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# How to assess the fit of multilevel logit models with Stata?

## A project in progress

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**“Models should not be true but it is important that  
they are applicable.” (John W. Tukey)**

Dr. Wolfgang Langer  
Martin-Luther-Universität  
Halle-Wittenberg  
Institut für Soziologie



Associate Assistant Professor  
Université du  
Luxembourg



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# 1. What is the problem ?

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Current situation in applied research:

- An increasing number of people use multilevel logistic models for qualitative dependent variables with binary and ordinal outcome
- But users often complain that there are no fit measures for these models
- Neither Stata 16 nor SPSS 26 offer any fit measure for these models
- Let me demonstrate how to generalize the Pseudo  $R^2$ s for binary and ordinal logit model for the multilevel analysis

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# Which solutions does Stata provide?

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- Indeed Stata estimates multilevel logit models for binary, ordinal and multinomial outcomes (melogit, meologit, gllamm) but it does not calculate any Pseudo  $R^2$ . It provides only the Akaike- (AIC) and Schwarz-Bayesian-Information Criteria (BIC)
- Stata provides a Wald test for the fixed effects and a Likelihood-Ratio- $\chi^2$  test for the random effects of the exogenous variables
- Even special purpose programs like HLM, MlwiN, MPLUS or SuperMix do not calculate any Pseudo  $R^2$

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# What can we learn from multilevel literature?

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- Raudenbush & Bryk (2002), Heck & Thomas (2009) and Rabe-Hesketh & Skrondal (2013) do not mention Pseudo  $R^2$ s at all
- Snijder & Bosker (2012) propose a variation of McKelvey & Zavoina Pseudo  $R^2$  for random-intercept and intercept-as-outcome logit models. It is not implemented in any program
- Hox (2010) discusses the McFadden, Cox & Snell, Nagelkerke and McKelvey & Zavoina Pseudo  $R^2$ . He recommends the last one to assess the model fit

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## 2. Summary of the econometric Monte-Carlo studies for testing Pseudo $R^2$ s

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- Econometricians made a lot of Monte-Carlo studies in the 1990s:
  - ▶ Hagle & Mitchell 1992
  - ▶ Veall & Zimmermann 1992, 1993, 1994
  - ▶ Windmeijer 1995
  - ▶ DeMaris 2002
- They systematically tested the most common Pseudo- $R^2$ s for binary and ordinal probit / logit models

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## Which Pseudo R<sup>2</sup>s were tested in these studies?

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- Likelihood-based measures:
  - ▶ Maddala / Cox & Snell Pseudo R<sup>2</sup> (1983 / 1989)
  - ▶ Cragg & Uhler / Nagelkerke Pseudo R<sup>2</sup> (1970 / 1992)
- Log-Likelihood-based measures:
  - ▶ McFadden Pseudo R<sup>2</sup> (1974)
  - ▶ Aldrich & Nelson Pseudo R<sup>2</sup> (1984)
  - ▶ Aldrich & Nelson Pseudo R<sup>2</sup> with the Veall & Zimmermann correction (1992)
- Basing on the estimated probabilities:
  - ▶ Efron / Lave Pseudo R<sup>2</sup> (1970 / 1978)
- Basing on the variance decomposition of the estimated Probits / Logits:
  - ▶ McKelvey & Zavoina Pseudo R<sup>2</sup> (1975)

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# Results of the Monte-Carlo-Studies for binary and ordinal logits or probits

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- The McKelvey & Zavoina Pseudo  $R^2$  is the best estimator for the “true  $R^2$ ” of the OLS regression
- The Aldrich & Nelson Pseudo  $R^2$  with the Veall & Zimmermann correction is the best approximation of the McKelvey & Zavoina Pseudo  $R^2$
- Lave / Efron, Aldrich & Nelson, McFadden and Cragg & Uhler Pseudo  $R^2$  severely underestimate the “true  $R^2$ ” of the OLS regression
- My personal advice:
  - ▶ Use the McKelvey & Zavoina Pseudo  $R^2$  to assess the fit of binary and ordinal logit models



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### 3. The generalization of the McKelvey & Zavoina Pseudo $R^2$ for the binary and ordinal multilevel logit model

