Estimating the ATE of an endogenously assigned treatment from a sample with endogenous selection by regression adjustment using an extended regression models

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- Fictional data on wellness program from large company

```
. // use wprogram
. use wprogram2
. describe
Contains data from wprogram2.dta
  obs:          3,000
 vars:              8                              28 Jul 2017 07:13
 size:         96,000
─────────────────────────────────────────────────────────────────────
              storage   display    value
variable name   type    format     label     variable label
─────────────────────────────────────────────────────────────────────
wchange        float    %9.0g      changel    Weight change level
age            float    %9.0g                 Years over 50
over           float    %9.0g                 Overweight (tens of pounds)
phealth        float    %9.0g                 Prior health score
prog           float    %9.0g      yesno      Participate in wellness program
wtprog         float    %9.0g      yesno      Offered work time to participate
                                                in program
wtsamp         float    %9.0g                 Offered work time to participate
                                                in sample
insamp         float    %9.0g                 In sample: attended initial and
                                                final weigh in
─────────────────────────────────────────────────────────────────────
Sorted by:
```

- Three levels of wchange

```
. tabulate wchange prog
   Weight |    Participate in
   change |   wellness program
    level |        No        Yes |     Total
----------+----------------------+----------
     Loss |       154        960 |     1,114
No change |       251        299 |       550
     Gain |       184         36 |       220
----------+----------------------+----------
    Total |       589      1,295 |     1,884
```

- Data are observational
- Table does not account for
    - how observed covariates that affect program participation also affect the potential outcome variables
    - how observed unobserved error that affect program participation also affect the potential outcome variables
    - the possibility that unobserved errors in the process caused some of 3,000 individuals not to show for the final weigh in may also affect the potential outcome variables

Ordinal Potential outcomes

- Because the outcome wchange is ordinal, there are really three binary outcomes wchange=="Loss", wchange=="No Change", and wchange=="Gain"
- In the potential outcome framework, there is an outcome for each person when they participate and when the do not participate
- Thus, there are really three binary outcomes for each potential outcome

Participate			Not participate		
wchange_p	==	"Loss"	wchange_{np}	==	"Loss"
wchange_p	==	"No change"	wchange_{np}	==	"No change"
wchange_p	==	"Gain"	wchange_{np}	==	"Gain"

Potential outcome framework

- For each outcome (Loss, No change, and Gain), we only observe one of these two potential outcomes for each individual
- We estimate the parameters of a model and use the estimated parameters to predict what each person does in the unobserved potential outcome

Average treatment effects

- In the case of one outcome, the Average treatment effect (ATE) is

$$\mathbf{E}[y_p - y_{np}]$$

- As there are three outcomes, there are three ATEs
  There are three ATEs one for "Loss", one for "No Change", and one for "Gain"

$$ATE_{Loss} = \mathbf{E}[\texttt{wchange}_p == \text{"Loss"} - \texttt{wchange}_{np} == \text{"Loss"}]$$

$$ATE_{Nochange} = \mathbf{E}[\texttt{wchange}_p == \text{"No change"} -$$

$$\texttt{wchange}_{np} == \text{"No change"}]$$

$$ATE_{Gain} = \mathbf{E}[\texttt{wchange}_p == \text{"Gain"} - \texttt{wchange}_{np} == \text{"Gain"}]$$

Average treatment effects

- I will provide some details about the average treatment effect for "Loss"
- The details for the outcomes of "No change" and "Gain" are analogous

- the average treatment effect (ATE) of the program on the Loss outcome ATE_{Loss}

$$ATE_{Loss} = \mathbf{E}[\texttt{wchange}_p == \text{``Loss''} - \texttt{wchange}_{np} == \text{``Loss''}]$$
$$= \mathbf{E}[\texttt{wchange}_p == \text{``Loss''}] - \mathbf{E}[\texttt{wchange}_{np} == \text{``Loss''}]$$
$$= \text{Pr}[\texttt{wchange}_p == \text{``Loss''}] - \text{Pr}[\texttt{wchange}_{np} == \text{``Loss''}]$$

