

## Using Stata 9 to Model Complex Nonlinear Relationships with Restricted Cubic Splines

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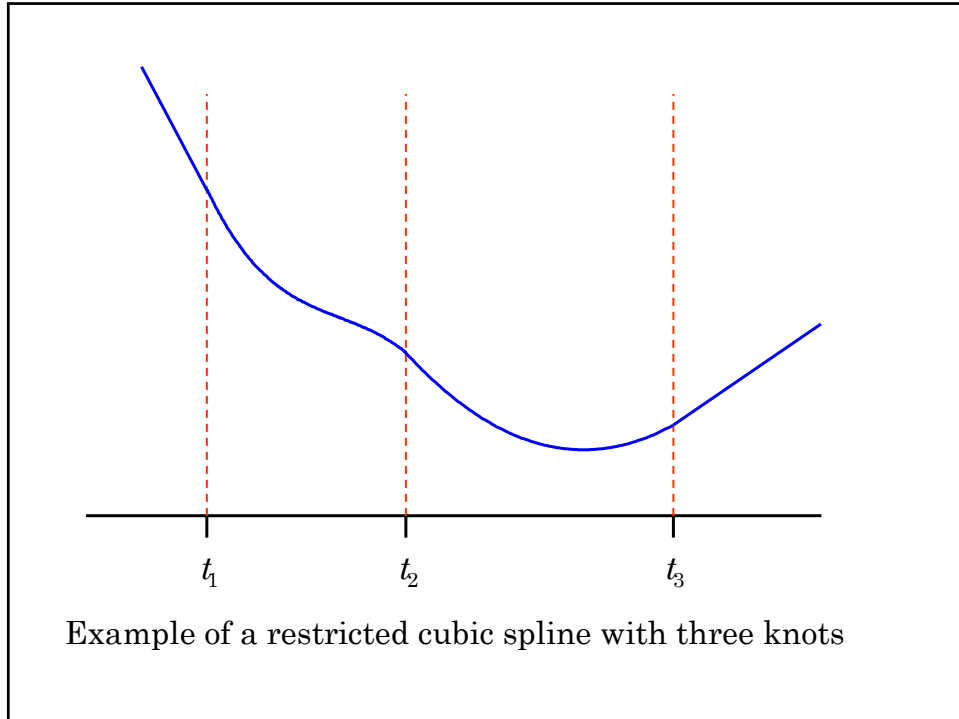
### Restricted Cubic Splines (Natural Splines)

Given  $\{(x_i, y_i) : i = 1, \dots, n\}$

We wish to model  $y_i$  as a function of  $x_i$  using a flexible non-linear model.

In a **restricted cubic spline model** we introduce  $k$  knots on the  $x$ -axis located at  $t_1, t_2, \dots, t_k$ . We select a model of the expected value of  $y$  given  $x$  that is

- ❖ linear before  $t_1$  and after  $t_k$ .
- ❖ consists of piecewise cubic polynomials between adjacent knots (i.e. of the form  $ax^3 + bx^2 + cx + d$ )
- ❖ continuous and smooth at each knot, with continuous first and second derivatives.



Given  $x$  and  $k$  knots, a restricted cubic spline can be defined by

$$y = \alpha + x_1\beta_1 + x_2\beta_2 + \dots + x_{k-1}\beta_{k-1}$$

where

$$x_1 = x$$

$$x_j = (x - t_{j-1})_+^3 - \frac{(x - t_{k-1})_+^3 (t_k - t_{j-1})}{(t_k - t_{k-1})} + \frac{(x - t_k)_+^3 (t_{k-1} - t_{j-1})}{(t_k - t_{k-1})}$$

for  $j = 2, \dots, k - 1$

$$(u)_+ = \begin{cases} u & : u > 0 \\ 0 & : u \leq 0 \end{cases}$$

These covariates are

- ❖ functions of  $x$  and the knots but are independent of  $y$ .
- ❖  $x_1 = x$  and hence the linear hypothesis is tested by  $\beta_2 = \beta_3 = \dots = \beta_{k-1} = 0$ .
- ❖ Stata programs to calculate  $x_1, \dots, x_{k-1}$  are available on the web.  
(Run `findit spline` from within Stata.)
- ❖ One of these is `rc_spline`

```
rc_spline xvar [fweight] [if exp] [in range]
           [,nknots(#) knots(numlist)]
```

generates the covariates  $x_1, \dots, x_{k-1}$  corresponding to  $x = \mathbf{xvar}$

`nknots(#)` option specifies the number of knots  
(5 by default)

`knots(numlist)` option specifies the knot locations

This program generates the spline covariates named

```
  _Sxvar1 = xvar
  _Sxvar2
  _Sxvar3
  .
  .
  .
```

Default knot locations are placed at the quantiles of the  $x$  variable given in the following table (Harrell 2001).

