## Predictive Margins and Marginal Effects in Stata

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### 11th German Stata Users Group meeting Potsdam, June 7, 2013

### Outline

- Motivation
- margins and marginsplot
- regplot

## Motivation

- For a long time, regression tables have been the preferred way of communicating results from statistical models.
- However, interpretation of regression tables can be very challenging in the case of interaction effects, categorical variables, or nonlinear functional forms.
- Moreover, interpretational difficulties can be overwhelming in nonlinear models such as logistic regression. In these models the raw coefficients are often not of much interest; what we want to see for interpretation are effects on outcomes such as probabilities, not on "latent" variables such as log odds.
- Fortunately, Stata has a number of handy commands such as margins, contrasts, and marginsplot for making sense of regression results.

## Example: Factorial Survey on Just Incomes

- Mail survey among a random sample of the Swiss population (N = 1945). Written questionnaire in German, French and Italian.
  - Data collected in fall 2010 as part of a follow-up survey to the "Swiss Environmental Survey 2007"

(see http://www.socio.ethz.ch/research/umweltsurvey/umweltsurvey2007)

- Respondents were asked to judge short text descriptions of (fictional) individuals (so called "vignettes"), in which certain elements are varied at random.
- For our research objective, we used vignettes describing men and women employing the following 2 × 2 × 2 × 3 design :
  - male vs. female
  - single without children vs. married without children
  - average work effort vs. above-average work effort
  - ▶ income levels: 5000 CHF, 5500 CHF, 6000 CHF

## Example: The Vignette

In letzter Zeit wird viel über die Höhe von Löhnen in verschiedenen Berufen gesprochen. Wir interessieren uns für Ihre persönliche Einschätzung zu diesem Thema.

Stellen Sie sich die folgende Situation vor:

{Herr | Frau} Müller, 25-jährig, {allein stehend und ohne Kinder | verheiratet in kinderloser Ehe}, arbeitet als kaufmännische{r|} Angestellte{r|} im Rechnungswesen eines mittleren Dienstleistungsbetriebs und erbringt dort {überdurchschnittliche | durchschnittliche} Leistungen. {Sein | Ihr} monatliches Bruttoeinkommen beträgt {5'000 | 5'500 | 6'000} Franken.

Wie bewerten Sie das Einkommen dieser Person? Ist das Einkommen Ihrer Meinung nach gerecht oder ist es ungerechterweise zu hoch oder zu niedrig?

viel zu niedrig		gerecht						viel zu hoch		
-5	-4	-3	-2	-1	0	+1	+2	+3	+4	+5

### Example: The Data

. use vignette (2010 Vignette		n Just Incom	nes)	
. d				
Contains data	from vig	nettes.dta		
obs:	1,482			2010 Vignette Study on Just Incomes
vars:	13			10 Jun 2013 11:16
size:	51,870			(_dta has notes)
	storage	display	value	
variable name	type	format	label	variable label
vrating	double	%13.0g	vrating	Vignette: rating (-5=much too low, 5=much too high)
vmale	byte	%8.0g	vmale	Vignette: male
vmarried	byte	%8.0g	vmarried	Vignette: married
veffort	byte	%25.0g	veffort	Vignette: above-average work effort
vinc	int	%10.0g		Vignette: income (CHF per month)
rmale	byte	%8.0g	rmale	Respondent: male
rage	byte	%8.0g		Respondent: age
reducyrs		%10.0g		Respondent: years of education
rright	byte	%8.0g		Respondent: political orientation (0=left, 10=right)
rmarstat	2	%8.0g	rmarstat	Respondent: marital status
rinc		%13.0g	rinc	Respondent: income (CHF per month)
wt	double	%10.0g		sampling weights
strata	byte	%8.0g		sampling strata

Sorted by:

### Example: Analysis of the Vignette Data

• A simple linear regression model with the vignette responses as dependent variable – have fun interpreting!