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- The multilevel logit model is a systematic extension of the classical binary and ordinal logit model for clustered subsamples (contextual units  $j$ )
  - ▶ The variance of the estimated logits is decomposed into
    - ▶ Fixed effects, ▶ Random effects and ▶ Level-1 Error variance  $\sigma^2(r_{ij})$
  - ▶ The variance of level 1 residua  $\sigma^2(r_{ij})$  can not be estimated because of its own heteroscedasticity. It is replaced by the variance of the logistic density function  $(\pi^2 / 3)$  multiplied with the sample size

# Let's have a short look at the lucky winner

- McKelvey & Zavoina Pseudo  $R^2$  (M & Z Pseudo  $R^2$ )

$$M \ \& \ Z \ Pseudo \ R^2 = \frac{\sum_{i=1}^n \left( \hat{y}_i^* - \overline{\hat{y}^*} \right)^2}{\sum_{i=1}^n \left( \hat{y}_i^* - \overline{\hat{y}^*} \right)^2 + n \times \frac{\pi^2}{3}}$$

Range:  $0 \leq M \ \& \ Z\text{-Pseudo } R^2 \leq 1$

Legend:

$\sum_{i=1}^n \left( \hat{y}_i^* - \overline{\hat{y}^*} \right)^2$  : Sum of squares of the estimated logits (latent variable  $Y^*$ )

$n$  : Sample size

$\frac{\pi^2}{3}$  : Variance of the logistic density function

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# Generalization to the 2-level logit model 2

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- Prediction of the latent variable  $Y^*$  (estimated binary or cumulative logit) in two ways
  - ▶ 1. Population-Average Prediction with the fixed effects of the exogenous variables (all random effects hold at zero)
    - Stata-command: `predict newvar1 if e(sample), xb`
  - ▶ 2. Unit-Specific Prediction of the fixed and random effects of the exogenous variable
    - Stata-command: `predict newvar2 if e(sample), eta`
- Therefore, the variation of the estimated logits ( $Y^*$ ) can be calculated in two different ways
  - ▶ 1. Only for the fixed effects of the exogenous variables
  - ▶ 2. For the fixed and random effects of the exogenous variables

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# Generalization to the 2-level logit model 3

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- Therefore we get two different McKelvey & Zavoina Pseudo  $R^2$ s
  - ▶ 1. “Population-Average” M & Z Pseudo  $R^2$  (fixed effects)
  - ▶ 2. “Unit-Specific” M & Z Pseudo  $R^2$  (fixed & random effects)
- The “Unit-Specific” M & Z Pseudo  $R^2$  uses all estimated fixed and random effects for prediction. Therefore it assesses the fit more realistically as its “Population-Average” counterpart

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# Let's have a short look at the lucky loser

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- McFadden Pseudo  $R^2$  (1974)

$$McFadden\ Pseudo\ R^2\ (\rho^2) = 1 - \left[ \frac{\ln L_A}{\ln L_0} \right]$$

Range:  $0 \leq$  McFadden Pseudo  $R^2 < 1$

but  $\rho^2$  does not reach the maximum of 1.0

Rule of thumb:  $0.20 \leq$  McFadden Pseudo  $R^2 \leq 0.40$  marks an excellent fit (McFadden 1979: 307)

Legend:  $\ln L_A$ : Log-Likelihood of the actual model  
 $\ln L_0$ : Log-Likelihood of the zero model

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# Generalization to the 2-level logit model 4

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- Conditions of application
  - ▶ Maximum-Likelihood estimation of the fixed and random effects of the exogenous variables
  - ▶ Actual and zero model have to use the same sample
  - ▶ Choice of the “appropriate zero model” ( $M_0$ ) depends on our knowledge to which contextual unit the respondent belongs
    - **Membership known:** Random-Intercept-Only Logit model estimates the proportion of  $Y^*$  which can be maximally explained by the contextual units (= ANOVA model)
    - **Membership unknown:** Fixed-Intercept-Only Logit model estimates only the marginal distribution of  $Y^*$  (= true zero model)

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# Generalization to the 2-level logit model 5

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- Calculation of McFadden Pseudo  $R^2$  is possible in two different ways using the following ones as zero model
  - ▶ 1. Random-Intercept-Only Logit-Model (RIOM)
    - It measures the proportional reduction of the log likelihood of the actual model in comparison with the RIOM caused by the fixed effects of the exogenous variables
    - Its Likelihood-Ratio  $\chi^2$  test refers to all fixed effects of the exogenous level 1 and level 2 variables
  - ▶ 2. Fixed-Intercept-Only Logit-Model (FIOM)
    - It measures the proportional reduction of the log likelihood of the actual model in comparison with the FIOM caused by fixed and random effects of all exogenous variables
    - Its Likelihood-Ratio  $\chi^2$  test refers to all fixed and random effects of the exogenous level 1 and level 2 variables