Number of knots $k$	Knot locations expressed in quantiles of the $x$ variable							
3	0.1	0.5	0.9					
4	0.05	0.35	0.65	0.95				
5	0.05	0.275	0.5	0.725	0.95			
6	0.05	0.23	0.41	0.59	0.77	0.95		
7	0.03	0.183	0.342	0.5	0.658	0.817	0.98	

### SUPPORT Study

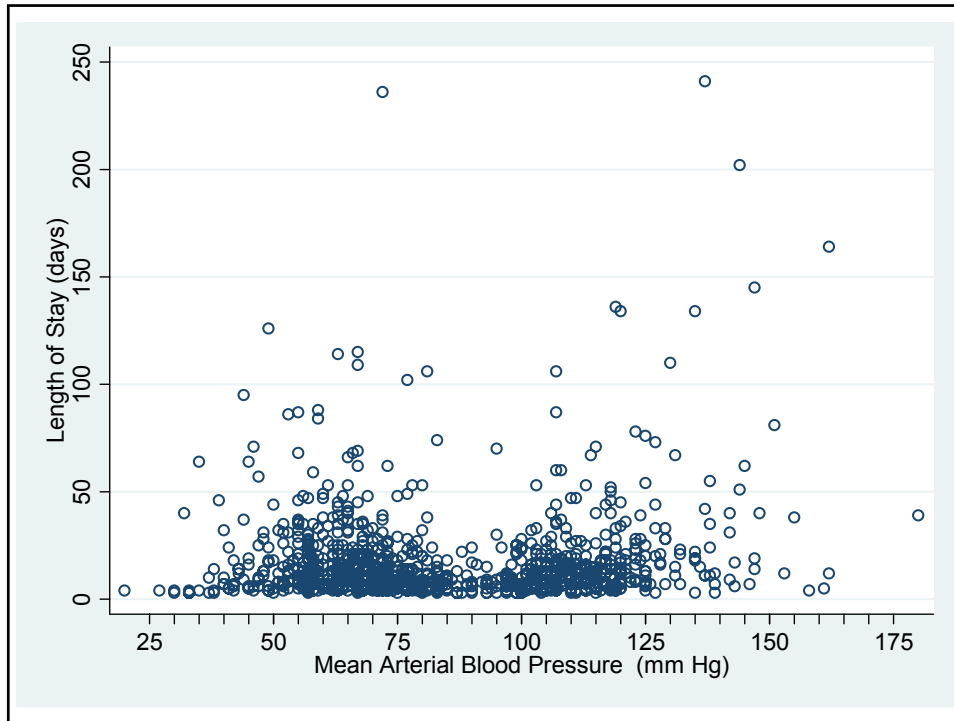
A prospective observational study of hospitalized patients

Lynn & Knauss: "Background for SUPPORT."  
*J Clin Epidemiol* 1990; 43: 1S - 4S.

los = length of stay in days.

meanbp = baseline mean arterial blood pressure

hospdead =  $\begin{cases} 1: \text{Patient died in hospital} \\ 0: \text{Patient discharged alive} \end{cases}$



```
. gen log_los = log(los)
```

```
. rc_spline meanbp
```

```
number of knots = 5  
value of knot 1 = 47  
value of knot 2 = 66  
value of knot 3 = 78  
value of knot 4 = 106  
value of knot 5 = 129
```

Define 4 spline covariates associated with 5 knots at their default locations.

The covariates are named

```
_Smeanbp1  
_Smeanbp2  
_Smeanbp3  
_Smeanbp4
```

```
. gen log_los = log(los)
```

```
. rc_spline meanbp  
number of knots = 5  
value of knot 1 = 47  
value of knot 2 = 66  
value of knot 3 = 78  
value of knot 4 = 106  
value of knot 5 = 129
```

Regress **log\_los** against all variables that start with the letters **\_S**. That is, against  
**\_Smeanbp1**  
**\_Smeanbp2**  
**\_Smeanbp3**  
**\_Smeanbp4**

```
. regress log_los _S*
```

Source	SS	df	MS	Number of obs =	996
Model	60.9019393	4	15.2254848	F( 4, 991) =	24.70
Residual	610.872879	991	.616420665	Prob > F =	0.0000
Total	671.774818	995	.675150571	R-squared =	0.0907
				Adj R-squared =	0.0870
				Root MSE =	.78512

log_los	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
_Smeanbp1	.0296009	.0059566	4.97	0.000	.017912	.0412899
_Smeanbp2	-.3317922	.0496932	-6.68	0.000	-.4293081	-.2342762
_Smeanbp3	1.263893	.1942993	6.50	0.000	.8826076	1.645178
_Smeanbp4	-1.124065	.1890722	-5.95	0.000	-1.495092	-.7530367
_cons	1.03603	.3250107	3.19	0.001	.3982422	1.673819

```
. test _Smeanbp2 _Smeanbp3 _Smeanbp4
```

```
( 1) _Smeanbp2 = 0  
( 2) _Smeanbp3 = 0  
( 3) _Smeanbp4 = 0
```

```
F( 3, 991) = 30.09  
Prob > F = 0.0000
```

