Estimating effects from extended regression models

David M. Drukker

Executive Director of Econometrics Stata

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Extended regression models

Extended regression model (ERM) is a Stata term for a class of regression models

- The outcome can be continuous (linear), probit, orderded probit, or censored (tobit)
- Some of the covariates may be endogenous
 - The endogenous covariates may be continuous, probit, or ordered probit
- Endogenous sample-selection may be modeled
- Exogenous or endogenous treatment assignment may be modeled
- The new-in-Stata-15 commands eregress, eprobit, eoprobit, and eintreg fit ERMs

- Some of the covariates may be endogenous
 - The endogenous covariates may be continuous, binary, or ordinal
 - Polynomial terms and interaction terms constructed from the endogenous covariates are allowed
 - Interactions among the endogenous covariates and interactions between the endogenous covariates and the exogenous covariates are allowed

- I cannot do justice to ERMs in this short talk
- I discuss examples in which I
 - define some of the terms that I have already used
 - illustrate some command syntax
 - illustrate how to estimate some effects using postestimation commands

• Fictional data on wellness program from large company

. use wprogra . describe	m			
Contains data obs: vars: size:	from wpro 3,000 6 72,000	ogram.dta		28 Jul 2017 07:13
variable name	storage type	display format	value label	variable label
wchange age over phealth prog wtprog	float float float float float float	%9.0g %9.0g %9.0g %9.0g %9.0g %9.0g	changel yesno yesno	Weight change level Years over 50 Overweight (tens of pounds) Prior health score Participate in wellness program Offered work time to participate in program

Sorted by:

• Three levels of wchange

. tabulate wchange prog

Weight change	Participa wellness p	rogram	T .+.]
level	No	Yes	Total
Loss No change Gain	239 468 593	909 605 186	1,148 1,073 779
Total	1,300	1,700	3,000

- Data are observational
- Table does not account for how observed covariates and/or unobserved errors that affect program participation also affect the outcome variable

- I want a model that
 - allows observed covariates to affect both wchange and assignment to prog
 - allows the errors that affect assignment to prog to be correlated with the errors that affect wchange
 - I suspect that unobservables that increase program participation are negatively correlated with unobservables that affect weight gain
- In other words, I want allow prog to be endogenous

If prog is endogenous, I must model the dependence. Consider

$$\begin{aligned} & \text{wchange} = \begin{cases} \text{``Loss''} & \text{if} & \beta_1 \text{prog} + \mathbf{x}\beta + \epsilon \leq cut1 \\ \text{``No change''} & \text{if } cut1 < \beta_1 \text{prog} + \mathbf{x}\beta + \epsilon \leq cut2 \\ \text{``Gain''} & \text{if } cut2 < \beta_1 \text{prog} + \mathbf{x}\beta + \epsilon \end{cases} \\ & \text{prog} = (\mathbf{x}\gamma + \gamma_1 \text{wtime} + \eta > 0) \\ & \epsilon \text{ and } \eta \text{ are correlated and joint normal} \\ & \mathbf{x}\beta = \beta_2 \text{age} + \beta_3 \text{over} + \beta_4 \text{phealth} \\ & \mathbf{x}\gamma = \gamma_2 \text{age} + \gamma_3 \text{over} + \gamma_4 \text{phealth} \end{aligned}$$

• wtime is an instrumental variable

- It is included in the model for treatment
- It is excluded from the model for the potential outcomes of wchange

$$\begin{aligned} & \text{wchange} = \begin{cases} \text{"Loss"} & \text{if} \qquad \beta_1 \text{prog} + \mathbf{x}\beta + \epsilon \leq cut1 \\ \text{"No change"} & \text{if } cut1 < \beta_1 \text{prog} + \mathbf{x}\beta + \epsilon \leq cut2 \\ \text{"Gain"} & \text{if } cut2 < \beta_1 \text{prog} + \mathbf{x}\beta + \epsilon \end{cases} \\ & \text{prog} = (\mathbf{x}\gamma + \gamma_1 \text{wtime} + \eta > 0) \\ & \epsilon \text{ and } \eta \text{ are correlated and joint normal} \\ & \mathbf{x}\beta = \beta_2 \text{age} + \beta_3 \text{over} + \beta_4 \text{phealth} \\ & \mathbf{x}\gamma = \gamma_2 \text{age} + \gamma_3 \text{over} + \gamma_4 \text{phealth} \end{aligned}$$