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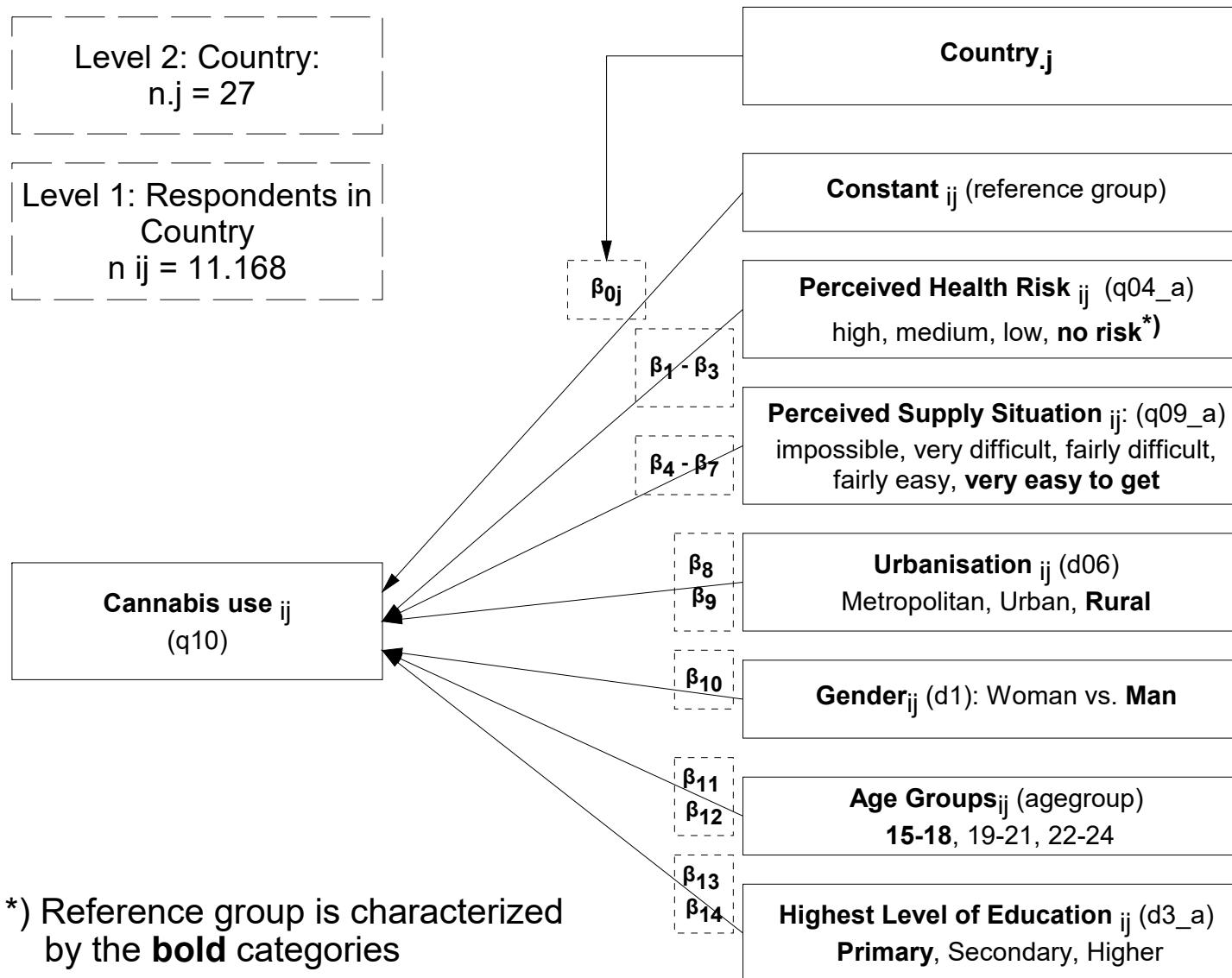
## 4. Example of application

