0.000	.0260022	.0382899	
5	-.0160521	.0064432	-2.49	0.013	-.028682	-.0034222	
6	-.0642503	.010517	-6.11	0.000	-.0848657	-.0436349	

Adding a Cubic Term

- The same process can be used with higher order polynomials, here we add a cubic term for age

```
. regress bmi c.age##c.age##c.age i.female i.region
```

Source	SS	df	MS	Number of obs	=	10,351
				F(7, 10343)	=	64.27
Model	10422.3157	7	1488.90224	Prob > F	=	0.0000
Residual	239601.846	10,343	23.1656044	R-squared	=	0.0417
				Adj R-squared	=	0.0410
Total	250024.162	10,350	24.1569239	Root MSE	=	4.8131

	bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
	age	.5056311	.0927387	5.45	0.000	.3238453 .6874169
	c.age#c.age	-.0077683	.0020967	-3.70	0.000	-.0118782 -.0036583
	c.age#c.age#c.age	.0000383	.0000149	2.57	0.010	9.07e-06 .0000675
	female					
	0	0 (base)				
	1	.0449127	.0947522	0.47	0.636	-.1408201 .2306454
	region					
	NE	0 (base)				
	MW	.0274302	.1393783	0.20	0.844	-.2457782 .3006386
	S	.025305	.1384883	0.18	0.855	-.2461589 .2967689
	W	-.1172832	.1410312	-0.83	0.406	-.3937317 .1591653
	_cons	15.6426	1.268785	12.33	0.000	13.15554 18.12967

- As before we can predict slopes at specified values of age

```
. margins, dydx(age) at(age=(20(10)70)) vsquish
```

Average marginal effects	Number of obs	=	10,351
Model VCE : OLS			

Expression : Linear prediction, `predict()`

$$dy/dx \text{ w.r.t. : age}$$

1._at	: age	=	20
2._at	: age	=	30
3._at	: age	=	40
4._at	: age	=	50
5._at	: age	=	60
6._at	: age	=	70

		Delta-method					
		dy/dx	Std. Err.	t	P> t	[95% Conf. Interval]	
age	_at						
	1	.2408252	.0275901	8.73	0.000	.1867432	.2949071
	2	.1428662	.0094709	15.08	0.000	.1243014	.161431
	3	.06787	.0062529	10.85	0.000	.055613	.0801269
	4	.0158363	.0070791	2.24	0.025	.0019598	.0297128

5		-.0132346	.0065341	-2.03	0.043	-.0260428	-.0004265
6		-.0193429	.0203971	-0.95	0.343	-.0593252	.0206395

- Or predict bmi at different values of age

```
. margins, at(age=(20(9)74)) vsquish
```

```
Predictive margins                                Number of obs    =    10,351
Model VCE      : OLS
```

```
Expression   : Linear prediction, predict()
1._at        : age                =        20
2._at        : age                =        29
3._at        : age                =        38
4._at        : age                =        47
5._at        : age                =        56
6._at        : age                =        65
7._at        : age                =        74
```

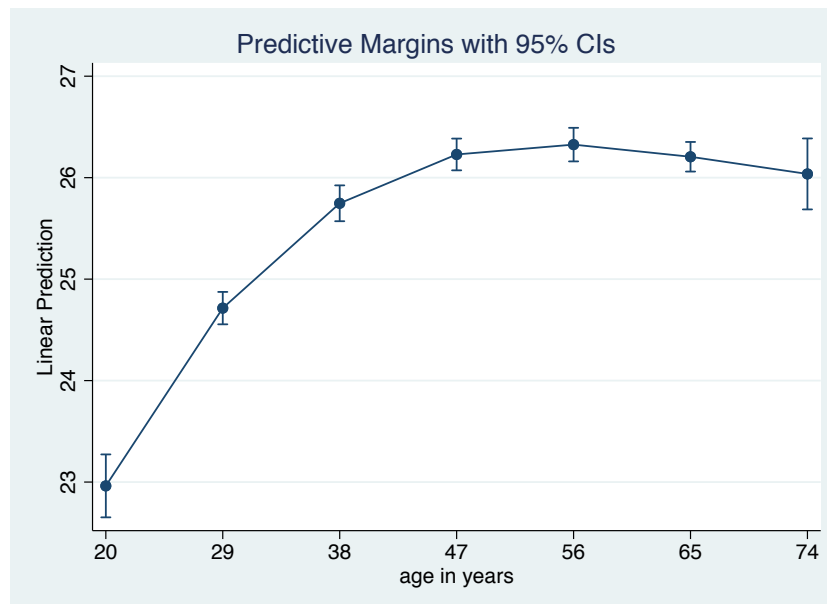
		Delta-method					
		Margin	Std. Err.	t	P> t	[95% Conf. Interval]	
	_at						
1		22.96222	.1582057	145.14	0.000	22.6521	23.27233
2		24.71431	.0814708	303.35	0.000	24.55461	24.87401
3		25.74733	.0900762	285.84	0.000	25.57077	25.9239
4		26.22869	.0798098	328.64	0.000	26.07225	26.38513
5		26.32577	.0843705	312.03	0.000	26.16039	26.49115
6		26.20598	.0744024	352.22	0.000	26.06013	26.35182
7		26.03671	.1783785	145.96	0.000	25.68705	26.38636

