

Marginal Effects in Multiply Imputed Datasets

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Multiple Imputation

The `mi` commands

`mi`

- introduced in Stata11
- imputes missing values
 - `mi impute`
- fits models to imputed data
 - `mi estimate`
- some postestimation commands
 - `mi test`
- performs data management tasks
 - `mi xeq`, `mi passive`, ...

Marginal effects and adjusted predictions

The margins command

margins

- introduced in Stata11
- estimates marginal effects
 - as derivatives or (semi-)elasticities
- estimates adjusted predictions
- replaces old `mf` and `adjust`

Factor variable notation

- introduced with and required by `margins`
- creates indicator variables, higher order terms, interactions
- replaces old `xi`

Subsequent enhancements

`mi`

- chained equations
- user imputation methods
- predictions

`margins`

- `marginsplot` visualizes results
- pairwise comparisons
 - `pwcompare`
- contrasts

So, what is missing?

Setting up example data

```
version 12.1

webuse nhanes2d , clear

svyset , clear
drop if (hlthstat > 5)

set seed 42

mi set mlong
mi register imputed vitaminc zinc
mi impute chained ///
    (pmm , knn(5)) vitaminc zinc ///
    = hlthstat i.sex i.ageg i.psu finalwgt i.strata ///
    , add(5) noisily

mi svyset psu [pweight = finalwgt] , strata(strata)

save nhanes2d_imputed5.dta , replace
```

So, what is missing?

Combining mi ...

```
. mi estimate : svy : mlogit hlthstat i.sex vitaminc i.agegrp zinc
Multiple-imputation estimates      Imputations      =          5
Survey: Multinomial logistic regression  Number of obs    =       10335
Number of strata =          31      Population size   =  116997257
Number of PSUs   =          62
                                     Average RVI       =       0.2784
                                     Largest FMI       =       0.1845
                                     Complete DF      =          31
DF adjustment:   Small sample      DF:   min        =       21.03
                                     avg            =       27.89
                                     max            =       29.17
Model F test:    Equal FMI          F( 32, 28.1)    =       197.60
Within VCE type: Linearized        Prob > F         =       0.0000
```

	hlthstat	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
1							
	2.sex	-.2535739	.0757789	-3.35	0.002	-.4085658	-.098582
	vitaminc	.3783082	.0836045	4.52	0.000	.207102	.5495144
	agegrp						
	2	-.2417474	.1129682	-2.14	0.041	-.4727367	-.0107581
	3	-.5148524	.116651	-4.41	0.000	-.75337	-.2763347
	4	-1.158049	.0944494	-12.26	0.000	-1.351177	-.9649211
	5	-1.464487	.1040847	-14.07	0.000	-1.677314	-1.251659
	6	-1.236679	.1044329	-11.84	0.000	-1.45025	-1.023107
	zinc	.0069707	.0029981	2.33	0.028	.0008261	.0131153
	_cons	-.3710475	.2991727	-1.24	0.225	-.9837576	.2416626

(output omitted)

So, what is missing?

...and margins

(output omitted)

		(base outcome)					
3		(output omitted)					
5							
2.sex		-.0199766	.1179985	-0.17	0.867	-.2613745	.2214214
vitaminC		-.9080216	.1067354	-8.51	0.000	-1.129973	-.6860703
agegrp							
2		.2911128	.3417395	0.85	0.401	-.4076501	.9898758
3		1.109931	.2787963	3.98	0.000	.5398748	1.679988
4		1.577982	.2523318	6.25	0.000	1.062038	2.093925
5		2.092809	.2107573	9.93	0.000	1.66184	2.523779
6		2.483614	.2154178	11.53	0.000	2.043079	2.924149
zinc		-.0107241	.0038114	-2.81	0.010	-.0186396	-.0028087
_cons		-1.316412	.3994441	-3.30	0.003	-2.139385	-.493439

```
.
. margins , vce(unconditional) dydx(*) predict(outcome(1))
last estimates not found
r(301);
```


General considerations

How to?

