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# Introduction to Bayesian Analysis in Stata

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Porto, Portugal



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    - Stata 14: The `bayesmh` command
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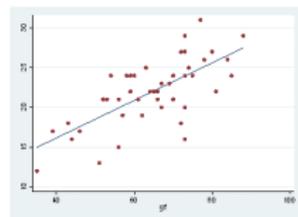
References

# Frequentist

. list in\_wage union hours who\_work tenure race grade cpi\_1mp in 1/15, noobs

in_wage	union	hours	who_work	tenure	race	grade	cpi_1mp
1.451214	-	20	27	.0833333	black	12	1.083333
1.12882	-	44	18	.0833333	black	12	1.271645
1.589977	1	65	51	.0848887	black	12	2.216451
1.198273	-	40	3	.0833333	black	12	2.384182
1.779512	-	10	24	.1088887	black	12	2.771645
1.778481	0	32	52	1.5	black	12	2.771645
2.493974	-	32	4	.0833333	black	12	3.861784
2.551715	1	65	75	1.833333	black	12	5.298872
2.42281	1	49	101	.0888887	black	12	5.298872
2.614312	1	42	91	3.916667	black	12	7.183256
2.536374	1	45	95	3.916667	black	12	8.98718
2.662927	1	49	79	5.933333	black	12	10.13331
1.180748	0	40	13	-.75	black	12	-7.151384
1.204598	-	40	22	2	black	12	1.188815
1.588863	-	40	11	.5833333	black	12	3.488338

Theoretical Model



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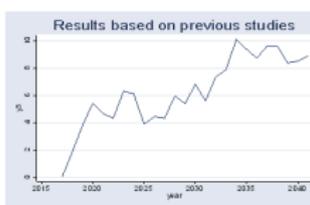
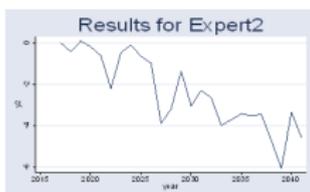
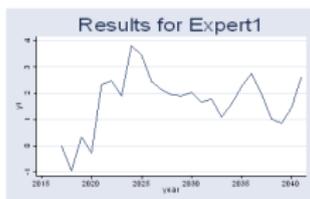
# Bayesian



```

. list _id_>age when born his_mach house rank grade his_esp in 1/10, mode
    
```

_id_>	age	when	born	his_mach	house	rank	grade	his_esp
1.	102219	-	22	27	.000000	0	1	1.000000
2.	102862	-	60	10	.000000	0	1	1.000000
3.	102877	1	62	81	.000000	0	1	1.000000
4.	102919	-	62	9	.000000	0	1	1.000000
5.	102919	-	10	24	.000000	0	1	1.000000
6.	102961	0	68	68	0	0	1	1.000000
7.	102978	-	68	4	.000000	0	1	1.000000
8.	103157	1	65	10	1.000000	0	1	1.000000
9.	103261	1	68	101	.000000	0	1	1.000000
10.	103312	1	65	87	1.000000	0	1	1.000000
11.	103374	1	65	80	1.000000	0	1	1.000000
12.	103549	1	68	10	1.000000	0	1	1.000000
13.	103549	0	65	13	0	0	1	1.000000
14.	103549	-	62	58	0	0	1	1.000000
15.	103993	-	62	17	.000000	0	1	1.000000



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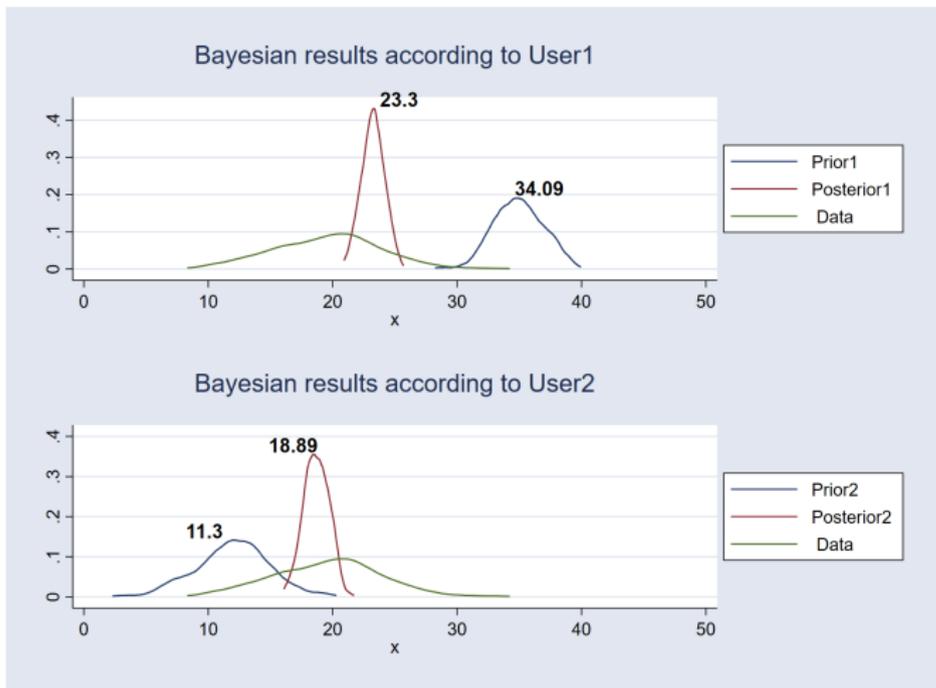
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# Bayesian Analysis vs Frequentist Analysis

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## Frequentist Analysis

- Results are based on estimations for unknown fixed parameters.
- The data are considered to be a (hypothetical) repeatable random sample.
- Uses the data to obtain estimates about the unknown fixed parameters.
- Depends on whether the data satisfies the assumptions for the specified model.

"Frequentists base their conclusions on the distribution of statistics derived from random samples, assuming that the parameters are unknown but fixed."

## Bayesian Analysis

- Results are based on probability distributions about unknown random parameters
- The data are considered to be fixed.
- The results are produced by combining the data with prior beliefs about the parameters.
- The posterior distribution is used to make explicit probabilistic statements

"Bayesian analysis answers questions based on the distribution of parameters conditional on the observed sample."

## Some Advantages

- Based on the Bayes rule, which applies to all parametric models.
- Inference is exact, estimation and prediction are based on posterior distribution.
- Provides more intuitive interpretation in terms of probabilities (e.g Credible intervals).
- It is not limited by the sample size.

## Some Disadvantages

- Subjectivity in specifying prior beliefs.
- Computationally challenging.
- Setting up a model and performing analysis could be an involving task.

## Some Examples (Taken from Hahn, 2014)

- TranScan Medical use small dataset and priors based on previous studies to determine the efficacy of its 2000 device for mammography (FDA 1999).
- homeprice.com.hk used Bayesian analysis for pricing information on over a million real state properties in Hong Kong and surrounding areas (Shamdasany, 2011).
- Researchers in the energy industry have used Bayesian analysis to understand petroleum reservoir parameters (Glinsky and Gunning, 2011).