. regress vrating vinc i.vmale i.vmarried i.veffort ///

> vmale##rmale##c.reducyrs##c.reducyrs, vsquish noheader

[95% Conf.	P> t	t	Std. Err.	Coef.	vrating
.0008225	0.000	10.71	.000094	.0010069	vinc
-7.548853	0.659	-0.44	3.140572	-1.388362	1.vmale
3312517	0.020	-2.33	.0771521	1799115	1.vmarried
8177084	0.000	-8.63	.0771985	6662772	1.veffort
-6.962975	0.910	-0.11	3.355975	3799532	1.rmale
					vmale#rmale
-5.46665	0.408	0.83	4.818208	3.984662	1 1
5476238	0.752	0.32	.3327842	.1051598	reducyrs
					vmale#c.reducyrs
877102	0.881	0.15	.4839479	.0722016	1
					rmale#c.reducyrs
9920235	0.980	0.03	.5123413	.0129762	1
					vmale#rmale#c.reducyrs
-1.882497	0.545	-0.61	.7331748	4443146	1 1
030612	0.607	-0.51	.0123638	0063593	c.reducyrs#c.reducyrs
					vmale#c.reducyrs#c.reducyrs
0340855	0.955	0.06	.0178947	.0010164	1
					rmale#c.reducyrs#c.reducyrs
0355169	0.942	0.07	.0187994	.0013596	1
					vmale#rmale#c.reducyrs#c.reducyrs
0420033	0.693	0.39	.0268027	.0105725	1 1
-9.189904	0.030	-2.17	2.224811	-4.825754	_cons
	.008225 -7.548853 3312517 8177084 -6.962975 -5.46665 5476238 877102 9920235 -1.882497 030612 0340855 0355169 0420033	$\begin{array}{ccccc} 0.000 & .0008225 \\ 0.659 & -7.548853 \\ 0.020 &3312517 \\ 0.000 &8177084 \\ 0.910 & -6.962975 \\ 0.408 & -5.46665 \\ 0.752 &5476238 \\ 0.881 &877102 \\ 0.980 &9920235 \\ 0.545 & -1.882497 \\ 0.607 &030612 \\ 0.955 &0340855 \\ 0.942 &0355169 \\ 0.693 &0420033 \\ \end{array}$	$\begin{array}{cccccccccccccccccccccccccccccccccccc$	$\begin{array}{c ccccccccccccccccccccccccccccccccccc$	.0010069         .000094         10.71         0.000           .1383862         3.140572         -0.44         0.659         -7.548853           -1799115         .0771521         -2.33         0.020        3312517          6662772         .0771985         -8.63         0.000        8177084          3799532         3.355975         -0.11         0.910         -6.962975           3.984662         4.818208         0.83         0.408         -5.46665           .1051598         .3327842         0.32         0.752        5476238           .0722016         .4839479         0.15         0.881        877102           .0129762         .5123413         0.03         0.980        9920235          4443146         .7331748         -0.61         0.545         -1.882497           .0010164         .0178947         0.06         0.955        0340855           .0013596         .0187994         0.07         0.942        0355169           .0105725         .0268027         0.39         0.693        0420033

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## Example: Analysis of the Vignette Data

- Questions we might have about the regression output:
  - What are the overall effects of the vignette factors?
  - How does the effect of vignette factor "sex" depend on education and sex of the respondent?

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- What is the shape of the effect of education depending on sex?
- Can we express effects in CHF?

### Example: Binary Dependent Variable

• A logistic regression of whether income in vignette was judged as "too low" or not:

generate	byte	toolow	=	vrating<0	if	vrating<.

. logit toolow vinc i.vmale i.vmarried i.veffort

Iteration	0:	log likelihood = -726.94882
Iteration	1:	log likelihood = -660.31413
Iteration	2:	log likelihood = -656.56237
Iteration	3:	log likelihood = -656.55323
Iteration	4:	log likelihood = -656.55323

Logistic regression

=	1482
=	140.79
=	0.0000
=	0.0968
	=

Log likelihood = -656.55323

toolow	Coef.	Std. Err.	z	P> z	[95% Conf.	Interval]
vinc	0013663	.0001789	-7.64	0.000	0017169	0010157
1.vmale	.3994001	.1393606	2.87	0.004	.1262584	.6725417
1.vmarried	.2886296	.1389959	2.08	0.038	.0162027	.5610565
1.veffort	1.164184	.1466717	7.94	0.000	.8767125	1.451655
_cons	4.939563	.9606374	5.14	0.000	3.056748	6.822378

## Example: Binary Dependent Variable

• Questions we might have about the logit output:

- What the hell do these coefficients mean?
- What is the conditional probability of "too low" depending on different levels of the factor variables?
- What is the marginal effect of the vignette factors on the probability of "too low"?

### Stata tools to answer these questions

- Stata commands margins and marginsplot can help us answer these questions.
- There's another useful command called contrast, but I am not going to talk about that.
- However, I will also show marginscontplot by Patrick Royston that will appear in one of the next issues of the Stata Journal.

## What can margins do?

- margins computes so-called margins of responses.
  - ► A "margin" is a statistic computed from predictions from a model while manipulating the values of the covariates.
    - "conditional margin": a prediction from a model where all covariates are set to fixed values
    - ★ "predictive margin": if some covariates are not fixed
  - Computed are *levels* of margins for different covariate values or *differences* in levels of margins if covariate values are changed (or even differences in differences). The later is often called *marginal effects*.
  - Continuous vs. discrete marginal effects
    - ★ For a continuous covariate, margins computes the first derivative of the response with respect to the covariate.
    - ★ For a discrete covariate, margins computes the effect of a discrete change of the covariate (discrete change effects).
  - MEM: marginal effects at the mean, AME: average marginal effects, MER: marginal effects at representative values

### Technical note

- You must use Stata's factor variable notation in the estimation command for margins to be able to compute correct results (see help fvvarlist).
  - Use the i. operator for discrete variables.
  - ▶ Use the *#* and *##* operators for interactions.
  - ▶ Use the c. for continuous variables involved in an interaction.