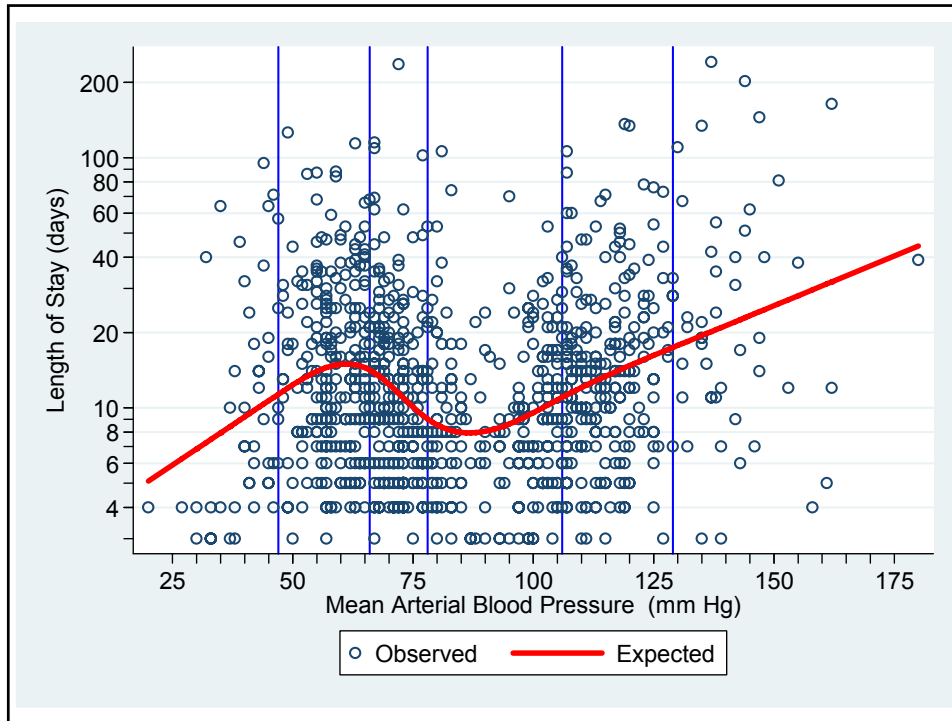
Test the null hypothesis that there is a linear relationship between **meanbp** and **log\_los**.

```
. predict y_hat, xb
```

**y\_hat** is the estimated expected value of **log\_los** under this model.

Graph a scatterplot of **log\_los** vs. **meanbp** together with a line plot of the expected **log\_los** vs. **meanbp**.

```
. scatter log_los meanbp ,msymbol(Oh) ///  
> || line y_hat meanbp ///  
> , xlabel(25 (25) 175) xtick(30 (5) 170) clcolor(red) ///  
> clwidth(thick) xline(47 66 78 106 129, lcolor(blue)) ///  
> ylabel('yloglabel', angle(0)) ytick('ylogtick') ///  
> ytitle("Length of Stay (days)") ///  
> legend(order(1 "Observed" 2 "Expected")) name(knot5, replace) ///
```



```

. drop _S* y_hat

. rc_spline meanbp, nknots(7)
  number of knots = 7
  value of knot 1 = 41
  value of knot 2 = 60
  value of knot 3 = 69
  value of knot 4 = 78
  value of knot 5 = 101.3251
  value of knot 6 = 113
  value of knot 7 = 138.075

. regress log_los _S*

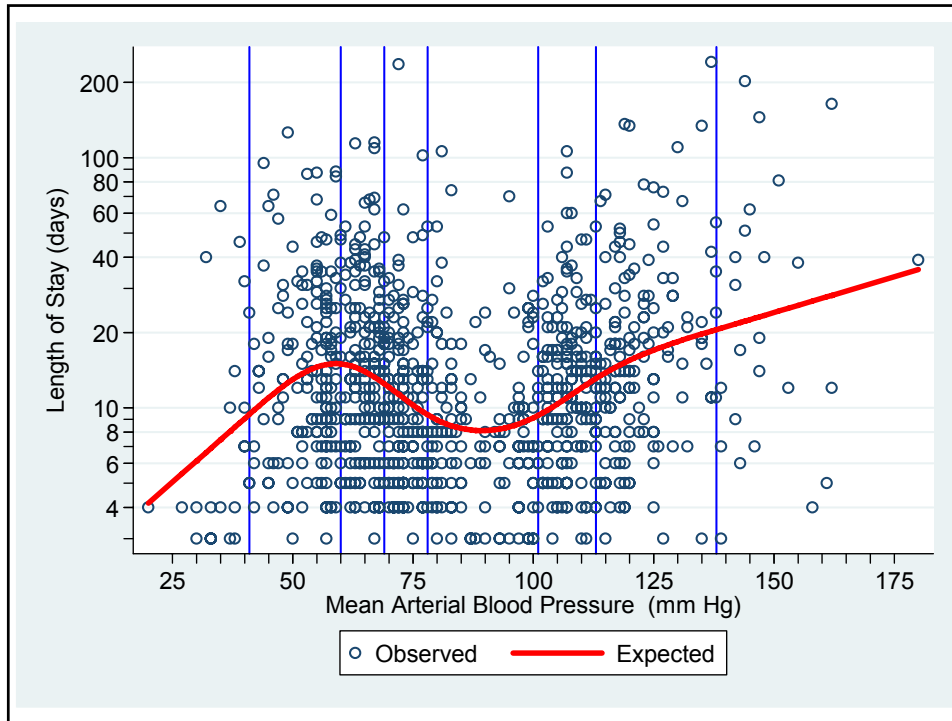
{ Output omitted }

. predict y_hat, xb

. scatter log_los meanbp ,msymbol(Oh)
> || line y_hat meanbp
> , xlabel(25 (25) 175) xtick(30 (5) 170) ccolor(red)
> clwidth(thick) xline(41 60 69 78 101 113 138, lcolor(blue))
> ylabel(`yloglabel', angle(0)) ytick(`ylogtick')
> ytitle("Length of Stay (days)")
> legend(order(1 "Observed" 2 "Expected")) name(setknots, replace)