- The first line says that ATE_{Loss} is the mean diffence in the outcomes when everyone participates instead of no one participates
- The second line says that the mean of the differences is the difference in the means
- The third line uses the fact that mean of binary outcome is the probability that the event is true, to say that the ATE_{Loss} is the difference in the the probability of being in the state of "Loss" when everyone participates instead of no one participates

- I am going to use the ERM comand `eoprobit` to estimate the parameters of $\Pr[\text{wchange}_p == \text{``Loss''} | \mathbf{x}]$ and $\Pr[\text{wchange}_{np} == \text{``Loss''} | \mathbf{x}]$ and
- Then I use `margins` to estimate

$$\mathbf{E}[\Pr[\text{wchange}_p == \text{``Loss''} | \mathbf{x}]] - \mathbf{E}[\Pr[\text{wchange}_{np} == \text{``Loss''} | \mathbf{x}]]$$
$$= \Pr[\text{wchange}_p == \text{``Loss''}] - \Pr[\text{wchange}_{np} == \text{``Loss''}]$$
$$= ATE_{Loss}$$

- The ATE_{Loss} is the mean difference in the the probability of being in the state of "Loss" when everyone participates instead of no one participates

Models for the ordinal outcome

- For exogenous treatment, we do a one-step equivalent to fitting two separate ordinal probit models
  - One fit to partipants
  - Another fit to not partipants

$$wchange = \begin{cases} \text{``Loss''} & \text{if} \quad \mathbf{x}\boldsymbol{\beta}_0 + \epsilon_0 \leq cut1_0 \\ \text{``No change''} & \text{if } cut1_0 < \mathbf{x}\boldsymbol{\beta}_0 + \epsilon_0 \leq cut2_0 \\ \text{``Gain''} & \text{if } cut2_0 < \mathbf{x}\boldsymbol{\beta}_0 + \epsilon_0 \end{cases}$$

$$\mathbf{x}\boldsymbol{\beta}_0 = \beta_{1,0}\texttt{age} + \beta_{2,0}\texttt{over} + \beta_{3,0}\texttt{phealth}$$

for the observations at which `prog`=0, and

$$wchange = \begin{cases} \text{``Loss''} & \text{if} \quad \mathbf{x}\boldsymbol{\beta}_1 + \epsilon_1 \leq cut1_1 \\ \text{``No change''} & \text{if } cut1_1 < \mathbf{x}\boldsymbol{\beta}_1 + \epsilon_1 \leq cut2_1 \\ \text{``Gain''} & \text{if } cut2_1 < \mathbf{x}\boldsymbol{\beta}_1 + \epsilon_1 \end{cases}$$

$$\mathbf{x}\boldsymbol{\beta}_1 = \beta_{1,1}\texttt{age} + \beta_{2,1}\texttt{over} + \beta_{3,1}\texttt{phealth}$$

for the observations at which `prog`=1

ϵ_0, and ϵ_1 are normal

$corr(\epsilon_0, \epsilon_1)$ is not identified or estimated

Exogenous treatment

```
. eoprobit wchange age over phealth, extreat(prog) vsquish nolog
Extended ordered probit regression              Number of obs    =      1,884
                                                Wald chi2(6)     =      99.08
Log likelihood = -1434.5465                     Prob > chi2      =     0.0000
```

wchange	Coef.	Std. Err.	z	P>|z|	[95% Conf. Interval]	
prog#c.age						
No	.2180787	.1464522	1.49	0.136	-.0689623	.5051196
Yes	-.2356064	.1196215	-1.97	0.049	-.4700603	-.0011526
prog#c.over						
No	.2156394	.0784599	2.75	0.006	.0618609	.3694179
Yes	-.0352986	.0781835	-0.45	0.652	-.1885355	.1179383
prog#						
c.phealth						
No	-.0746153	.0844652	-0.88	0.377	-.2401641	.0909334
Yes	-.6229527	.0669733	-9.30	0.000	-.7542181	-.4916874
/wchange						
prog#c.cut1						
No	-.4960282	.0978731			-.6878559	-.3042005
Yes	.0712884	.0810525			-.0875716	.2301484
prog#c.cut2						
No	.642945	.0988945			.4491153	.8367747
Yes	1.421407	.0984319			1.228484	1.61433

