

Estimating the ATE of an endogenously assigned treatment from a sample with endogenous selection by regression adjustment using an extended regression models

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- Fictional data on wellness program from large company

```
. // use wprogram
. use wprogram2
. describe
```

Contains data from wprogram2.dta

```
obs:      3,000
vars:      8                28 Jul 2017 07:13
size:     96,000
```

variable name	storage type	display format	value label	variable label
wchange	float	%9.0g	changel	Weight change level
age	float	%9.0g		Years over 50
over	float	%9.0g		Overweight (tens of pounds)
phealth	float	%9.0g		Prior health score
prog	float	%9.0g	yesno	Participate in wellness program
wtprog	float	%9.0g	yesno	Offered work time to participate in program
wtsamp	float	%9.0g		Offered work time to participate in sample
insamp	float	%9.0g		In sample: attended initial and final weigh in

Sorted by:

- Three levels of wchange

```
. tabulate wchange prog
```

Weight change level	Participate in wellness program		Total
	No	Yes	
Loss	154	960	1,114
No change	251	299	550
Gain	184	36	220
Total	589	1,295	1,884

- Data are observational
- Table does not account for
 - how observed covariates that affect program participation also affect the potential outcome variables
 - how observed unobserved error that affect program participation also affect the potential outcome variables
 - the possibility that unobserved errors in the process caused some of 3,000 individuals not to show for the final weigh in may also affect the potential outcome variables

Ordinal Potential outcomes

- Because the outcome w_{change} is ordinal, there are really three binary outcomes $w_{change} == \text{"Loss"}$, $w_{change} == \text{"No Change"}$, and $w_{change} == \text{"Gain"}$
- In the potential outcome framework, there is an outcome for each person when they participate and when they do not participate
- Thus, there are really three binary outcomes for each potential outcome

Participate		Not participate	
w_{change}_p	$==$ "Loss"	w_{change}_{np}	$==$ "Loss"
w_{change}_p	$==$ "No change"	w_{change}_{np}	$==$ "No change"
w_{change}_p	$==$ "Gain"	w_{change}_{np}	$==$ "Gain"

Potential outcome framework

- For each outcome (Loss, No change, and Gain), we only observe one of these two potential outcomes for each individual
- We estimate the parameters of a model and use the estimated parameters to predict what each person does in the unobserved potential outcome

Average treatment effects

- In the case of one outcome, the Average treatment effect (ATE) is

$$\mathbf{E}[y_p - y_{np}]$$

- As there are three outcomes, there are three ATEs
 - one for “Loss”, one for “No Change”, and one for “Gain”

$$ATE_{Loss} = \mathbf{E}[wchange_p == \textit{“Loss”} - wchange_{np} == \textit{“Loss”}]$$

$$ATE_{Nochange} = \mathbf{E}[wchange_p == \textit{“No change”} - wchange_{np} == \textit{“No change”}]$$

$$ATE_{Gain} = \mathbf{E}[wchange_p == \textit{“Gain”} - wchange_{np} == \textit{“Gain”}]$$

Average treatment effects

- I will provide some details about the average treatment effect for “Loss”
- The details for the outcomes of “No change” and “Gain” are analogous

- the average treatment effect (ATE) of the program on the Loss outcome ATE_{Loss}

$$\begin{aligned}
 ATE_{Loss} &= \mathbf{E}[\text{wchange}_p == \text{"Loss"} - \text{wchange}_{np} == \text{"Loss"}] \\
 &= \mathbf{E}[\text{wchange}_p == \text{"Loss"}] - \mathbf{E}[\text{wchange}_{np} == \text{"Loss"}] \\
 &= \text{Pr}[\text{wchange}_p == \text{"Loss"}] - \text{Pr}[\text{wchange}_{np} == \text{"Loss"}]
 \end{aligned}$$

- The first line says that ATE_{Loss} is the mean difference in the outcomes when everyone participates instead of no one participates
- The second line says that the mean of the differences is the difference in the means
- The third line uses the fact that mean of binary outcome is the probability that the event is true, to say that the ATE_{Loss} is the difference in the the probability of being in the state of "Loss" when everyone participates instead of no one participates