Technically,

- run both, the estimation command and `margins` on each imputed dataset
- post estimates to $e(b)$ and $e(V)$
- combine results according to Rubin's rules
 - better yet: let `mi estimate` do the work

Rubin's rules

A brief review

Point estimate

$$\bar{q}_{mi} = \frac{1}{M} \sum_{i=1}^M \hat{q}_i$$

Variance-covariance matrix

$$\text{VCE}_{mi} = \bar{W} + \left(1 + \frac{1}{M}\right) \mathbf{B}$$

with

$$\bar{W} = \frac{1}{M} \sum_{i=1}^M \hat{W}_i$$

$$\mathbf{B} = \left(\frac{1}{M-1}\right) \sum_{i=1}^M (\hat{q}_i - \bar{q}_{mi}) (\hat{q}_i - \bar{q}_{mi})'$$

Requirements

Applicability of Rubin's rules

Statistically, $q \dots$

- estimates population parameter
- does not depend on sample size
- is (asymptotic) normal

Satisfied by, e.g.,

- regression coefficients
- linear predictor
- average marginal effects (in large samples)

The proposed solution

Isabel Cañette and Yulia Marchenko

```
program emargins , eclass properties(mi)
    version 11
    args outcome
    svy : mlogit hlthstat i.sex vitaminc i.ageg zinc
    margins , vce(unconditional) dydx(*) ///
        predict(outcome(`outcome`)) post
end
```

Let's try

```
. forvalues j = 1/5 {
  2.      mi estimate : emargins `j'
  3. }
```

```
Multiple-imputation estimates          Imputations      =          5
Average marginal effects              Number of obs    =       10335

Number of strata =                   31
Number of PSUs  =                   62

                                     Average RVI      =    0.0209
                                     Largest FMI      =    0.0906
                                     Complete DF     =         31
DF adjustment:  Small sample         DF:   min       =    25.74
                                     avg           =    28.64
                                     max           =    29.17

Within VCE type:  Linearized
```

	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
2.sex	-.0506149	.011062	-4.58	0.000	-.073242	-.0279878
vitaminC	.0746023	.0103436	7.21	0.000	.0534397	.0957648
agegrp						
2	-.0283339	.0185412	-1.53	0.137	-.0662458	.0095781
3	-.0703962	.0191397	-3.68	0.001	-.1095317	-.0312607
4	-.1800159	.0154733	-11.63	0.000	-.2116561	-.1483758
5	-.2412945	.0155433	-15.52	0.000	-.2730762	-.2095128
6	-.2335327	.0144687	-16.14	0.000	-.2631215	-.203944
zinc	.0010539	.0004302	2.45	0.021	.0001692	.0019386

(output omitted)

Let's try

(output omitted)

```

Multiple-imputation estimates      Imputations      =      5
Average marginal effects          Number of obs    =     10335

Number of strata =                31
Number of PSUs  =                62

                                Average RVI        =     0.0683
                                Largest FMI         =     0.2394
                                Complete DF         =      31
DF adjustment: Small sample      DF: min          =     18.37
                                avg                =     26.98
                                max                =     29.17

Within VCE type: Linearized

```

	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
2.sex	-.0006929	.0044593	-0.16	0.878	-.0098181	.0084322
vitaminC	-.0396292	.0052561	-7.54	0.000	-.0504941	-.0287643
agegrp						
2	.005949	.0042071	1.41	0.168	-.0026535	.0145515
3	.0314414	.0075778	4.15	0.000	.015947	.0469358
4	.0666996	.0083927	7.95	0.000	.0495386	.0838605
5	.103145	.0093551	11.03	0.000	.0840132	.1222768
6	.1286984	.0152025	8.47	0.000	.0976045	.1597924
zinc	-.0005411	.0001548	-3.50	0.003	-.0008658	-.0002164

Adapting the general approach

A simple logit model

```
mi xeq : generate byte goodhlth = (hlthstat < 3)

program emargins2 , eclass properties(mi)
    version 11
    svy : logit goodhlth i.sex vitaminc i.ageg zinc
    margins , vce(unconditional) dydx(*) post
end

. mi estimate : emargins2
varlist specification required
r(198);
```