• Predictive margins / adjusted predictions (levels)

```
. quietly regress vrating vinc i.vmale i.vmarried i.veffort ///
```

```
> vmale##rmale##c.reducyrs##c.reducyrs, vsquish noheader
```

. margins vmale vmarried veffort

Predictive margins

Number of obs = 1482

Model VCE : OLS

Expression : Linear prediction, predict()

	Margin	Delta-method Std. Err.	z	P> z	[95% Conf.	Interval]
vmale						
	5707044	0544000	40 57	0 000	4004045	0700507
0	.5727241	.0541998	10.57	0.000	.4664945	.6789537
1	.2916397	.0547894	5.32	0.000	.1842544	.399025
vmarried						
0	. 5250302	.054352	9.66	0.000	.4185023	.6315581
1	.3451187	.0546468	6.32	0.000	.238013	.4522245
veffort						
0	.7664507	.0543315	14.11	0.000	.6599629	.8729385
1	.1001736	.0547003	1.83	0.067	0070371	.2073843

- Interpretation of predictive margins for vmale:
  - If all respondents would have answered the *female* vignette (keeping the other vignette factors and the repondent's sex and education as they happen to be), then the average response would have been 0.57.
  - If all respondents would have answered the *male* vignette (keeping the other vignette factors and the repondent's sex and education as they happen to be), then the average response would have been 0.29.
  - This means that, keeping everything else constant, the same income is more likely to be judged as too low in the male vignette than in the female vignette.
  - ► To find out whether the difference is significant, we can use margins to compute contrasts or marginal effects.
    - ★ An alternative would be to specify the post option in the above command and then apply the test command (see below).

### • Contrasts (differences in levels): use the r. operator

. margins r.vmale r.vmarried r.veffort

```
Contrasts of predictive margins
Model VCE : OLS
```

Expression : Linear prediction, predict()

	df	chi2	P>chi2
vmale	1	13.30	0.0003
vmarried	1	5.44	0.0197
veffort	1	74.49	0.0000

	Contrast	Delta-method Std. Err.	[95% Conf.	Interval]
vmale (1 vs 0)	2810844	.0770733	4321454	1300234
vmarried (1 vs 0)	1799115	.0771521	3311268	0286961
veffort (1 vs 0)	6662772	.0771985	8175835	5149709

• Marginal effects for discrete variables (discrete change effects): use the dydx() option

. margins, dydx(vmale vmarried veffort)			
Average marginal effects Model VCE : OLS	Number of obs	=	1482
Expression : Linear prediction, predict() dy/dx w.r.t. : 1.vmale 1.vmarried 1.veffort			

		Delta-method Std. Err.	Z	P> z	[95% Conf.	Interval]
1.vmale	2810844	.0770733	-3.65	0.000	4321454	1300234
1.vmarried	1799115	.0771521	-2.33	0.020	3311268	0286961
1.veffort	6662772	.0771985	-8.63	0.000	8175835	5149709

Note: dy/dx for factor levels is the discrete change from the base level.

### • Interpretation contrasts / marginal effects

- We see that both commands yield the same results.
- ▶ The effect of male vs. female sex in the vignette is an average decrease of 0.28 points on the response scale. This is simply the difference in the predictive margins computed above.
- ► The difference is highly significant with a *z*-value of 3.65 or a χ<sup>2</sup>(1)-value of 13.3 (which is simply the square of the *z*-value because the test has 1 degree of freedom).
- ► The 95% confidence interval of the effect is -0.43 to -0.13.

- Let's start with the interaction with respondent's sex.
- Predictive margins for the vignette sex by sex of the respondent can be computed as follows:

. margins vm	ale, over(rmale)			
Predictive m Model VCE	0	Number of obs	=	1482
Expression over	: Linear prediction, predict() : rmale			

	I Margin	Delta-method Std. Err.	z	P> z	[95% Conf.	Interval]
rmale#vmale						
0 0	.580111	.0744694	7.79	0.000	.4341537	.7260683
0 1	.2069453	.0758263	2.73	0.006	.0583285	.3555621
1 0	.5645165	.0790644	7.14	0.000	.4095531	.7194799
1 1	.3857447	.0792884	4.87	0.000	.2303422	.5411472

- We see that for female respondents, the difference in predictive margins for vignette sex is larger (0.21 0.58 = -0.37) than for male respondents (0.39 0.56 = -0.17).
  - Note that the difference could be due to differential educational levels of female and male respondents, because an interaction with education was included in the regression model and the predictive margins are averaged over female and male respondents as is. We could, for example, type

. margins vmale rmale, at((omean) reducyrs)

to find out whether controlling for education changes the picture (it does a bit, but not much) (omean sets the respondent's education to the overall mean across all observations).

• To find out whether effects for female and males are significant we can again resort to the r. contrast operator or the dydx() option.

• Using the r. contrast operator:

. margins	r.vma	ale, ove	er(rmale)	
Contrasts	of p	redictiv	ve margins	
Model VCE	:	OLS		
Expression	ı :	Linear	prediction,	predict()
over	:	rmale		

	df	chi2	P>chi2
vmale@rmale (1 vs 0) 0 (1 vs 0) 1 Joint	1 1 2	12.32 2.55 14.88	0.0004 0.1105 0.0006

	Contrast	Delta-method Std. Err.	[95% Conf.	Interval]
vmale@rmale (1 vs 0) 0 (1 vs 0) 1	3731657 1787718	.1063105 .1120037	5815304 3982952	164801 .0407515

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• Using the dydx() option:

. margins, dydx(vmale) over(rmale)	
Average marginal effects Model VCE : OLS	Number of obs = 1482
Expression : Linear prediction, pred dy/dx w.r.t. : 1.vmale over : rmale	ict()

	l dy/dx	Delta-method Std. Err.	z	P> z	[95% Conf.	Interval]
1.vmale						
rmale						
0	3731657	.1063105	-3.51	0.000	5815304	164801
1	1787718	.1120037	-1.60	0.110	3982952	.0407515

Note: dy/dx for factor levels is the discrete change from the base level.