- I am going to use the ERM comand eoprobit to estimate the parameters of $\Pr[\text{wchange}_p == \text{"Loss"} | \mathbf{x}]$ and $\Pr[\text{wchange}_{np} == \text{"Loss"} | \mathbf{x}]$ and
- Then I use margins to estimate

$$\begin{aligned} & \mathbf{E}[\Pr[\text{wchange}_p == \text{"Loss"} | \mathbf{x}]] - \mathbf{E}[\Pr[\text{wchange}_{np} == \text{"Loss"} | \mathbf{x}]] \\ &= \Pr[\text{wchange}_p == \text{"Loss"}] - \Pr[\text{wchange}_{np} == \text{"Loss"}] \\ &= ATE_{Loss} \end{aligned}$$

- The ATE_{Loss} is the mean difference in the the probability of being in the state of "Loss" when everyone participates instead of no one participates

Models for the ordinal outcome

- For exogenous treatment, we do a one-step equivalent to fitting two separate ordinal probit models
 - One fit to participants
 - Another fit to non participants

$$wchange = \begin{cases} \text{"Loss"} & \text{if } \mathbf{x}\beta_0 + \epsilon_0 \leq cut1_0 \\ \text{"No change"} & \text{if } cut1_0 < \mathbf{x}\beta_0 + \epsilon_0 \leq cut2_0 \\ \text{"Gain"} & \text{if } cut2_0 < \mathbf{x}\beta_0 + \epsilon_0 \end{cases}$$

$$\mathbf{x}\beta_0 = \beta_{1,0}age + \beta_{2,0}over + \beta_{3,0}phealth$$

for the observations at which prog=0, and

$$wchange = \begin{cases} \text{"Loss"} & \text{if } \mathbf{x}\beta_1 + \epsilon_1 \leq cut1_1 \\ \text{"No change"} & \text{if } cut1_1 < \mathbf{x}\beta_1 + \epsilon_1 \leq cut2_1 \\ \text{"Gain"} & \text{if } cut2_1 < \mathbf{x}\beta_1 + \epsilon_1 \end{cases}$$

$$\mathbf{x}\beta_1 = \beta_{1,1}age + \beta_{2,1}over + \beta_{3,1}phealth$$

for the observations at which prog=1

ϵ_0 , and ϵ_1 are normal

$corr(\epsilon_0, \epsilon_1)$ is not identified or estimated

Exogenous treatment

```
. eoprobit wchange age over phealth, extreat(prog) vsquish nolog
```

```
Extended ordered probit regression
```

```
Number of obs = 1,884
```

```
Wald chi2(6) = 99.08
```

```
Log likelihood = -1434.5465
```

```
Prob > chi2 = 0.0000
```

wchange	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
prog#c.age						
No	.2180787	.1464522	1.49	0.136	-.0689623	.5051196
Yes	-.2356064	.1196215	-1.97	0.049	-.4700603	-.0011526
prog#c.over						
No	.2156394	.0784599	2.75	0.006	.0618609	.3694179
Yes	-.0352986	.0781835	-0.45	0.652	-.1885355	.1179383
prog# c.phealth						
No	-.0746153	.0844652	-0.88	0.377	-.2401641	.0909334
Yes	-.6229527	.0669733	-9.30	0.000	-.7542181	-.4916874
/wchange						
prog#c.cut1						
No	-.4960282	.0978731			-.6878559	-.3042005
Yes	.0712884	.0810525			-.0875716	.2301484
prog#c.cut2						
No	.642945	.0988945			.4491153	.8367747
Yes	1.421407	.0984319			1.228484	1.61433

ATE

```
. generate prog_original = prog
. replace prog = 0
(1,700 real changes made)
. predict double pr_loss_0 , outlevel("Loss")
(option pr assumed; predicted probabilities)
. replace prog = 1
(3,000 real changes made)
. predict double pr_loss_1 , outlevel("Loss")
(option pr assumed; predicted probabilities)
. replace prog = prog_original
(1,300 real changes made)
. drop prog_original
. mean pr_loss_0 pr_loss_1
```

Mean estimation Number of obs = 3,000

	Mean	Std. Err.	[95% Conf. Interval]	
pr_loss_0	.2721432	.0009077	.2703634	.273923
pr_loss_1	.7096007	.0020206	.7056388	.7135625

ATE

```
. estimates restore oprobit
(results oprobit are active now)

. margins prog,          ///
>   predict(outlevel("Loss"))    ///
>   predict(outlevel("No change"))  ///
>   predict(outlevel("Gain")) noesample

Predictive margins                    Number of obs      =      3,000
Model VCE      : OIM

1._predict    : Pr(wchange==Loss), predict(outlevel("Loss"))
2._predict    : Pr(wchange==No change), predict(outlevel("No change"))
3._predict    : Pr(wchange==Gain), predict(outlevel("Gain"))
```