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# The Method

## The Method

- Let's start by writing the Bayes' Rule:

$$p(B|A) = \frac{p(A|B)p(B)}{p(A)}$$

Where:

$p(A|B)$ : conditional probability of A given B

$p(B|A)$ : conditional probability of B given A

$p(B)$ : marginal probability of B

$p(A)$ : marginal probability of A

## The Method

- If we have a probability model for a vector of observations  $\mathbf{y}$  and a vector of unknown parameters  $\theta$ , we can represent the model with a likelihood function:

$$L(\theta; \mathbf{y}) = f(\mathbf{y}; \theta) = \prod_{i=1}^n f(y_i | \theta)$$

Where:

$f(\mathbf{y}; \theta)$ : conditional probability of  $\mathbf{y}$  give  $\theta$

- Let's assume that  $\theta$  has a probability distribution  $\pi(\theta)$ , and that denote  $\mathbf{m}(\mathbf{y})$  denote the marginal distribution of  $\mathbf{y}$ , such that:

$$m(\mathbf{y}) = \int f(\mathbf{y}; \theta) \pi(\theta) d\theta$$

## The Method

- Let's now write the inverse law of probability (Bayes' Theorem):

$$f(\theta|y) = \frac{f(y; \theta) \pi(\theta)}{f(y)}$$

- But notice that the marginal distribution of  $y$ ,  $f(y)$ , does not depend on  $(\theta)$
- Then, we can write the fundamental equation for Bayesian analysis:

$$p(\theta|y) \propto L(y|\theta) \pi(\theta)$$

# Let's go back to our initial example

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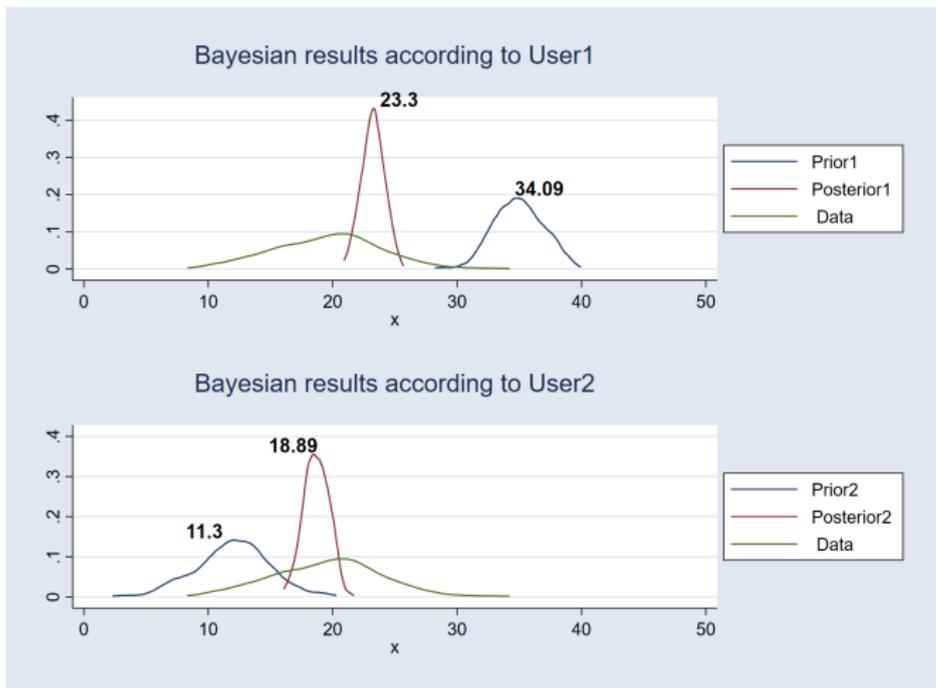
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## The Method

- In the example we have the data (the likelihood component)
- We also have the experts belief (the prior component)
- Then, how do we get the posterior distribution?
- We use the fundamental equation

$$p(\theta|y) \propto L(y|\theta) \pi(\theta)$$

## The Method

- Let's assume that both, the data and the prior beliefs, are normally distributed:
  - **The data:**  $y \sim N(\theta, \sigma_d^2)$
  - **The prior:**  $\theta \sim N(\mu_p, \sigma_p^2)$
- **Homework...:** Doing the algebra with the fundamental equation we find that the posterior distribution would be normal with:
  - **The posterior:**  $\theta|y \sim N(\mu, \sigma^2)$

Where:

$$\begin{aligned}\mu &= \sigma^2 (N\bar{y}/\sigma_d^2 + \mu_p/\sigma_p^2) \\ \sigma^2 &= (N/\sigma_d^2 + 1/\sigma_p^2)^{-1}\end{aligned}$$

## The Method

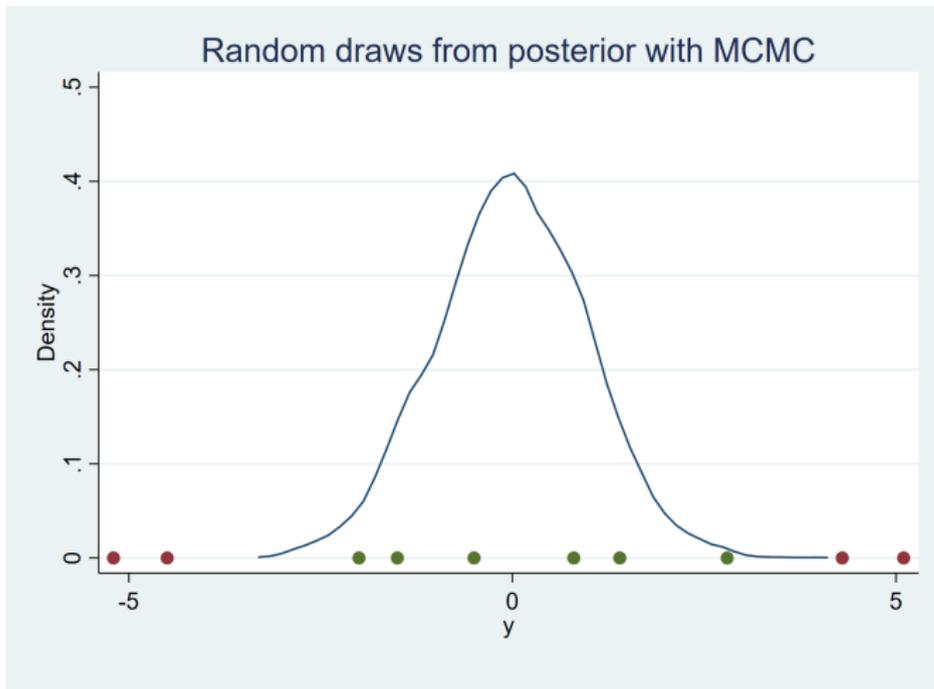
- Doing the algebra was relatively straightforward in the previous case.
- What about more complex distributions?
  - Integration is performed via simulation
  - We need to use intensive computational simulation tools to find the posterior distribution in most cases.
  - Markov chain Monte Carlo (MCMC) methods are the current standard in most software. Stata implement two alternatives:
    - Metropolis-Hastings (MH) algorithm
    - Gibbs sampling

## The Method

- Metropolis-Hastings (MH) algorithm
  - ① Specify a proposal probability distribution  $q(\cdot)$
  - ② Set an initial state within the domain of the posterior distribution  $\theta_0$
  - ③ Propose a new state for the posterior distribution  $\theta_t$  ;  $t=1,2,\dots$
  - ④ Compute an acceptance rate based on the ratio of the posterior distribution evaluated at the proposed state  $\theta_t$  and at the previous state  $\theta_{t-1}$ .
  - ⑤ If the ratio is:
    - Greater than 1  $\rightarrow$  keep the proposed value (state)
    - Less than one  $\rightarrow$  draw a random number from  $U(0,1)$  and keep  $\theta_t$  if the ratio is greater than the random draw.
  - ⑥ Repeat the process from 3 with the selected  $\theta_t$

# The Method

- Green points represent accepted proposal states and red points represent rejected proposal states.