Adapting the general approach

Pairwise comparisons

```
capture program drop emargins2
program emargins2 , eclass properties(mi)
    version 11
    svy : logit goodhlth i.sex vitaminc i.ageg zinc
    margins i.ageg , vce(unconditional) ///
    post pwcompare
end
```

```
. mi estimate : emargins2 whatever
```

```
Multiple-imputation estimates          Imputations      =          5
Pairwise comparisons of predictive margins
Number of strata =          31
Number of PSUs   =          62

Average RVI = 0.0017
Largest FMI = 0.0078
Complete DF = 31
DF: min     = 29.05
      avg    = 29.14
      max    = 29.16

DF adjustment: Small sample
Within VCE type: Linearized
```

	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
agegrp						
1	.7059828	.0106223	66.46	0.000	.684263	.7277026
2	.6568665	.0195888	33.53	0.000	.6168121	.6969209
3	.5459818	.0152304	35.85	0.000	.5148394	.5771241
4	.4056276	.014571	27.84	0.000	.375833	.4354221
5	.3318254	.01227	27.04	0.000	.3067364	.3569145
6	.3026305	.0155862	19.42	0.000	.2707557	.3345053

Generalizing the proposed solution

`mimrgns`

Goals

- syntax similar to `margins`
- output similar to `margins`
- no hard coded estimation command

Replicating the original example

```
. mi estimate : svy : mlogit hlthstat i.sex vitaminc i.agegrp zinc
```

(output omitted)

```
. mimrgns , vce(unconditional) dydx(*) predict(outcome(1))
```

```
Multiple-imputation estimates      Imputations      =          5
Average marginal effects          Number of obs    =       10335
```

```
Number of strata =          31
Number of PSUs  =          62
```

```
Average RVI      =    0.0209
Largest FMI      =    0.0906
Complete DF      =           31
DF:  min         =    25.74
     avg         =    28.64
     max         =    29.17
```

```
DF adjustment:  Small sample
```

```
Within VCE type:  Linearized
```

```
Expression : Pr(hlthstat==1), predict(outcome(1))
```

```
dy/dx w.r.t. : 2.sex vitaminc 2.agegrp 3.agegrp 4.agegrp 5.agegrp 6.agegrp zinc
```

	dy/dx	Std. Err.	t	P> t	[95% Conf. Interval]	
2.sex	-.0506149	.011062	-4.58	0.000	-.073242	-.0279878
vitaminc	.0746023	.0103436	7.21	0.000	.0534397	.0957648
agegrp						
2	-.0283339	.0185412	-1.53	0.137	-.0662458	.0095781
3	-.0703962	.0191397	-3.68	0.001	-.1095317	-.0312607
4	-.1800159	.0154733	-11.63	0.000	-.2116561	-.1483758
5	-.2412945	.0155433	-15.52	0.000	-.2730762	-.2095128
6	-.2335327	.0144687	-16.14	0.000	-.2631215	-.203944
zinc	.0010539	.0004302	2.45	0.021	.0001692	.0019386

Note: dy/dx for factor levels is the discrete change from the base level.

The logit model

```
. mi estimate : svy : logit goodhlth i.sex vitaminc i.ageg zinc
```

```
Multiple-imputation estimates      Imputations      =          5
Survey: Logistic regression        Number of obs    =       10335
Number of strata =                  31      Population size  =  116997257
Number of PSUs   =                   62

                                Average RVI       =       0.0340
                                Largest FMI        =       0.1144
                                Complete DF       =          31
DF adjustment:  Small sample      DF:      min     =       24.57
                                           avg         =       27.90
                                           max         =       29.16
Model F test:      Equal FMI      F( 8, 29.0)    =       93.99
Within VCE type:  Linearized      Prob > F       =       0.0000
```