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- Flash Eurobarometer No 330 about youth attitudes on drugs (2011)
  - ▶ WebCATI-Survey of  $n_{ij} = 12.313$  respondents (aged 15 -24) in  $n_{.j} = 27$  EU member states (contextual units  $j$ )
  - ▶ My focus:
    - prevalence of cannabis use by juveniles and young adults (q10): Have you used cannabis by yourself?
      - 1) never
      - 2) more than 12 months ago
      - 3) less than 12 months ago
      - 4) in the last 30 days
  - ▶ Let us have a look at the exogenous variables in the following diagram



# Theoretical 2-level-model: RIM



# Stata-Output Version 16

Mixed-effects ologit regression  
Group variable: country

Number of obs = 11,168  
Number of groups = 27  
Obs per group:  
min = 211  
avg = 413.6  
max = 490

Integration method: mvaghermite

Integration pts. = 7

Log pseudolikelihood = -7424.2751

Wald chi2(14) = 3363.39  
Prob > chi2 = 0.0000

(Std. Err. adjusted for 27 clusters in country)

q10ord	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
-----						
q4_a						
high risk	-2.656545	.1498799	-17.72	0.000	-2.950304	-2.362785
medium risk	-1.668222	.1015971	-16.42	0.000	-1.867348	-1.469095
low risk	-.7527713	.0463199	-16.25	0.000	-.8435567	-.6619859
q9_a						
impossible	-2.976197	.1910924	-15.57	0.000	-3.350731	-2.601663
very difficult	-2.132899	.1642871	-12.98	0.000	-2.454896	-1.810903
fairly difficult	-1.527717	.1004747	-15.21	0.000	-1.724644	-1.330791
fairly easy	-.6241175	.0843823	-7.40	0.000	-.7895038	-.4587312
d6						
metropolitan zone	.3950294	.1038006	3.81	0.000	.191584	.5984749
other town/urban centre	.2082467	.0751678	2.77	0.006	.0609206	.3555729
d1						
female	-.4777186	.0525238	-9.10	0.000	-.5806634	-.3747738
agegroup						
19 - 21	.4967552	.0607537	8.18	0.000	.3776801	.6158303
22 - 24	.6850326	.0788467	8.69	0.000	.5304958	.8395694
d3_a						
secondary education	-.0122673	.0702398	-0.17	0.861	-.1499348	.1254003
higher education	-.0268133	.1227001	-0.22	0.827	-.2673011	.2136745
/cut1	-.3882657	.1392505			-.6611918	-.1153396
/cut2	.7153171	.1438259			.4334236	.9972106
/cut3	1.902703	.1602572			1.588605	2.216802
-----						
country						
var(_cons)	.2617043	.0850279			.1384379	.4947284
-----						

► Fixed effects

► Thresholds

► Random effect

# What does Stata offer to assess the fit?

- Akaike (AIC) and Schwarz Bayesian Information Criterion (BIC)
  - ▶ Decision rule: Choose the model with the lowest AIC or BIC

Akaike's information criterion and Bayesian information criterion

Model	N	ll(null)	ll(model)	df	AIC	BIC
fiom_w4	11,168	.	-9343.799	3	18693.60	18715.56
riom_w4	11,168	.	-9036.540	4	18081.08	18110.36
rim_w4	11,168	.	-7424.275	18	14884.55	15016.32

Note: BIC uses N = number of observations. See [R] BIC note.

- ▶ Looking at AIC and BIC, the RIM fits best of all bad models
- ▶ But we do not know how well the RIM really fits !

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# Output of my fit\_meologit\_2lev.ado

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1

- Assessing the fit by the McKelvey & Zavoina-, McFadden-Pseudo  $R^2$ s for the fixed & random effects

```
. fit_meologit_2lev
```

```
Fit-measures for the MELOGIT/MEOLOGIT Model:
```

```
McKelvey&Zavoina-Pseudo R2 (fixed & random effects)= 0.5097
```

```
McKelvey&Zavoina-Pseudo R2 (fixed effects only)= 0.4728
```

```
Just estimating the Fixed-/ Random-Intercept-Only-Logit Model
```

```
McFadden Pseudo R2 (fixed effects only) = 0.1784
```

```
McFadden Pseudo R2 (fixed & random effects) = 0.2054
```