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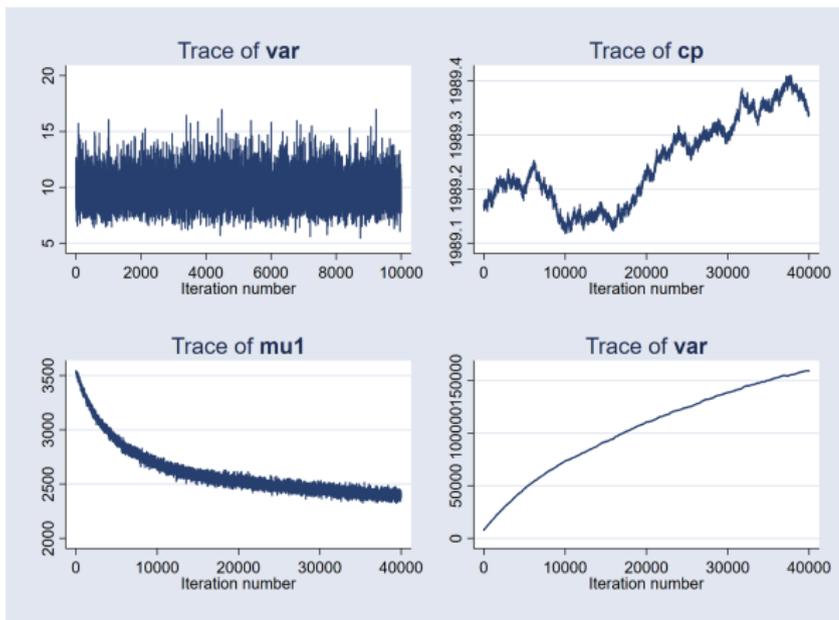
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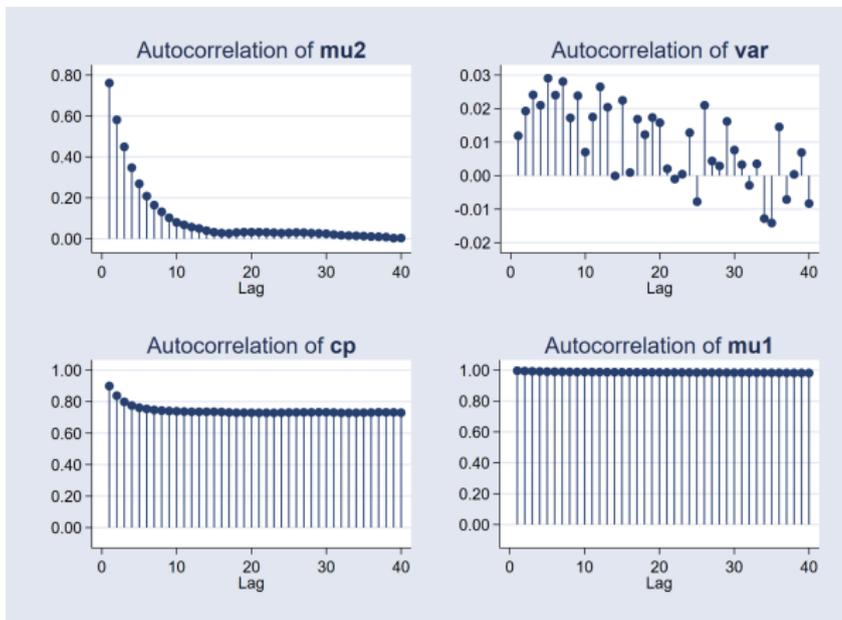
# The Method

- The trace plot illustrates the sequence of accepted proposal states.
- We expect to obtain a stationary sequence when convergence is achieved.



# The Method

- An efficient MCMC should have small autocorrelation.
- We expect autocorrelation to become negligible after a few lags.



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# The Stata tools for Bayesian regression

## The Stata tools: `bayesmh`

- In Stata 14 we introduce `bayesmh`.
- This is a general purpose command to perform Bayesian analysis using MCMC (MH or Gibbs).
- We are going to work with a few examples to show different facilities available in Stata for the analysis.
- Let's look at our first example:

- We have stats on number of wins by the Porto soccer team.
- We fit a linear regression for yearly wins.
- Let's consider three specifications:

$$\text{wins} = \alpha_1 + \beta_{gs} * \text{goals\_scored} + \epsilon_1$$

$$\text{wins} = \alpha_2 + \beta_{ga} * \text{goals\_against} + \epsilon_2$$

$$\text{wins} = \alpha_3 + \beta_{gs2} * \text{goals\_scored} + \beta_{ga2} * \text{goals\_against} + \epsilon_3$$

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## The Stata tools: Regression with `bayesmh`

- Here is one syntax with `bayesmh` to fit this model:

```
bayesmh wins gs, likelihood(normal({sigma2})) ///  

prior({wins:gs _cons}, normal(0,10000)) ///  

prior({sigma2}, igamma(.01,.01)) ///  

rseed(123)
```

- But let's use the Graphical User Interface (GUI) (Menus and dialog boxes):

## The Stata tools: Menu for Bayesian regression

- 1 Make the following sequence of selection from the main menu:  
Statistics > Bayesian analysis  
> General estimation and regression
- 2 Select 'Univariate linear models'
- 3 Specify the dependent variable (wins) and the explanatory variable (gs)
- 4 Select the 'Likelihood model' (Normal regression)
  - For 'Variance' click on 'Create' and select 'Specify as a model parameter'
  - Type 'sigma2' in 'Parameter name'

## The Stata tools: Menu for Bayesian regression

- For "'Priors of model parameters' click on 'Create'
  - Select wins:gs and wins:\_cons
  - Select the 'Normal distribution'
  - write '0' for the mean and '10000' for the variance.
- Next, create the prior for the variance of the likelihood sigma2
  - Select the Inverse gamma distribution
  - Specify .01 and .01 for the 'Shape' and 'Scale' parameters.
- Click on the 'Simulation' tab and set the 'Random-number seed' to 123

```
. bayesmh wins gs, likelihood(normal({sigma2})) ///  
>          prior({wins:gs _cons}, normal(0,10000)) ///  
>          prior({sigma2}, igamma(.01,.01)) ///  
>          rseed(123)
```

**Burn-in ...**

**Simulation ...**

**Model summary**

---

**Likelihood:**

wins ~ normal(xb\_wins, {sigma2})

**Priors:**

{wins:gs \_cons} ~ normal(0,10000)

{sigma2} ~ igamma(.01,.01)

---

(1) Parameters are elements of the linear form `xb_wins`.

## The Stata tools: Regression output

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```
. bayesmh wins gs, likelihood(normal({sigma2})) ///
> prior({wins:gs _cons}, normal(0,10000)) ///
> prior({sigma2}, igamma(.01, .01)) ///
> rseed(123)
```

```
Bayesian normal regression                MCMC iterations =      12,500
Random-walk Metropolis-Hastings sampling  Burn-in           =       2,500
                                           MCMC sample size =     10,000
                                           Number of obs    =       47
                                           Acceptance rate  =      .2222
                                           Efficiency: min  =      .04521
                                           avg             =      .06161
                                           max             =      .07185
```

Log marginal likelihood = -135.77023

		Mean	Std. Dev.	MCSE	Median	Equal-tailed [95% Cred. Interval]	
<b>wins</b>	gs	.2360223	.0365801	.001405	.2363132	.162386	.3096086
	_cons	6.711756	2.417745	.090197	6.704956	1.923868	11.53032
	<b>sigma2</b>	9.380877	2.089641	.098277	9.040789	6.262636	14.55403

The Stata tools: `bayesstats ess`

- Let's use the postestimation command `bayesstats ess` to evaluate the effective sample size

```
. bayesstats ess
```

Efficiency summaries      MCMC sample size =      10,000

	ESS	Corr. time	Efficiency
<b>wins</b>			
<b>gs</b>	677.68	14.76	0.0678
<b>_cons</b>	718.51	13.92	0.0719
<b>sigma2</b>	452.10	22.12	0.0452

- We expect to have an acceptance rate (see previous slide) that is neither too small nor too large.
- We also expect to have low correlation
- Efficiencies over 10% are considered good for MH. Efficiencies under 1% would be a source of concern.

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## The Stata tools: Blocking of parameters

- **Blocking of parameters**
  - The update steps for MH are performed simultaneously for all parameters.
  - For high dimensional models this may result in poor mixing.
  - Blocking of parameters helps improving mixing efficiency

## The Stata tools: Blocking of parameters

- Blocking of parameters
  - How it works?
    - It separates the model parameters into two or more subsets of blocks.
    - MH updates are applied to each block separately
    - Computations are performed in the order the blocks are specified

```
bayesmh wins gs,likelihood(normal({sigma2})) ///
prior({wins:gs _cons}, normal(0,10000)) ///
prior({sigma2}, igamma(.01,.01)) ///
block({wins:gs _cons}) block({sigma2}) ///
rseed(123)
```

## The Stata tools: Menu for Blocking of parameters

Let's go back to our previous example:

- 1 Click on the 'Blocking' tab
- 2 Select 'Display block summary'
- 3 Click on 'Create'
- 4 Select wins:gs and wins:\_cons and click 'OK'
- 5 Click on 'Create'
- 6 Select sigma2 and click 'OK'

# The Stata tools: Blocking of parameters

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```
. bayesmh wins gs,likelihood(normal({sigma2})) //
> prior({wins:gs _cons},normal(0,10000)) prior({sigma2},igamma(.01,.01)) //
> block({wins:gs _cons}) block({sigma2}) rseed(123) blocksummary
```

Burn-in ...

