

# Comparing coefficients of nested nonlinear probability models using `khb`

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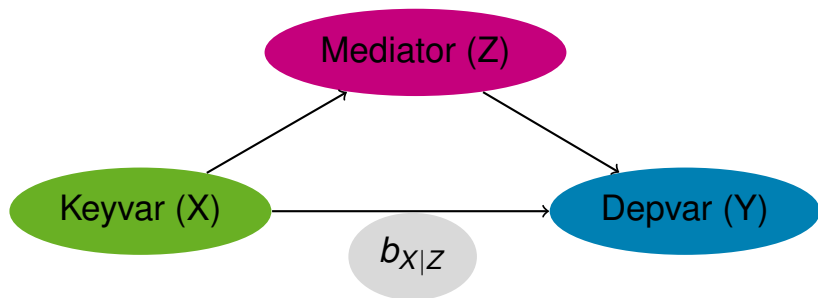
# Outline

- 1 Introduction
- 2 The KHB-method
- 3 The command khb
- 4 Application
- 5 References

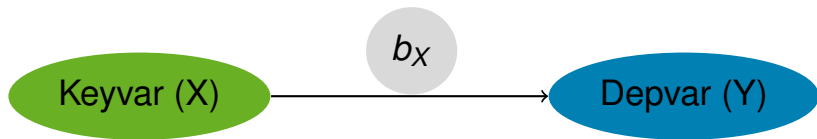
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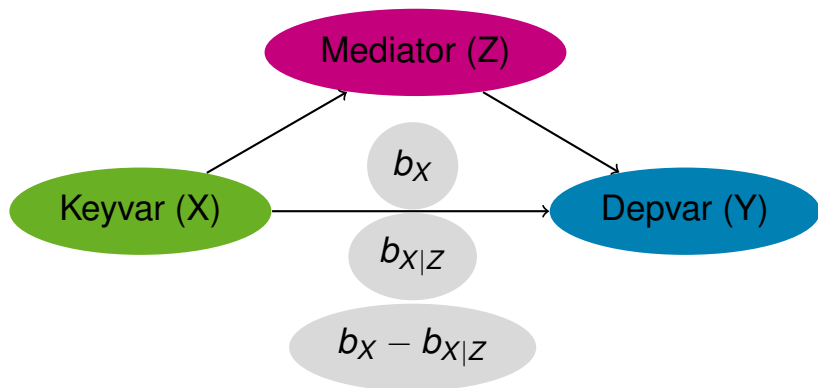
# Reasons to compare



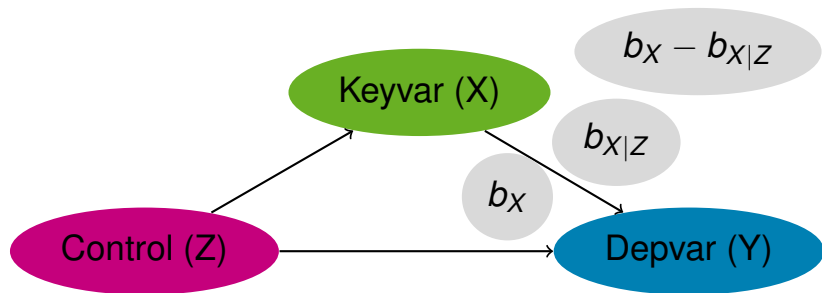
# Reasons to compare



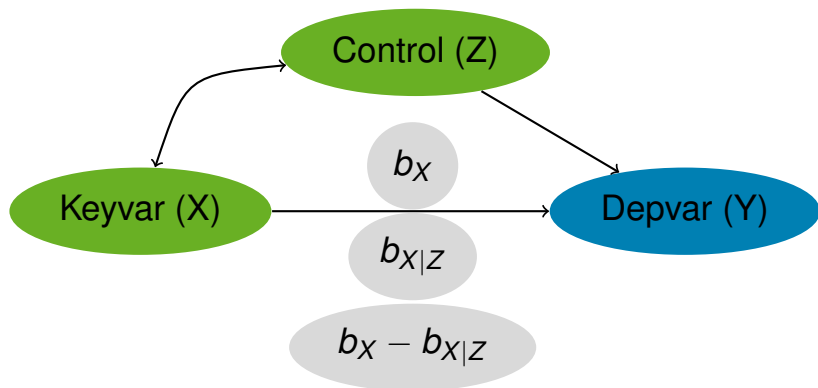
# Reasons to compare



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# The problem

We are interested in obtaining  $\beta_R - \beta_F$  from the following models for latent  $Y^*$ :

$$Y^* = \alpha_F + \beta_F X + \gamma_F Z + \delta_F C + \epsilon \quad (1)$$

$$Y^* = \alpha_R + \beta_R X + \delta_R C + \varepsilon \quad (2)$$

Having observed  $Y$  with value 0 if  $Y^* < \tau$  and 1 if  $Y^* \geq \tau$  we can obtain the logit/probit estimates with

$$b_F = \frac{\beta_F}{\sigma_F} \quad \text{and} \quad b_R = \frac{\beta_R}{\sigma_R} \quad (3)$$

**Note:** We identify the underlying coefficients of interest relative to a scale unknown to us.

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# General idea

The KHB-method extracts from  $Z$  the information that is not contained in  $X$ . This is done by calculating the residuals of a linear regression of  $Z$  on  $X$ , i.e.,

$$R = Z - (a + bX) \quad , \quad (4)$$

where  $a$  and  $b$  are the estimated regression parameters of a linear regression.

