

Multinomial logit models for longitudinal data

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The multinomial logit model

- Primary use case: unordered categorical outcome variable.
- Panel data model in utility-maximization form: $U_{ijt} = x_{it}\beta_j + u_{ij} + \epsilon_{ijt}$
- u_{ij} is a panel-level error term, also referred to as panel-level unobservables, that captures (time-invariant) unobserved heterogeneity.
- Assuming a type-I extreme value distribution for ϵ_{ijt} and normalizing with respect to a base category gives rise to the logit model:

$$\Pr(y_{it} = m \mid x_{it}, \beta_j, u_{ij}) = F(y_{it} = m, x_{it}\beta_j + u_{ij}) =$$
$$\frac{1}{1 + \sum_{j=2}^J \exp(x_{it}\beta_j + u_{ij})} \quad \text{if } m = 1$$
$$\frac{\exp(x_{it}\beta_m + u_{im})}{1 + \sum_{j=2}^J \exp(x_{it}\beta_j + u_{ij})} \quad \text{if } m > 1$$

The random-effects estimator

- `xtmlogit` implements a random-effects and conditional fixed-effects estimator.
- The random-effects estimator assumes that the panel-level unobservables are uncorrelated with the covariates.
- The fixed-effects estimator relaxes this assumption and the unobservables can be arbitrarily correlated with the covariates.
- Random-effects estimator:

$$l_i = \int_{-\infty}^{\infty} \cdots \int_{-\infty}^{\infty} \left\{ \prod_{t=1}^{T_i} F(y_{it} = m, \mathbf{x}_{it}\beta_j + u_{ij}) \right\} \phi(u_i, \Sigma_u) du_i$$

- This integral has no closed form and we approximate it using quadrature.

The fixed-effects estimator

- Fixed-effects estimator:

$$\log l_i = \sum_{t=1}^{T_i} \sum_{j=2}^J Y_{ijt} \mathbf{x}_{it} \beta_j - \log \sum_{\tilde{Y}_{ijt} \in \Psi(c_i)} \exp \left(\sum_{t=1}^{T_i} \sum_{j=2}^J \tilde{Y}_{ijt} \mathbf{x}_{it} \beta_j \right)$$

- $\Psi(c_i)$ is a set of all permutations of individual i 's observed sequence of outcomes that satisfies the condition $\sum_{t=1}^{T_i} \tilde{Y}_{it} = c_i$. That is,

$$\Psi(c_i) = \left\{ \tilde{Y}_i = (\tilde{Y}_{i1}, \dots, \tilde{Y}_{iT_i}) \mid \sum_{t=1}^{T_i} \tilde{Y}_{it} = c_i \right\}$$

- where $\tilde{Y}_{it} = (\tilde{Y}_{i1t}, \dots, \tilde{Y}_{iJt})$ is a vector of indicators with respect to the permutations of the observed outcome sequence Y_i .
- For example, the sequence $Y_i = (3, 2, 3)$ has a total of three permutations (including the original sequence): $(2, 3, 3)$, $(3, 2, 3)$, and $(3, 3, 2)$

Curse of dimensionality: random-effects estimator

- For the random-effects estimator, the curse is rooted in J , the number of outcomes, because we have a $J - 1$ dimensional integral.
- Computation time can be high for more than just three or four outcomes.
- For example, with six outcomes, we have to approximate a five-dimensional integral.
- Using seven quadrature integration points, we would end up with $7^5 = 16,807$ integration points.
- If computation time becomes infeasible one might consider using a single, shared variance component, if appropriate (option `covariance(shared)`).

Curse of dimensionality: fixed-effects estimator

- The curse of dimensionality in case of the fixed-effects estimator is rooted mainly in T_i , the number of repeated observations, and potentially in J .
- The number of permutations in $\Psi(c_i)$ grows exponentially with T_i and can become infeasibly large.
- The number of permutations of panel i 's observed vector of outcomes is:

$$K_i = \frac{T_i!}{c_{i1}! \cdots c_{ij}! \cdots c_{iJ}!}$$

- For example, for an individual with 15 repeated observations in a dataset with 6 outcomes, $j = 1, 2, \dots, 6$, with the sequence of outcomes $Y_i = (3, 3, 3, 2, 4, 1, 1, 5, 4, 6, 6, 1, 1, 2, 4)$ the size of the set of permutations of this outcome vector is

$$K_i = \frac{15!}{4! 2! 3! 3! 1! 2!} = 378,378,000$$

- Realistically, T_i should not be larger than around 9 or 10.