ates store oprobit

ATE

```
. generate prog_original = prog
. replace  prog = 0
(1,700 real changes made)
. predict double pr_loss_0 , outlevel("Loss")
(option pr assumed; predicted probabilities)
. replace  prog = 1
(3,000 real changes made)
. predict double pr_loss_1 , outlevel("Loss")
(option pr assumed; predicted probabilities)
. replace  prog = prog_original
(1,300 real changes made)
. drop prog_original
. mean pr_loss_0 pr_loss_1
Mean estimation                     Number of obs   =      3,000
```

	Mean	Std. Err.	[95% Conf. Interval]	
pr_loss_0	.2721432	.0009077	.2703634	.273923
pr_loss_1	.7096007	.0020206	.7056388	.7135625

```
. estimates restore oprobit
(results oprobit are active now)
. margins prog,                          ///
>         predict(outlevel("Loss"))       ///
>         predict(outlevel("No change"))  ///
>         predict(outlevel("Gain")) noesample
Predictive margins                              Number of obs    =      3,000
Model VCE     : OIM
1._predict    : Pr(wchange==Loss), predict(outlevel("Loss"))
2._predict    : Pr(wchange==No change), predict(outlevel("No change"))
3._predict    : Pr(wchange==Gain), predict(outlevel("Gain"))
```

	Margin	Delta-method Std. Err.	z	P>|z|	[95% Conf. Interval]	
_predict#prog						
1#No	.2721432	.0191116	14.24	0.000	.2346853	.3096012
1#Yes	.7096007	.0142655	49.74	0.000	.6816407	.7375606
2#No	.4260522	.0203869	20.90	0.000	.3860947	.4660097
2#Yes	.25725	.0133175	19.32	0.000	.2311483	.2833518
3#No	.3018046	.0191367	15.77	0.000	.2642973	.3393118
3#Yes	.0331493	.0055184	6.01	0.000	.0223334	.0439652

```
. margins r.prog,                              ///
>         predict(outlevel("Loss"))            ///
>         predict(outlevel("No change"))       ///
>         predict(outlevel("Gain"))            ///
>         contrast(nowald)                     ///
>         noesample
Contrasts of predictive margins
Model VCE   : OIM
1._predict  : Pr(wchange==Loss), predict(outlevel("Loss"))
2._predict  : Pr(wchange==No change), predict(outlevel("No change"))
3._predict  : Pr(wchange==Gain), predict(outlevel("Gain"))
```

| | | Delta-method | | |
	Contrast	Std. Err.	[95% Conf.	Interval]
prog@_predict				
(Yes vs No) 1	.4374574	.0238486	.390715	.4841999
(Yes vs No) 2	-.1688022	.0243512	-.2165296	-.1210748
(Yes vs No) 3	-.2686552	.0199165	-.3076908	-.2296196

- When everyone joins the program instead of when no one participants in the program,
    - On average, the probablity of "Loss" goes up by .44
    - On average, the probablity of "No change" goes down by .17
    - On average, the probablity of "Gain" goes down .27

Endogenous Treatment model

The potential-outcome model for an endogenous treatment

- Allows the coefficients to differ for the treated and not-treated state
- Allows the cut offs to differ for the treated and not-treated state
- Allows for distinct (nonzero) correlations between the errors driving treatment assignment and the errors driving the ordinal outcomes for the treated and not-treated states

$$prog = (\mathbf{x}\boldsymbol{\gamma} + \gamma_1 \texttt{wtprog} + \eta > 0)$$

$$wchange = \begin{cases} \text{``Loss''} & \text{if} \quad \mathbf{x}\boldsymbol{\beta}_0 + \epsilon_0 \leq cut1_0 \\ \text{``No change''} & \text{if } cut1_0 < \mathbf{x}\boldsymbol{\beta}_0 + \epsilon_0 \leq cut2_0 \\ \text{``Gain''} & \text{if } cut2_0 < \mathbf{x}\boldsymbol{\beta}_0 + \epsilon_0 \end{cases}$$

