ddml: Double/debiased machine learning in Stata

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Package website: https://statalasso.github.io/ Latest version available here

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Introduction

A rich and growing literature exploits machine learning to facilitate causal inference.

A central focus: *high-dimensional* controls and/or instruments, which can arise if

- ▶ we observe many controls/instruments
- ► controls/instruments enter through an unknown function

Belloni, Chernozhukov, and Hansen (2014) and Belloni et al. (2012) propose estimators *relying on the Lasso* that allow for high-dimensional controls/instruments.

 \Rightarrow Available via pdslasso in Stata (Ahrens, Hansen, and Schaffer, 2020)

Introduction

What if we don't want to use the lasso?

- The Lasso might not be the best-performing machine learner for a particular problem.
- ► The Lasso relies on the *approximate sparsity assumption*, which might not be appropriate in some settings.

Chernozhukov et al. (2018) propose *Double/Debiased Machine Learning* (DDML) which allow to exploit machine learners other than the Lasso.

Our contribution:

- ▶ We introduce ddml, which implements DDML for Stata.
- We provide simulation evidence on the finite sample performance of DDML.
- Our recommendation is to use DDML in combination with Stacking.

Motivating example. The partial linear model:



How do we account for confounding factors \mathbf{x}_i ? — The standard approach is to assume linearity $g(\mathbf{x}_i) = \mathbf{x}'_i \beta$ and consider alternative combinations of controls.

Problems:

- Non-linearity & unknown interaction effects
- ► High-dimensionality: we might have "many" controls
- ► We don't know which controls to include

Motivating example. The partial linear model:



Post-double selection (Belloni, Chernozhukov, and Hansen, 2014) and *post-regularization* (Chernozhukov, Hansen, and Spindler, 2015) provide data-driven solutions for this setting.

Both "double" approaches rely on the *sparsity assumption* and use two auxiliary lasso regressions: $y_i \rightsquigarrow \mathbf{x}_i$ and $d_i \rightsquigarrow \mathbf{x}_i$. Lasso PDS

Related approaches exist for *optimal IV* estimation (Belloni et al., 2012) and/or *IV with many controls* (Chernozhukov, Hansen, and Spindler, 2015).

These methods have been implemented for Stata in pdslasso (Ahrens, Hansen, and Schaffer, 2020), dsregress (StataCorp) and R (hdm; Chernozhukov, Hansen, and Spindler, 2016).

Example syntax:

. pdslasso \$Y \$D (c.(\$X)#c.(\$X)), robust

Example 1 (pdslasso) allows for high-dimensional controls.

Example 2 (ivlasso) treats avexpr as endogenous and exploits
logem4 as an instrument. (More details in the pds/ivlasso help
file.)

There are **advantages** of relying on lasso:

- ▶ intuitive assumption of (approximate) sparsity
- computationally relatively cheap (due to plugin lasso penalty; no cross-validation needed)
- Linearity has its advantages (e.g. extension to fixed effects; Belloni et al., 2016)

But there are also drawbacks:

- What if the sparsity assumption is not plausible?
- There is a wide set of machine learners at disposable—Lasso might not be the best choice.
- Lasso requires careful feature engineering to deal with non-linearity & interaction effects.

Review of DDML

The partial linear model:

$$y_i = \theta d_i + g(\mathbf{x}_i) + \varepsilon_i$$
$$d_i = m(\mathbf{x}_i) + v_i$$

Naive idea: We estimate conditional expectations $\ell(\mathbf{x}_i) = E[y_i | \mathbf{x}_i]$ and $m(\mathbf{x}_i) = E[d_i | \mathbf{x}_i]$ using ML and partial out the effect of \mathbf{x}_i (in the style of Frisch-Waugh-Lovell):

$$\hat{\theta}_{DDML} = \left(\frac{1}{n}\sum_{i}\hat{v}_{i}^{2}\right)^{-1}\frac{1}{n}\sum_{i}\hat{v}_{i}(y_{i}-\hat{\ell}),$$

where $\hat{v}_i = d_i - \hat{m}_i$.

