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Introduction to Bayesian Analysis in Stata

Gustavo Sánchez

StataCorp LLC

October 24 , 2018 Barcelona, Spain

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Bayesian Analysis vs Frequentist Analysis

Frequentist Analysis

- Estimate unknown fixed parameters.
- Data for a (hypothetical) repeatable random sample.
- Uses data to estimate unknown fixed parameters.
- Data expected to satisfy the assumptions for the specified model.

"Conclusions are based on the distribution of statistics derived from random samples, assuming unknown but fixed parameters." **Bayesian Analyis**

- Probability distributions for unknown random parameters
- The data is fixed.
- Combines data with prior beliefs to get probability distributions for the parameters.
- Posterior distribution is used to make explicit probabilistic statements.

"Bayesian analysis answers questions based on the distribution of parameters conditional on the observed sample."

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Stata's simple syntax: bayes:

regress y x1 x2 x3

bayes: regress y x1 x2 x3

logit y x1 x2 x3

bayes: logit y x1 x2 x3

mixed y x1 x2 x3 || region: bayes: mixed y x1 x2 x3 || region:

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• Inverse law of probability (Bayes' Theorem):

$$f(\theta|y) = \frac{f(y; \theta) \pi(\theta)}{f(y)}$$

- Marginal distribution of y, f(y), does not depend on (θ)
- We can then write the fundamental equation for Bayesian analysis:

 $\boldsymbol{\rho}(\theta|\boldsymbol{y}) \propto \boldsymbol{L}(\boldsymbol{y}|\theta) \pi(\theta)$

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- Let's assume that both, the data and the prior beliefs, are normally distributed:
 - The data: $y \sim N(\theta, \sigma_d^2)$

• The prior:
$$\theta \sim N\left(\mu_p, \sigma_p^2\right)$$

- Homework...: Doing the algebra with the fundamental equation we find that the posterior distribution would be normal with (see for example Cameron & Trivedi 2005):
 - The posterior: $\theta | \mathbf{y} \sim \mathbf{N} \left(\mu, \sigma^2 \right)$

Where:

$$\mu = \sigma^2 \left(N \bar{y} / \sigma_d^2 + \mu_p / \sigma_p^2 \right)$$

$$\sigma^2 = \left(N / \sigma_d^2 + 1 / \sigma_p^2 \right)^{-1}$$

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Example (Posterior distributions)

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100

150

Number of days



200

250

300

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- The Method
 - The previous example has a closed form solution.
 - What about the cases with non-closed solutions, or more complex distributions?
 - Integration is performed via simulation
 - We need to use intensive computational simulation tools to find the posterior distribution in most cases.
 - Markov chain Monte Carlo (MCMC) methods are the current standard in most software. Stata implement two alternatives:

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- Metropolis-Hastings (MH) algorithm
- Gibbs sampling

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- Links for Bayesian analysis and MCMC on our youtube channel:
 - Introduction to Bayesian statistics, part 1: The basic concepts

https://www.youtube.com/watch?v=0F0QoMCSKJ4&feature=youtu.be

• Introduction to Bayesian statistics, part 2: MCMC and the Metropolis Hastings algorithm.

https://www.youtube.com/watch?v=OTO1DygELpY&feature=youtu.be

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Monte Carlo Simulation



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Markov Chain Monte Carlo Simulation



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- Metropolis Hastings intuitive idea
 - Green points represent accepted proposal states and red points represent rejected proposal states.



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- Metropolis Hastings simulation
 - The trace plot illustrates the sequence of accepted proposal states.



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We expect to obtain a stationary sequence when convergence is achieved.



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- An efficient MCMC should have small autocorrelation.
- We expect autocorrelation to become negligible after a few lags.



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The Stata tools: bayesmh & bayes:

- bayesmh General purpose command for Bayesian analysis
 - You need to specify all the components for the Bayesian regression: Likelihood, priors, hyperpriors, blocks, etc
- bayes: Simple syntax for Bayesian regressions
 - Estimation command defines the likelihood for the model.
 - Default priors are assumed to be "noninformative"'.
 - Other model specifications are set by default depending on the model defined by the estimation command.
 - Alternative specifications may need to be evaluated.