When to use which estimator

- If the assumptions of the random-effects estimator hold, both the random-effects and fixed-effects estimators are consistent, but the random-effects estimator is more efficient. If the assumptions do not hold, the random-effects estimator becomes inconsistent.
- If all we care about is estimating β_j , we should use the fixed-effects estimator because we have to make fewer assumptions.
- Practical consideration: beyond the gain in efficiency, if we want to perform marginal predictions or include time-constant covariates, the random-effects estimator is preferable.
- With the fixed-effects estimator, we cannot make predictions that account for the panel-level unobservables because these are not parameterized.
- A potential way to decide between the two is to perform a Hausman test.

Some notable features

- In the context of the **random-effects** estimator, we can decide what restrictions to place on the variance-covariance matrix of the random effects:
 - ▶ `covariance(independent)`; distinct variances, the default
 - ▶ `covariance(unstructured)`; no restrictions
 - ▶ `covariance(exchangeable)`; equal variances and covariances
 - ▶ `covariance(identity)`; equal variances, covariance 0
 - ▶ `covariance(shared)`; one common random effect
- With the **fixed-effects** estimator, the curse of dimensionality can be alleviated to some degree by taking random samples from the permutation sets using option `rsample()`.
- `xtmlogit` can be used with `svy`
- `bayes:` `xtmlogit` (random-effects estimator)

Example data (1)

```
. webuse estatus  
(Fictional employment status data)
```

```
. describe
```

Contains data from <https://www.stata-press.com/data/r17/estatus.dta>

```
Observations:    4,761    Fictional employment status data  
Variables:        8        26 Feb 2021 15:17
```

Variable name	Storage type	Display format	Value label	Variable label
id	int	%9.0g		Respondent ID
year	int	%9.0g		Year of survey
estatus	byte	%18.0g	alt	Employment status
hhchild	byte	%9.0g	noyes	Children <18 years old in household
hhincome	int	%10.0g		Annual household income (in \$1,000s)
hhsigno	byte	%10.0g	noyes	Significant other living in household
bwinner	byte	%10.0g	noyes	Primary/sole breadwinner in household
age	byte	%10.0g		Age (in years)

```
Sorted by: id year
```

```
. tab estatus
```

Employment status	Freq.	Percent	Cum.
Out of labor force	1,682	35.33	35.33
Unemployed	703	14.77	50.09
Employed	2,376	49.91	100.00
Total	4,761	100.00	

Example data (2)

```
. list in 1/10, compress noobs
```

id	year	estatus	hhc_d	hhi_e	hhs_o	bwi_r	age
1	2002	Out of labor force	Yes	56	Yes	Yes	34
1	2004	Out of labor force	Yes	62	No	Yes	36
1	2006	Employed	Yes	68	Yes	No	38
1	2008	Employed	No	73	No	Yes	40
1	2010	Out of labor force	No	85	Yes	No	42
2	2002	Employed	No	28	Yes	No	19
2	2004	Out of labor force	Yes	26	Yes	Yes	21
2	2006	Employed	Yes	31	No	No	23
2	2008	Employed	Yes	36	Yes	Yes	25
2	2010	Unemployed	Yes	42	Yes	No	27

Using mlogit

```
. mlogit estatus i.hhchild age hhincome i.hhsigno i.bwinner
```

```
Iteration 0: log likelihood = -4746.2288  
Iteration 1: log likelihood = -4586.1414  
Iteration 2: log likelihood = -4581.5114  
Iteration 3: log likelihood = -4581.496  
Iteration 4: log likelihood = -4581.496
```

Multinomial logistic regression

```
Number of obs = 4,761  
LR chi2(10) = 329.47  
Prob > chi2 = 0.0000  
Pseudo R2 = 0.0347
```