# Output of my fit\_meologit\_2lev.ado

2

- Intra-Class Correlation and corresponding Likelihood-Ratio- $\chi^2$  tests for fixed & random effects

```
ICC of Random-Intercept-Only-Logit Model (Sample M(A))  
Intra-Class-Correlation (Level 2) = 0.1431
```

```
H0: ICC of Level 2 is zero in the population
```

```
LR-chi2 test statistic ( 1) = 614.52 Prob > chi2 = 0.0000
```

```
LR-chi2 test: H0: all fixed effects are zero in the population
```

```
LR-chi2 test statistic ( 14) = 3224.53 Prob > chi2 = 0.0000
```

```
LR-chi2 test: H0: all fixed & random effects are zero in the population
```

```
LR-chi2 test statistic ( 15) = 3839.05 Prob > chi2 = 0.0000
```

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# What you get afterwards

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- r-containers of fit\_meologit\_2lev.ado

```
. return list
```

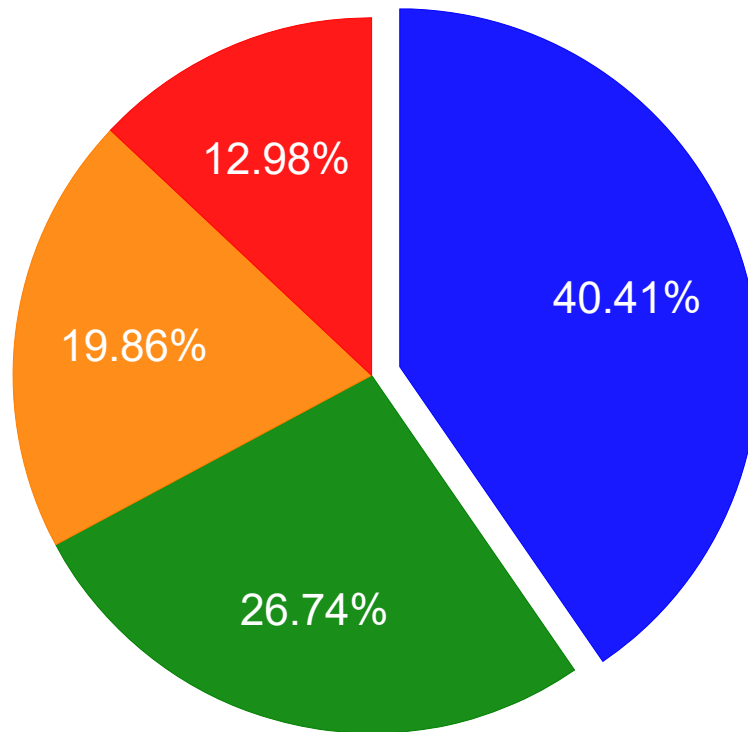
```
scalars:
```

```
    r(p_chi2_fr) = 0
      r(df_fr) = 15
    r(chi2_fr) = 3839.047042018567
    r(p_chi2_f) = 0
      r(df_f) = 14
    r(chi2_f) = 3224.530363760103
    r(p_icc) = 0
    r(df_icc) = 1
    r(chi2_icc) = 614.5166782584638
    r(icc_riom) = .1431007329841504
    r(mcr2_fiom) = .2054328872219874
    r(mcr2_riom) = .1784162011068596
      r(mzr2f) = .4728218379947466
    r(mzr2fr) = .5097289342586476
```