```

Define 6 spline covariates associated with 7 knots at their default locations.



```

. drop _S* y_hat

. rc_spline meanbp, nknots(7) knots(40(17)142)
number of knots = 7
value of knot 1 = 40
value of knot 2 = 57
value of knot 3 = 74
value of knot 4 = 91
value of knot 5 = 108
value of knot 6 = 125
value of knot 7 = 142

. regress log_los _S*
{ Output omitted }

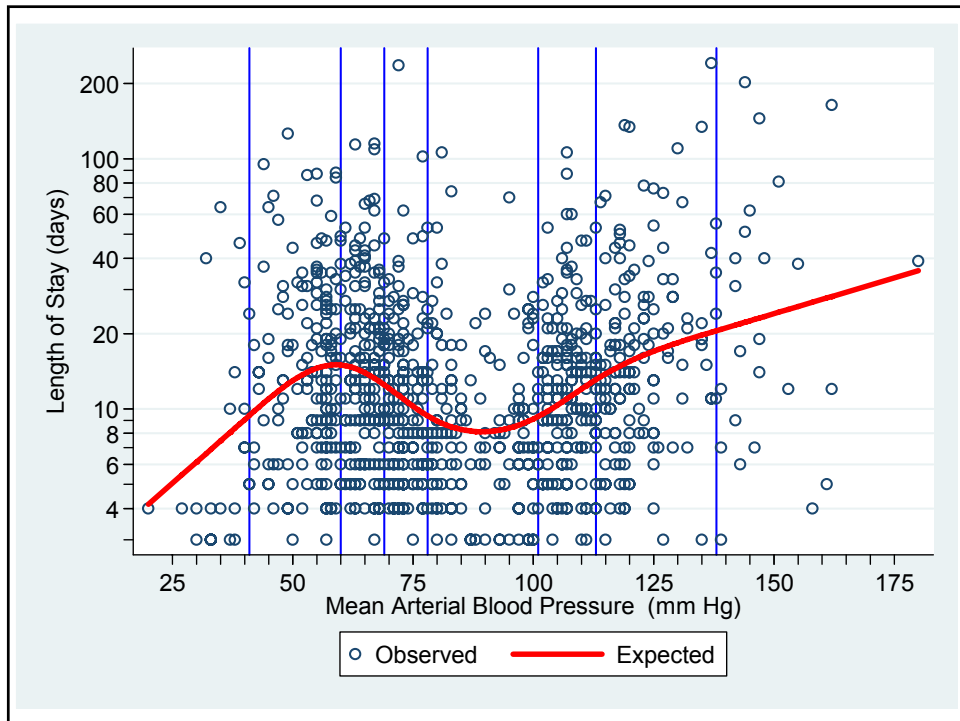
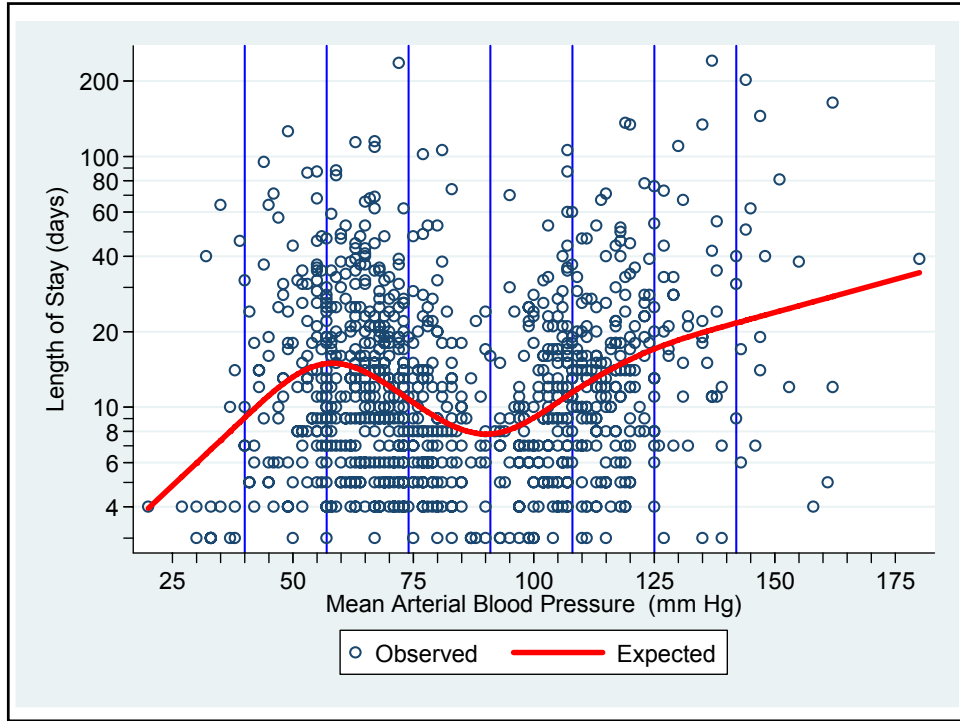
. predict y_hat, xb

. scatter log_los meanbp ,msymbol(Oh)
> || line y_hat meanbp
> , xlabel(25 (25) 175) xtick(30 (5) 170) clcolor(red)
> clinewidth(thick) xline(40(17)142, lcolor(blue))
> ylabel(`yloglabel', angle(0)) ytick(`ylogtick')
> ytitle("Length of Stay (days)")
> legend(order(1 "Observed" 2 "Expected")) name(setknots, replace)

