Control-function models in Stata

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Control Functions

Outline

cfregress y w (x = z1 z2),
$$vce(robust)$$

To discuss:

- The idea behind control functions
- The simple linear case: how it works
- Variations: theory, syntax, example output
- VCE/Standard errors
- Postestimation

Regression models often suffer from endogeneity.

- e.g. $y = \beta_0 + \beta_1 x + u$: we want β_1 , but when x moves in our data, so does u

Our workhorses for these models are instrumental-variables (IV) methods.

- Idea: find an *instrument z* that moves *x* but not *u*. Then "move" *z* enough to move *x* one unit, and see what happens to *y*.

Control function (CF) methods are a variation on plain IV.

- Idea: model the part of x that z cannot explain, call it a control function, v, then include an estimate \hat{v} in our regression.

Why use CF if we have IV/2SLS?

Short answer: more flexibility (at the cost of stronger assumptions)

Long answer:

- Built-in tests of endogeneity
- Easy estimation of some correlated random coefficient models
- Simplified handling of endogenous variables entering as interactions
- Exploit discreteness of binary endogenous variables
- And more ...
- See Wooldridge (JHR, 2015)

Control Functions

They've been there all along

Stata commands that already used CF methods:

- etregress, cfunction
- eteffects
- ivprobit, twostep
- ivtobit, twostep
- ivpoisson cfunction

The idea of cfregress and cfprobit: control function regression commands that let users manipulate the CF specification and exploit the distinctive features of control functions, while taking care of standard errors.

Plain linear IV setup with one endogenous regressor (exogenous regressors partialled out):

$$y = \beta x + u,$$

$$x = Z\pi + v,$$

$$E(Zu) = 0,$$

$$\pi \neq 0,$$

$$E(Zv) = 0.$$

Note that the endogeneity of x, $E(xu) \neq 0$, implies that u and v are correlated. Let $\rho = E(uv)/E(v^2)$, and let $\varepsilon = u - \rho v$.

Substituting $u = \rho v + \varepsilon$ into our main equation, we have

$$y = \beta x + \rho v + \varepsilon.$$

We have that $E(x\varepsilon) = E(v\varepsilon) = E(uv) - \rho E(v^2) = 0$, so x is uncorrelated with ε and we can estimate β if we observe v.

We do not observe ν , so in practice we use $\hat{\nu} = x - Z\hat{\pi}$. Because $\hat{\pi}$ is a consistent estimator of π , we still get a consistent estimate of β (note that $\nu - \hat{\nu} = Z(\hat{\pi} - \pi)$).

Control Functions

The linear case

$$y = \beta x + \rho \hat{\mathbf{v}} + \varepsilon.$$

Note: a test of $\rho = 0$ is a valid test for endogeneity.

Another existing use of control functions: estat endogenous after ivregress 2sls, vce(robust).

$$y = \beta x + u,$$

$$x = Z\pi + v.$$

In this linear model, $\hat{\beta}_{CF} = \hat{\beta}_{2SIS}$.

- Intuition: CF uses x along with its first-stage residuals $x Z\hat{\pi}$ while 2SLS uses fitted values $Z\hat{\pi}$, but both contain the same information about $Z\pi$.
- In other models, there is generally not an IV method using fitted values that is equivalent to CF.

The linear case

$$y = \beta x + \rho \hat{\mathbf{v}} + \varepsilon,$$

Note the above can be rewritten:

$$y = \beta(\hat{x} + \hat{v}) + \rho\hat{v} + \varepsilon,$$

$$y = \beta \hat{x} + (\beta + \rho)\hat{v} + \varepsilon,$$

And since \hat{x} and \hat{v} are orthogonal, we get the same estimate of β by running:

$$y = \beta \hat{x} + error.$$

Note: \hat{x} is not an estimator of x, which is known; it is an estimator of $Z\pi$.