# How do the effects look like?

## ● The baseline

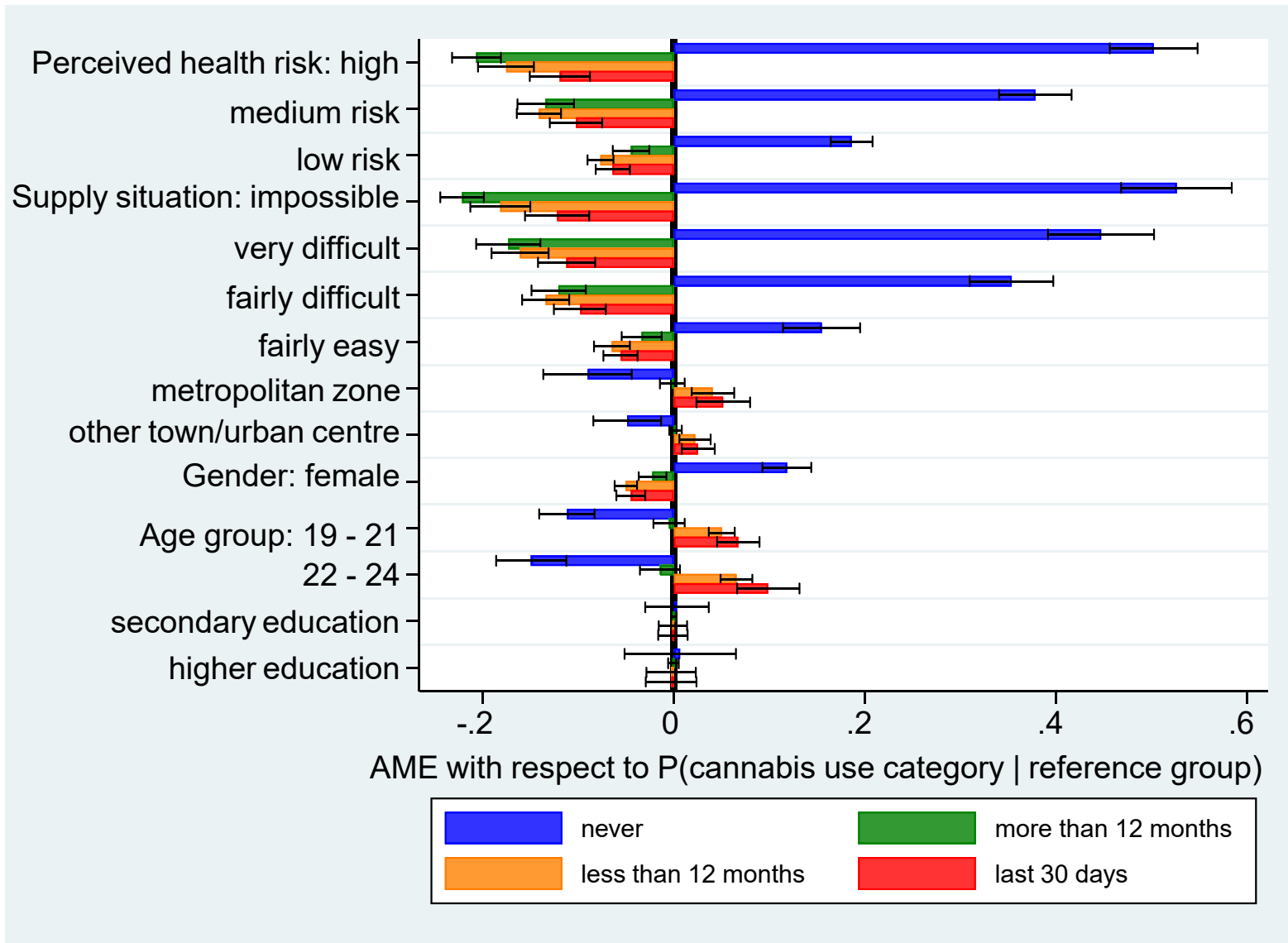
Estimated probabilities of cannabis use for the reference group



## ▶ Reference Group

- Men
- Age 15-18
- Location: rural
- Education: primary
- Perceived health risk: no risk
- Perceived supply situation: very easy to get

# Ben Jann's Coefplot for the 4 categories





- What's known
  - ▶ The Monte-Carlo-simulation studies show that the McKelvey & Zavoina Pseudo  $R^2$  is the best fit measure for binary and ordinal logit models
- What's new
  - ▶ Generalization of the M & Z-Pseudo  $R^2$  to binary and ordinal multilevel logit models. The prediction of estimated logits bases upon the fixed effects only or upon fixed and random effects of exogenous variables
  - ▶ The McFadden-Pseudo  $R^2$  bases upon the fixed effects only or upon fixed and random effects of the exogenous variables using a context-independent zero model

- What's new
  - ▶ Simultaneous Likelihood-Ratio- $\chi^2$  test for the estimated fixed effects using the Random-Intercept-Only Model (RIOM) as the zero model
  - ▶ Simultaneous Likelihood-Ratio- $\chi^2$  test for the estimated fixed and random effects using the Fixed-Intercept-Only Model (FIOM) as the zero model
  - ▶ Use of probability weights for each level  $j$
  - ▶ You get all Pseudo-R2s and tests in r-containers
- That's why
  - ▶ I suggest to use my `fit_meologit_2lev.ado` and `fit_meologit_3lev.ado` to assess the fit of 2- and 3-level logit models with binary and ordinal outcome

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# Closing words

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- Thank you for your attention
- Do you have some questions?