Log likelihood = -4581.496

estatus	Coefficient	Std. err.	z	P> z	[95% conf. interval]	
Out_of_lab_e						
hhchild						
Yes	.3603056	.0767921	4.69	0.000	.2097958	.5108155
age	-.0051251	.0050107	-1.02	0.306	-.0149459	.0046958
hhincome	-.0041118	.0012464	-3.30	0.001	-.0065546	-.001669
hhsigno						
Yes	.4251824	.085001	5.00	0.000	.2585835	.5917814
bwinner						
Yes	-.4341223	.0646731	-6.71	0.000	-.5608793	-.3073654
_cons	-.3054547	.2219434	-1.38	0.169	-.7404558	.1295464
Unemployed						
hhchild						
Yes	-.0421006	.1026229	-0.41	0.682	-.2432378	.1590366
age	.0027158	.0066839	0.41	0.685	-.0103845	.015816
hhincome	-.0278535	.0020728	-13.44	0.000	-.0319161	-.0237909
hhsigno						
Yes	.1070661	.1100785	0.97	0.331	-.1086839	.3228161
bwinner						
Yes	-.2185759	.0879215	-2.49	0.013	-.3908989	-.046253
_cons	.0142404	.293933	0.05	0.961	-.5618577	.5903385
Employed						
	(base outcome)					

Panel data model with random effects (1)

```
. xtset id
Panel variable: id (unbalanced)
. xtmlogit estatus i.hhchild age hhincome i.hhsigno i.bwinner
Fitting comparison model ...
Refining starting values:
Grid node 0:   log likelihood = -4483.1721
Grid node 1:   log likelihood = -4516.6753
Fitting full model:
Iteration 0:   log likelihood = -4483.1721
Iteration 1:   log likelihood = -4474.3849
Iteration 2:   log likelihood = -4468.9353
Iteration 3:   log likelihood = -4468.8415
Iteration 4:   log likelihood = -4468.8413
Random-effects multinomial logistic regression
Group variable: id
Random effects u_i ~ Gaussian
Number of obs   = 4,761
Number of groups = 800
Obs per group:
    min = 5
    avg = 6.0
    max = 7
Integration method: mvaghermite
Integration pts. = 7
Wald chi2(10)   = 239.26
Prob > chi2     = 0.0000
Log likelihood = -4468.8413
```

estatus	Coefficient	Std. err.	z	P> z	[95% conf. interval]
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Panel data model with random effects (2)

estatus	Coefficient	Std. err.	z	P> z	[95% conf. interval]	
Out_of_lab_e						
hhchild						
Yes	.4628125	.0962758	4.81	0.000	.2741154	.6515096
age	-.004825	.0066428	-0.73	0.468	-.0178446	.0081946
hhincome	-.0046922	.001839	-2.55	0.011	-.0082965	-.0010879
hhsigno						
Yes	.4967056	.0946442	5.25	0.000	.3112063	.6822049
bwinner						
Yes	-.4740919	.0727992	-6.51	0.000	-.6167756	-.3314082
_cons	-.4787579	.2845139	-1.68	0.092	-1.036395	.0788792
Unemployed						
hhchild						
Yes	-.0401989	.119596	-0.34	0.737	-.2746027	.1942049
age	.0042644	.0081818	0.52	0.602	-.0117716	.0203004
hhincome	-.0308468	.0026529	-11.63	0.000	-.0360463	-.0256473
hhsigno						
Yes	.0968	.1192659	0.81	0.417	-.1369568	.3305568
bwinner						
Yes	-.2252587	.0951984	-2.37	0.018	-.4118441	-.0386733
_cons	-.0953821	.3508736	-0.27	0.786	-.7830817	.5923175
Employed						
	(base outcome)					
var (u1)	.8587807	.1090216			.6696113	1.101392
var (u2)	.7370366	.1388917			.5094287	1.066338

LR test vs. multinomial logit: chi2(2) = 225.31

Prob > chi2 = 0.0000

Note: LR test is conservative and provided only for reference.



Displaying relative-risk ratios

```
. xtmlogit, rrr
Random-effects multinomial logistic regression
Group variable: id
Random effects u_i ~ Gaussian

Number of obs   - 4,761
Number of groups - 800
Obs per group:
    min   - 5
    avg   - 6.0
    max   - 7

Integration method: mvaghermite
Integration pts. - 7

Log likelihood - -4468.8413
Wald chi2(10) - 239.26
Prob > chi2 - 0.0000
```