Review of DDML

Yet, there is a problem: The estimation error $\ell(\mathbf{x}_i) - \hat{\ell}$ and v_i may be correlated due to **over-fitting**, leading to poor performance.

DDML, thus, relies on **cross-fitting** (sample splitting with swapped samples).

DDML for the partial linear model (DML 2)

We split the sample in K random folds of equal size denoted by I_k :

- For k = 1,..., K, estimate ℓ(x_i) and m(x_i) using sample I^c_k and form out-of-sample predictions ℓ̂_i and m̂_i for all i in I_k.
- Construct estimator $\hat{\theta}$ as

$$\left(\frac{1}{n}\sum_{i}\hat{v}_{i}^{2}\right)^{-1}\frac{1}{n}\sum_{i}\hat{v}_{i}(y_{i}-\hat{\ell}),$$

where $\hat{v}_i = d_i - \hat{m}_i$. \hat{m}_i and $\hat{\ell}_i$ are the cross-fitted predicted values.

The DDML framework can be applied to other models (all implemented in ddml):

Interactive model

$$y_i = g(d_i, \mathbf{x}_i) + u_i \qquad E[u_i | \mathbf{x}_i, d_i] = 0$$

$$z_i = m(\mathbf{x}_i) + v_i \qquad E[u_i | \mathbf{x}_i] = 0$$

As in the Partial Linear Model, we are interested in the ATE, but do not assume that d_i (a binary treatment variable) and x_i are separable.

We estimate the conditional expectations $E[y_i|\mathbf{x}_i, d_i = 0]$ and $E[y_i|\mathbf{x}_i, d_i = 1]$ as well as $E[d_i|\mathbf{x}_i]$ using a supervised machine learner.

The DDML framework can be applied to other models (all implemented in ddml):

Partial linear IV model

$$y_i = d_i\theta + g(\mathbf{x}_i) + u_i \qquad E[u_i|\mathbf{x}_i, z_i] = 0$$

$$z_i = m(\mathbf{x}_i) + v_i \qquad E[v_i|\mathbf{x}_i] = 0$$

where the aim is to estimate the average treatment effect θ using observed instrument z_i in the presence of controls x_i . We estimate the conditional expectations $E[y_i|x_i]$, $E[d_i|x_i]$ and $E[z_i|x_i]$ using a supervised machine learner.

The DDML framework can be applied to other models (all implemented in ddml):

High-dimensional IV model

$$y_i = d_i\theta + g(\mathbf{x}_i) + u_i$$

$$d_i = h(\mathbf{z}_i) + m(\mathbf{x}_i) + v_i$$

where the parameter of interest is θ . The instruments and controls enter the model through unknown functions g(), h() and f().

We estimate the conditional expectations $E[y_i|\mathbf{x}_i]$, $E[\hat{d}_i|\mathbf{x}_i]$ and $\hat{d}_i := E[d_i|\mathbf{z}_i, \mathbf{x}_i]$ using a supervised machine learner. The instrument is then formed as $\hat{d}_i - \hat{E}[\hat{d}_i|\mathbf{x}_i]$ where $\hat{E}[\hat{d}_i|\mathbf{x}_i]$ denotes the estimate of $E[\hat{d}_i|\mathbf{x}_i]$.

The DDML framework can be applied to other models (all implemented in ddml):

Interactive IV model

$$y_{i} = \mu(\mathbf{x}_{i}, \mathbf{z}_{i}) + u_{i} \qquad E[u_{i}|\mathbf{x}_{i}, z_{i}] = 0$$

$$d_{i} = m(z_{i}, \mathbf{x}_{i}) + v_{i} \qquad E[v_{i}|\mathbf{x}_{i}, z_{i}] = 0$$

$$z_{i} = p(\mathbf{x}_{i}) + \xi_{i} \qquad E[\xi_{i}|\mathbf{x}_{i}] = 0$$

where the aim is to estimate the local average treatment effect.