◇ Here, we get predictions across the full range of ages in the dataset (i.e. 20-74)

Graphing the Cubic Term

- And we can easily graph this as well

```
. marginsplot
```

4 Conclusion

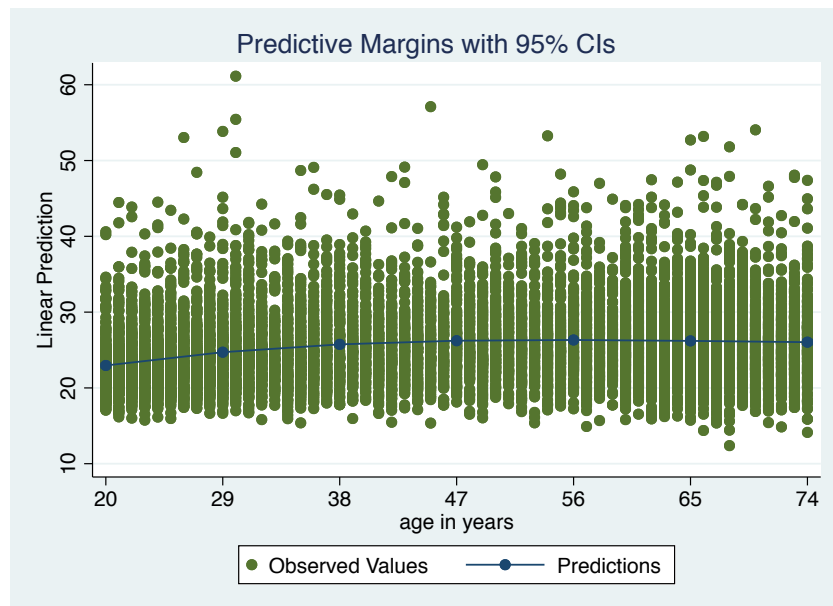
4.1 Graphing Extras

Adding Additional Plots

- We can add other types of twoway plots to the plots drawn by `marginsplots`
 - ◊ Continuing with our cubic example
- The `addplot` option allows us to add additional plots to our `marginsplots`
- We do want to be careful about the order in which graphs are drawn, we usually want the most dense graphs, for example individual data points, drawn first
 - ◊ Specifying `addplot(..., below)` draws the added plot below the `marginsplot`

Adding Observed Data

```
. marginsplot, addplot(scatter bmi age, below ///  
    legend(order(3 "Observed Values" 2 "Predictions"))) ///  
    xlabel(20(9)74))
```



- Note: The confidence intervals are in the plot, they're just small relative to the scale of the y-axis, so they're hard to see.

Changing the Plot Type

- We can change the plots drawn by `marginsplot` to another twoway plot type
 - ◊ See `help twoway` for a list
- The `recast()` option changes the plot for the predictions
 - ◊ `recastci()` changes how the CIs are plotted
- Let's run a simple model to demonstrate

```
. regress bmi i.region
. margins region
```

Source	SS	df	MS	Number of obs	=	10,351
Model	64.491028	3	21.4970093	F(3, 10347)	=	0.89
Residual	249959.671	10,347	24.1576951	Prob > F	=	0.4455
Total	250024.162	10,350	24.1569239	R-squared	=	0.0003
				Adj R-squared	=	-0.0000
				Root MSE	=	4.915