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Example 1: Linear Regression

- Let's look at our first example:
 - We have stats on the average number of days tourists spend in Cataluña and their average per capita expenditure.
 - We fit a linear regression for the average number of days.
 - Let's consider two specifications:

tripdays = $\alpha_1 + \beta_{day} * capexp_day + \epsilon_1$ tripdays = $\alpha_2 + \beta_{avg} * avgexp_cap + \epsilon_2$

Where:

tripdays : Number of days tourists spend in Cataluña. capexp_day: Tourists' daily per capita expenditure. avgexp_cap: Tourists' total per capita expenditure.

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Example 1: Linear Regression

• Linear regression with the bayes: prefix

bayes ,rseed(123): regress tripdays capex_day

• Equivalent model with <code>bayesmh</code>

| bayesmh tripdays capexp_day, rseed(123) | |
|--|--|
| likelihood(normal(sigma2)) | |
| <pre>prior(tripdays:capexp_day, normal(0,10000))</pre> | |
| <pre>prior(tripdays:_cons, normal(0,10000))</pre> | |
| prior(sigma2, igamma(.01,.01)) | |
| <pre>block(tripdays:capexp_day _cons)</pre> | |
| block(sigma2) | |

Example 1: Menu for Bayesian regression

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| | SEM (structural equation modeling) LCA (latent class analysis) FMM (finite mixture models) IRT (item response theory) Survey data analysis | Ъ Ъ Ъ | Continuous outcomes Binary outcomes Ordinal outcomes Categorical outcomes | |
| Command | Multiple imputation Nonparametric analysis Multivariate analysis Exact statistics Resampling Power and sample size Bayesian analysis Postestimation | Regression models General estimation and regression Graphical summaries Effective sample sizes Summary statistics Information orienta Henderbeits testina usine model costerior probability | Court outcomes Fractional outcomes Generalized linear model (GLM) Survival models Selection models Censored and funcated models Zaro-inflation court models ties Multiveer models | |
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Censored and truncated models

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- bayes: bayesmi Postestimation

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1- Linear regression

- bayesstats ess bayesgraph thinning() bayestestmodel
- 2- Random effects probit bayesgraph bayestest interval
- 3- Change point model Gibbs sampling
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Example 1: Menu for Bayesian regression

- Make the following sequence of selection from the main menu:
 - Statistics > Bayesian analysis > Regression models

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- 2 Select 'Continuous outcomes'
- 3 Select 'Linear regression'
- 4 Click on 'Launch'
- Specify the dependent variable (tripdays) and the explanatory variable (capex_day)
- 6 Click on 'OK'

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Example 1: bayes: prefix

. bayes ,rseed(123) blocksummary:regress tripdays capexp_day

Burn-in ... Simulation ... Model summary

Likelihood: tripdays ~ regress(xb_tripdays,{sigma2}) Priors: {tripdays:capexp_day _cons} ~ normal(0,10000) {sigma2} ~ igamma(.01,.01)