• Mechanics of the over() option:

 Specifying over is equivalent to running margins on subpopulations. That is, the results above could also be computed by typing

. margins if rmale==0, dydx(vmale)

```
. margins if rmale==1, dydx(vmale)
```

Beware, however, that using if can lead to biased standard errors in complex samples. A saver approach is to use the subpop() option:

. margins, dydx(vmale) subpop(if rmale==0)

. margins, dydx(vmale) subpop(if rmale==1)

- What we really want to know is whether the effect of the vignette sex is different for female respondents and for male respondents.
- We could test this, for example, as follows:

. estimates s	tore lin				
. margins, dy	dx(vmale) ove	er(rmale) coeflegend pos	st		
Average margi Model VCE			Number of obs	=	1482
Expression dy/dx w.r.t. over	: 1.vmale	liction, predict()			
	dy/dx	Legend			
1.vmale	dy/dx	Legend			
rmale					
	3731657	Legend _b[1.vmale:0bn.rmale] _b[1.vmale:1.rmale]			

Note: dy/dx for factor levels is the discrete change from the base level.

- A more direct approach is to have margins compute an estimate for the difference in differences by adding the r. operator within the over() option.
- Either type ...

```
. margins r.vmale, over(r.rmale)
Contrasts of predictive margins
Model VCE : 0LS
Expression : Linear prediction, predict()
over : rmale
```

	df	chi2	P>chi2
rmale#vmale	1	1.58	0.2083

	Contrast	Delta-method Std. Err.	[95% Conf.	Interval]
rmale#vmale (1 vs 0) (1 vs 0)	. 1943939	.1544914	1084037	.4971914

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#### • ... or type

. margins, dy	yd:	x(vmale) over(r.rmale)
Contrasts of Model VCE		verage marginal effects OLS
Expression dy/dx w.r.t. over	:	Linear prediction, predict() 1.vmale rmale

	df	chi2	P>chi2
Ob.vmale rmale	(omitted)		
1.vmale rmale	1	1.58	0.2083

		Delta-method Std. Err.	[95% Conf.	Interval]
1.vmale rmale				
(1 vs 0)	. 1943939	.1544914	1084037	.4971914

Note: dy/dx for factor levels is the discrete change from the base level.

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Predictive Margins and Marginal Effects

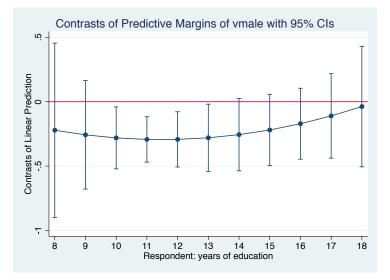
- Our conclusion from the above results would be that the effects of the vignette sex appears to be a bit stronger for female, but the difference is not significant.
- Again, note that there is a third variable involved in the interaction (education), so that part of the differences between the effects for female and male respondents might be due to different educational level.
- To see how the effect of the vignette sex changes by education, we could type

```
. margins r.vmale, at(reducyrs=(8(1)18))
```

. marginsplot, yline(0)

or

- . margins, dydx(vmale) at(reducyrs=(8(1)18))
- . marginsplot, yline(0)



- We see that, after an initial increase, the effect of the vignette sex diminishes (i.e. gets closer to zero) as education increases.
- Still we are looking only at a two-way interaction (vignette sex by education). To explore the full three-way interaction specified in the model we have to go one step further.
- For example, to see how the effect of vignette sex depends on education by sex of respondent, we could type

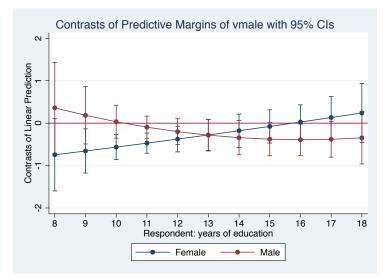
```
. margins r.vmale, over(rmale) at(reducyrs=(8(1)18))
```

```
. marginsplot, yline(0)
```

```
or
```

```
. margins, dydx(vmale) over(rmale) at(reducyrs=(8(1)18))
```

```
. marginsplot, yline(0)
```

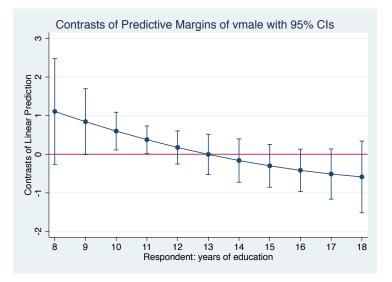


- The pattern of the effect of the vignette sex appears to be somewhat different for female respondents and male respondents.
- To better see how the difference in the effect of the vignette sex between female and male respondents changes with education, we could type

```
. margins r.vmale, over(r.rmale) at(reducyrs=(8(1)18))
. marginsplot, yline(0)
```