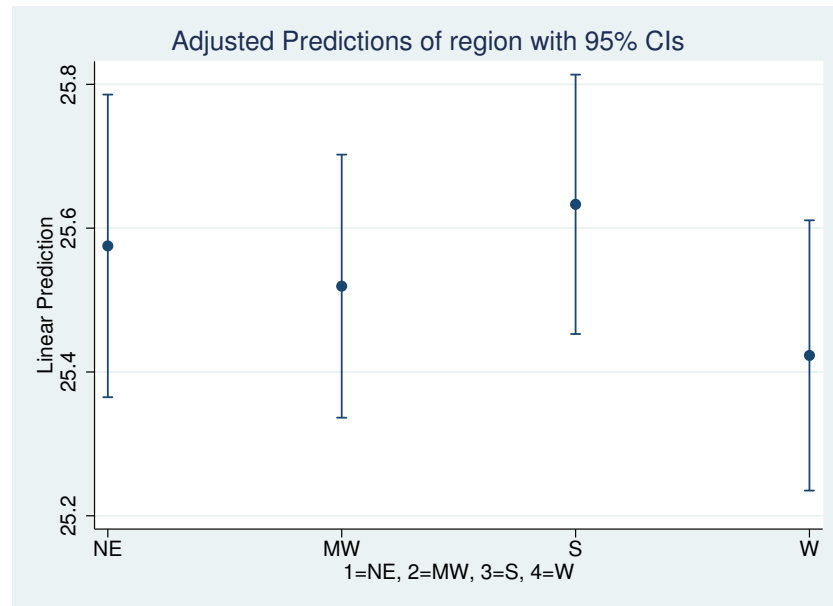
bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
region						
NE	0	(base)				
MW	-.055989	.1422471	-0.39	0.694	-.3348208	.2228428
S	.0578207	.1413969	0.41	0.683	-.2193446	.334986
W	-.1523645	.1439376	-1.06	0.290	-.4345101	.1297811
_cons	25.57535	.1073574	238.23	0.000	25.36491	25.78579

Adjusted predictions
Model VCE : OLS
Expression : Linear prediction, predict()
Number of obs = 10,351

		Delta-method					
		Margin	Std. Err.	t	P> t	[95% Conf. Interval]	
region							
NE		25.57535	.1073574	238.23	0.000	25.36491	25.78579
MW		25.51936	.09332	273.46	0.000	25.33644	25.70229
S		25.63317	.0920189	278.56	0.000	25.4528	25.81355
W		25.42299	.0958771	265.16	0.000	25.23505	25.61092

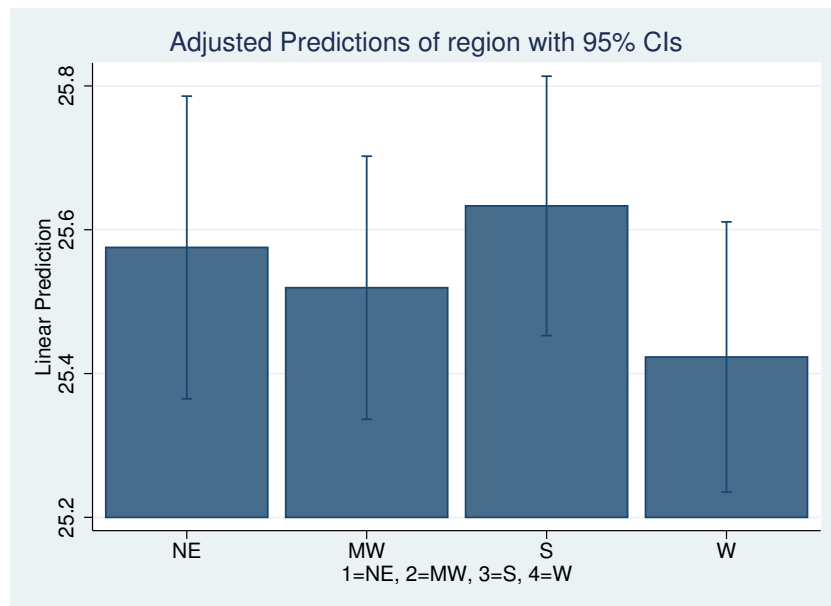
Estimates as a Scatterplot

```
. marginsplot, recast(scatter)
```



Estimates as a Bar plot

```
. marginsplot, recast(bar) plotopts(barwidth(.9))
```



- ◇ The `plotopts()` option allows you to specify options for the plots
- ◇ `barwidth()` specifies the width of the bars in units of the x variable

4.2 Conclusion

Conclusion

- We've seen how to fit models that include interactions
 - We've learned how to use Stata's postestimation tools to explore the resulting models
 - We've learned how to graph predictions and how to modify those graphs
-

Index