goodhlth	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
2.sex	-.21535	.061994	-3.47	0.002	-.3421322	-.0885678
vitaminc	.5040865	.0639858	7.88	0.000	.3727249	.6354481
agegrp						
2	-.2307692	.0919848	-2.51	0.018	-.4188591	-.0426792
3	-.7052038	.0684112	-10.31	0.000	-.8450875	-.5653201
4	-1.284035	.0787597	-16.30	0.000	-1.445081	-1.122989
5	-1.608576	.073101	-22.00	0.000	-1.758052	-1.459101
6	-1.746134	.0925375	-18.87	0.000	-1.935379	-1.556888
zinc	.0082014	.0020081	4.08	0.000	.0040671	.0123356
_cons	-.2285908	.1915549	-1.19	0.244	-.6234565	.166275

Pairwise comparisons

```
. mimrgns i.ageg , predict(pr) vce(unconditional) pwcompare
```

```
Multiple-imputation estimates      Imputations      =          5
Pairwise comparisons of predictive margins  Number of obs    =       10335

Number of strata =          31
Number of PSUs  =          62

                                Average RVI      =       0.0017
                                Largest FMI       =       0.0078
                                Complete DF        =          31
DF adjustment:  Small sample      DF:  min       =       29.05
                                avg               =       29.14
                                max               =       29.16

Within VCE type:  Linearized

Expression   : Pr(goodhlth), predict(pr)
```

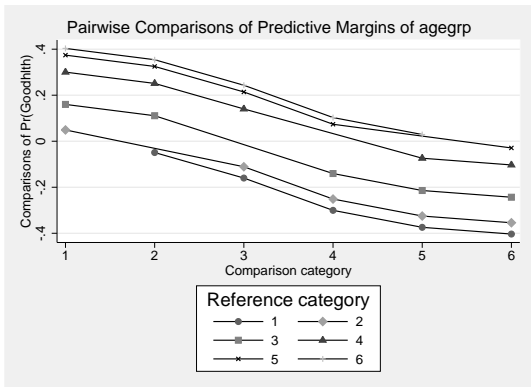
	Contrast	Std. Err.	[95% Conf. Interval]	
agegrp				
2 vs 1	-.0491163	.0201039	-.0902249	-.0080076
3 vs 1	-.1600011	.0156144	-.191928	-.1280741
4 vs 1	-.3003552	.0176029	-.3363501	-.2643604
5 vs 1	-.3741574	.0152539	-.4053472	-.3429676
6 vs 1	-.4033523	.0189403	-.4420859	-.3646187
3 vs 2	-.1108848	.0276638	-.1674514	-.0543182
<i>(output omitted)</i>				
6 vs 4	-.1029971	.019452	-.1427737	-.0632204
6 vs 5	-.0291949	.0156898	-.0612829	.002893

Plotting results

- ```
. mimrgns i.ageg , predict(pr) vce(unconditional) pwcompare cmdmargins
(output omitted)
. marginsplot , noci
```

Variables that uniquely identify margins: `_pw1` `_pw0`

`_pw` enumerates all pairwise comparisons; `_pw0` enumerates the reference categories; `_pw1` enumerates the comparison categories.



# What is the catch?

## Unresolved issues

### Statistical

- non-linear (adjusted) predictions
  - Rubin's rules applicable?
- higher order terms and interactions
  - passive imputation vs. JAV
- ...

# What is the catch?

## Unresolved issues

### Technical

- non-linear (adjusted) predictions
  - transformations not easily implemented
- higher order terms and interactions
  - `margins` relies on factor variables
  - JAV approach not feasible
- execution time
  - need to run estimation command and `margins`  $M$  times
  - cf. Miles (2015)
- ...

# Summary

- some of `margins'` results can be combined using Rubin's rules
  - e.g., linear predictions, average marginal effects
- `mimrgns` makes this easy
- there are still unresolved statistical and technical issues
- produced results are not guaranteed to be valid



## References

- Cañette, I and Marchenko, Y. (2010). Re: st: Average marginal effects for a multiply imputed complex survey. Statalist.
- Miles, A. (2015). Obtaining Predictions from Models Fit to Multiply-Imputed Data. Sociological Methods and Research 45(1):175-185.