	Delta-method					
	Margin	Std. Err.	z	P> z	[95% Conf. Interval]	
__predict#prog						
1#No	.2721432	.0191116	14.24	0.000	.2346853	.3096012
1#Yes	.7096007	.0142655	49.74	0.000	.6816407	.7375606
2#No	.4260522	.0203869	20.90	0.000	.3860947	.4660097
2#Yes	.25725	.0133175	19.32	0.000	.2311483	.2833518
3#No	.3018046	.0191367	15.77	0.000	.2642973	.3393118
3#Yes	.0331493	.0055184	6.01	0.000	.0223334	.0439652

```

. margins r.prog,          ///
>   predict(outlevel("Loss"))    ///
>   predict(outlevel("No change"))  ///
>   predict(outlevel("Gain"))    ///
>   contrast(nowald)            ///
>   noesample

```

Contrasts of predictive margins

Model VCE : OIM

```

1._predict : Pr(wchange==Loss), predict(outlevel("Loss"))
2._predict : Pr(wchange==No change), predict(outlevel("No change"))
3._predict : Pr(wchange==Gain), predict(outlevel("Gain"))

```

	Delta-method		
	Contrast	Std. Err.	[95% Conf. Interval]
prog@_predict			
(Yes vs No) 1	.4374574	.0238486	.390715 .4841999
(Yes vs No) 2	-.1688022	.0243512	-.2165296 -.1210748
(Yes vs No) 3	-.2686552	.0199165	-.3076908 -.2296196

- When everyone joins the program instead of when no one participants in the program,
 - On average, the probability of “Loss” goes up by .44
 - On average, the probability of “No change” goes down by .17
 - On average, the probability of “Gain” goes down .27

Endogenous Treatment model

The potential-outcome model for an endogenous treatment

- Allows the coefficients to differ for the treated and not-treated state
- Allows the cut offs to differ for the treated and not-treated state
- Allows for distinct (nonzero) correlations between the errors driving treatment assignment and the errors driving the ordinal outcomes for the treated and not-treated states

$$prog = (\mathbf{x}\gamma + \gamma_1 wtprog + \eta > 0)$$

$$wchange = \begin{cases} \text{"Loss"} & \text{if } \mathbf{x}\beta_0 + \epsilon_0 \leq cut1_0 \\ \text{"No change"} & \text{if } cut1_0 < \mathbf{x}\beta_0 + \epsilon_0 \leq cut2_0 \\ \text{"Gain"} & \text{if } cut2_0 < \mathbf{x}\beta_0 + \epsilon_0 \end{cases}$$

$$\mathbf{x}\beta_0 = \beta_{1,0}age + \beta_{2,0}over + \beta_{3,0}phealth$$

for the observations at which $prog=0$, and

$$wchange = \begin{cases} \text{"Loss"} & \text{if } \mathbf{x}\beta_1 + \epsilon_1 \leq cut1_1 \\ \text{"No change"} & \text{if } cut1_1 < \mathbf{x}\beta_1 + \epsilon_1 \leq cut2_1 \\ \text{"Gain"} & \text{if } cut2_1 < \mathbf{x}\beta_1 + \epsilon_1 \end{cases}$$

$$\mathbf{x}\beta_1 = \beta_{1,1}age + \beta_{2,1}over + \beta_{3,1}phealth$$

for the observations at which $prog=1$

ϵ_0 , ϵ_1 , and η are correlated and joint normal

ρ_0 correlation between ϵ_0 and η

ρ_1 correlation between ϵ_1 and η

Endogenous treatment model

```

. eoprobit wchange age over phealth , ///
>      entreat(prog = age over phealth wtprog, pocorr ) ///
>      vce(robust) vsquish nolog
Extended ordered probit regression      Number of obs      =      1,884
                                          Wald chi2(6)        =      137.27
Log pseudolikelihood = -2335.2213      Prob > chi2         =      0.0000