$$\mathbf{x}\boldsymbol{\beta}_0 = \beta_{1,0}\texttt{age} + \beta_{2,0}\texttt{over} + \beta_{3,0}\texttt{phealth}$$

for the observations at which $\texttt{prog}=0$, and

$$wchange = \begin{cases} \text{``Loss''} & \text{if} \quad \mathbf{x}\boldsymbol{\beta}_1 + \epsilon_1 \leq cut1_1 \\ \text{``No change''} & \text{if } cut1_1 < \mathbf{x}\boldsymbol{\beta}_1 + \epsilon_1 \leq cut2_1 \\ \text{``Gain''} & \text{if } cut2_1 < \mathbf{x}\boldsymbol{\beta}_1 + \epsilon_1 \end{cases}$$

$$\mathbf{x}\boldsymbol{\beta}_1 = \beta_{1,1}\texttt{age} + \beta_{2,1}\texttt{over} + \beta_{3,1}\texttt{phealth}$$

for the observations at which $\texttt{prog}=1$

ϵ_0, ϵ_1, and η are correlated and joint normal

ρ_0 correlation between ϵ_0 and η

ρ_1 correlation between ϵ_1 and η

Endogenous treatment model

```
. eoprobit wchange age over phealth ,                              ///
>         entreat(prog = age over phealth wtprog, pocorr ) ///
>         vce(robust) vsquish nolog
Extended ordered probit regression              Number of obs    =       1,884
                                                Wald chi2(6)     =      137.27
Log pseudolikelihood = -2335.2213               Prob > chi2      =      0.0000
```

	Coef.	Robust Std. Err.	z	P>|z|	[95% Conf. Interval]	
wchange						
prog#c.age						
No	.4919782	.1357859	3.62	0.000	.2258427	.7581137
Yes	-.1111304	.1183412	-0.94	0.348	-.3430749	.1208142
prog#c.over						
No	.4659558	.0789709	5.90	0.000	.3111757	.6207359
Yes	.0458895	.0794788	0.58	0.564	-.109886	.2016651
prog# c.phealth						
No	-.3162974	.0872579	-3.62	0.000	-.4873198	-.145275
Yes	-.6880971	.0713535	-9.64	0.000	-.8279474	-.5482467
prog						
age	-.9224146	.1057226	-8.72	0.000	-1.129627	-.7152021
over	-.9957274	.0675412	-14.74	0.000	-1.128106	-.863349
phealth	.7483889	.0604543	12.38	0.000	.6299007	.8668771
wtprog	1.718043	.1160706	14.80	0.000	1.490549	1.945537

	Coef.	Robust Std. Err.	z	P>|z|	[95% Conf. Interval]	
wchange						
prog#c.age						
No	.4919782	.1357859	3.62	0.000	.2258427	.7581137
Yes	-.1111304	.1183412	-0.94	0.348	-.3430749	.1208142
prog#c.over						
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Yes	-.6880971	.0713535	-9.64	0.000	-.8279474	-.5482467
prog						
age	-.9224146	.1057226	-8.72	0.000	-1.129627	-.7152021
over	-.9957274	.0675412	-14.74	0.000	-1.128106	-.863349
phealth	.7483889	.0604543	12.38	0.000	.6299007	.8668771
wtprog	1.718043	.1160706	14.80	0.000	1.490549	1.945537
_cons	.3398047	.0690413	4.92	0.000	.2044863	.475123
/wchange						
prog#c.cut1						
No	.1953761	.1544741			-.1073875	.4981397
Yes	-.133868	.0985578			-.3270377	.0593017
prog#c.cut2						
No	1.193014	.111908			.9736779	1.412349
Yes	1.170747	.1289195			.9180695	1.423425
corr(e.prog, e.wchange)						