The linear case

$$y = \beta x + \rho \hat{\mathbf{v}} + \varepsilon.$$

To estimate β and ρ in cfregress, we use syntax familiar from ivregress:

We can include exogenous variables:

Example output

```
. cfregress rent pcturban (hsngval = faminc i.region), vce(robust)
Control-function linear regression
                                                      Number of obs =
                                                      Wald chi2(2) =
                                                                       44.98
                                                     Prob > chi2
                                                                   = 0.0000
```

R-squared = 0.5989 Root MSE = 22.1656

Endogenous variable model: Linear: hsngval

rent	Coefficient	Robust std. err.	z	P> z	[95% conf	. interval]
rent						
hsngval	.0022398	.000672	3.33	0.001	.0009227	.0035569
pcturban	.081516	.4445938	0.18	0.855	789872	.9529039
_cons	120.7065	15.25546	7.91	0.000	90.80636	150.6067
e.rent cf(hsngval)	0015889	.000806	-1.97	0.049	0031687	-9.10e-06

Instruments for hsngval: faminc 2.region 3.region 4.region



In a conditional mean sense, CF methods can be thought of as saying

$$E(u|v,Z)=\rho v.$$

But suppose it does not hold.

For example, we could have $E(u|v,Z) = \rho_1 v + \rho_2 v w$, for an exogenous variable w. We may think of including an interaction term in our estimating equation:

$$y = \beta_0 + \beta_1 x + \gamma w + \rho_1 v + \rho_2 vw + \eta$$

$$y = \beta_0 + \beta_1 x + \gamma w + \rho_1 v + \rho_2 vw + \eta$$

With this setup, you can show we need $u - \rho_2 vw$ uncorrelated with Z. Under our original assumption E(Zu) = 0, this means we need $\rho_2 vw$ uncorrelated with Z. Because w is part of Z, we can get this by assuming E(v|Z) = 0.

Or, we can say we don't need E(Zu)=0, but rather $E(Z(u-\rho_2vw))=0$. In other words, we only need Z to be exogenous to whatever is left over after partialling out vw. This condition is implied by our CF assumption $E(u|Z,v)=\rho_1v+\rho_2vw$.

If we can impose independence then either way is fine, but here we can see that there exist DGPs where the interacted CF approach will give valid results and the regular CF approach won't!

Nonlinearity of the error in *v*

$$y = \beta_0 + \beta_1 x + \gamma w + \rho_1 \hat{v} + \rho_2 \hat{v} w + \eta$$

Command:

Example output

```
. cfregress rent pcturban (hsngval = faminc i.region, interact(pcturban)), vce(robust)
Control-function linear regression
                                                    Number of obs =
                                                                       50
                                                    Wald chi2(2) = 44.83
                                                    Prob > chi2 = 0.0000
                                                    R-squared
                                                                = 0.5574
                                                    Root MSE
                                                                 = 23.2829
```

Endogenous variable model: Linear: hsngval

	rent	Coefficient	Robust std. err.	z	P> z	[95% conf.	interval]
rent							
	hsngval	.0024082	.0006391	3.77	0.000	.0011556	.0036608
	pcturban	.1459889	.4308807	0.34	0.735	6985218	.9904997
	_cons	108.2288	17.36071	6.23	0.000	74.20243	142.2552
e.rent							
	cf(hsngval)	.0015522	.0019371	0.80	0.423	0022444	.0053488
cf(hsngval)#pcturban		0000419	.0000236	-1.78	0.075	0000881	4.26e-06

Instruments for hsngval: faminc 2.region 3.region 4.region



Suppose x_1 is endogenous, z_1 and z_2 are exogenous and

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_1 z_2 + u.$$

The IV way to approach this would be to treat x_1 and $x_1 z_2$ as two endogenous regressors that share two instruments z_1 and $z_1 z_2$.

A control function approach is to model a control function only for x_1 , with instrument z_1 , and estimate the regression

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_1 z_2 + \hat{v} \rho_1 + error,$$

or even

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_1 z_2 + \hat{v} \rho_1 + \hat{v} z_2 \rho_2 + error.$$

Commands:

Variables to appear only in the main equation

Note we use the option mainonly() to specify a variable that should be treated as exogenous, but should not appear in the first stage.

ivregress includes all exogenous variables in the first stage. cfregress does too (except for those specified in mainonly()), because there is seldom good reason for doing otherwise.