Instead of using equation (2) we then use

$$Y^* = \tilde{\alpha}_R + \tilde{\beta}_R X + \tilde{\gamma}_R R + \tilde{\delta}_R C + \epsilon \quad . \quad (5)$$

## Difference of coefficients

As  $R$  and  $Z$  differ only in the component in  $Z$  that is correlated with  $X$ , model (1) is no more predictive than model (5), and consequently the residuals have the same standard deviation so that

$$\tilde{\sigma}_R = \sigma_F \quad (6)$$

As  $\tilde{\beta}_R = \beta_R$  we can write

$$\tilde{b}_R - b_F = \frac{\tilde{\beta}_R}{\tilde{\sigma}_R} - \frac{\beta_F}{\sigma_F} = \frac{\beta_R - \beta_F}{\sigma_F} \quad (7)$$

Hence, the difference obtained reflects the difference searched divided by some common scale.

# Derived statistics

## Confounding ratio

$$\frac{\tilde{b}_R}{b_F} = \frac{\frac{\beta_R}{\sigma_F}}{\frac{\beta_F}{\sigma_F}} = \frac{\beta_R}{\beta_F} \quad , \quad (8)$$

## Counfounding percentage

$$100 \cdot \frac{\tilde{b}_R - b_F}{\tilde{b}_R} = 100 \cdot \frac{\frac{\beta_R}{\sigma_F} - \frac{\beta_F}{\sigma_F}}{\frac{\beta_R}{\sigma_F}} = 100 \cdot \frac{\beta_R - \beta_F}{\beta_R} \quad , \quad (9)$$

# Significance test for the difference in effects

- Analytically derived standard errors for the difference in effects exist.
- Based on the delta method (Sobel, 1982).
- Simple for one  $X$  and one  $Z$  but fairly complicated for situations with more than one  $X$ ,  $Z$ .
- Karlson et al. (2010) has more details; also see our Stata Journal publications (in Press)

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# Syntax

```
khb model-type depvar key-vars || mediator-vars [if][in][, options ]
```

*model-type* can be any of `regress`, `logit`, `ologit`, `probit`, `oprobit`, `cloglog`, `slogit`, `scobit`, `rologit`, `clogit`, and `mlogit`.

*key-vars* may contain factor variables

`aweight`s, `fweight`s, `iweight`s, and `pweight`s are allowed if they are allowed for the specified model type.



## Options (most important ones)

<i>options</i>	description
<code>concomitant (varlist)</code>	concomitants
<code>disentangle</code>	disentangle difference of effects
<code>summary</code>	summary of decomposition
<code>vce (vcetype)</code>	robust or cluster <code>clustvar</code>
<code>ape</code>	decomposition using avg. partial effects
<code>verbose</code>	show restricted and full model
<code>keep</code>	keep residuals of mediators

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# Preliminaries

- Examples from educational sociology
- Subset of Danish National Longitudinal Survey (DLSY).
- Reproduce analysis presented by Karlson and Holm (2011).

```
. use dlsy_khb, clear
. describe
Contains data from dlsy_khb.dta
  obs:      1,896
  vars:      8                               17 Jan 2011 10:26
  size:      49,296 (99.9% of memory free)
```

variable name	storage type	display format	value label	variable label
edu	byte	%20.0g	edu	Educational attainment
upsec	byte	%10.0g	yesno	Complete upper secondary education (Gymnasium)
univ	byte	%13.0g	yesno	Complete University education
fgroup	byte	%9.0g	fgroup	Father's social group/class
fses	float	%9.0g		Father's SES, standardized with mean 0 and sd 1
abil	double	%10.0g		Standardized ability measure, with mean 0 and sd 1
intact	byte	%9.0g	yesno	Intact family
boy	byte	%9.0g	yesno	Boy

Sorted by:



# Basic use

```
. khb logit univ fses || abil, c(intact boy)
```

Decomposition using the KHB-Method

```
Model-Type: logit                Number of obs   =   1896
Variables of Interest: fses      Pseudo R2       =    0.19
Z-variable(s): abil
Concomitant: intact boy
```

	univ	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
fses						
	Reduced	.5459815	.0779806	7.00	0.000	.3931424 .6988206
	Full	.3817324	.0778061	4.91	0.000	.2292353 .5342295
	Diff	.1642491	.0293249	5.60	0.000	.1067734 .2217247

# Confounding ratio/percentage

```
. khb logit univ fses || abil, c(intact boy) summary notable
```

Decomposition using the KHB-Method

```
Model-Type: logit                                Number of obs   =   1896
```

```
Variables of Interest: fses                      Pseudo R2       =    0.19
```

```
Z-variable(s): abil
```

```
Concomitant: intact boy
```