Simulation ...

## Block summary

```
1: {wins:gs _cons}
2: {sigma2}
```

Bayesian normal regression  
Random-walk Metropolis-Hastings sampling

```
MCMC iterations = 12,500
Burn-in = 2,500
MCMC sample size = 10,000
Number of obs = 47
Acceptance rate = .3426
Efficiency: min = .09882
              avg = .1156
              max = .1464
```

Log marginal likelihood = -135.7408

	Mean	Std. Dev.	MCSE	Median	Equal-tailed [95% Cred. Interval]	
<b>wins</b>						
gs	.2363963	.0373595	.001188	.2366527	.1626758	.3109461
_cons	6.690619	2.452853	.076957	6.69004	1.683672	11.63661
<b>sigma2</b>	9.392034	2.136876	.05585	9.09526	6.170093	14.36234

The Stata tools: `bayesstats ess`

- Let's evaluate again the effective sample size

```
. bayesstats ess
Efficiency summaries      MCMC sample size =      10,000
```

	ESS	Corr. time	Efficiency
<b>wins</b>			
<b>gs</b>	988.18	10.12	0.0988
<b>_cons</b>	1015.90	9.84	0.1016
<b>sigma2</b>	1463.91	6.83	0.1464

- The efficiency is now around 10% or more for all the parameters.
- Correlation was reduced
- The effective sample size is also higher for all the parameters.

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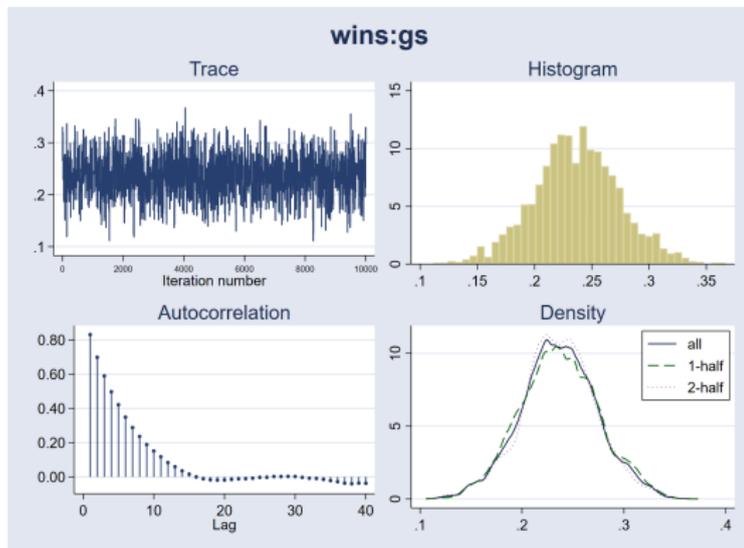
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## The Stata tools: bayesgraph

- We can use `bayesgraph` to look at the trace, the correlation, and the density. For example:

**. bayesgraph diagnostic {gs}**



- The trace indicates that convergence was achieved
- Correlation becomes negligible after 10 periods

## The Stata tools: `bayes : prefix`

- In Stata 15 we introduce the prefix command `bayes :`
- This is a simple syntax to perform Bayesian analysis.
- You specify the prefix followed by your estimation command.
- The specified estimation defines the likelihood for the model.
- The default priors are assumed to be noninformative in many cases.
- But the priors may become informative due to the scale of the parameters.
- The default priors could be consider a starting point.
- However, alternative priors may need to be considered.
- Postestimation commands would help decide on the final model.
- Let's use `bayes :` to fit our previous model:

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```
. bayes, rseed(123) nomodelsummary: regress wins gs
```

```
Burn-in ...
```

```
Simulation ...
```

```
Bayesian linear regression
```

```
Random-walk Metropolis-Hastings sampling
```

```
MCMC iterations = 12,500
Burn-in = 2,500
MCMC sample size = 10,000
Number of obs = 47
Acceptance rate = .3426
Efficiency: min = .09882
              avg = .1156
              max = .1464
```

```
Log marginal likelihood = -135.7408
```

		Mean	Std. Dev.	MCSE	Median	Equal-tailed [95% Cred. Interval]	
<b>wins</b>	<b>gs</b>	.2363963	.0373595	.001188	.2366527	.1626758	.3109461
	<b>_cons</b>	6.690619	2.452853	.076957	6.69004	1.683672	11.63661
	<b>sigma2</b>	9.392034	2.136876	.05585	9.09526	6.170093	14.36234

```
Note: Default priors are used for model parameters.
```

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The Stata tools: `bayesstats ic`

- Let's fit now the other two models that we specify at the beginning of this example.
- We will store the results for the three models and we will use the postestimation command `bayesstats ic` to select one of them.

```
quietly {  
    bayes , rseed(123): regress wins gs  
    estimates store m_gs  
  
    bayes , rseed(123): regress wins ga  
    estimates store m_ga  
  
    bayes , rseed(123): regress wins gs ga  
    estimates store m_full  
}  
bayesstats ic m_gs m_ga m_full,basemodel(m_ga)
```

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# The Stata tools: `bayesstats ic`

- `bayesstats ic` reports three statistics
  - Log of the marginal likelihood
  - DIC:
    - It is designed for Bayesian estimation involving MCMC simulations.
    - It Has a penalty term based on the difference between the expected log likelihood and the likelihood at the posterior mean point.
    - You should select the model with the lowest DIC.
  - Bayes factors (BF)
    - Incorporates information about model priors.
    - Ratio of the marginal likelihood of two models (fit on the same sample).
    - It can be used to compare nested and nonnested models.
    - Not applicable to models with improper priors.

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The Stata tools: `bayesstats ic`

- Here is the output for `bayesstats ic`

```
. quietly {
. bayesstats ic m_gs m_ga m_full, basem(m_full) bayesf
Bayesian information criteria
```

	DIC	log (ML)	BF
<code>m_gs</code>	240.6314	-135.7408	<b>5.015833</b>
<code>m_ga</code>	268.5267	-148.5384	<b>.0000139</b>
<code>m_full</code>	230.9162	-137.3534	.

Note: Marginal likelihood (ML) is computed using Laplace-Metropolis approximation.

- Interpretation for Bayes Factors (Jeffreys 1961)

<code>log10(BF_jb)</code>	<code>BF_jb</code>	Evidence against <code>M_b</code>
0 to 1/2	1 to 3.2	Bare mention
1/2 to 1	3.2 to 10	Substantial
1 to 2	10 to 100	Strong
> 2	> 100	Decisive

The Stata tools: `bayestest model`

- `bayestest model` is another postestimation command to compare different models.
- We can again store the results for our alternative models, and then use `bayestest model`.