(1) Parameters are elements of the linear form xb_tripdays. Block summary

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```
1: {tripdays:capexp_day _cons}
2: {sigma2}
```

Fundamental MCMC

bayes: - bayesmh

1- Linear regression

bayestestmodel

bayesgraph

Example 1: bayes: prefix

. bayes , rseed(123) blocksummary:regress tripdays capexp_day

| Bayesian linear regression | MCMC iterations = | 12,500 |
|--|--------------------|--------|
| Random-walk Metropolis-Hastings sampling | Burn-in = | 2,500 |
| | MCMC sample size = | 10,000 |
| | Number of obs = | 5 |
| | Acceptance rate = | . 3799 |
| | Efficiency: min = | .03477 |
| | avg = | .08801 |
| Log marginal likelihood = -16.207649 | max = | .1146 |
| | | |

| | | | | | Equal-tailed | | |
|------------|----------|-----------|---------|----------|--------------|-----------|--|
| | Mean | Std. Dev. | MCSE | Median | [95% Cred. | Interval] | |
| tripdays | | | | | | | |
| capexp_day | 0383973 | .0128253 | .000379 | 0377857 | 0647811 | 0122725 | |
| _cons | 12.64544 | 2.484331 | .073378 | 12.52498 | 7.610854 | 17.84337 | |
| sigma2 | .0926729 | .0928459 | .004979 | .0616775 | .0151486 | .3563017 | |

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Note: Default priors are used for model parameters.

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Example 1: bayesstats ess

· Let's evaluate the effective sample size

. bayesstats ess

| Efficiency | summaries | MCMC | sample | size | = | 10,000 |
|------------|-----------|------|--------|------|---|--------|
|------------|-----------|------|--------|------|---|--------|

| | ESS | Corr. time | Efficiency | |
|------------|---------|------------|------------|--|
| ripdavs | | | | |
| capexp_day | 1146.26 | 8.72 | 0.1146 | |
| _cons | 1146.27 | 8.72 | 0.1146 | |
| sigma2 | 347.72 | 28.76 | 0.0348 | |

- We expect to have an acceptance rate (see previous slide) that is neither to small nor too large.
- We also expect to have low correlation
- Efficiencies over 10% are considered good for MH. Efficiencies under 1% would be a source of concern.

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Example 1: bayesgraph

- We can use <code>bayesgraph</code> to look at the trace, the correlation, and the density. For example:
 - . bayesgraph diagnostic {capex_day}



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- · The trace indicates that convergence was achieved
- Correlation becomes negligible after 10 periods

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Example 1: bayesgraph

• We can use <code>bayesgraph</code> to look at the trace, the correlation, and the density. For example:

. bayesgraph diagnostic {sigma2}



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· Correlation is still persistent after 10 periods

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Example 1: thinning()

- We can reduce autocorrelation by using thinning
- Save the random draws skipping a prespecified number of simulated values in the MCMC iteration process.
- Use the option 'thinning(#)' to indicate that Stata should save simulated values from every (1+k*#)th iteration (k=0,1,2,...).

bayes ,nomodelsummary nodots rseed(123) /// thinning(4): regress tripdays capexp_day

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Example 1: thining()

. bayes ,rseed(123) nomodelsummary thinning(4): ///
> regress tripdays capexp_day

note: discarding every 3 sample observations; using observations 1,5,9,...

Burn-in ... Simulation ...

Bayesian linear regression Random-walk Metropolis-Hastings sampling

| Loa | marginal | likelihood | = | -16.191209 |
|-----|----------|------------|---|------------|
| | | | | |

| | | | | Equal- | tailed | |
|------------|----------|-----------|---------|----------|------------|-----------|
| | Mean | Std. Dev. | MCSE | Median | [95% Cred. | Interval] |
| tripdays | | | | | | |
| capexp_day | 0384152 | .0126655 | .000196 | 0383658 | 0636837 | 0126029 |
| _cons | 12.64972 | 2.455834 | .037984 | 12.62602 | 7.628034 | 17.57862 |
| sigma2 | .0917518 | .0951007 | .002932 | .0605151 | .0151486 | .3519349 |

Note: Default priors are used for model parameters.

MCMC iterations

Acceptance rate

Number of obs

Efficiency:

MCMC sample size =

Burn-in

42,497

10,000

2,500

.3773

.1052

.313

.418

5

=

=

=

=

min =

avg =

max =

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Example 1: bayesstats ess

• Let's evaluate again the effective sample size

. bayesstats ess

Efficiency summaries MCMC sample size = 10,000

| | ESS | Corr. time | Efficiency |
|------------|---------|------------|------------|
| tripdays | | | |
| capexp_day | 4159.44 | 2.40 | 0.4159 |
| _cons | 4180.27 | 2.39 | 0.4180 |
| sigma2 | 1051.71 | 9.51 | 0.1052 |

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- The efficiency improved for all the parameters.
- Correlation time was significantly reduced.

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Example 1: bayestest model

- bayestest model is another postestimation command to compare different models.
- bayestest model computes the posterior probabilities for each model.
- The result indicates which model is more likely.
- It requires that the models use the same data and that they have proper posterior.
- It can be used to compare models with:
 - Different priors and/or different posterior distributions.
 - Different regression functions.
 - Different covariates
- MCMC convergence should be verified before comparing the models.

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Example 1: bayestest model

- Let's fit now two other models and compare them to the one we already fitted.
- We store the results for the three models and we use the postestimation command <code>bayestest model</code> to select one of them.

quietly {

bayes , rseed(123) saving(pcap,replace): /// regress tripdays capexp_day estimates store daily

bayes , rseed(123) saving(total,replace): /// regress tripdays avgexp_cap estimates store total

bayes , rseed(123) saving(media,replace) /// prior(tripdays:_cons, normal(9,.4)): /// regress tripdays estimates store mean

bayestest model daily total mean

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Example 1: bayestest model

- Here is the output for bayestest model
 - . quietly {
 - . bayestest model daily total mean

Bayesian model tests

| | log(ML) | P (M) | P (M y) |
|---------------|----------------------|------------------|------------------|
| daily | -16.2076 | 0.3333 | 0.4997 |
| total mean | -18.6705 -16.2955 | 0.3333 0.3333 | 0.0426 0.4577 |

Note: Marginal likelihood (ML) is computed using Laplace-Metropolis approximation.

• We could also assign different priors for the models:

```
. bayestest model daily total mean, //,
> prior(.15 0.75 0.1)
```

Bayesian model tests

| log(ML) | P (M) | P(M y) |
|----------|-------|--------|
| -16.2076 | | 0.4910 |
| -18.6705 | | 0.2092 |
| -16.2955 | | 0.2998 |

Note: Marginal likelihood (ML) is computed using Laplace-Metropolis approximation.

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Example 1: bayestest model

- Here is the output for bayestest model
 - . quietly {
 - . bayestest model daily total mean

Bayesian model tests

| | log(ML) | P (M) | P (M y) |
|-------|----------|--------|---------|
| daily | -16.2076 | 0.3333 | 0.4997 |
| total | -18.6705 | 0.3333 | 0.0426 |
| mean | -16.2955 | 0.3333 | 0.4577 |

Note: Marginal likelihood (ML) is computed using Laplace-Metropolis approximation.

· We could also assign different priors for the models:

```
. bayestest model daily total mean, ///
> prior(.15 0.75 0.1)
```

Bayesian model tests

| | log(ML) | P (M) | P(M y) |
|---------------|----------------------|------------------|------------------|
| daily | -16.2076 | 0.1500 | 0.4910 |
| total mean | -18.6705 -16.2955 | 0.7500 0.1000 | 0.2092 0.2998 |

Note: Marginal likelihood (ML) is computed using Laplace-Metropolis approximation.

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Example 2: Random Effects Probit model

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Example 2: Random effects probit model

• Let's use bayes: to fit a random effects for a binary variable, whose values depend on a linear latent variable.

$$y_{it} = \beta_0 + \beta_1 x \mathbf{1}_{it} + \beta_2 x \mathbf{2}_{it} + \dots + \beta_k x k_{it} + \alpha_i + \epsilon_{it}$$

Where:

$$y_{it} = \begin{cases} 1 & \text{if } y_{it} * > 0 \\ 0 & \text{otherwise} \end{cases}$$

 $\alpha_i \sim N(0, \sigma_{\alpha}^2)$ is the individual random panel effect $\epsilon_{it} \sim N(0, \sigma_e^2)$ is the idiosyncratic error term

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- This is also referred as a two-level random intercept model.
- We can also fit this model with meprobit or xtprobit, re

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Example 2: Random effects probit model

- This time we are going to work with simulated data.
- Here is the code to simulate the panel dataset:

```
clear
set obs 100
set seed 1
* Panel level *
generate id=_n
generate alpha=rnormal()
expand 5
* Observation level *
bysort id:generate year=_n
xtset id year
generate x1=rnormal()
generate x2=runiform()>.5
```

```
generate x3=uniform()
generate u=rnormal()
```

* Generate dependent variable *
 generate y=.5+1*x1+(-1)*x2+1*x3+alpha+u>0

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Example 2: Random effects probit model

Let's show the results with meprobit:

| x2 x3 cons | .9426958 .5220418 | .2187448 | 0.017 | .0933098 | .9507738 |
|--------------------|----------------------|-----------|---------|-------------|-----------|
| x2 x3 cons | .9426958 | 2187448 | 0 017 | 0933098 | 9507738 |
| x2 x3 | . 9426958 | | | | |
| x 2 | | 2941061 | 0.001 | 3662584 | 1.519133 |
| | 9896286 | .1853433 | 0.000 | -1.352895 | 6263625 |
| x 1 | .9769118 | .1143889 | 0.000 | .7527138 | 1.20111 |
| У | Coef. | Std. Err. | P> z | [95% Conf. | Interval] |
| Log likelihood | 1 = -236.8858 | 9 | Prob > | chi2 = | 0.0000 |
| | | | Weldeh | +2 (2) - | 00.00 |
| Integration me | thod: mvaghe: | rmite | Integra | tion pts. = | 7 |
| | | | | max = | 5 |
| | | | | avg = | 5.0 |
| | | | - | min = | 5 |
| | | | Obs per | group: | |
| Group variable: id | | id | Number | of groups = | 100 |
| Group wariable | probit regres | ssion | Number | of obs = | 500 |
| Mixed-effects | | | | | |

LR test vs. probit model: chibar2(01) = 67.24 Prob >= chibar2 = 0.0000

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Example 2: Random effects probit model

• We now fit the model with bayes:

bayes , nodots rseed(123) thinning(5) blocksummary: /// meprobit y x1 x2 x3 || id:

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• Equivalent model with bayesmh

| bayesmh y x1 x2 x3, thinning(5) rseed(123) | |
|--|-----|
| likelihood(probit) | |
| prior(y:i.id, normal(0,y:var)) | /// |
| prior(y:x1 x2 x3 _cons, normal(0,10000)) | /// |
| prior(y:var, igamma(.01,.01)) | /// |
| block(y:var) | /// |
| blocksummary dots | |

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Example 2: Random effects probit model