We estimate, using a supervised machine learner, the following conditional expectations: $E[y_i | \mathbf{x}_i, z_i = 0]$ and $E[y_i | \mathbf{x}_i, z_i = 1]$; $E[D|\mathbf{x}_i, z_i = 0]$ and $E[D|\mathbf{x}_i, z_i = 1]$; $E[z_i | \mathbf{x}_i]$.

The ddml package

We introduce ddml for Stata:

- Compatible with various ML programs in Stata (e.g. lassopack, pylearn, randomforest).
 - \rightarrow Any program with the classical "reg y x" syntax and post-estimation predict will work.
- ▶ Short (one-line) and flexible multi-line version
- ► 5 models supported: partial linear model, interactive model, interactive IV model, partial IV model, optimal IV.

Stacking regression

Which machine learner should we use?

ddml supports a range of ML programs: pylearn, lassopack, randomforest. — Which one should we use?

We don't know whether we have a sparse or dense problem; linear or non-linear. We don't know whether, e.g., lasso or random forests will perform better.

Stacking, as implemented in pystacked, provides a solution: We use an 'optimal' combination of base learners.

Stacking regression

Which machine learner should we use?

We don't know whether we have a sparse or dense problem; linear or non-linear; etc.

Stacking is an ensemble method that combines multiple base learners into one model. As the default, we use *constrained least squares*:

$$\boldsymbol{w} = \arg\min_{w_j} \sum_{i=1}^n \left(y_i - \sum_{j=1}^J w_j \hat{y}_i^{(j)} \right)^2, \quad w_j \ge 0, \quad \sum_j w_j = 1$$

where $\hat{y}_i^{(j)}$ are cross-validated predictions of base learner j.

Voting regression is a special case with unweighted (or user-specified) weights.

Extended ddml syntax

Step 1: Initialise ddml and select model:

ddml init model [, kfolds(integer) reps(integer) ...] where model is either 'partial', 'iv', 'interactive', 'ivhd', 'late'.

Step 2: Add supervised ML programs for estimating conditional expectations:

```
ddml eq newvarname [, eqopt]: command depvar indepvars [,
cmdopt]
```

where *eq* selects the conditional expectations to be estimated. *command* is a ML program that supports the standard reg y x-type syntax. *cmdopt* are specific to that program.

Multiple estimation commands per equation are allowed.

Extended ddml syntax

```
Step 3: Cross-fitting
```

ddml crossfit [, shortstack]

```
Step 4: Estimation of causal effects
ddml estimate [, robust ...]
```

Additional auxiliary commands:

ddml describe (describe current model set up), ddml save & ddml use (to import/save ddml objects), ddml extract (to retrieve objects), ddml export (export in csv format).

Extended ddml syntax: Example

We demonstrate the use of ddml using the partially linear model by extending the analysis of 401(k) eligibility and total financial wealth of Poterba, Venti, and Wise (1995). The data consists of n = 9915 households from the 1991 SIPP.

Step 0: Load data, define globals

```
. use "sipp1991.dta", clear
```

- . global Y net_tfa
- . global X age inc educ fsize marr twoearn db pira hown
- . global D e401

Step 1: Initialise ddml and select model:

- . set seed 42
- . ddml init partial, kfolds(4)

Extended ddml syntax: Example (cont'd.)

Step 2: Add supervised ML programs for estimating conditional expectations. We use OLS, Lasso and Random Forest.

```
. *** add learners for E[Y|X]
. ddml E[Y|X]: reg $Y $X
Learner Y1_reg added successfully.
. ddml E[Y|X]: cvlasso $Y c.($X)#c.($X), lopt postresults
Learner Y2 cvlasso added successfully.
. ddml E[Y|X], vtype(none): rforest $Y $X, type(reg)
Learner Y3 rforest added successfully.
. *** add learners for E[D|X]
. ddml E[D|X]: reg $D $X
Learner D1 reg added successfully.
. ddml E[D|X]: cvlasso $D c.($X)#c.($X), lopt postresults
Learner D2 cvlasso added successfully.
. ddml E[D|X], vtype(none): rforest $D $X, type(reg)
Learner D3 rforest added successfully.
```

Extended ddml syntax: Example (cont'd.)