```
quietly {  
    bayes , rseed(123): regress wins gs  
    estimates store m_gs  
  
    bayes , rseed(123): regress wins ga  
    estimates store m_ga  
  
    bayes , rseed(123): regress wins gs ga  
    estimates store m_full  
  
    bayes , prior({wins:gs _cons}, normal(20,10)) ///  
        rseed(123): regress wins  
    estimates store m_meanonly  
}  
bayestest model m_gs m_ga m_full m_meanonly
```

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## The Stata tools: `bayestest model`

- `bayestest model` computes the posterior probabilities for each model.
- The result indicates which model is more likely.
- It requires that the models use the same data and that they have proper posterior.
- It can be used to compare models with:
  - Different priors and/or different posterior distributions.
  - Different regression functions.
  - Different covariates
- MCMC convergence should be verified before comparing the models.

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The Stata tools: `bayestest` model

- Here is the output for `bayestest` model

```
. bayestest model m_gs m_ga m_full m_meanonly
```

	log(ML)	P (M)	P (M y)
<code>m_gs</code>	-135.7408	0.2500	0.8211
<code>m_ga</code>	-148.5384	0.2500	0.0000
<code>m_full</code>	-137.3534	0.2500	0.1637
<code>m_meanonly</code>	-139.7326	0.2500	0.0152

Note: ML is computed using Laplace-Metropolis approximation.

- We could also assign different priors for the models:

```
. bayestest model m_gs m_ga m_full m_meanonly, ///
      prior(.2 .1 .4 .3)
```

	log(ML)	P (M)	P (M y)
<code>m_gs</code>	-135.7408	0.2000	0.7010
<code>m_ga</code>	-148.5384	0.1000	0.0000
<code>m_full</code>	-137.3534	0.4000	0.2795
<code>m_meanonly</code>	-139.7326	0.3000	0.0194

Note: ML is computed using Laplace-Metropolis approximation.

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# Random Effects Probit model

## The Stata tools: Random effects probit model

- We are going to use `bayes`: to fit a random effects probit model for a binary variable  $y_{it}$ , which depends on the latent variable .

$$y_{it}^* = \beta_0 + \beta_1 x_{it1} + \beta_2 x_{it2} + \dots + \beta_k x_{itk} + \alpha_j + \epsilon_{it}$$

Where:

$$y_{it} = \begin{cases} 1 & \text{if } y_{it}^* > 0 \\ 0 & \text{otherwise} \end{cases}$$

$\alpha_j \sim N(0, \sigma_\alpha^2)$  is the individual random panel effect  
 $\epsilon_{it} \sim N(0, \sigma_\epsilon^2)$  is the idiosyncratic error term

- This is also referred as a two-level random intercept model.
- We can also fit this model with `meprobit` or `xtprobit, re`

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# The Stata tools: Random effects probit model

- This time we are going to work with simulated data.
- Here is the code to simulate the panel dataset:

```
clear
set obs 100
set seed 1
```

```
* Panel level *
```

```
generate id=_n
generate alpha=rnormal()
expand 5
```

```
* Observation level *
```

```
bysort id:generate year=_n
xtset id year
generate x1=rnormal()
generate x2=runiform().5
generate x3=uniform()
generate u=rnormal()
```

```
* Generate dependent variable *
```

```
generate y=.5+1*x1+(-1)*x2+1*x3+alpha+u>0
```

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## The Stata tools: Random effects probit model

Let's show the results with `meprobit`:

```
. meprobit y x1 x2 x3 || id:,nolog
Mixed-effects probit regression      Number of obs      =      500
Group variable:                      id                 Number of groups   =      100
                                                           Obs per group:
                                      min =              5
                                      avg =              5.0
                                      max =              5
Integration method: mvaghermite      Integration pts.    =        7
Log likelihood = -236.88589           Wald chi2(3)       =      82.83
                                      Prob > chi2        =      0.0000
```

	y	Coef.	Std. Err.	P> z	[95% Conf. Interval]	
	x1	.9769118	.1143889	0.000	.7527138	1.20111
	x2	-.9896286	.1853433	0.000	-1.352895	-.6263625
	x3	.9426958	.2941061	0.001	.3662584	1.519133
	_cons	.5220418	.2187448	0.017	.0933098	.9507738
id	var(_cons)	1.31	.3835866		.7379508	2.325494

LR test vs. probit model: `chibar2(01) = 67.24`    **Prob >= chibar2 = 0.0000**

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## The Stata tools: Random effects probit model

We now fit the model with `bayes` :

```
. bayes , dryrun: meprobit y x1 x2 x3 || id:
```

Multilevel structure

---

```
id
  {U0}: random intercepts
```

---

Model summary

---

**Likelihood:**

```
y ~ meprobit(xb_y)
```

**Priors:**

```
{y:x1 x2 x3 _cons} ~ normal(0,10000) (1)
```

```
. {U0} ~ normal(0, {U0:sigma2}) (1)
```

**Hyperprior:**

```
{U0:sigma2} ~ igamma(.01, .01)
```

---

(1) Parameters are elements of the linear form `xb_y`.

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We now fit the model with `bayes`:

```
. bayes , nomodelsummary nodots rseed(123): meprobit y x1 x2 x3 || id:
Burn-in ...
Simulation ...
Multilevel structure
```

---

```
id
      {U0}: random intercepts
```

---

```
Bayesian multilevel probit regression      MCMC iterations =      12,500
Random-walk Metropolis-Hastings sampling  Burn-in           =       2,500
                                           MCMC sample size =     10,000
Group variable: id                       Number of groups  =       100
                                           Obs per group:
                                           min =              5
                                           avg =             5.0
                                           max =              5
Family : Bernoulli                        Number of obs     =       500
Link   : probit                          Acceptance rate   =      .3247
                                           Efficiency: min =  .01333
                                           avg =            .02736
                                           max =            .04012
```

---

```
Log marginal likelihood
```

---

		Mean	Std. Dev.	MCSE	Median	Equal-tailed [95% Cred. Interval]	
<b>y</b>	x1	.9866518	.1129356	.006316	.9850336	.7789124	1.215904
	x2	-1.005328	.1793814	.009673	-1.003398	-1.357951	-.6617393
	x3	.9856235	.2968089	.014819	.9666234	.4282133	1.591159
	_cons	.5051288	.2055344	.017802	.5032979	.0933563	.889766
<b>id</b>	U0:sigma2	1.432124	.4234419	.032504	1.388553	.7326054	2.388284

Note: Default priors are used for model parameters.

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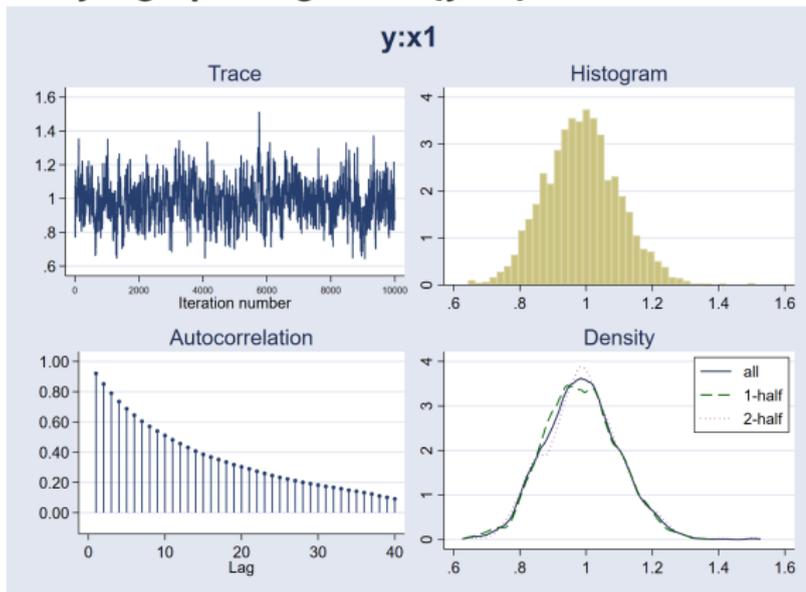
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# The Stata tools: `bayesgraph` `diagnostic`

- We can look at the diagnostic graph for a couple of variables:

## . `bayesgraph diagnostic {y:x1}`

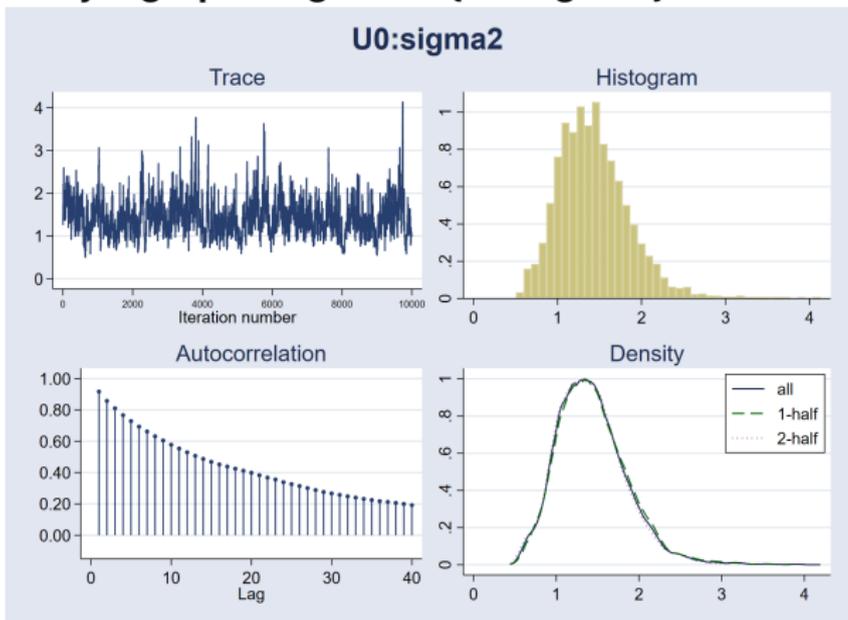


- The trace shows periods with trends.
- Correlation is persistent for around 25 periods.

## The Stata tools: `bayesgraph` `diagnostic`

- Look now at the diagnostic graphs for `U0:sigma2`