prog

Yes	−.1111304	.1183412	−0.94	0.348	−.3430749	.1208142
prog#c.over						
No	.4659558	.0789709	5.90	0.000	.3111757	.6207359
Yes	.0458895	.0794788	0.58	0.564	−.109886	.2016651
prog# **c.phealth**						
No	−.3162974	.0872579	−3.62	0.000	−.4873198	−.145275
Yes	−.6880971	.0713535	−9.64	0.000	−.8279474	−.5482467
prog						
age	−.9224146	.1057226	−8.72	0.000	−1.129627	−.7152021
over	−.9957274	.0675412	−14.74	0.000	−1.128106	−.863349
phealth	.7483889	.0604543	12.38	0.000	.6299007	.8668771
wtprog	1.718043	.1160706	14.80	0.000	1.490549	1.945537
_cons	.3398047	.0690413	4.92	0.000	.2044863	.475123
/wchange						
prog#c.cut1						
No	.1953761	.1544741			−.1073875	.4981397
Yes	−.133868	.0985578			−.3270377	.0593017
prog#c.cut2						
No	1.193014	.111908			.9736779	1.412349
Yes	1.170747	.1289195			.9180695	1.423425
corr(e.prog, **e.wchange)**						
prog						
No	−.6325687	.1073524	−5.89	0.000	−.7992197	−.3755982
Yes	−.4199058	.1042067	−4.03	0.000	−.6015292	−.1970056

```
. estat teffects
Predictive margins                              Number of obs    =      1,884
ATE_Pr0      : Pr(wchange=0=Loss)
ATE_Pr1      : Pr(wchange=1=No change)
ATE_Pr2      : Pr(wchange=2=Gain)
```

	Margin	Unconditional Std. Err.	z	P>|z|	[95% Conf. Interval]	
ATE_Pr0						
prog (Yes vs No)	.0998676	.0620691	1.61	0.108	-.0217856	.2215208
ATE_Pr1						
prog (Yes vs No)	-.013337	.0452574	-0.29	0.768	-.1020399	.0753658
ATE_Pr2						
prog (Yes vs No)	-.0865306	.0222063	-3.90	0.000	-.1300542	-.043007

- When everyone joins the program instead of when no one participants in the program,
  - On average, the probablity of "Loss" goes up by .1
  - On average, the probablity of "No change" does not change by much
  - On average, the probablity of "Gain" goes down .09

```
. margins r.prog,                                          ///
>         predict(fix(prog) outlevel("Loss"))              ///
>         predict(fix(prog) outlevel("No change"))         ///
>         predict(fix(prog) outlevel("Gain"))              ///
>         contrast(nowald) vce(unconditional) noesample
Contrasts of predictive margins
1._predict   : Pr(wchange==Loss), predict(fix(prog) outlevel("Loss"))
2._predict   : Pr(wchange==No change), predict(fix(prog) outlevel("No
                 change"))
3._predict   : Pr(wchange==Gain), predict(fix(prog) outlevel("Gain"))
```

| | | Unconditional | | |
	Contrast	Std. Err.	[95% Conf. Interval]	
prog@_predict				
(Yes vs No) 1	.0998676	.0620691	-.0217856	.2215208
(Yes vs No) 2	-.013337	.0452574	-.1020399	.0753658
(Yes vs No) 3	-.0865306	.0222063	-.1300542	-.043007

Endogenous sample selection

- Reconsider our fictional weight-loss program
  - Some program participants and some nonparticipants will not show up for the final weigh in
    This is commonly known as lost to follow up
  - If unobservables that affect whether someone is lost to follow up
    - are independent of the unobservables that affect program participantion
    - and they are independent of the unobservables that affect the outcomes with and without the program,
  - the previously discussed estimator consistently estimates the effects
- Any dependence among the unobservables must be modeled

$$insamp = (\mathbf{x}\boldsymbol{\alpha} + \alpha_1\mathtt{wtsamp} + \xi > 0)$$

$$prog = (\mathbf{x}\boldsymbol{\gamma} + \gamma_1\mathtt{wtprog} + \eta > 0)$$

$$wchange = \begin{cases} \text{``Loss''} & \text{if} \quad \mathbf{x}\boldsymbol{\beta}_0 + \epsilon_0 \leq cut1_0 \\ \text{``No change''} & \text{if } cut1_0 < \mathbf{x}\boldsymbol{\beta}_0 + \epsilon_0 \leq cut2_0 \\ \text{``Gain''} & \text{if } cut2_0 < \mathbf{x}\boldsymbol{\beta}_0 + \epsilon_0 \end{cases}$$