```

Define 6 spline covariates associated with 7 knots at evenly spaced locations.





```

. drop _S* y_hat
. rc_spline meanbp, nknots(7)
{ Output omitted }
. regress log_los _S*
{ Output omitted }
. predict y_hat, xb
. predict se, stdp
. generate lb = y_hat - invttail(_N-7, 0.025)*se
. generate ub = y_hat + invttail(_N-7, 0.025)*se

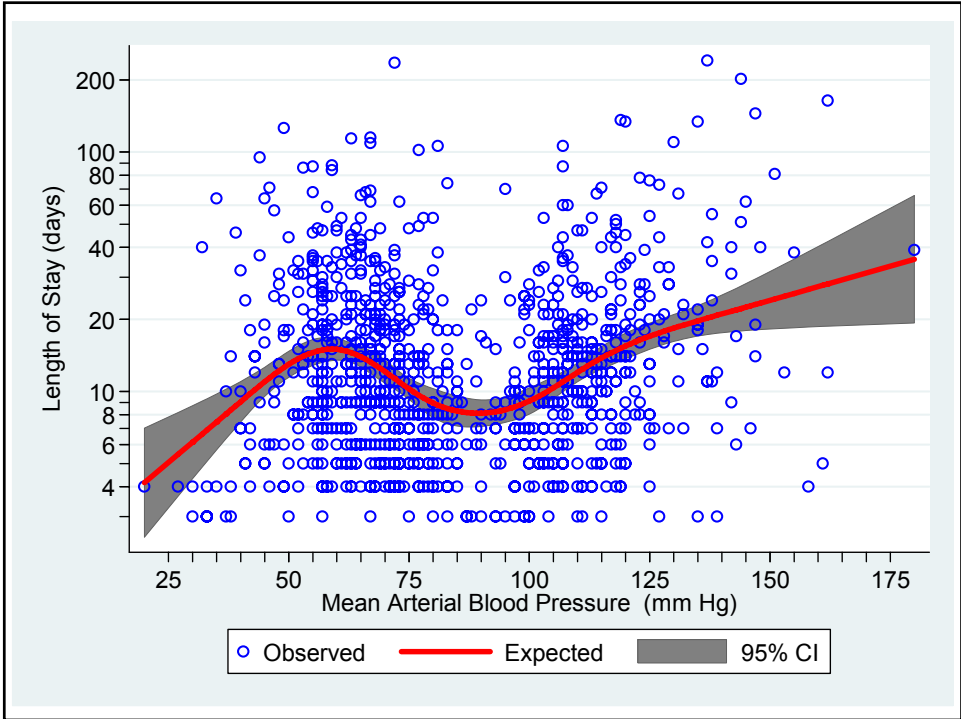
. twoway rarea lb ub meanbp , bcolor(gs6) lwidth(none)
> || scatter log_los meanbp ,msymbol(Oh) mcolor(blue)
> || line y_hat meanbp, xlabel(25 (25) 175) xtick(30 (5) 170)
> clcolor(red) clwidth(thick) ytitle("Length of Stay (days)")
> ylabel('yloglabel', angle(0)) ytick('ylogtick') name(ci,replace)
> legend(rows(1) order(2 "Observed" 3 "Expected" 1 "95% CI" ))

```

Define **se** to be the standard error of **y\_hat**.