```

	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
wchange						
prog#c.age						
No	.4919782	.1357859	3.62	0.000	.2258427	.7581137
Yes	-.1111304	.1183412	-0.94	0.348	-.3430749	.1208142
prog#c.over						
No	.4659558	.0789709	5.90	0.000	.3111757	.6207359
Yes	.0458895	.0794788	0.58	0.564	-.109886	.2016651
prog#						
c.phealth						
No	-.3162974	.0872579	-3.62	0.000	-.4873198	-.145275
Yes	-.6880971	.0713535	-9.64	0.000	-.8279474	-.5482467
prog						
age	-.9224146	.1057226	-8.72	0.000	-1.129627	-.7152021
over	-.9957274	.0675412	-14.74	0.000	-1.128106	-.863349
phealth	.7483889	.0604543	12.38	0.000	.6299007	.8668771
wtprog	1.718043	.1160706	14.80	0.000	1.490549	1.945537
pocorr	.3388047	.0690413	4.92	0.000	.2044863	.475123

	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
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prog#c.age						
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No	-.3162974	.0872579	-3.62	0.000	-.4873198	-.145275
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prog						
age	-.9224146	.1057226	-8.72	0.000	-1.129627	-.7152021
over	-.9957274	.0675412	-14.74	0.000	-1.128106	-.863349
phealth	.7483889	.0604543	12.38	0.000	.6299007	.8668771
wtprog	1.718043	.1160706	14.80	0.000	1.490549	1.945537
_cons	.3398047	.0690413	4.92	0.000	.2044863	.475123
/wchange						
prog#c.cut1						
No	.1953761	.1544741			-.1073875	.4981397
Yes	-.133868	.0985578			-.3270377	.0593017
prog#c.cut2						
No	1.193014	.111908			.9736779	1.412349
Yes	1.170747	.1289195			.9180695	1.423425
corr(e.prog, wchange)						
prog						

	Yes	-.1111304	.1183412	-0.94	0.348	-.3430749	.1208142
prog#c.over	No	.4659558	.0789709	5.90	0.000	.3111757	.6207359
	Yes	.0458895	.0794788	0.58	0.564	-.109886	.2016651
prog#c.phealth	No	-.3162974	.0872579	-3.62	0.000	-.4873198	-.145275
	Yes	-.6880971	.0713535	-9.64	0.000	-.8279474	-.5482467
prog	age	-.9224146	.1057226	-8.72	0.000	-1.129627	-.7152021
	over	-.9957274	.0675412	-14.74	0.000	-1.128106	-.863349
	phealth	.7483889	.0604543	12.38	0.000	.6299007	.8668771
	wtprog	1.718043	.1160706	14.80	0.000	1.490549	1.945537
	_cons	.3398047	.0690413	4.92	0.000	.2044863	.475123
/wchange	prog#c.cut1						
	No	.1953761	.1544741			-.1073875	.4981397
	Yes	-.133868	.0985578			-.3270377	.0593017
	prog#c.cut2						
	No	1.193014	.111908			.9736779	1.412349
	Yes	1.170747	.1289195			.9180695	1.423425
corr(e.prog, e.wchange)	prog						
	No	-.6325687	.1073524	-5.89	0.000	-.7992197	-.3755982
	Yes	-.4199058	.1042067	-4.03	0.000	-.6015292	-.1970056

```
. estat teffects
```

```
Predictive margins
```

```
Number of obs      =      1,884
```

```
ATE_Pr0      : Pr(wchange=0=Loss)  
ATE_Pr1      : Pr(wchange=1=No change)  
ATE_Pr2      : Pr(wchange=2=Gain)
```

	Unconditional Margin	Std. Err.	z	P> z	[95% Conf. Interval]	
ATE_Pr0 prog (Yes vs No)	.0998676	.0620691	1.61	0.108	-.0217856	.2215208
ATE_Pr1 prog (Yes vs No)	-.013337	.0452574	-0.29	0.768	-.1020399	.0753658
ATE_Pr2 prog (Yes vs No)	-.0865306	.0222063	-3.90	0.000	-.1300542	-.043007

- When everyone joins the program instead of when no one participants in the program,
 - On average, the probability of “Loss” goes up by .1
 - On average, the probability of “No change” does not change by much
 - On average, the probability of “Gain” goes down .09

```

. margins r.prog,                               ///
> predict(fix(prog) outlevel("Loss"))         ///
> predict(fix(prog) outlevel("No change"))    ///
> predict(fix(prog) outlevel("Gain"))         ///
> contrast(nowald) vce(unconditional) noesample

```

Contrasts of predictive margins

```

1._predict   : Pr(wchange==Loss), predict(fix(prog) outlevel("Loss"))
2._predict   : Pr(wchange==No change), predict(fix(prog) outlevel("No
change"))
3._predict   : Pr(wchange==Gain), predict(fix(prog) outlevel("Gain"))