$$\mathbf{x}\boldsymbol{\beta}_0 = \beta_{1,0}\mathtt{age} + \beta_{2,0}\mathtt{over} + \beta_{3,0}\mathtt{phealth}$$

for the observations at which prog=0, and

$$wchange = \begin{cases} \text{``Loss''} & \text{if} \quad \mathbf{x}\boldsymbol{\beta}_1 + \epsilon_1 \leq cut1_1 \\ \text{``No change''} & \text{if } cut1_1 < \mathbf{x}\boldsymbol{\beta}_1 + \epsilon_1 \leq cut2_1 \\ \text{``Gain''} & \text{if } cut2_1 < \mathbf{x}\boldsymbol{\beta}_1 + \epsilon_1 \end{cases}$$

$$\mathbf{x}\boldsymbol{\beta}_1 = \beta_{1,1}\mathtt{age} + \beta_{2,1}\mathtt{over} + \beta_{3,1}\mathtt{phealth}$$

for the observations at which prog=1

$\xi, \epsilon_0, \ \epsilon_1,$ and η are correlated and joint normal

distinct correlations between each treatment error and others

```
. eoprobit wchange age over phealth ,                                    ///
>         entreat(prog = age over phealth wtprog, pocorr ) ///
>         select(insamp = age over phealth wtsamp )  ///
>         vce(robust) vsquish nolog
Extended ordered probit regression          Number of obs    =      3,000
                                            Selected         =      1,884
                                            Nonselected      =      1,116
                                            Wald chi2(6)     =     163.70
Log pseudolikelihood = -4483.9683           Prob > chi2      =     0.0000
```

	Coef.	Robust Std. Err.	z	P>|z|	[95% Conf. Interval]
wchange					
prog#c.age					
No	.4174575	.1335097	3.13	0.002	.1557832	.6791318
Yes	-.0779536	.1120819	-0.70	0.487	-.2976301	.141723
prog#c.over					
No	.5046857	.0836683	6.03	0.000	.3406989	.6686725
Yes	.1930521	.0973183	1.98	0.047	.0023118	.3837924
prog# c.phealth					
No	-.4250361	.091857	-4.63	0.000	-.6050726	-.2449996
Yes	-.8098627	.0753678	-10.75	0.000	-.9575809	-.6621444
insamp					
age	-.0231005	.0805424	-0.29	0.774	-.1809607	.1347597
over	-.7639994	.0450909	-16.94	0.000	-.852376	-.6756229
phealth	.7765721	.0467569	16.61	0.000	.6849303	.8682139
wtsamp	2.611108	.2660121	9.82	0.000	2.089734	3.132483
_cons	.2832551	.0516926	5.48	0.000	.1819395	.3845707

c.phealth						
No	-.4250361	.091857	-4.63	0.000	-.6050726	-.2449996
Yes	-.8098627	.0753678	-10.75	0.000	-.9575809	-.6621444
insamp						
age	-.0231005	.0805424	-0.29	0.774	-.1809607	.1347597
over	-.7639994	.0450909	-16.94	0.000	-.852376	-.6756229
phealth	.7765721	.0467569	16.61	0.000	.6849303	.8682139
wtsamp	2.611108	.2660121	9.82	0.000	2.089734	3.132483
_cons	.2832551	.0516926	5.48	0.000	.1819395	.3845707
prog						
age	-.9371024	.0818803	-11.44	0.000	-1.097585	-.7766199
over	-1.060975	.0492229	-21.55	0.000	-1.15745	-.9645
phealth	.890558	.0494954	17.99	0.000	.7935487	.9875673
wtprog	1.644504	.0731516	22.48	0.000	1.501129	1.787878
_cons	.0153225	.0527572	0.29	0.771	-.0880796	.1187247
/wchange						
prog#c.cut1						
No	-.2754667	.1708586			-.6103433	.05941
Yes	-.4323606	.1401249			-.7070003	-.1577208
prog#c.cut2						
No	.6797857	.1534354			.3790578	.9805137
Yes	.7803365	.2260056			.3373737	1.223299
corr(e.ins~p,						
e.wchange)						
prog						
No	-.5779184	.1004465	-5.75	0.000	-.7420068	-.3484981
Yes	-.5355424	.1948537	-2.75	0.006	-.81217	-.0623165
corr(e.prog,						
~change)						