Correlated random coefficients

We may be interested in a model with correlated random coefficients:

$$y = \beta_0 + \beta_1(\omega)x + u$$
; with $\beta_1(\omega) = \beta_1 + \omega$.

where ω is a random variable with mean zero. We can write this as

$$y = \beta_0 + \beta_1 x + \omega x + u,$$

$$x = Z\pi + v.$$

We take the error to be $\omega x + u$ and project both ω and u onto v. We then estimate

$$y = \beta_0 + \beta_1 x + \rho_1 x \hat{\mathbf{v}} + \rho_2 \hat{\mathbf{v}} + \text{error}.$$

Correlated random coefficients

$$y = \beta_0 + \beta_1 x + \rho_1 x \hat{\mathbf{v}} + \rho_2 \hat{\mathbf{v}} + \text{error}.$$

Command:

cfregress
$$y (x = z, interact(x))$$

Conveniently, we can test the heterogeneity of $\beta_1(\omega)$ in x, as a test of the null $\rho_1 = 0$. See Wooldridge (JHR, 2015) for a discussion.

We may have binary x_1 and believe it is probit conditional on Z:

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + u,$$

$$x_1 = \mathbb{1}(Z\pi_1 + \pi_2 x_2 + v > 0).$$

The CF approach involves estimating

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \hat{r}\rho + \varepsilon,$$

where \hat{r} is the score from the first-stage probit. Under appropriate assumptions, this is valid. The two-step IV approach of plugging in fitted values is generally not.

Command:

Example output

. cfregress lndrug age lninc (ins = i.married i.work, probit), mainonly(i.chron) vce(robust)

Control-function linear regression

Number of obs = 6,000 Wald chi2(4) = 2833.77 Prob > chi2 = 0.0000 R-squared = 0.2393 Root MSE = 1.2203

Endogenous variable model: Probit: 1.ins

Robust Coefficient std. err. P>|z| [95% conf. interval] lndrug lndrug 1.ins -.8992025 .3399829 -2.640.008 -1.565557 -.2328483 1.chron .4675479 .0319717 14.62 0.000 .4048845 .5302113 age .1011597 .0027163 37.24 0.000 .0958359 .1064836 lninc .0505756 .0217621 2.32 0.020 .0079228 .0932285 1.827957 10.24 0.000 1.478126 2.177787 cons .1784883 e.lndrug cf(1.ins) .6157838 .1991464 3.09 0.002 .225464 1.006104

Instruments for 1.ins: 1.married 1.work



Multiple endogenous regressors

We can have multiple endogenous regressors and multiple control functions:

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + u$$

$$x_1 = \pi_{10} + \pi_{11} z_1 + \pi_{12} z_2 + v_1$$

$$x_2 = \pi_{20} + \pi_{21} z_1 + \pi_{22} z_2 + v_2$$

and estimate the equation

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \rho_1 \hat{v}_1 + \rho_2 \hat{v}_2 + \epsilon.$$

Command:

cfregress y (x1 x2 = z1 z2),
$$vce(robust)$$

Note: this is still equivalent to 2SLS despite multiple endogenous variables.

Multiple endogenous regressors

We typically assume:

$$E(u|v_1,v_2,z_1,z_2)=\rho_1v_1+\rho_2v_2.$$

However, this may feel like a strong assumption. We can consider adding an interaction term and estimating:

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \rho_1 \hat{v}_1 + \rho_2 \hat{v}_2 + \rho_3 \hat{v}_1 \hat{v}_2 + error.$$

Command:

cfregress y (x1 x2 = z1 z2),
$$vce(robust)$$
 cfinteract.

The cfinteract option works using "##" interaction logic, where all combinations of control functions are interacted and included. Note: no effect with only one endogenous variable.