Define **lb** and **ub** to be the lower and upper bound of a 95% confidence interval for **y\_hat**.

This **twoway** plot includes an **rarea** plot of the shaded 95% confidence interval for **y\_hat**.



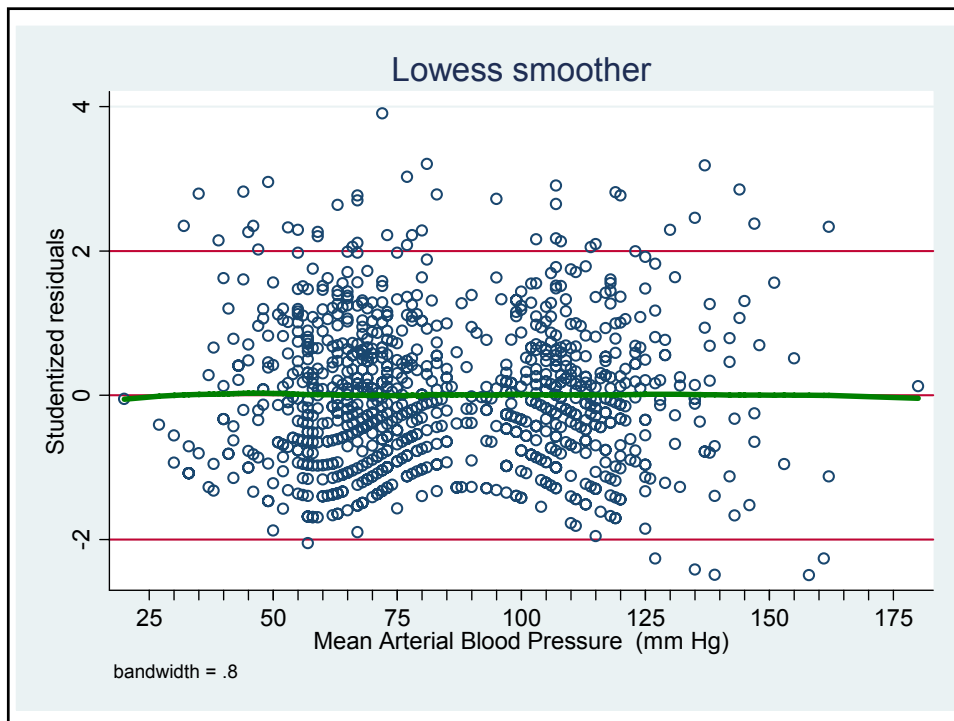
```
. predict rstudent, rstudent
```

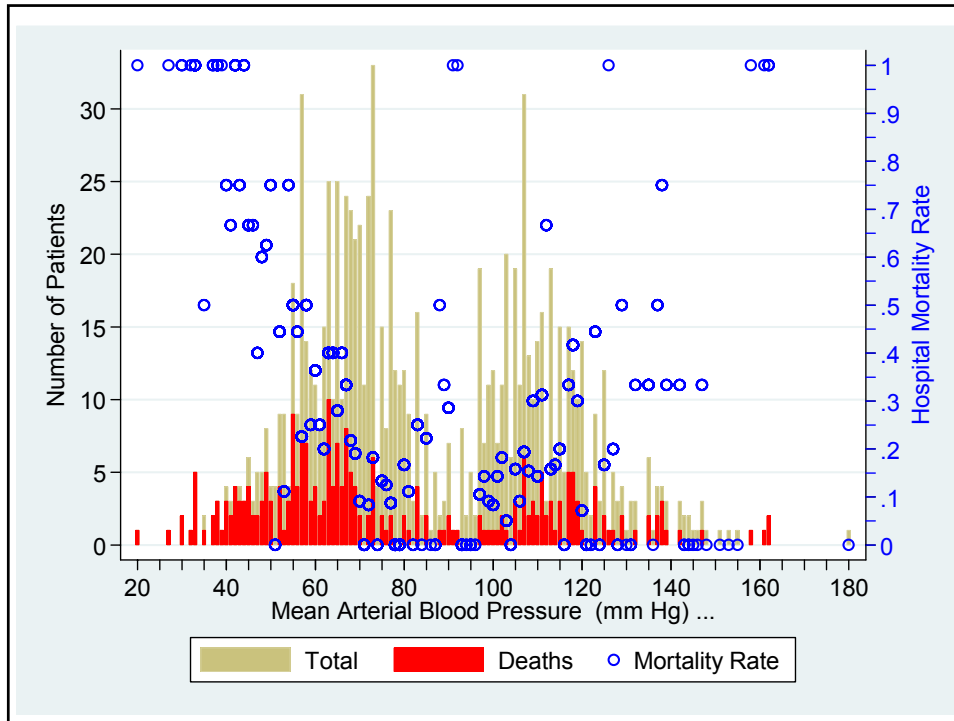
Define **rstudent** to be the studentized residual.

```
. lowess rstudent meanbp
```

```
> , yline(-2 0 2) msymbol(Oh) rlopts(clcolor(green) clwidth(thick)) ///  
> xlabel(25 (25) 175) xtick(30 (5) 170) ///
```