prog#c.cut1						
No	-.2754667	.1708586			-.6103433	.05941
Yes	-.4323606	.1401249			-.7070003	-.1577208
prog#c.cut2						
No	.6797857	.1534354			.3790578	.9805137
Yes	.7803365	.2260056			.3373737	1.223299
corr(e.ins~p, e.wchange)						
prog						
No	-.5779184	.1004465	-5.75	0.000	-.7420068	-.3484981
Yes	-.5355424	.1948537	-2.75	0.006	-.81217	-.0623165
corr(e.prog, e.wchange)						
prog						
No	-.6031412	.1119322	-5.39	0.000	-.7790275	-.3392526
Yes	-.4940044	.0934446	-5.29	0.000	-.6547774	-.2904625
corr(e.prog, e.insamp)	.4745668	.0298397	15.90	0.000	.4140283	.5309257

- Nonzero correlations between e.insamp and e.wchange imply endogenous sample selection for outcomes
- Nonzero correlations between e.prog and e.wchange imply endogenous treatment assignment

```
. estat teffects
Predictive margins                                    Number of obs    =      3,000
ATE_Pr0      : Pr(wchange=0=Loss)
ATE_Pr1      : Pr(wchange=1=No change)
ATE_Pr2      : Pr(wchange=2=Gain)
```

	Margin	Unconditional Std. Err.	z	P>|z|	[95% Conf. Interval]	
ATE_Pr0						
prog						
(Yes vs No)	.1406344	.0785061	1.79	0.073	-.0132346	.2945035
ATE_Pr1						
prog						
(Yes vs No)	.0210902	.0369635	0.57	0.568	-.0513569	.0935372
ATE_Pr2						
prog						
(Yes vs No)	-.1617246	.0642328	-2.52	0.012	-.2876187	-.0358305

- When everyone joins the program instead of when no one participants in the program,
  - On average, the probablity of "Loss" goes up by .14
  - On average, the probablity of "No change" does not change
  - On average, the probablity of "Gain" goes down .16

```
. margins r.prog,                                           ///
>         predict(fix(prog) outlevel("Loss"))               ///
>         predict(fix(prog) outlevel("No change"))          ///
>         predict(fix(prog) outlevel("Gain"))               ///
>         contrast(nowald) vce(unconditional) noesample
Contrasts of predictive margins
1._predict    : Pr(wchange==Loss), predict(fix(prog) outlevel("Loss"))
2._predict    : Pr(wchange==No change), predict(fix(prog) outlevel("No
                change"))
3._predict    : Pr(wchange==Gain), predict(fix(prog) outlevel("Gain"))
```

	Contrast	Unconditional Std. Err.	[95% Conf. Interval]	
prog@_predict				
(Yes vs No) 1	.1406344	.0785061	-.0132346	.2945035
(Yes vs No) 2	.0210902	.0369635	-.0513569	.0935372
(Yes vs No) 3	-.1617246	.0642328	-.2876187	-.0358305

- When everyone joins the program instead of when no one participants in the program,

  - On average, the probablity of "Loss" goes up by .14
  - On average, the probablity of "No change" does not change
  - On average, the probablity of "Gain" goes down .16

More about ERM commands

- The commands `eregress`, `eprobit`, and `eintreg` fit ERMs handle continuous-and-unbounded, binary, and censored/corner outcomes
- Look at

    http://www.stata.com/manuals/erm.pdf

    for more examples and a wealth of details