estatus	RRR	Std. err.	z	P> z	[95% conf. interval]	
Out_of_lab_e						
hhchild						
Yes	1.588535	.1529375	4.81	0.000	1.315367	1.918435
age	.9951866	.0066108	-0.73	0.468	.9823137	1.008228
hhincome	.9953188	.0018303	-2.55	0.011	.9917379	.9989127
hhsigno						
Yes	1.643299	.1555288	5.25	0.000	1.365071	1.978235
bwinner						
Yes	.6224501	.0453138	-6.51	0.000	.5396818	.7179121
_cons	.6195525	.1762713	-1.68	0.092	.3547312	1.082074
Unemployed						
hhchild						
Yes	.9605983	.1148837	-0.34	0.737	.7598739	1.214345
age	1.004274	.0082168	0.52	0.602	.9882974	1.020508
hhincome	.9696241	.0025723	-11.63	0.000	.9645956	.9746788
hhsigno						
Yes	1.10164	.1313881	0.81	0.417	.8720079	1.391743
bwinner						
Yes	.7983097	.0759978	-2.37	0.018	.6624275	.9620649
_cons	.9090255	.3189531	-0.27	0.786	.4569955	1.808174
Employed (base outcome)						
var(u1)	.8587807	.1090216			.6696113	1.101392
var(u2)	.7370366	.1388917			.5094287	1.066338

Note: Estimates are transformed only in the first 3 equations to relative-risk ratios.
Note: **_cons** estimates baseline relative risk (conditional on zero random effects).
LR test vs. multinomial logit: chi2(2) = 225.31 Prob > chi2 = 0.0000
Note: LR test is conservative and provided only for reference.

Unstructured covariance estimation (1)

```
. xtlogit estatus i.hhchild age hhincome i.hhsigno i.bwinner, covariance(unstr  
> uctured)
```

```
Fitting comparison model ...
```

```
Refining starting values:
```

```
Grid node 0: log likelihood = -4483.1721
```

```
Grid node 1: log likelihood = -4516.6753
```

```
Fitting full model:
```

```
Iteration 0: log likelihood = -4483.1721
```

```
Iteration 1: log likelihood = -4454.5313
```

```
Iteration 2: log likelihood = -4438.9076
```

```
Iteration 3: log likelihood = -4438.2905
```

```
Iteration 4: log likelihood = -4438.2887
```

```
Iteration 5: log likelihood = -4438.2887
```

```
Random-effects multinomial logistic regression
```

```
Group variable: id
```

```
Random effects u_i ~ Gaussian
```

```
Number of obs = 4,761
```

```
Number of groups = 800
```

```
Obs per group:
```

```
min = 5
```

```
avg = 6.0
```

```
max = 7
```

```
Integration method: mvaghermite
```

```
Integration pts. = 7
```

```
Log likelihood = -4438.2887
```

```
Wald chi2(10) = 242.93
```

```
Prob > chi2 = 0.0000
```

estatus	Coefficient	Std. err.	z	P> z	[95% conf. interval]
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Unstructured covariance estimation (2)

estatus	Coefficient	Std. err.	z	P> z	[95% conf. interval]	
Out_of_lab_e						
hhchild						
Yes	.4924799	.1002988	4.91	0.000	.295898	.6890619
age	-.004219	.0070064	-0.60	0.547	-.0179513	.0095133
hhincome	-.006046	.001992	-3.04	0.002	-.0099503	-.0021417
hhsigno						
Yes	.5036976	.0966982	5.21	0.000	.3141726	.6932225
bwinner						
Yes	-.489057	.0745454	-6.56	0.000	-.6351632	-.3429507
_cons	-.3930378	.298386	-1.32	0.188	-.9778636	.191788
Unemployed						
hhchild						
Yes	.0399687	.1238417	0.32	0.747	-.2027565	.2826939
age	.0045538	.0085081	0.54	0.592	-.0121219	.0212294
hhincome	-.0315377	.0027426	-11.50	0.000	-.0369131	-.0261624
hhsigno						
Yes	.1495817	.1214242	1.23	0.218	-.0884053	.3875687
bwinner						
Yes	-.2552257	.0968165	-2.64	0.008	-.4449826	-.0654689
_cons	-.0417024	.3633406	-0.11	0.909	-.7538368	.670432
Employed						
	(base outcome)					
var(u1)	1.132081	.1331468			.899012	1.425572
var(u2)	1.102612	.1698422			.8152803	1.49121
cov(u1,u2)	.7871916	.1222148	6.44	0.000	.547655	1.026728

LR test vs. multinomial logit: $\chi^2(3) = 286.41$

Prob > $\chi^2 = 0.0000$

Note: LR test is conservative and provided only for reference.