```

	Unconditional			
	Contrast	Std. Err.	[95% Conf. Interval]	
prog@_predict				
(Yes vs No) 1	.0998676	.0620691	-.0217856	.2215208
(Yes vs No) 2	-.013337	.0452574	-.1020399	.0753658
(Yes vs No) 3	-.0865306	.0222063	-.1300542	-.043007

Endogenous sample selection

- Reconsider our fictional weight-loss program
 - Some program participants and some nonparticipants will not show up for the final weigh in
This is commonly known as lost to follow up
 - If unobservables that affect whether someone is lost to follow up
 - are independent of the unobservables that affect program participation
 - and they are independent of the unobservables that affect the outcomes with and without the program,
 - the previously discussed estimator consistently estimates the effects
- Any dependence among the unobservables must be modeled

$$insamp = (\mathbf{x}\alpha + \alpha_1 wtsamp + \xi > 0)$$

$$prog = (\mathbf{x}\gamma + \gamma_1 wtprog + \eta > 0)$$

$$wchange = \begin{cases} \text{"Loss"} & \text{if } \mathbf{x}\beta_0 + \epsilon_0 \leq cut1_0 \\ \text{"No change"} & \text{if } cut1_0 < \mathbf{x}\beta_0 + \epsilon_0 \leq cut2_0 \\ \text{"Gain"} & \text{if } cut2_0 < \mathbf{x}\beta_0 + \epsilon_0 \end{cases}$$

$$\mathbf{x}\beta_0 = \beta_{1,0}age + \beta_{2,0}over + \beta_{3,0}phealth$$

for the observations at which prog=0, and

$$wchange = \begin{cases} \text{"Loss"} & \text{if } \mathbf{x}\beta_1 + \epsilon_1 \leq cut1_1 \\ \text{"No change"} & \text{if } cut1_1 < \mathbf{x}\beta_1 + \epsilon_1 \leq cut2_1 \\ \text{"Gain"} & \text{if } cut2_1 < \mathbf{x}\beta_1 + \epsilon_1 \end{cases}$$

$$\mathbf{x}\beta_1 = \beta_{1,1}age + \beta_{2,1}over + \beta_{3,1}phealth$$

for the observations at which prog=1

ξ , ϵ_0 , ϵ_1 , and η are correlated and joint normal

distinct correlations between each treatment error and others


```

. eoprobit wchange age over phealth , ///
>     entreat(prog = age over phealth wtprog, pocorr ) ///
>     select(insamp = age over phealth wtsamp ) ///
>     vce(robust) vsquish nolog

```

```

Extended ordered probit regression           Number of obs   =       3,000
                                           Selected        =       1,884
                                           Nonselected     =       1,116
                                           Wald chi2(6)    =       163.70
                                           Prob > chi2     =       0.0000
Log pseudolikelihood = -4483.9683

```

	Coef.	Robust Std. Err.	z	P> z	[95% Conf. Interval]	
wchange						
prog#c.age						
No	.4174575	.1335097	3.13	0.002	.1557832	.6791318
Yes	-.0779536	.1120819	-0.70	0.487	-.2976301	.141723
prog#c.over						
No	.5046857	.0836683	6.03	0.000	.3406989	.6686725
Yes	.1930521	.0973183	1.98	0.047	.0023118	.3837924
prog#						
c.phealth						
No	-.4250361	.091857	-4.63	0.000	-.6050726	-.2449996
Yes	-.8098627	.0753678	-10.75	0.000	-.9575809	-.6621444

insamp						
age	-.0231005	.0805424	-0.29	0.774	-.1809607	.1347597
over	-.7639994	.0450909	-16.94	0.000	-.852376	-.6756229
phealth	.7765721	.0467569	16.61	0.000	.6849303	.8682139
wtsamp	2.611108	.2660121	9.82	0.000	2.089734	3.132483
cons	.2832551	.0516926	5.48	0.000	.1819395	.3845707