$ \begin{array}{c ccccccccccccccccccccccccccccccccccc$
Log likelihood = -4401.0952 Prob > chi2 = 0.0000 Coef. Std. Err. z P> z [95% Conf. Interval] wchange .2155906 .0705048 3.06 0.002 .0774037 .3537776 wchange .4349946 .0387185 11.23 0.000 .3591078 .5108814 phealth 4933361 .0411866 -11.98 0.000 5740603 412612 prog Yes 3624996 .1031408 -3.51 0.000 5646519 1603473 prog age 9341234 .0840002 -11.12 0.000 -1.098761 7694861
Coef. Std. Err. z P> z [95% Conf. Interval] wchange .2155906 .0705048 3.06 0.002 .0774037 .3537776 over .4349946 .0387185 11.23 0.000 .3591078 .5108814 phealth 4933361 .0411866 -11.98 0.000 5740603 412612 prog Yes 3624996 .1031408 -3.51 0.000 5646519 1603473 prog age 9341234 .0840002 -11.12 0.000 -1.098761 7694861
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prog age9341234 .0840002 -11.12 0.000 -1.0987617694861
age9341234 .0840002 -11.12 0.000 -1.0987617694861
age9341234 .0840002 -11.12 0.000 -1.0987617694861
phealth .9001108 .0504804 17.83 0.000 .801171 .9990507
wtprog 1.631615 .0780834 20.90 0.000 1.478574 1.784656
_cons .0090842 .0535434 0.17 0.865095859 .1140274
/wchange
cut15897304 .078162674292644365345
cut2 .5029323 .068292 .3690825 .6367821
corr(e.prog, e.wchange)3478179 .0604422 -5.75 0.00046032822243109

Log likelihood	d = -4401.0952	2		Prob >	chi2 =	0.0000
	Coef.	Std. Err.	Z	P> z	[95% Conf.	Interval]
wchange						
age	.2155906	.0705048	3.06	0.002	.0774037	.3537776
over	.4349946	.0387185	11.23	0.000	.3591078	.5108814
phealth	4933361	.0411866	-11.98	0.000	5740603	412612
prog						
Yes	3624996	.1031408	-3.51	0.000	5646519	1603473
prog						
age	9341234	.0840002	-11.12	0.000	-1.098761	7694861
over	-1.058621	.0514252	-20.59	0.000	-1.159412	9578294
phealth	.9001108	.0504804	17.83	0.000	.801171	.9990507
wtprog	1.631615	.0780834	20.90	0.000	1.478574	1.784656
_cons	.0090842	.0535434	0.17	0.865	095859	.1140274
/wchange						
cut1	5897304	.0781626			7429264	4365345
cut2	.5029323	.068292			.3690825	.6367821
					·····	·
corr(e.prog,	2470470	0004400	F 75	0 000	4602000	0040400
e.wchange)	3478179	.0604422	-5.75	0.000	4603282	2243109

• The coefficient on wtprog and its standard error give the impression that the instrument is relevant

						.3690825	.6367821
corr(e.p e.wcha	rog, nge)	3478179	.0604422	-5.75	0.000	4603282	2243109

- The nonzero correlation between e.prog and e.wchange indicates that prog is endogenous
- Those who are more likely to participate are more likely to lose weight

. margins r.prog, ///
<pre>> predict(fix(prog) outlevel("Loss")) ///</pre>
<pre>> predict(fix(prog) outlevel("No change")) ///</pre>
<pre>> predict(fix(prog) outlevel("Gain")) ///</pre>
> contrast(nowald)
Contrasts of predictive margins Model VCE : OIM
1predict : Pr(wchange==Loss), predict(fix(prog) outlevel("Loss"))
<pre>2predict : Pr(wchange==No change), predict(fix(prog) outlevel("No change"))</pre>
3predict : Pr(wchange==Gain), predict(fix(prog) outlevel("Gain"))
Delta-method
Contrast Std. Err. [95% Conf. Interval]