```
. bayes , nodots rseed(123) thinning(5) blocksummary:
    meprobit y x1 x2 x3 || id:
```

note: discarding every 4 sample observations; using observations 1,6,11,... Burn-in ... Simulation ...

Multilevel structure

id

{U0}: random intercepts

Model summary

Likelihood:

y _ meprobit(xb_y)

Priors:

```
{y:x1 x2 x3 _cons} _ normal(0,10000) (1)
{U0} _ normal(0,{U0:sigma2}) (1)
```

Hyperprior:

{U0:sigma2} _ igamma(.01,.01)

(1) Parameters are elements of the linear form xb_y.

Block summary

1: {y:x1 x2 x3 _cons} 2: {U0:sigma2} 3: {U0[id]:1 2 3 4 5 6 7 8 9 10 11 12 13 14 15 16 17 18 19 20 21 22 23 24 > 25 26 27 28 29 30 31 32 33 34 35 36 37 38 39 40 41 42 43 44 45 46 47 48 49 50 > 51 52 53 54 55 56 57 58 59 60 61 62 63 64 65 66 67 68 69 70 71 72 73 74 75 76 > 77 78 79 80 81 82 83 84 85 86 87 88 89 90 91 92 93 94 95 96 97 98 99 100}

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Example 2: Random effects probit model

. bayes , nodots rseed(123) thinning(5) blocksummary: meprobit y x1 x2 x3 || id:

| Bayesian multilevel probit regression Random-walk Metropolis-Hastings sampling Group variable: id | MCMC iterations Burn-in MCMC sample size Number of groups | = = = | 52,496 2,500 10,000 100 |
|---|--|-------------|----------------------------------|
| | Obs per group: | | |
| | min | = | 5 |
| | avg | = | 5.0 |
| | max | = | 5 |
| Family : Bernoulli | Number of obs | = | 500 |
| Link : probit | Acceptance rate | = | .3268 |
| | Efficiency: min | = | .05399 |
| | avg | = | .102 |
| Log marginal likelihood | max | = | .1628 |

| | | | | | | Equal- | tailed |
|----|------------|-----------|-----------|---------|-----------|------------|-----------|
| | | Mean | Std. Dev. | MCSE | Median | [95% Cred. | Interval] |
| У | | | | | | | |
| | x 1 | . 9977099 | .1181726 | .003773 | .9936143 | .7810441 | 1.242439 |
| | x 2 | -1.018063 | .1892596 | .00557 | -1.012598 | -1.396798 | 6509636 |
| | x 3 | .9539304 | .2936949 | .007279 | .9514395 | .3823801 | 1.52913 |
| | _cons | . 5433822 | .2205077 | .00949 | . 5398387 | .1216346 | .9847166 |
| id | | | | | | | |
| | U0:sigma2 | 1.456558 | .4384163 | .015537 | 1.401461 | .7611919 | 2.463175 |

Note: Default priors are used for model parameters.

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Example 2: bayesgraph diagnostic

• We can look at the diagnostic graph for a couple of variables:

. bayesgraph diagnostic {y:x1}



- The trace seems to indicate convergence this time.
- Autocorrelation decays quicker and becomes negligible after about 15 periods.

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Example 2: bayesgraph diagnostic

• We now look now at the diagnostic graphs for {U0:sigma2}

. bayesgraph diagnostic U0:sigma2



- The trace seems to indicate convergence this time.
- Autocorrelation decays quicker and becomes negligible after about 15 periods.

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Example 2: bayestest interval

- We can perform interval testing with the postestimation command bayestest interval.
- It estimates the probability that a model parameter lies in a particular interval.
- For continuous parameters the hypothesis is formulated in terms of intervals.
- We can perform point hypothesis testing only for parameters with discrete posterior distributions.
- bayestest interval estimates the posterior distribution for a null interval hypothesis.
- bayestest interval reports the estimated posterior mean probability for Ho.

bayestest interval ({y:x1},lower(.9) upper(1.02)) /// ({y:x2},lower(-1.1) upper(-.8))

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Example 2: bayestest interval

• We can, for example, perform separate tests for different parameters:

| I | Mean | Std. Dev. | MCSE |
|-------|-------|-----------|----------|
| prob1 | .3888 | 0.48750 | .0077073 |
| prob2 | .5474 | 0.49777 | .0097517 |

We can also perform a joint test:

| . bayeste | st inter | <pre>val (({y:x1},</pre> | lower(.9) | upper(1.02)) /// | / |
|-----------|----------|--------------------------|-------------|------------------|---|
| > | | ({y:x2},lower | (-1.1) up | per(8)),joint) | |
| Interval | tests | MCMC sample | size = | 10,000 | |
| prob1 | : .9 < { | y:x1 < 1.02, | $-1.1 < \{$ | y:x2} < −.8 | |

| Mean | Std. Dev. | MCSE |
|------|-----------|----------|
| | 0.41754 | .0066399 |

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Example 2: bayestest interval

• We can, for example, perform separate tests for different parameters:

| | Mean | Std. Dev. | MCSE |
|-------|--------|-----------|----------|
| prob1 | . 3888 | 0.48750 | .0077073 |
| prob2 | . 5474 | 0.49777 | .0097517 |

We can also perform a joint test:

| . bayest | est inter | <pre>rval (({y:x1}</pre> | },lower(.9) | upper(1.02)) | /// |
|----------|-----------|--------------------------|---------------|-----------------------------|-----|
| > | | ({y:x2},lowe | er(-1.1) up | per(8)), <mark>joi</mark> : | nt) |
| Interval | tests | MCMC sampl | le size = | 10,000 | |
| prob1 | : .9 < { | $\{y:x1\} < 1.02$ | $2, -1.1 < {$ | y:x2} < −.8 | |

| | | Mean | Std. Dev. | MCSE |
|-------|---|--------|-----------|----------|
| prob1 | I | . 2249 | 0.41754 | .0066399 |

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Example 3: Change-point model

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Example 3: Change-point model

- Let's work now with an example where we write our model using a substitutable expression.
- We have yearly data on fertility for Spain:



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- The series has a significant change around 1980.
- We may consider fitting a change-point model.

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Example 3: Gibbs sampling

Change point model specification with blocking

| payesmh fertil = ({mu1}*sign(year<{cp}) | |
|---|-----|
| + {mu2}*sign(year>={cp})), | |
| likelihood(normal({var})) | /// |
| <pre>prior({mu1}, normal(1,5))</pre> | /// |
| prior({mu2}, normal(5,5)) | /// |
| prior({cp}, uniform(1960,2015)) | /// |
| prior({var}, igamma(2,1)) | /// |
| initial({mu1} 5 {mu2} 1 {cp} 1960) | /// |
| | |
| rseed(123) mcmcsize(40000) | |
| dots(500,every(5000)) | |
| title(Change-point analysis) | |

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Example 3: Gibbs sampling

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|--|-----|
| + {mu2}*sign(year>={cp})), | /// |
| likelihood(normal({var})) | /// |
| prior({mu1}, normal(1,5)) | /// |
| prior({mu2}, normal(5,5)) | /// |
| prior({cp}, uniform(1960,2015)) | /// |
| prior({var}, igamma(2,1)) | /// |
| initial({mu1} 5 {mu2} 1 {cp} 1960) | /// |
| block(var, gibbs) block(cp) blocksummary | /// |
| rseed(123) mcmcsize(40000) | /// |
| dots(500,every(5000)) | /// |
| title(Change-point analysis) | |

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Example 3: Gibbs sampling

Change point model specification with blocking

| | bayesmh | <pre>fertil=({mu1}*sign(year<{cp})+{mu2}*sign(year>={cp})), //</pre> | | | | |
|---|---------|--|-----|--|--|--|
| > | | <pre>likelihood(normal({var}))</pre> | 111 | | | |
| > | | <pre>prior({mu1}, normal(0,5))</pre> | 111 | | | |
| > | | <pre>prior({mu2}, normal(5,5))</pre> | 111 | | | |
| > | | <pre>prior({cp}, uniform(1960,2015))</pre> | 111 | | | |
| > | | <pre>prior({var}, igamma(2,1))</pre> | 111 | | | |
| > | | initial({mu1} 5 {mu2} 1 {cp} 1960) | 111 | | | |
| > | | block(var, gibbs) block(cp) blocksummary | 111 | | | |
| > | | <pre>rseed(123) mcmcsize(40000) dots(500, every(5000))</pre> | 111 | | | |
| > | | title (Modelo de Cambio de Punto) | | | | |

Burn-in 2500 aaaaa done

| Si | mulation | 40000 | 10000 | 15000 | 20000 |
|----|----------|-------|-----------|-----------|-------|
| > | | 25000 | 35000 | 40000 doi | ıe |

Model summary

Likelihood:

```
fertility ~ normal({mul}*sign(year<{cp})+{mu2}*sign(year>={cp}), {var})
Priors:
{var} ~ igamma(2,1)
{mul} ~ normal(0,5)
{mul} ~ normal(5,5)
{cp} ~ uniform(1960,2015)
```

Block summary

| 1: | {var} | (Gibbs) |
|----|-------------|---------|
| 2: | {cp} | |
| 3: | {mu1} {mu2} | |

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Example 3: Gibbs sampling

Change point model specification with blocking

| . bayesmh | <pre>fertil=({mu1}*sign(year<{cp})+{mu2}*sign</pre> | gn(year>={cp})), | 111 | |
|-----------|--|------------------|-----|--------|
| > | <pre>likelihood(normal({var}))</pre> | | 111 | |
| > | <pre>prior({mu1}, normal(0,5))</pre> | | 111 | |
| > | <pre>prior({mu2}, normal(5,5))</pre> | | 111 | |
| > | <pre>prior({cp}, uniform(1960,2015))</pre> | | /// | |
| > | <pre>prior({var}, igamma(2,1))</pre> | | 111 | |
| > | initial({mu1} 5 {mu2} 1 {cp} 1960) | | 111 | |
| > | <pre>block(var, gibbs) block(cp) blocksumma</pre> | ry | 111 | |
| > | rseed(123) mcmcsize(40000) dots(500, e | very(5000)) | 111 | |
| > | title(Modelo de Cambio de Punto) | | | |
| Modelo de | Cambio de Punto | MCMC iterations | = | 42,500 |
| Metropoli | s-Hastings and Gibbs sampling | Burn-in | = | 2,500 |
| | | MCMC sample size | | 40,000 |
| | | Number of obs | = | 56 |
| | | Acceptance rate | = | .5704 |
| | | Efficiency: min | . = | .08572 |
| | | avg | r = | .2629 |
| Log margi | nal likelihood = -16.240692 | max | . = | .7203 |

| | | | | | Equal-tailed | | |
|-----|----------|-----------|---------|----------|--------------|-----------|--|
| | Mean | Std. Dev. | MCSE | Median | [95% Cred. | Interval] | |
| ср | 1980.87 | .7407595 | .010454 | 1980.772 | 1979.439 | 1982.517 | |
| mu1 | 2.771024 | .0654542 | .001118 | 2.770196 | 2.64247 | 2.897339 | |
| mu2 | 1.376056 | .0489823 | .000706 | 1.375648 | 1.281815 | 1.472107 | |
| var | .078699 | .0152773 | .00009 | .0768054 | .0541305 | .1136579 | |

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Example 3: bayesgraph trace

Use bayesgraph trace to look at the trace for all the parameters.

. bayesgraph trace



The plots indicate that convergence seems to be achieved.

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Example 3: bayesgraph ac

Use bayesgraph ac to look at the autocorrelation for all the parameters.

. bayesgraph ac



 Autocorrelation decays and becomes negligible quickly for almost all the parameters.

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Summing up

- Bayesian analysis: An statistical approach that can be used to answer questions about unknown parameters in terms of probability statements.
- It can be used when we have prior information on the distribution of the parameters involved in the model.
- Alternative approach or complementary approach to classic/frequentist approach?

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