Standard deviation and correlation

```
. estat sd
```

estatus	Coefficient	Std. err.	z	P> z	[95% conf. interval]	
sd(u1)	1.063993	.0625694			.9481624	1.193973
sd(u2)	1.050053	.0808731			.9029287	1.221151
corr(u1,u2)	.7045801	.0632646	11.14	0.000	.5581225	.8084624

Panel data model with conditional fixed effects (1)

```
. xtlogit estatus i.hhchild age hhincome i.hhsigno i.bwinner, fe
note: 80 groups (451 obs) omitted because of no variation in the outcome
      variable over time.
```

```
Computing initial values ...
```

```
Setting up 26,168 permutations:
```

```
....10%....20%....30%....40%....50%....60%....70%....80%....90%....100%
```

```
Fitting full model:
```

```
Iteration 0:  log likelihood = -2136.5919
```

```
Iteration 1:  log likelihood = -2136.2728
```

```
Iteration 2:  log likelihood = -2136.2728
```

```
Fixed-effects multinomial logistic regression
```

```
Group variable: id
```

```
Number of obs   = 4,310
```

```
Number of groups = 720
```

```
Obs per group:
```

```
min = 5
```

```
avg = 6.0
```

```
max = 7
```

```
LR chi2(10)    = 103.29
```

```
Prob > chi2    = 0.0000
```

```
Log likelihood = -2136.2728
```

estatus	Coefficient	Std. err.	z	P> z	[95% conf. interval]
---------	-------------	-----------	---	------	----------------------

Panel data model with conditional fixed effects (2)

estatus	Coefficient	Std. err.	z	P> z	[95% conf. interval]	
Out_of_lab_e						
hhchild						
Yes	.5881852	.1258696	4.67	0.000	.3414854	.834885
age	-.0003842	.0147741	-0.03	0.979	-.0293409	.0285725
hhincome	-.0122043	.0088464	-1.38	0.168	-.029543	.0051344
hhsigno						
Yes	.5090034	.100111	5.08	0.000	.3127893	.7052174
bwinner						
Yes	-.4655745	.0782841	-5.95	0.000	-.6190085	-.3121406
Unemployed						
hhchild						
Yes	.163612	.1638934	1.00	0.318	-.1576132	.4848372
age	.0063355	.019404	0.33	0.744	-.0316957	.0443667
hhincome	-.029742	.0120031	-2.48	0.013	-.0532676	-.0062164
hhsigno						
Yes	.1173192	.1301364	0.90	0.367	-.1377435	.3723819
bwinner						
Yes	-.2489958	.1030027	-2.42	0.016	-.4508773	-.0471142
Employed	(base outcome)					

Fixed-effects model with permutation sampling (1)

```
. xtset id year
Panel variable: id (unbalanced)
Time variable: year, 2002 to 2014, but with gaps
Delta: 1 unit

. xtmlogit estatus i.hhchild age hhincome i.hhsigno i.bwinner, fe rsample(10, r
> seed(123))
note: option vce() set to vce(robust) because of permutation sampling.
note: 80 groups (451 obs) omitted because of no variation in the outcome
variable over time.

Computing initial values ...
Setting up 3,495 permutations:
....10%....20%....30%....40%....50%....60%....70%....80%....90%....100%
Fitting full model:
Iteration 0: log pseudolikelihood = -908.26163
Iteration 1: log pseudolikelihood = -906.4585
Iteration 2: log pseudolikelihood = -906.45801
Iteration 3: log pseudolikelihood = -906.45801

Fixed-effects multinomial logistic regression
Group variable: id
Number of obs = 4,310
Number of groups = 720
Obs per group:
    min = 5
    avg = 6.0
    max = 7
Wald chi2(10) = 72.91
Prob > chi2 = 0.0000
Log pseudolikelihood = -906.45801
(Std. err. adjusted for 720 clusters in id)
```

estatus	Robust				
	Coefficient	std. err.	z	P> z	[95% conf. interval]