- prog@_predict (Yes vs No) 1 .0356631 .0560914 .1958883 .1259899 (Yes vs No) 2 -.0293965 -.0185024.0055583 -.0076084(Yes vs No) 3 -.1074874.0306512 -.1675628-.0474121
 - When everyone joins the program instead of when no one participants in the program,
 - On average, the probablity of "Loss" goes up by .13
 - On average, the probablity of "No change" goes down by .02
 - \bullet On average, the probablity of "Gain" goes down by .11

- fix(prog) gets us the effect of the program that is not contaminated by the selection effect/correlation between ϵ and η that increases the participation among people more likely to lose weight
- predict(fix(prog)) tells margins to specify fix(prog) to predict when computing each predicted probability
- fix(prog) causes the value of prog not to affect ϵ , even though they are correlated
 - fix(prog) specifies that the part of ϵ that is correlated with y2 be integrated out

- This type of prediction is sometimes called the structural prediction or an average structural function; see Blundell and Powell (2003), Blundell and Powell (2004), Wooldridge (2010), and Wooldridge (2014),
- The difference between the mean of the average of the structural predictions when prog=1 and the mean of the average of the structural predictions when prog=0 is an average treatment effect (Blundell and Powell (2003) and Wooldridge (2014))

Standard errors for population versus sample

- The delta-method standard errors reported by margins hold the covariates fixed at their sample values
 - The delta-method standard errors are for a sample-average treatment effect instead of a population-averaged treatment effect
 - The sample-averaged treatment effect is for those individuals that showed up in that run of the treatment
 - The population-averaged treatment effect is for a random draw of individuals from the population
- To get standard errors for the population-average treatment effect, specify vce(robust) to the estimation command and specify vce(unconditional) to margins

<pre>. quietly eoprobit wchange age over phealth , /// > endog(prog = age over phealth wtprog, probit) /// > vce(robust)</pre>
. margins r.prog, ///
<pre>> predict(fix(prog) outlevel("Loss")) ///</pre>
<pre>> predict(fix(prog) outlevel("No change")) ///</pre>
<pre>> predict(fix(prog) outlevel("Gain")) ///</pre>
<pre>> contrast(nowald) vce(unconditional)</pre>
Contrasts of predictive margins
<pre>1predict : Pr(wchange==Loss), predict(fix(prog) outlevel("Loss"))</pre>
<pre>2predict : Pr(wchange==No change), predict(fix(prog) outlevel("No change"))</pre>
3predict : Pr(wchange==Gain), predict(fix(prog) outlevel("Gain"))

	Contrast	Unconditional Std. Err.	[95% Conf.	Interval]
prog@_predict (Yes vs No) 1 (Yes vs No) 2 (Yes vs No) 3	.1259899 0185024 1074874	.0349061 .0054389 .0300866	.0575753 0291624 1664561	.1944045 0078424 0485188

. matrix b = r(b)

- The commands eregress, eprobit, and eintreg fit ERMs handle continuous-and-unbounded, binary, and censored/corner outcomes
- Look at

http://www.stata.com/manuals/erm.pdf

for more examples and a wealth of details

Blundell, R. W., and J. L. Powell. 2003. Endogeity in nonparametric and semiparametric regression models. In Advances in Economics and Econometrics: Theory and Applications, Eighth World Congress, ed. M. Dewatripont, L. P. Hansen, and S. J. Turnovsky, vol. 2, 312–357. Cambridge: Cambridge University Press.

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Wooldridge, J. M. 2010. *Econometric Analysis of Cross Section and Panel Data*. 2nd ed. Cambridge, Massachusetts: MIT Press.

——. 2014. Quasi-maximum likelihood estimation and testing for nonlinear models with endogenous explanatory variables. *Journal of Econometrics* 182: 226–234.