#### or

```
. margins, dydx(vmale) over(r.rmale) at(reducyrs=(8(1)18))
. marginsplot, yline(0)
```



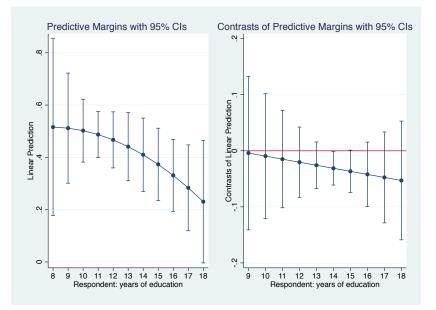
• In this graph, a difference-in-differences estimate is displayed for each educational level. We see that the difference in effects between female and male respondents is first positive, but then declines and becomes negative after 13 years of education

## Answers: Shape of the effect of education

- Up to now, we looked at how the effect of the vignette sex changes depending on education. However, we might also be interested in the main effect of education on the responses.
- To see how the response level changes with education we can simply type
  - . margins, at(reducyrs=(8(1)18))
  - . marginsplot
- Furthermore, for a picture of how the effect of education changes with educational level, type
  - . margins, at(reducyrs=(8(1)18)) contrast(atcontrast(ar.\_at))
  - . marginsplot, yline(0)

The contrast(atcontrast(ar.\_at)) option causes margins to compute contrasts between predictive margins across adjacent educational levels.

## Answers: Shape of the effect of education

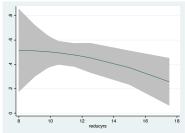


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Predictive Margins and Marginal Effects

## Answers: Shape of the effect of education

- Respondents with lower educational level are more likely to judge the income in the vignette as too high than respondents with higher educational level. Furthermore, the (negative) effect of education is getting steeper with additional education.
- A variant of the plot on the left can also be quickly produced by Patrick Royston's marginscontplot:

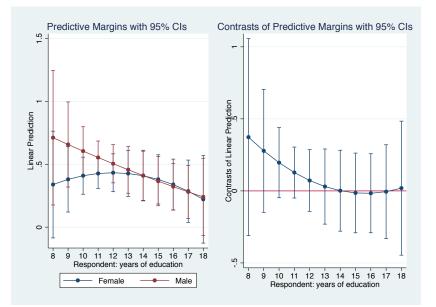


. marginscontplot reducyrs, ci

# Answers: Shape of the effect of education

- To compare the shape of the effect of education between female and male respondents, we could type something like
  - . margins, at(reducyrs=(8(1)18)) over(rmale)
  - . marginsplot, name(a, replace)
  - . margins, at(reducyrs=(8(1)18)) over(r.rmale)
  - . marginsplot, yline(0) nodraw name(b, replace)
  - . graph combine a b

# Answers: Shape of the effect of education



# Answers: Shape of the effect of education

- In the left plot, the predictive margins by educational level are shown for female respondents and male respondents.
- In the right plot, the difference between the two curves is plotted. This shows how the effect of the sex of the respondent changes with educational level.

- As we can see in the regression output on slide 7, the vignette income has an effect of 0.0010069 on the judgement. An increase in income by 1 CHF increases the expected judgement by 0.0010069 points.
- This means, that 1 point on the judgement scale is worth about 993 CHF:

```
. display 1/_b[vinc] 993.15639
```

- You can use the expression() option in margins to compute predictive margins and marginal effects with respect to a rescaled outcome so that, in our case, all effects are expressed in CHF.
- margins will take care of the details and also provide consistent standard errors.

#### • Example: Effects of vignette factors

. margins, dydx(vmale vmarried veffort) expression(xb()/-\_b[vinc])

 Average marginal effects
 Number of obs = 1482

 Model VCE
 : OLS

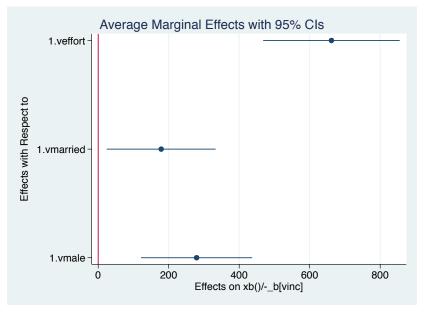
 Expression
 : xb()/-\_b[vinc]

 dy/dx w.r.t.
 : 1.vmale 1.vmarried 1.veffort

_	l dy/dx	Delta-method Std. Err.	z	P> z	[95% Conf.	Interval]
1.vmale	279.1608	80.52689	3.47	0.001	121.331	436.9906
1.vmarried	178.6802	78.74228	2.27	0.023	24.3482	333.0123
1.veffort	661.7174	98.82199	6.70	0.000	468.0299	855.405

Note: dy/dx for factor levels is the discrete change from the base level.

• Interpretation: Males "should" get 279 CHF more per month than females, the married should get 179 CHF more, the hard-working should get 661 CHF more.