```
. bayesgraph diagnostic {U0:sigma2}
```



- The trace also shows periods with trends.
- Correlation is persistent for around 30 periods.

## The Stata tools: thinning

- We can reduce autocorrelation by using thinning
- This would save the random draws skipping a prespecified number of simulated values in the iteration process for the MCMC.
- We can use the option 'thinning(#)' to indicate that Stata should save simulated values from every  $(1+k*\#)$ th iteration ( $k=0,1,2,\dots$ ).
- Let's try using 'thinning(5)'

```
bayes ,nomodelsummary nodots rseed(123) ///  
thinning(5): meprobit y x1 x2 x3 || id:
```

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## The Stata tools: thinning

Let's show the results with 'thinning(5)'

```
. bayes,nomodelsummary nodots rseed(123) thinning(5):meprobit y x1 x2 x3 || id:
note: discarding every 4 sample observations; using observations 1,6,11,...
```

Burn-in ...

Simulation ...

Multilevel structure

id

{U0}: random intercepts

```
Bayesian multilevel probit regression      MCMC iterations =      52,496
Random-walk Metropolis-Hastings sampling   Burn-in           =       2,500
                                           MCMC sample size =     10,000
```

Group variable: id

```
Number of groups =      100
Obs per group:
```

```
min =      5
avg  =     5.0
max  =      5
```

Family : Bernoulli

Number of obs = 500

Link : probit

Acceptance rate = .3268

Efficiency: min = .05399

avg = .102

max = .1628

Log marginal likelihood

		Mean	Std. Dev.	MCSE	Median	Equal-tailed [95% Cred. Interval]	
<b>y</b>							
	x1	.9977099	.1181726	.003773	.9936143	.7810441	1.242439
	x2	-1.018063	.1892596	.00557	-1.012598	-1.396798	-.6509636
	x3	.9539304	.2936949	.007279	.9514395	.3823801	1.52913
	_cons	.5433822	.2205077	.00949	.5398387	.1216346	.9847166
<b>id</b>							
	U0:sigma2	1.456558	.4384163	.015537	1.401461	.7611919	2.463175

Note: Default priors are used for model parameters.

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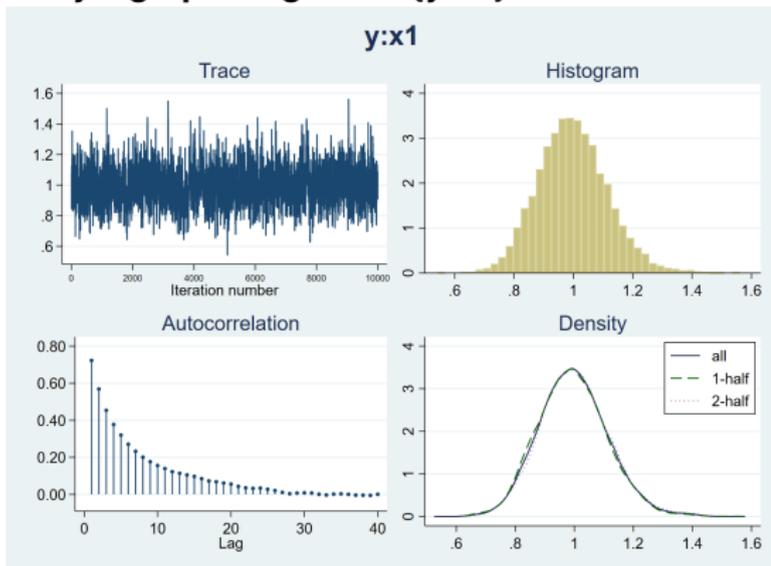
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## bayesgraph diagnostic

- We now look at the diagnostic graph for the same two variables:

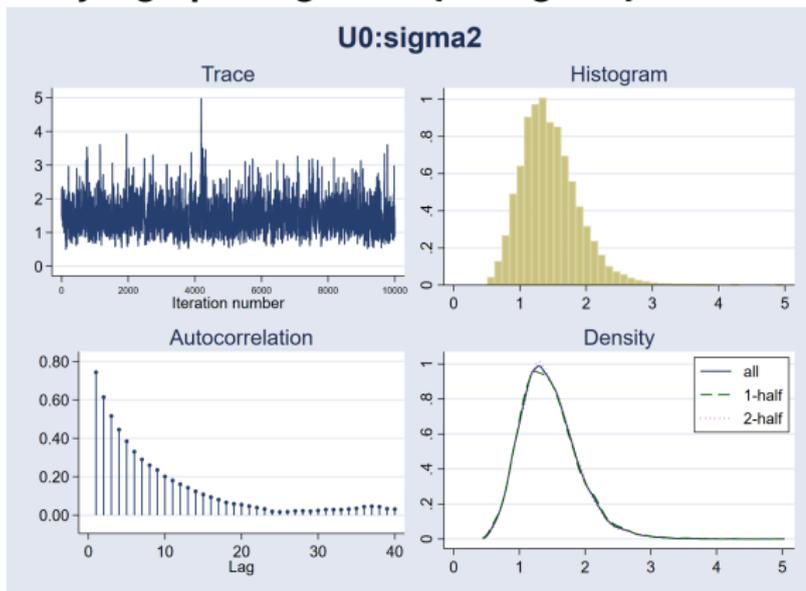
### . bayesgraph diagnostic {y:x1}



- The trace seems to indicate convergence this time.
- Autocorrelation decays quicker and becomes negligible after about 15 periods.

- We now look now at the diagnostic graphs for `U0:sigma2`

## . `bayesgraph diagnostic {U0:sigma2}`



- The trace seems to indicate convergence this time.
- Autocorrelation decays quicker and becomes negligible after about 15 periods.

## The Stata tools: `bayestest interval`

- We can perform interval testing with the postestimation command `bayestest interval`.
- It estimates the probability that a model parameter lies in a particular interval.
- For continuous parameters the hypothesis is formulated in terms of intervals.
- We can perform point hypothesis testing only for parameters with discrete posterior distributions.
- `bayestest interval` estimates the posterior distribution for a null interval hypothesis.
- `bayestest interval` reports the estimated posterior mean probability for  $H_0$ .