Fixed-effects model with permutation sampling (2)

estatus	Robust				
	Coefficient	std. err.	z	P> z	[95% conf. interval]
Out_of_lab_e					
hhchild					
Yes	.5827048	.1487376	3.92	0.000	.2911845 .8742251
age	-.0055092	.0168589	-0.33	0.744	-.0385521 .0275337
hhincome	-.0142493	.0100457	-1.42	0.156	-.0339386 .0054399
hhsigno					
Yes	.4441511	.1213382	3.66	0.000	.2063325 .6819696
bwinner					
Yes	-.4613157	.0978061	-4.72	0.000	-.6530122 -.2696193
Unemployed					
hhchild					
Yes	.1714144	.1831194	0.94	0.349	-.187493 .5303217
age	-.0046656	.0217006	-0.21	0.830	-.0471979 .0378668
hhincome	-.0344681	.0131707	-2.62	0.009	-.0602821 -.0086541
hhsigno					
Yes	-.0760533	.1396542	-0.54	0.586	-.3497705 .197664
bwinner					
Yes	-.2889772	.1178606	-2.45	0.014	-.5199797 -.0579747
Employed	(base outcome)				

Performing a Hausman test

```
. qui xtlogit estatus i.hhchild age hhincome i.hhsigno i.bwinner, fe
. estimates store FE
. qui xtlogit estatus i.hhchild age hhincome i.hhsigno i.bwinner
. estimates store RE
. hausman FE RE, alleqs
```

	Coefficients			
	(b) FE	(B) RE	(b-B) Difference	sqrt(diag(V_b-V_B)) Std. err.
Out_of_lab_e				
1.hhchild	.5881852	.4628125	.1253727	.0810809
age	-.0003842	-.004825	.0044408	.0131965
hhincome	-.0122043	-.0046922	-.0075122	.0086532
1.hhsigno	.5090034	.4967056	.0122977	.0326296
1.bwinner	-.4655745	-.4740919	.0085173	.0287868
Unemployed				
1.hhchild	.163612	-.0401989	.203811	.1120618
age	.0063355	.0042644	.0020711	.0175947
hhincome	-.029742	-.0308468	.0011048	.0117062
1.hhsigno	.1173192	.0968	.0205192	.0520686
1.bwinner	-.2489958	-.2252587	-.0237371	.0393297

b = Consistent under H0 and Ha; obtained from **xtlogit**.

B = Inconsistent under Ha, efficient under H0; obtained from **xtlogit**.

Test of H0: Difference in coefficients not systematic

$$\begin{aligned}\text{chi2}(10) &= (b-B)' [(V_b-V_B)^{-1}] (b-B) \\ &= 8.05\end{aligned}$$

Prob > chi2 = 0.6238

Using margins after xtmlogit, re

```
. qui xtmlogit estatus i.hhchild age hhincome i.hhsigno i.bwinner
. margins i.hhchild
Predictive margins                                Number of obs = 4,761
Model VCE: OIM
1. _predict: Pr(estatus==Out_of_labor_force), predict(pr outcome(1))
2. _predict: Pr(estatus==Unemployed), predict(pr outcome(2))
3. _predict: Pr(estatus==Employed), predict(pr outcome(3))
```

	Delta-method				
	Margin	std. err.	z	P> z	[95% conf. interval]
_predict#					
hhchild					
1#No	.3021986	.0131047	23.06	0.000	.2765138 .3278834
1#Yes	.3912783	.0119865	32.64	0.000	.3677852 .4147714
2#No	.1630791	.0101239	16.11	0.000	.1432367 .1829216
2#Yes	.139782	.0079417	17.60	0.000	.1242167 .1553474
3#No	.5347223	.0136504	39.17	0.000	.507968 .5614766
3#Yes	.4689397	.0116018	40.42	0.000	.4462006 .4916787

- Note: using margins after xtmlogit, fe is not advisable.

predict

- `xtmlogit, re`
 - ▶ `pr` marginal probability of the specified outcome; the default
 - ▶ `pcr` conditional probability of the specified outcome
 - ▶ `pu0` probability of the specified outcome, assuming zero random effects
 - ▶ `xb` linear prediction for the specified outcome, including random effects
 - ▶ `xb0` linear prediction for the specified outcome, assuming zero random effects
- `xtmlogit, fe`
 - ▶ `pu0` probability of the specified outcome, assuming zero fixed effects; the default
 - ▶ `xb` linear prediction for the specified outcome, assuming zero random effects
- Option `outcome()` to specify a particular outcome.
- Or use `stub*` to get predictions for all outcomes.

Thank You!