Plot a lowess regression curve of **rstudent** against **meanbp**





Simple logistic regression of hospdead against meanbp

```

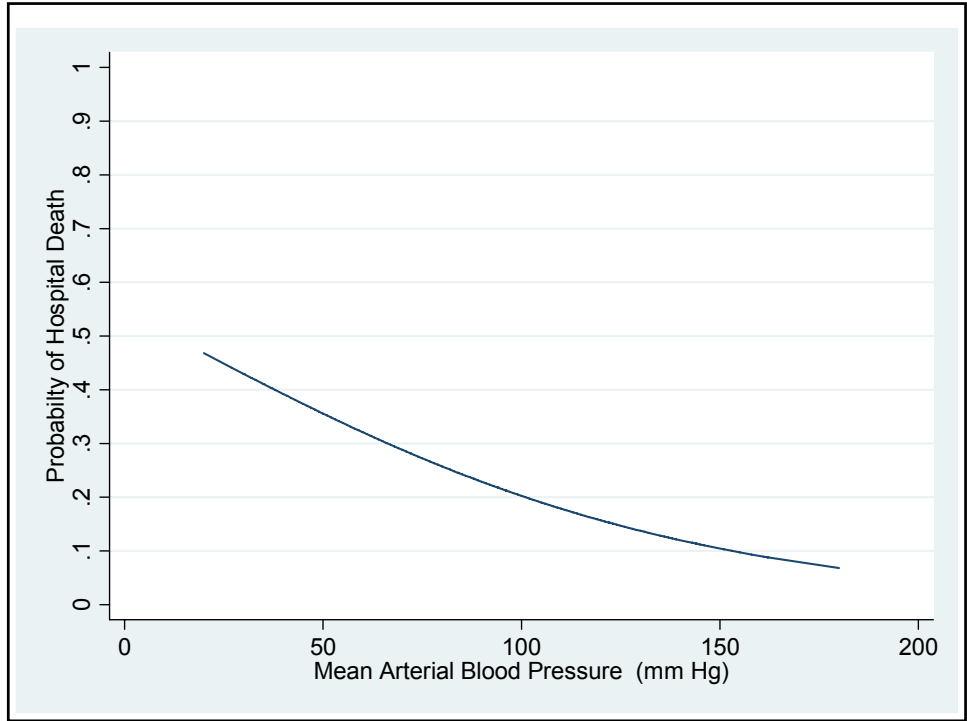
. logistic hospdead meanbp

Logistic regression                               Number of obs   =       996
                                                  LR chi2(1)      =       29.66
                                                  Prob > chi2     =       0.0000
Log likelihood = -545.25721                    Pseudo R2      =       0.0265

-----+-----
 hospdead | Odds Ratio   Std. Err.      z    P>|z|    [95% Conf. Interval]
-----+-----
 meanbp |   .9845924   .0028997    -5.27  0.000    .9789254   .9902922
-----+-----

. predict p,p
. line p meanbp, ylabel(0 (.1) 1) ytitle(Probability of Hospital Death)

```



```

. drop _S* p
. rc_spline meanbp
  number of knots = 5
  value of knot 1 = 47
  value of knot 2 = 66
  value of knot 3 = 78
  value of knot 4 = 106
  value of knot 5 = 129
. logistic hospdead _S*, coef

```

Logistic regression of **hospdead** against spline covariates for **meanbp** with 5 knots.

Spline covariates are significantly different from zero

```

Logistic regression
Log likelihood = -498.65571
Number of obs = 996
LR chi2(4) = 122.86
Prob > chi2 = 0.0000
Pseudo R2 = 0.1097

```

hospdead	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
__Smeanbp1	-.1055538	.0203216	-5.19	0.000	-.1453834	-.0657241
__Smeanbp2	.1598036	.1716553	0.93	0.352	-.1766345	.4962418
__Smeanbp3	.0752005	.6737195	0.11	0.911	-1.245265	1.395666
__Smeanbp4	-.4721096	.6546662	-0.72	0.471	-1.755232	.8110125
__cons	5.531072	1.10928	4.99	0.000	3.356923	7.705221

```
. test _Smeanbp2 _Smeanbp3 _Smeanbp4
```

```
( 1) _Smeanbp2 = 0  
( 2) _Smeanbp3 = 0  
( 3) _Smeanbp4 = 0
```

```
      chi2( 3) = 80.69  
      Prob > chi2 = 0.0000
```

We reject the null hypothesis that the log odds of death is a linear function of mean BP.