```
bayestest interval ( {y:x1} ,lower(.9) upper(1.02)) ///  
                    ( {y:x2} ,lower(-1.1) upper(-.8))
```

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# The Stata tools: bayestest interval

- We can, for example, perform separate tests for different parameters:

```
. bayestest interval ({y:x1},lower(.9) upper(1.02)) ///  
> ({y:x2},lower(-1.1) upper(-.8))
```

```
Interval tests      MCMC sample size =    10,000
```

```
prob1 : .9 < {y:x1} < 1.02  
prob2 : -1.1 < {y:x2} < -.8
```

	Mean	Std. Dev.	MCSE
prob1	.3888	0.48750	.0077073
prob2	.5474	0.49777	.0097517

- We can also perform a joint test:

```
. bayestest interval (({y:x1},lower(.9) upper(1.02)) ///  
> ({y:x2},lower(-1.1) upper(-.8)), joint)
```

```
Interval tests      MCMC sample size =    10,000
```

```
prob1 : .9 < {y:x1} < 1.02, -1.1 < {y:x2} < -.8
```

	Mean	Std. Dev.	MCSE
prob1	.2249	0.41754	.0066399

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# Change-point model

## The Stata tools: Change-point model

- Let's work now with an example where we write our model using a substitutable expression.
- We have data on yearly trademark applications in portugal:



- The series has a significant change around 1990.
- We may consider fitting a change-point model.

## The Stata tools: Change-point model

## Change point model specification

```
bayesmh //  
    trdmark=({mu1}*sign(year<{cp}))+{mu2}*sign(year>={cp})), ///  
    likelihood(normal({var})) ///  
    prior({mu1}, normal(3000,2000000)) ///  
    prior({mu2}, normal(16000,2000000)) ///  
    prior({cp}, uniform(1960,2016)) ///  
    prior({var}, igamma(2,1)) ///  
    initial({mu1} 5000 {mu2} 10000 {cp} 1960) ///  
    rseed(123) mcmcsize(40000) ///  
    dots(500, every(5000)) ///  
    title(Change-point analysis)
```

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```
. bayesmh trdmark=({mu1}*sign(year<{cp})+{mu2}*sign(year>={cp})), ///
> likelihood(normal({var})) ///
> prior({mu1}, normal(3000,2000000)) ///
> prior({mu2}, normal(16000,2000000)) ///
> prior({cp}, uniform(1960,2016)) ///
> prior({var}, igamma(2,1)) ///
> initial({mu1} 5000 {mu2} 10000 {cp} 1960) ///
> rseed(123) mcmcsize(40000) dots(500, every(5000)) ///
> title(Change-point analysis)
```

Burn-in 2500 aaaaa done

```
Simulation 40000 .....5000.....10000.....15000.....20000
> .....25000.....30000.....35000.....40000 done
```

## Model summary

## Likelihood:

```
trdmark ~ normal({mu1}*sign(year<{cp})+{mu2}*sign(year>={cp}), {var})
```

## Priors:

```
{var} ~ igamma(2,1)
{mu1} ~ normal(3000,2000000)
{mu2} ~ normal(16000,2000000)
{cp} ~ uniform(1960,2016)
```

## The Stata tools: Change-point model

## Change point model specification

```

. bayesmh trdmark=({mu1}*sign(year<{cp})+{mu2}*sign(year>={cp})), ///
> likelihood(normal({var})) ///
> prior({mu1}, normal(3000,2000000)) ///
> prior({mu2}, normal(16000,2000000)) ///
> prior({cp}, uniform(1960,2016)) ///
> prior({var}, igamma(2,1)) ///
> initial({mu1} 5000 {mu2} 10000 {cp} 1960) ///
> rseed(123) mcmcsize(40000) dots(500, every(5000)) ///
> title(Change-point analysis)

```

```

Change-point analysis                                MCMC iterations =      42,500
Random-walk Metropolis-Hastings sampling            Burn-in           =       2,500
                                                    MCMC sample size =    40,000
                                                    Number of obs     =       55
                                                    Acceptance rate   =     .4117
                                                    Efficiency: min   =    .001033
                                                    avg               =     .03796
                                                    max               =     .1362

```

Log marginal likelihood = -621.28408

	Mean	Std. Dev.	MCSE	Median	Equal-tailed [95% Cred. Interval]	
cp	1989.492	.2891978	.003918	1989.492	1989.023	1989.972
mu1	3754.837	153.0364	11.9209	3761.923	3468.338	4015.751
mu2	17448.84	144.531	7.04777	17448.23	17170.98	17736.22
var	463983.1	144106.8	22418.1	487445.9	89224.3	621052.3

Note: There is a high autocorrelation after 500 lags.

Note: Adaptation tolerance is not met.

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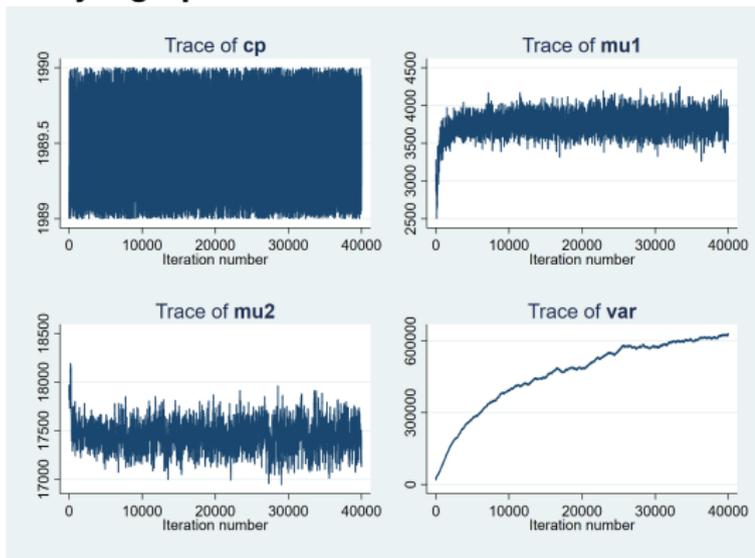
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## The Stata tools: `bayesgraph trace`

- We can use `bayesgraph trace` to look at the trace for all the parameters.
- This helps in determining convergence.

### . `bayesgraph trace`

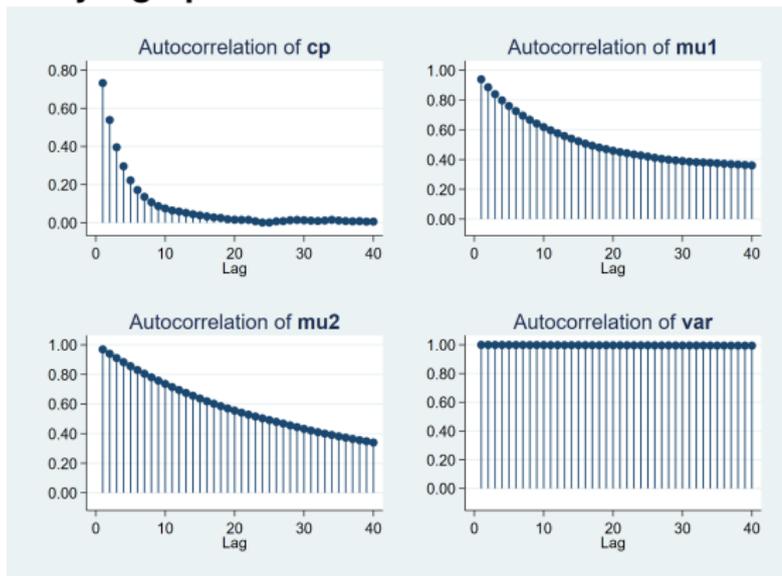


- We observe signs of lack of convergence, particularly for the variance.

## The Stata tools: `bayesgraph ac`

- We can use `bayesgraph ac` to look at the autocorrelation for all the parameters.
- This also helps in determining convergence.

### . `bayesgraph ac`



- The plot shows autocorrelation for almost all the parameters.