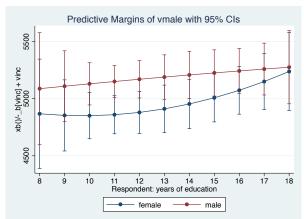


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- Example: "Just" income levels for females and males by education of respondent
  - . margins vmale, at(reducyrs = (8(1)18)) expression(xb()/-\_b[vinc] + vinc)
  - . marginsplot



# Answers: Interpretation of a logit model

• What is the conditional probability of "too low" depending on different levels of the factor variables?

```
. quietly logit toolow vinc i.vmale i.vmarried i.veffort

. margins vmale vmarried veffort

Predictive margins Number of obs = 1482

Model VCE : OIM

Expression : Pr(toolow), predict()
```

	Delta-method								
	Margin	Std. Err.	z	P> z	[95% Conf.	Interval]			
vmale									
0	.1652877	.012963	12.75	0.000	.1398807	.1906946			
1	.2215201	.0145671	15.21	0.000	.192969	.2500711			
vmarried									
0	.1725705	.0133054	12.97	0.000	.1464925	.1986485			
1	.2131927	.0142477	14.96	0.000	.1852677	.2411177			
veffort									
0	.1115456	.0113453	9.83	0.000	.0893093	.1337819			
1	.2749636	.015868	17.33	0.000	.2438629	.3060642			

# Answers: Interpretation of a logit model

• What is the (average) marginal effect of the vignette factors on the probability of "too low"?

. margins, dyd	lx(vmale vmar	ried veffort)					
Average margin Model VCE				Number	of obs	=	1482
Expression : dy/dx w.r.t. :			fort				
		Delta-method Std. Err.	z	P> z	[95% 0	Conf.	Interval]

2 88

2.08

8 38

0 004

0.037

0 000

0180066

.0024053

1251842

0944582

.0788392

2016517

Note: dy/dx for factor levels is the discrete change from the base level.

0195033

.0194988

0195074

• We see that, for example, the average marginal effect of the vignette sex is 5 percentage points. That is, everything else equal, we would expect a 5 percentage point increase in the proportion of respondents who judge the vignette income as too low if we change the vignette sex from female to male.

0562324

.0406222

1634179

1 vmale

1.vmarried

1 veffort

# What margins and marginsplot can't do

- The default for continuous variables is to compute marginal effects as first derivatives. Discrete change effects for continuous variables can only be computed for special cases (e.g. min to max, in steps across the scale).
  - It would be nice to be able to compute discrete change effects around observed values (e.g +/- half a standard deviation).
- Marginal effects for transformed covariates can only be computed in special cases (e.g. quadratic).
- margins can only deal with one equation at the time in models with multiple outcomes (e.g. mlogit).
- margins can be excessively slow on big datasets.
- marginsplot can only handle results from margins and can only display one set of results at the time.

### New command: regplot

• A new command called regplot provides a solution for the last problem.

regplot models, options

- regplot uses marginsplot internally, but it can be applied to any estimation results, be This commands has been r not.
- Multiple estimation results can be combined in a single graph.
- By default, regplot creates a horizontal "dot plot" of the coefficients found in e(b) and includes spikes for confidence intervals. I call this a "regression plot". Others sometimes call it an "airplane plot".

### New command: regplot

- How to specify models:
  - Models in separate plots:

```
(modelname, options) || (modelname, options) ...
```

Several models in one plot:

(modelname, options) ...

Append models:

(modelname, options  $\setminus$  modelname, options) ...

• A combination of the above.

### Plot of raw coefficients

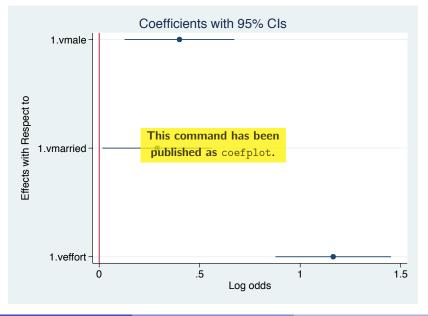
• regplot can be applied directly after an estimation command to produce a plot of the estimated coefficients.

. logit toolow vinc i.vmale i.vmarried i.veffort, nolog

Logistic regression				Number of obs LR chi2(4)			1482 140.79
					> chi2	=	0.0000
Log likelihood	1 = -656.553 <mark>2</mark>	This comm	and has	beenet	ıdo R2	=	0.0968
		published	as coefr	olot.			
toolow	Coef.	Std. Err.	Z	P> z	[95%	Conf.	Interval]
vinc	0013663	.0001789	-7.64	0.000	0017	169	0010157
1.vmale	.3994001	.1393606	2.87	0.004	. 1262	584	.6725417
1.vmarried	.2886296	.1389959	2.08	0.038	.0162	2027	.5610565
1.veffort	1.164184	.1466717	7.94	0.000	.8767	125	1.451655
_cons	4.939563	.9606374	5.14	0.000	3.056	748	6.822378

. regplot ., keep(1.\*) xline(0) xtitle(Log odds)