Step 3: Cross-fitting with 5 folds

. ddml crossfit, shortstack Cross-fitting E[Y|X] equation: net_tfa Cross-fitting fold 1 2 3 4 ...completed cross-fitting...completed short-stacking Cross-fitting E[D|X] equation: e401 Cross-fitting fold 1 2 3 4 ...completed cross-fitting...completed short-stacking

Extended ddml syntax: Example (cont'd.)

Step 4: Estimation of causal effects

. ddml estimate, robust DDML estimation results: spec Y learner D learner b SE r D1 reg 5964.151(1522.426) 1 1 Y1 reg 2 1 Y1_reg D2_cvlasso 8390.126(1356.633) 3 1 Y1_reg D3_rforest 8054.667(1271.281) 4 1 Y2 cvlasso D1 reg 9350.056(1381.641) 51 Y2 cvlasso D2 cvlasso 9570.601(1318.880) 1 [shortstack] [ss] 9401.724(1300.628) SS <-click or type ddml estimate, replay full to display full summary * = minimum MSE specification for that resample. Min MSE DDML model, specification 5 y-E[y|X] = Y2_cvlasso_1 Number of obs 9915 = D-E[D|X,Z]= D2_cvlasso_1

net_tfa	Coefficient	Robust std. err.	z	P> z	[95% conf.	interval]
e401	9570.601	1318.88	7.26	0.000	6985.644	12155.56

qddml example: Partial linear model

qddml is the one-line ('quick') version of ddml and uses a syntax similar to pds/ivlasso.

```
. qddml $Y $D (c.($X)#c.($X)), model(partial) ///
                               cmd(cvlasso) cmdopt(lopt postresults) ///
>
>
                               robust
DDML estimation results:
           Y learner
                         D learner
                                            b
                                                     SE
spec
    r
                             D1 reg 9504.777(1368.314)
      1
   1
              Y1 reg
  2 1
*
              Y1 reg D2 cvlasso 9512.796(1357.514)
   3 1
          Y2 cvlasso
                             D1 reg 9534.451(1373.390)
           Y2 cvlasso D2 cvlasso 9483.607(1361.398)
      1
* = minimum MSE specification for that resample.
Min MSE DDML model, specification 2
v-E[v|X] = Y1 reg 1
                                                   Number of obs
                                                                          9915
D-E[D|X,Z]= D2 cvlasso 1
```

net_tfa	Coefficient	Robust std. err.	z	P> z	[95% conf.	interval]
e401	9512.796	1357.514	7.01	0.000	6852.117	12173.47

Simulation I: Advantages of Stacking

Simulation set-up

We consider a *linear DGP* and a *non-linear DGP*, and compare performance of OLS, PDS-Lasso and various machine learners, including stacking.

We would expect that stacking performs well under both settings, while linear approaches only perform well if the DGP is linear.

Simulation I: Advantages of DDML+Stacking

Calibrated simulation based on Poterba, Venti, and Wise (1995), who estimate the causal effect of 401(k) eligibility on wealth.

- 1. Construct the partial residuals $y_i^{(r)} = y_i \hat{\tau}_{OLS} d_i$, $\forall i$ where $\hat{\tau}_{OLS}$ is the full sample OLS estiamte.
- 2. We predict $y_i^{(r)}$ with the controls x_i either using
 - ► linear regression (*Linear DGP*)

▶ gradient boosting (Non-Linear DGP) and call the fitted estimator *h*.

- 3. Similarly, predict d_i given x_i and call the estimator \tilde{g} .
- 3. We draw bootstrap sample \mathcal{D}_b of size n_s from the data
- 4. To generate $401(\mathsf{k})$ eligibility and log wealth, we calculate

$$egin{aligned} & ilde{d}_i^{(b)} = \mathbbm{1}\{ ilde{h}(x_i) +
u_i \geq 0.5\}, \quad
u_i \stackrel{iid}{\sim} \mathcal{N}(0,\kappa_1) \ & ilde{y}_i^{(b)} = au_0 ilde{d}_i^{(b)} + ilde{g}(x_i) + arepsilon_i, \quad arepsilon_i \stackrel{iid}{\sim} \mathcal{N}(0,\kappa_2), \quad orall i \in \mathcal{D}_b \end{aligned}$$

where κ_1 and κ_2 are chosen to match distributions of d_i and y_i .