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# Contact

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- Affiliation

- ▶ Dr. Wolfgang Langer  
University of Halle  
Institute of Sociology  
D-06099 Halle (Saale)

- ▶ Email: [wolfgang.langer@soziologie.uni-halle.de](mailto:wolfgang.langer@soziologie.uni-halle.de)

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# Appendix

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# Multilevel ordered logit model

1

## Equations of the 2-level-ordered logit model

Level 2: Between-Context Regression

2a) Logistic Intercept-as-Outcome Model:

$$\beta_{0j} = 0 + \gamma_{01} \times Z_{.j} + u_{0j}$$

2b) Logistic Slope-as-Outcome Model:

$$\beta_{1j} = \gamma_{10} + \gamma_{11} \times Z_{.j} + u_{1j}$$

Level 1: Within-Context Regression

$$1) \ln \left[ \frac{P(Y > k)}{P(Y \leq k)} \right] = \beta_{0j} + \beta_{1j} \times X_{ij} - \sum_{K=1}^{k-1} \delta_k \{+r_{ij}\}$$

Single equation notation: 2a) and 2b) in 1)

$$\ln \left[ \frac{P(Y > k)}{P(Y \leq k)} \right] = \left( 0 + \gamma_{01} \times Z_{.j} + u_{0j} \right) + \left( \gamma_{10} \times X_{ij} + \gamma_{11} \times X_{ij} \times Z_{.j} + u_{1j} \times X_{ij} \right) - \sum_{k=1}^{K-1} \delta_k \{+r_{ij}\}$$

Notation of Raudenbush & Bryk (2002):

$\gamma$ : fixed-effect estimator

$Z$ : exogenous level 2 variable

$\beta$ : random-effect estimator

$X$ : exogenous level 1 variable

$u_{0j}$ : residuum random-intercept

$u_{1j}$ : residuum random-slope

$r_{ij}$ : residuum of within-context-logistic regression

$\delta_k$ : threshold for category k of Y

- Interpretation of the residua of the Between-Context Regression

$$3a) u_{0j} = \beta_{0j} - [\gamma_{00} + \gamma_{01} \times Z_{.j}] = \beta_{0j} - \widehat{\beta}_{0j}$$

$$3b) u_{1j} = \beta_{1j} - [\gamma_{10} + \gamma_{11} \times Z_{.j}] = \beta_{1j} - \widehat{\beta}_{1j}$$

- Assumptions for the residua of the logistic 2-level logit model

Level 1:

1.1)  $r_{ij}$  is binomial distributed with an expected value of zero

and a variance  $\sigma_{r_{ij}}^2 = \widehat{P}_{ij}(Y=1) \times (1 - \widehat{P}_{ij}(Y=1))$

1.2) Heteroscedasticity of  $r_{ij}$  in all contextual units  $j$

- Implication for the level 1 residuum  $r_{ij}$ 
  - ▶ The variance  $\sigma^2(r_{ij})$  can not be estimated because of its own heteroscedasticity. It is replaced by the variance of the logistic density function ( $\pi^2 / 3$ )

## ● Residua of level 2

2.1)  $u_{kj}$  is normal distributed with an expected value of zero and a covariance matrix T of the residua

$$E \begin{bmatrix} u_{0j} \\ u_{1j} \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \end{bmatrix} \quad T = \begin{bmatrix} \tau_{00} & \tau_{01} \\ \tau_{10} & \tau_{11} \end{bmatrix} \quad \sigma_{u_{0j}}^2 = \tau_{00} \quad \sigma_{u_{1j}}^2 = \tau_{11}$$

$$\sigma_{u_{0j}, u_{1j}} = \tau_{10} = \tau_{01}$$

2.2) The residua of level 1 and level 2 are not correlated:

$$\sigma_{u_{0j}, r_{ij}} = \sigma_{u_{1j}, r_{ij}} = 0$$



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# Alternative in Stata: Information criteria

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- Calculation of Akaike- (AIC) and Schwarz Bayesian-Information-Criteria (BIC)

$$AIC = -2 \times \ln L_{M_A} + 2 \times k$$

$$BIC = -2 \times \ln L_{M_A} + \ln N \times k$$

deviance

complexity of the model

Range:  $0 < AIC \leq +\infty$

$0 < BIC \leq +\infty$

Legend:

$\ln L_{M_A}$  : Log Likelihood of actual model

$k$  : Number of estimated parameters

$N$  : *Sample size*

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