# Plot of raw coefficients



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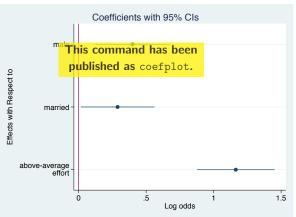
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### Label the coefficients

• Coefficients are positioned on the y-axis at 1, 2, 3, ... (from top). Hence, you can use ylabel() to define custom labels.

```
. regplot ., keep(1.*) xline(0) xtitle(Log odds) ///
> ylabel(1 "male" 2 "married" 3 `""above-average" "effort""') ///
> yscale(range(0.75 3.25))
```



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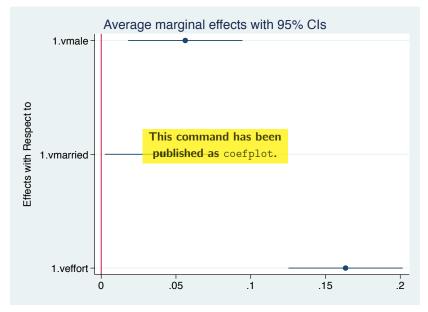
### Plot results from margins

• regplot can also be applied after margins, if the post option is specified with margins:

. margins, dy	dx(vmale vmar	ried veffort	) post			
Average marginal effects Model VCE : OIM				Numb	er of obs =	1482
Expression dy/dx w.r.t.	: Pr(toolow),. : 1.vmale 1.v	This comma published a				
		Delta-method	1			
	dy/dx	Std. Err.	Z	P> z	[95% Conf.	Interval]
1.vmale 1.vmarried 1.veffort	.0562324 .0406222 .1634179	.0195033 .0194988 .0195074	2.88 2.08 8.38	0.004 0.037 0.000	.0180066 .0024053 .1251842	.0944582 .0788392 .2016517

Note: dy/dx for factor levels is the discrete change from the base level. . regplot ., xline(0) keep(1.\*) title(Average marginal effects with 95% CIs)

# Plot results from margins



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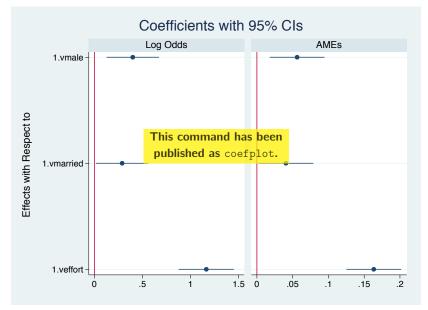
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### Combine raw coefficients and AMEs

- regplot can be applied to stored models. For example, we could combine raw coefficients from logit and corresponding average marginal effects from margins
  - . quietly logit toolow vinctive to the provident of the p
  - . estimates store raw **published as** coefplot.
  - . quietly margins, dydx(vmale vmarried veffort) post
  - . estimates store ame
  - . regplot (raw, bylabel(Log Odds)) || (ame, bylabel(AMEs)) ///
  - > , xline(0) keep(1.\*) byopts(xrescale)

# Combine raw coefficients and AMEs



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Predictive Margins and Marginal Effects

Combine predictive margins and contrasts

• We could also predictive margins and contrasts/marginal effects in one graph.

- . quietly logit toolow vinc i.vmale i.vmarried i.veffort
- . estimates store raw
- . quietly margins vmale vmarried veffort, post
- . estimates store margins This command has been

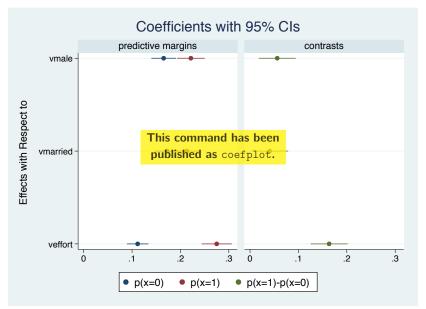
```
. estimates restore raw published as coefplot. (results raw are active now)
```

```
. quietly margins, dydx(vmale vmarried veffort) post
```

```
. estimates store contrasts
```

```
. regplot (margins, rename(0bn. "") drop(1.*) label(p(x=0))) ///
> (margins, rename(1. "") drop(0bn.*) label(p(x=1)) ///
> bylabel(predictive margins)) || ///
> (contrasts, rename(1. "") drop(0.*) label(p(x=1)-p(x=0)) ///
> bylabel(contrasts)), norecycle legend(rows(1))
```

# Combine predictive margins and contrasts



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### Append models

• Use \ to append models in one plot.