Summary of confounding

Variable	Conf_ratio	Conf_Pct	Resc_Fact
fses	1.4302727	30.08	1.0602422

# Option ape

```
. khb logit univ fses || abil, c(intact boy) ape summary
```

Decomposition using the APE-Method

```
Model-Type: logit                      Number of obs   =   1896
Variables of Interest: fses             Pseudo R2       =   0.19
Z-variable(s): abil
Concomitant: intact boy
```

	univ	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
fses							
	Reduced	.0384906	.0054429	7.07	0.000	.0278226	.0491585
	Full	.0269113	.0054476	4.94	0.000	.0162343	.0375884
	Diff	.0115792	.0020667	5.60	0.000	.0075286	.0156298

Note: Standard errors of difference not known for APE method

Summary of confounding

Variable	Conf_ratio	Conf_Pct	Dist_Sens
fses	1.4302727	30.08	.95931864

# Disentangle contributions of mediators

```
. khb logit univ fses || abil intact boy, s d not
```

Decomposition using the KHB-Method

```
Model-Type: logit                Number of obs   =   1896
Variables of Interest: fses      Pseudo R2       =   0.19
Z-variable(s): abil intact boy
```

Summary of confounding

Variable	Conf_ratio	Conf_Pct	Resc_Fact
fses	1.5207722	34.24	1.1317064

Components of Difference

Z-Variable	Coef	Std_Err	P_Diff	P_Reduced
fses				
abil	.1661177	.0301003	83.56	28.61
intact	.020142	.0144611	10.13	3.47
boy	.0125359	.011524	6.31	2.16

# More than one key variable

```
. khb logit univ boy intact || abil, c(fses) s
```

Decomposition using the KHB-Method

```
Model-Type: logit                Number of obs   =   1896
Variables of Interest: boy intact Pseudo R2       =   0.19
Z-variable(s): abil
Concomitant: fses
```

	univ	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
boy						
	Reduced	1.06178	.1848087	5.75	0.000	.6995613 1.423998
	Full	.9821406	.1848351	5.31	0.000	.6198704 1.344411
	Diff	.0796391	.133004	0.60	0.549	-.1810438 .3403221
intact						
	Reduced	1.129767	.7386976	1.53	0.126	-.3180536 2.577588
	Full	1.08391	.7386558	1.47	0.142	-.3638292 2.531648
	Diff	.0458575	.1328438	0.35	0.730	-.2145116 .3062266

Summary of confounding

Variable	Conf_ratio	Conf_Pct	Resc_Fact
boy	1.0810873	7.50	1.0033213
intact	1.0423075	4.06	1.03542



# Categorical variables

```
. xtile catabil = abil, n(4)
. tab catabil, gen(catabil)
. khb logit univ i.fgroup || catabil2-catabil4, c(intact boy) s d
```

# Ordered outcome

```
. forv i = 1/3 {
2.         quietly eststo: khb ologit edu fses || abil, out(`i') ape s
3. }
. esttab, scalars("ratio_fses Conf.-Ratio" "pct_fses Conf.-Perc.")
```

	(1) edu	(2) edu	(3) edu
fses			
Reduced	-0.103*** (-11.33)	0.0643*** (10.72)	0.0385*** (9.27)
Full	-0.0755*** (-8.02)	0.0472*** (7.76)	0.0283*** (7.23)
Diff	-0.0272*** (-6.50)	0.0170*** (6.44)	0.0102*** (5.95)
N	1896	1896	1896
Conf.-Ratio	1.360	1.360	1.360
Conf.-Perc.	26.48	26.48	26.48

t statistics in parentheses  
 \* p<0.05, \*\* p<0.01, \*\*\* p<0.001

# Multinomial outcome

```
. forv i = 2/3 {
2.         quietly eststo: khb mlogit edu fses || abil, out(`i´) base(1) s
3. }
. esttab, scalars("ratio_fses Conf.-Ratio" "pct_fses Conf.-Perc.")
```

	(1) edu	(2) edu
fses		
Reduced	0.423*** (7.63)	0.779*** (9.30)
Full	0.313*** (5.70)	0.552*** (6.68)
Diff	0.109*** (5.93)	0.227*** (6.04)
N	1896	1896
Conf.-Ratio	1.349	1.411
Conf.-Perc.	25.88	29.15

t statistics in parentheses

\* p<0.05, \*\* p<0.01, \*\*\* p<0.001

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# References

- Karlson, K. B. and A. Holm. 2011. Decomposing primary and secondary effects: A new decomposition method. *Research in Stratification and Social Mobility* 29: XXXX.
- Karlson, K. B., A. Holm, and R. Breen. 2010. Comparing regression coefficients between models using logit and probit: a new method. Unpublished paper (currently under review).
- Sobel, M. E. 1982. Asymptotic confidence intervals for indirect effects in structural equation models. In *Sociological Methodology 1982*, ed. L. S., 290–312. Washington D.C.: American Sociological Association.