c.phealth							
No		-.4250361	.091857	-4.63	0.000	-.6050726	-.2449996
Yes		-.8098627	.0753678	-10.75	0.000	-.9575809	-.6621444
insamp							
age		-.0231005	.0805424	-0.29	0.774	-.1809607	.1347597
over		-.7639994	.0450909	-16.94	0.000	-.852376	-.6756229
phealth		.7765721	.0467569	16.61	0.000	.6849303	.8682139
wtsamp		2.611108	.2660121	9.82	0.000	2.089734	3.132483
_cons		.2832551	.0516926	5.48	0.000	.1819395	.3845707
prog							
age		-.9371024	.0818803	-11.44	0.000	-1.097585	-.7766199
over		-1.060975	.0492229	-21.55	0.000	-1.15745	-.9645
phealth		.890558	.0494954	17.99	0.000	.7935487	.9875673
wtprog		1.644504	.0731516	22.48	0.000	1.501129	1.787878
_cons		.0153225	.0527572	0.29	0.771	-.0880796	.1187247
/wchange							
prog#c.cut1							
No		-.2754667	.1708586			-.6103433	.05941
Yes		-.4323606	.1401249			-.7070003	-.1577208
prog#c.cut2							
No		.6797857	.1534354			.3790578	.9805137
Yes		.7803365	.2260056			.3373737	1.223299
corr(e.ins~p, e.wchange)							
prog							
No		-.5779184	.1004465	-5.75	0.000	-.7420068	-.3484981
Yes		-.5355424	.1948537	-2.75	0.006	-.81217	-.0623165
corr(e.prog, e.wchange)							

prog#c.cut1						
No	-.2754667	.1708586			-.6103433	.05941
Yes	-.4323606	.1401249			-.7070003	-.1577208
prog#c.cut2						
No	.6797857	.1534354			.3790578	.9805137
Yes	.7803365	.2260056			.3373737	1.223299
corr(e.ins~p, e.wchange)						
prog						
No	-.5779184	.1004465	-5.75	0.000	-.7420068	-.3484981
Yes	-.5355424	.1948537	-2.75	0.006	-.81217	-.0623165
corr(e.prog, e.wchange)						
prog						
No	-.6031412	.1119322	-5.39	0.000	-.7790275	-.3392526
Yes	-.4940044	.0934446	-5.29	0.000	-.6547774	-.2904625
corr(e.prog, e.insamp)	.4745668	.0298397	15.90	0.000	.4140283	.5309257

- Nonzero correlations between e.insamp and e.wchange imply endogenous sample selection for outcomes
- Nonzero correlations between e.prog and e.wchange imply endogenous treatment assignment

```
. estat teffects
```

```
Predictive margins
```

```
Number of obs      =      3,000
```

```
ATE_Pr0      : Pr(wchange=0=Loss)  
ATE_Pr1      : Pr(wchange=1=No change)  
ATE_Pr2      : Pr(wchange=2=Gain)
```

	Unconditional Margin	Std. Err.	z	P> z	[95% Conf. Interval]	
ATE_Pr0 prog (Yes vs No)	.1406344	.0785061	1.79	0.073	-.0132346	.2945035
ATE_Pr1 prog (Yes vs No)	.0210902	.0369635	0.57	0.568	-.0513569	.0935372
ATE_Pr2 prog (Yes vs No)	-.1617246	.0642328	-2.52	0.012	-.2876187	-.0358305

- When everyone joins the program instead of when no one participants in the program,
 - On average, the probability of “Loss” goes up by .14
 - On average, the probability of “No change” does not change
 - On average, the probability of “Gain” goes down .16

```

. margins r.prog, ///
> predict(fix(prog) outlevel("Loss")) ///
> predict(fix(prog) outlevel("No change")) ///
> predict(fix(prog) outlevel("Gain")) ///
> contrast(nowald) vce(unconditional) noesample

```

Contrasts of predictive margins

```

1._predict : Pr(wchange==Loss), predict(fix(prog) outlevel("Loss"))
2._predict : Pr(wchange==No change), predict(fix(prog) outlevel("No
change"))
3._predict : Pr(wchange==Gain), predict(fix(prog) outlevel("Gain"))

```

	Unconditional			
	Contrast	Std. Err.	[95% Conf. Interval]	
prog@_predict				
(Yes vs No) 1	.1406344	.0785061	-.0132346	.2945035
(Yes vs No) 2	.0210902	.0369635	-.0513569	.0935372
(Yes vs No) 3	-.1617246	.0642328	-.2876187	-.0358305

- When everyone joins the program instead of when no one participants in the program,
 - On average, the probability of “Loss” goes up by .14
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More about ERM commands

- The commands `eregress`, `eprobit`, and `eintreg` fit ERMs handle continuous-and-unbounded, binary, and censored/corner outcomes
- Look at

<http://www.stata.com/manuals/erm.pdf>

for more examples and a wealth of details