- . quietly logit toolow vinc i.vmale i.vmarried i.veffort
- . estimates store logit
- . quietly margins vmale, post
- . estimates store vmale
- . estimates restore logit

(results logit are active now)

. quietly margins vmarried This command has been

```
. estimates store vmarried published as coefplot.
```

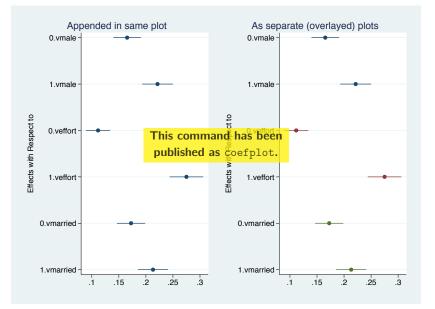
. estimates restore logit (results logit are active now)

- . quietly margins veffort, post
- . estimates store veffort
- . regplot (vmale \ veffort \ vmarried), title(Appended in same plot) ///
- > nodraw name(a, replace)

```
. regplot (vmale) (veffort) (vmarried), title(As separate (overlayed) plots) ///
> legend(off) nodraw name(b, replace)
```

```
. graph combine a b
```

# Append models



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Predictive Margins and Marginal Effects

# Plot coefficients from mlogit

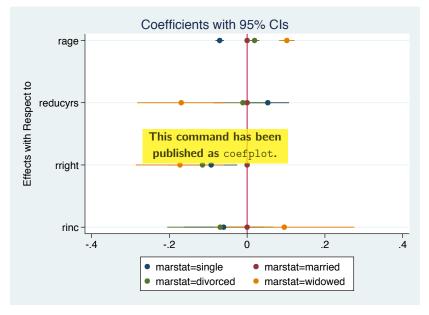
 In multiequation models, use the eq() option to select the equation to be plotted.

. mlogit rmars	stat rage red	ucyrs rright	rinc, n	olog			
Multinomial lo				LR ch	r of obs 12(12) > chi2	=	1482 477.71 0.0000 0.1475
rmarstat	Coef.	Std. Err.	z	P> z		onf.	Interval]
single							
rage reducyrs	0705039 .0531021	.0055962 .0278531	-12.60 1.91	0.000	08147 00148		0595355 .1076931
rright rinc _cons	0924911 0604214 2.533033	.0340842 .0522317 .4611503		-0.247	nman od <sup>1.6291</sup>		has bee
married	(base outco	ome)	pu	DIISH	eu as	00	efplot
divorced							
rage reducyrs rright rinc _cons	.0190981 0109678 1147781 0691778 -1.927327	.0064624 .0386512 .0449214 .0695684 .633678	2.96 -0.28 -2.56 -0.99 -3.04	0.003 0.777 0.011 0.320 0.002	.0064 08672 20282 20552 -3.1693	28 25 93	.0317642 .0647873 0267338 .0671736 6853408
widowed							
rage reducyrs rright rinc cons	.1021225 1693922 1726326 .0955222 -6.098965	.0101098 .0577359 .0580334 .0920018 .9666954	10.10 -2.93 -2.97 1.04 -6.31	0.000 0.003 0.003 0.299 0.000	.08230 28255 2863 0847 -7.9936	26 76 98	.1219373 0562319 0588892 .2758423 -4.204277

. estimates store marstat

- . regplot (marstat, eq(single)) ///
- > (marstat, eq(married)) ///
- > (marstat, eq(divorced)) ///
- > (marstat, eq(widowed)), nocons xline(0)

# Plot coefficients from mlogit



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### Plot marginal effects from mlogit

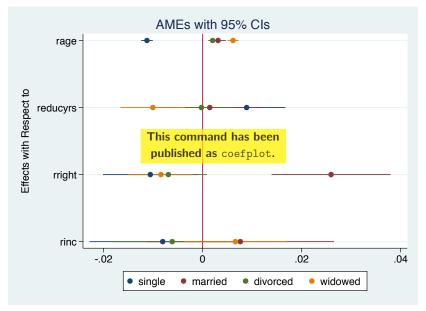
• To plot marginal effects for an mlogit model you have to run margins for each outcome.

. quietly margins, dydx(\*) predict(outcome(single)) post

- . estimates store single
- estimates restore marstat
- (results marstat are active now)
- . quietly margins, dydx(\*) predict(outcome(married)) post
- . estimates store married
- published as coefplot. . estimates restore marst (results marstat are active now)
- . quietly margins, dydx(\*) predict(outcome(divorced)) post
- . estimates store divorced
- estimates restore marstat
- (results marstat are active now)
- . quietly margins, dydx(\*) predict(outcome(widowed)) post
- estimates store widowed
- . regplot single married divorced widowed, legend(row(1)) xline(0) ///
- title(AMEs with 95% CIs) >

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# Plot marginal effects from mlogit



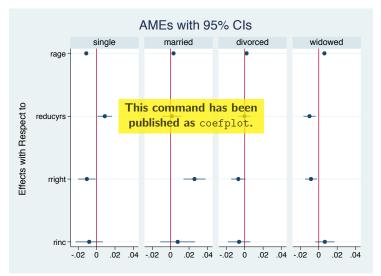
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# Plot marginal effects from mlogit

. regplot single || married || divorced || widowed, xline(0) ///
> byopt(rows(1) title(AMEs with 95% CIs))



#### References

- Rising, B. (2012). Working in the margins to plot a clear course.
   Presentation at the 10th German Stata Users Group Meeting in Berlin.
- Rising, B. (2012). How to get an edge with margins and marginsplot. Presentation at the 2012 UK Stata Users Group Meeting in London.
- Royston, P. (forthcoming). marginscontplot: plotting the marginal effects of continuous predictors. *The Stata Journal*.
- Williams, R. (2012). Using the margins command to estimate and interpret adjusted predictions and marginal effects. *The Stata Journal* 12(2):308–331.