Example output

. cfregress mpg (price foreign = weight length), cfinteract vce(robust)

Control-function linear regression

Number of obs = 74

Wald chi2(2) = 21.70

Prob > chi2 = 0.0000

Root MSE = 9.5680

Endogenous variable models: Linear: price foreign

mpg	Coefficient	Robust std. err.	z	P> z	[95% conf.	interval]
mpg						
price	.0005727	.0012578	0.46	0.649	0018926	.0030379
foreign	20.41465	7.523141	2.71	0.007	5.669567	35.15974
_cons	12.22066	9.600274	1.27	0.203	-6.595535	31.03685
e.mpg						
cf(price)	0005092	.0012874	-0.40	0.692	0030324	.002014
cf(foreign)	-21.18877	7.245094	-2.92	0.003	-35.38889	-6.988643
cf(foreign)#cf(price)	0010821	.0004913	-2.20	0.028	0020452	0001191

Instruments for price: weight length Instruments for foreign: weight length



In cfregress, we even allow users to specify endogenous regressors with different instrument sets:

$$y = \beta_1 x_1 + \beta_2 x_2 + u,$$

$$x_1 = \pi_{11} z_1 + \pi_{12} z_2 + v_1,$$

$$x_2 = \pi_{21} z_2 + v_2.$$

Command:

cfregress y (x1 = z1 z2) (x2 = z2),
$$vce(robust)$$
.

Example output

. cfregress mpg (price = weight length) (foreign = length), vce(robust)

Control-function linear regression Number of obs = 74Wald chi2(2) = 47.71

Wald chi2(2) = 47.71 Prob > chi2 = 0.0000 R-squared = 0.0035 Root MSE = 5.7361

Endogenous variable models: Linear: price foreign

mpg	Coefficient	Robust std. err.	z	P> z	[95% conf.	interval]
mpg price foreign _cons	0009205 13.07771 23.0844	.0005543 3.601915 4.048762	-1.66 3.63 5.70	0.097 0.000 0.000	002007 6.018082 15.14897	.000166 20.13733 31.01982
e.mpg cf(price) cf(foreign)	.0008831 -14.65215	.0005937 3.523867	1.49 -4.16	0.137 0.000	0002805 -21.5588	.0020467 -7.745497

Instruments for price: weight length Instrument for foreign: length



Multiple sets of instruments

We still need the same main control function assumption as in the previous example.

$$E(u|v_1, v_2, z_1, z_2) = \rho_1 v_1 + \rho_2 v_2,$$

but now we are implicitly allowing z_1 to be correlated with v_2 . But note this is an unusual combination of assumptions—usually it is safer to use both instruments for both endogenous variables.

Standard errors

Standard linear case when we have $y = \beta x + u$:

$$x = Z\pi + v,$$

 $y = \beta x + \rho \hat{v} + error.$

As will be familiar from 2SLS, the standard errors produced by running the two stages sequentially will be wrong.

One good option is to use option vce(bootstrap).

Ideally, however, we would like an analytic option.

Note that it is well known that GMM produces the 2SLS estimator of β , with appropriate standard errors, when used with the right weights.

$$\sum z_i'(y_i - \beta x_i) = 0,$$
with $W = \left(\sum z_i' z_i\right)^{-1}$.

This works because the GMM objective function then includes a projection matrix P_{7} with elements $z_i(\sum z_i'z_i)^{-1}z_k'$, which leads to the familiar estimator $\hat{\beta}_{2SIS} = (X'P_{7}X)^{-1}X'P_{7}V$

You can write a numerically equivalent exactly-identified GMM system without weights, by taking advantage of the fact that $P_Z P_Z = P_Z$. Specifically, the following sample moment condition gives the same estimates and variance estimator:

$$\sum \hat{x}_i(y_i - \beta x_i) = 0.$$

So, you can run

$$gmm (y - {b}*x), inst(xhat)$$

and get 2SLS estimates and standard errors.