Simulation I: Advantages of DDML+Stacking

	Linear DGP		Non-Li	Non-Linear DGP	
	Y X (1)	D X (2)	Y X (3)	D X (4)	
OLS	0.06	0.29	0.00	0.02	
cv-Lasso	0.35	0.17	0.01	0.00	
cv-Ridge	0.35	0.17	0.01	0.00	
Series Lasso (w/o interactions) Series Lasso (w/ interactions)	0.11 0.07	0.11 0.13	0.08 0.31	0.22 0.30	
Gradient boosting (low regularization)	0.03	0.05	0.30	0.23	
Gradient boosting (high regularization)	0.02	0.05	0.29	0.20	
Random forest (low regularization) Random forest (high regularization)	0.01 0.01	0.01 0.01	0.01 0.01	0.01 0.01	
Neural Network	0.00	0.00	0.00	0.00	

Table: Average Stacking Weights

Simulation I: Advantages of DDML+Stacking

	$n_s = 9$	915	$n_s = 99$	9150
Panel (A): Linear DGP	Bias	Rate	Bias	Rate
	(1)	(2)	(3)	(4)
OLS	-89.7	0.95	-0.3	0.94
DDML:				
cv-Lasso	-88.2	0.95	-0.4	0.94
Gradient boosting	-103.6	0.95	-7.2	0.94
Ensemble (stacking)	-112.5	0.94	-2.5	0.94
Panel (B): Non-Linear DGP	Bias	Rate	Bias	Rate
	(5)	(6)	(7)	(8)
OLS	-2580.0	0.54	-2599.4	0.00
DDML:				
cv-Lasso	-2615.7	0.53	-2604.0	0.00
Gradient boosting	-50.7	0.94	-0.4	0.98
Ensemble (stacking)	248.8	0.94	-1.2	0.98

Table: Coefficient Estimates

Wüthrich and Zhu (2021, henceforth WZ) demonstrate that PDS-Lasso suffers from a large finite sample bias and tends to underselect; again using the application of Poterba, Venti, and Wise (1995) and Belloni et al. (2017).

They use two specifications:

- two-way interactions (TWI) (as in Chernozhukov and Hansen, 2004); p = 167
- quadratic splines & interactions (QSI) (as in Belloni et al., 2017); p = 272

WZ run their simulations on bootstrap samples of the data $(n_b = \{200, 400, 800, 1600\})$ and calculate the bias as the mean difference to the full sample estimate (N = 9915).



(a) Bias (TWI specification)

(b) Bias (QSI specification)

Notes: The figures report the mean bias calculated as the mean difference to the full sample estimates. Following WZ, we draw 600 bootstrap samples of size $n_b = \{200, 400, 600, 800, 1200, 1600\}$. 'TWI' indicates that the predictors have been expanded by two-way interactions. 'QSI' refers to the quadratic spline & interactions specification of Belloni et al. (2017).

Figure: Replication of Figure 8 in WZ



(a) CV-Lasso

(b) CV-Ridge

Figure: Mean bias relative to full sample



(a) Boosted trees

(b) Stacking

Figure: Mean bias relative to full sample

The small sample bias of stacking stabilizes for $n_b > 600$, suggesting that stacking may perform well for 'moderate' sample sizes.

Summary

- ddml implements Double/Debiased Machine Learning for Stata:
 - Compatible with various ML programs in Stata
 - Short (one-line) and flexible multi-line version
 - Uses Stacking Regression as the default machine learner; implemented via separate program pystacked
 - 5 models supported
- The advantage to pdslasso is that we can make use of almost any machine learner.
- But which machine learner should we use?
 - We suggest stacking. We don't know which learner is best suited for a particular problem.
 - Stacking allows to consider multiple learners in a joint framework, and thus reduces the risk of misspecification.
- We are in the final phase of development; hopefully we can make ddml available soon (following your feedback)

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