## Change point model specification with blocking

```
bayesmh //
  trdmark=({mu1}*sign(year<{cp}))+{mu2}*sign(year>={cp})), //
  likelihood(normal({var})) //
  prior({mu1}, normal(3000,2000000)) //
  prior({mu2}, normal(10000,2000000)) //
  prior({cp}, uniform(1960,2016)) //
  prior({var}, igamma(2,1)) //
  initial{mu1} 5000 {mu2} 10000 {cp} 1960 //
  block{var}, gibbs block{cp} blocksummary //
  rseed(123) mcmcsize(40000) //
  dots(500, every(5000)) //
  title(Change-point analysis)
```

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## Change point model specification with blocking

```

. bayesmh trdmark=({mu1}*sign(year<{cp})+{mu2}*sign(year>={cp})), ///
> likelihood(normal({var})) ///
> prior({mu1}, normal(3000,2000000)) ///
> prior({mu2}, normal(16000,2000000)) ///
> prior({cp}, uniform(1960,2016)) ///
> prior({var}, igamma(2,1)) ///
> initial({mu1} 5000 {mu2} 10000 {cp} 1960) ///
> block({var}, gibbs) block({cp}) blocksummary ///
> rseed(123) mcmcsize(40000) dots(500,every(5000)) ///
> title(Change-point analysis)

```

Burn-in 2500 aaaaa done

```

Simulation 40000 .....5000.....10000.....15000.....20000
> .....25000.....30000.....35000.....40000 done

```

**Model summary****Likelihood:**

```
trdmark ~ normal({mu1}*sign(year<{cp})+{mu2}*sign(year>={cp}), {var})
```

**Priors:**

```
{var} ~ igamma(2,1)
{mu1} ~ normal(3000,2000000)
{mu2} ~ normal(16000,2000000)
{cp} ~ uniform(1960,2016)
```

**Block summary**

```

1: {var} (Gibbs)
2: {cp}
3: {mu1} {mu2}

```

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## Change point model specification with blocking

```

. bayesmh trdmark=({mu1}*sign(year<{cp})+{mu2}*sign(year>={cp})), ///
> likelihood(normal({var})) ///
> prior({mu1}, normal(3000,2000000)) ///
> prior({mu2}, normal(16000,2000000)) ///
> prior({cp}, uniform(1960,2016)) ///
> prior({var}, igamma(2,1)) ///
> initial({mu1} 5000 {mu2} 10000 {cp} 1960) ///
> block({var}, gibbs) block({cp}) blocksummary ///
> rseed(123) mcmcsize(40000) dots(500, every(5000)) ///
> title(Change-point analysis)

```

```

Change-point analysis                MCMC iterations =      42,500
Metropolis-Hastings and Gibbs sampling  Burn-in           =       2,500
                                         MCMC sample size =    40,000
                                         Number of obs    =       55
                                         Acceptance rate  =     .5288
                                         Efficiency: min =     .07912
                                         avg              =     .2638
                                         max              =     .6717

Log marginal likelihood = -533.33098

```

	Mean	Std. Dev.	MCSE	Median	Equal-tailed [95% Cred. Interval]	
cp	1989.496	.2944166	.003126	1989.496	1989.019	1989.975
mu1	3780.26	341.1711	6.06443	3783.149	3108.712	4446.395
mu2	17332.57	372.1327	6.47794	17344.63	16588.7	18068.87
var	3798272	739589.8	4512.18	3708037	2612399	5480970

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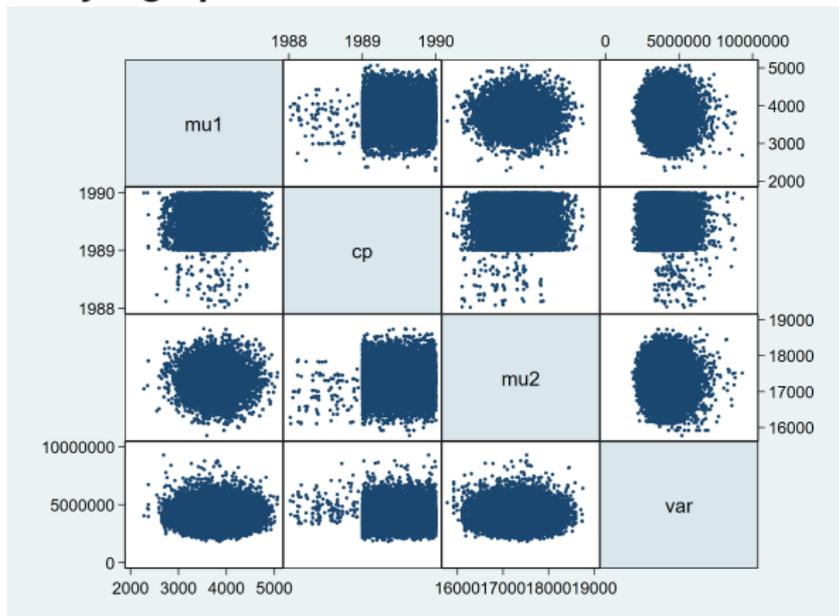
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## The Stata tools: `bayesgraph` `matrix`

- We check the scatterplots again for the simulated values of the coefficients and the variance.

### . `bayesgraph matrix`

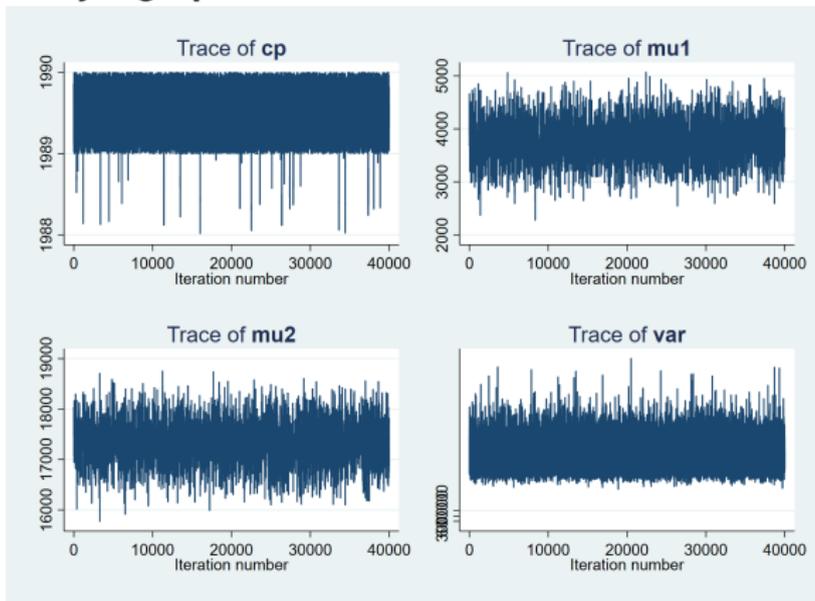


- We do not observe any pairwise correlations now.

## The Stata tools: `bayesgraph` `trace`

- We can use `bayesgraph trace` to look at the trace for all the parameters.

### . `bayesgraph trace`

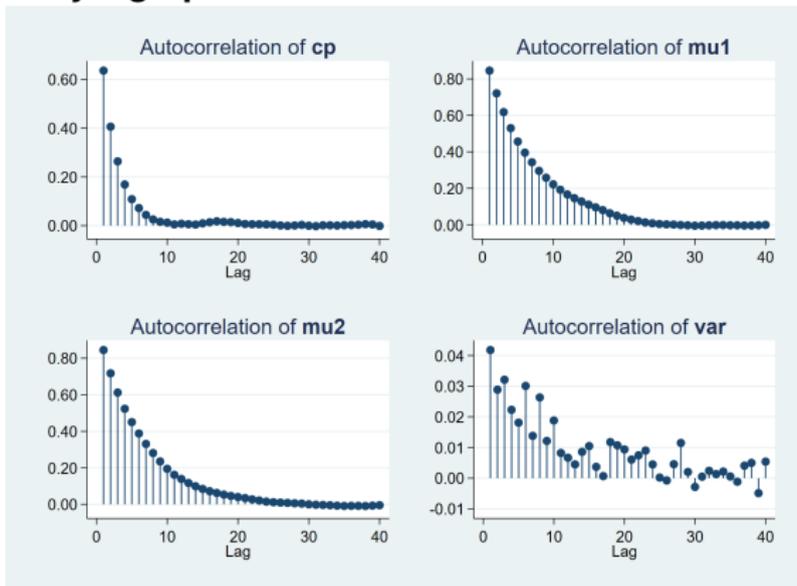


- The plots indicate that convergence seems to be achieved.

## The Stata tools: `bayesgraph ac`

- We can also use `bayesgraph ac` to look at the autocorrelation for all the parameters.

### . `bayesgraph ac`



- Autocorrelation decays and becomes negligible quickly for almost all the parameters.

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### 1 Bayesian analysis: The general idea

### 2 Basic Concepts

- The Method
- The tools
- Stata 14: The `bayesmh` command
- Stata 15: The `bayes` prefix
- Postestimation commands

### 3 A few examples

- Linear regression
- Panel data random effect probit model
- Change point model

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## References

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