Note this condition can be written as $\sum \hat{x}_i(y_i - \beta z_i \pi - \beta(x_i - z_i \pi)) = 0$. We can even write down the following and still get the same estimates and standard errors:

$$\sum \hat{x}_i(y_i - \beta z_i \pi) = 0,$$

$$\sum z'_i(x_i - z_i \pi) = 0.$$

(See Newey, 1984.) Finally, we can introduce the remaining component of the error, $(\rho + \beta)v$, and rearrange:

$$\sum \hat{x}_i(y_i - \beta x_i - \rho(x_i - z_i \pi)) = 0,$$

$$\sum z'_i(x_i - z_i \pi) = 0.$$

But we are short a constraint now that we have introduced ρ . We can use the fact that $\hat{x}_i = x_i - \hat{v}_i$ and set as our new conditions, which imply those previously,

$$\sum x_i(y_i - \beta x_i - \rho(x_i - z_i \pi)) = 0,$$

$$\sum \hat{v}_i(y_i - \beta x_i - \rho(x_i - z_i \pi)) = 0,$$

$$\sum z'_i(x_i - z_i \pi) = 0.$$

These conditions are intuitive, because x and \hat{v} are our regressors in our estimating equation.

$$\sum x_i(y_i - \beta x_i - \rho v(\pi; x_i, z_i)) = 0,$$

$$\sum \hat{v}_i(y_i - \beta x_i - \rho v(\pi; x_i, z_i)) = 0,$$

$$\sum z_i' v(\pi; x_i, z_i) = 0.$$

We compute GMM-style variance-covariance matrices using moment conditions based on the estimating equation and first stage.

Note the sample variance will thus depend on $G(\hat{\beta},\hat{\rho},\hat{\pi})$, the Jacobian with respect to the parameters, evaluated at the optimum. Since π appears in the error function of the main equation, we account for dependence between the two equations.

Using the GMM framework lets us easily allow for clustering and HAC VCEs.



Computation

Note: we do not run gmm under the hood. We get estimates using the regular two-step procedure. Then GMM standard errors are computed directly in Mata, making the procedure quite fast.

Postestimation

- predict: allowed with statistics xb, xbv, e, ve
 - xb returns linear prediction without the control function
 - xbv returns linear prediction with control function
 - e returns the residual not including the control function
 - ve returns the residual including the control function
- margins: allowed with xb and xbv
- estat endogenous: Translates readily from the corresponding postestimation command for ivregress.

Postestimation example

. cfregress rent pcturban (hsngval = faminc i.region, interact(pcturban)), vce(robust)

Control-function linear regression

Number of obs = 50
Wald chi2(2) = 44.83
Prob > chi2 = 0.0000
R-squared = 0.5574
Root MSE = 23.2829

Endogenous variable model: Linear: hsngval

	rent	Coefficient	Robust std. err.	z	P> z	[95% conf.	interval]
rent							
	hsngval	.0024082	.0006391	3.77	0.000	.0011556	.0036608
	pcturban	.1459889	.4308807	0.34	0.735	6985218	.9904997
	_cons	108.2288	17.36071	6.23	0.000	74.20243	142.2552
e.rent							
	cf(hsngval)	.0015522	.0019371	0.80	0.423	0022444	.0053488
cf(hsng	val)#pcturban	0000419	.0000236	-1.78	0.075	0000881	4.26e-06

Instruments for hsngval: faminc 2.region 3.region 4.region

. estat endogenous

Tests of endogeneity

HO: Variables are exogenous

- (1) [e.rent]cf(hsngval) = 0
- (2) [e.rent]cf(hsngval)#c.pcturban = 0

chi2(2) = 7.67 Prob > chi2 = 0.0216



- 1. When $\hat{\beta}_{CF} = \hat{\beta}_{IV}$, meaning you have no endogenous interactions or fancy first-stage modeling, use plain IV, unless you want the convenient endogeneity test.
- 2. When you have information about the form of the endogeneity, use CF.
- 3. When you have endogenous variables entering as interactions, use CF unless you think the IV assumptions are preferable.
- 4. When you want a nonlinear first stage, use CF.
- 5. When you have a model that IV commands won't let you run (different instrument sets, exogenous variables excluded from the first stage, etc.), feel free to use CF but make sure you can justify the appropriate assumptions.

Further reading

- Kim, K., and A. K. Petrin, 2011. A new control function approach for non-parametric regressions with endogenous variables. Working Paper 16679, National Bureau of Economic Research.
- Newey, W. K. 1984. A method of moments interpretation of sequential estimators. Fconomics Letters 14: 201-206.
- Wooldridge, J. M. 2015. Control function methods in applied econometrics. Journal of Human Resources 50: 420–445