### Estimated Statistics at Given Mean BP

**p** = probability of death  
**logodds** = log odds of death  
**stderr** = standard error of logodds  
**(lodds\_lb, lodds\_ub)** = 95% CI for **logodds**  
**(ub\_p, lb\_p)** = 95% CI for **p**

```
. predict p,p
```

```
. predict logodds, xb
```

```
. predict stderr, stdp
```

```
. generate lodds_lb = logodds - 1.96*stderr
```

```
. generate lodds_ub = logodds + 1.96*stderr
```

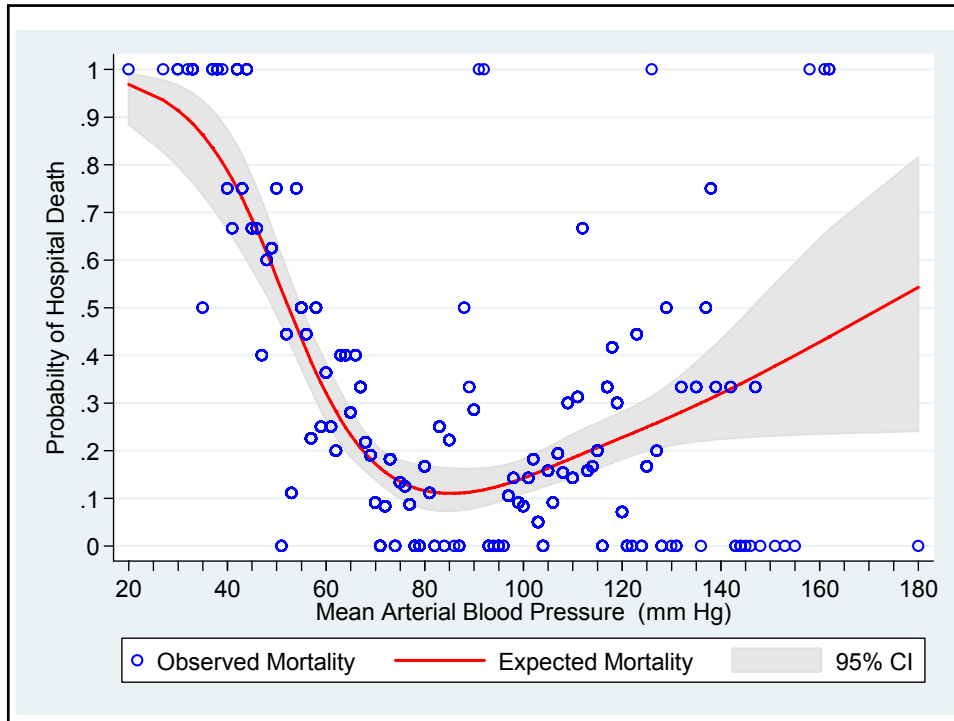
```
. generate ub_p = exp(lodds_ub)/(1+exp(lodds_ub))
```

```
. generate lb_p = exp(lodds_lb)/(1+exp(lodds_lb))
```

```
. by meanbp: egen rate = mean(hospdead)
```

**rate** = proportion of deaths at each blood pressure

```
. twoway rarea lb_p ub_p meanbp, bcolor(gs14) ///  
> || line p meanbp, ccolor(red) clwidth(medthick) ///  
> || scatter rate meanbp, msymbol(Oh) mcolor(blue) ///  
> , ylabel(0 (.1) 1, angle(0)) xlabel(20 (20) 180) ///  
> xtick(25 (5) 175) ytitle(Probability of Hospital Death) ///  
> legend(order(3 "Observed Mortality" ///  
> 2 "Expected Mortality" 1 "95% CI") rows(1))
```



We can use this model to calculate mortal odds ratios for patients with different baseline blood pressures.

```
. list _S*
> if (meanbp==60 | meanbp==90 | meanbp==120) & meanbp == meanbp[_n-1] ///
```

	_Smean-1	_Smean-2	_Smean-3	_Smean-4
178.	60	.32674	0	0
575.	90	11.82436	2.055919	.2569899
893.	120	56.40007	22.30039	10.11355

Logodds of death for patients with **meanbp = 60**

```
. lincom (5.531072 + 60*_Smeanbp1 + .32674*_Smeanbp2) } ///
> + 0*_Smeanbp3 + 0*_Smeanbp4) } ///
> - (5.531072 + 90*_Smeanbp1 + 11.82436*_Smeanbp2) } ///
> + 2.055919*_Smeanbp3 + .2569899*_Smeanbp4) } ///
```

Logodds of death for patients with **meanbp = 90**

```
( 1) - 30 _Smeanbp1 - 11.49762 _Smeanbp2 - 2.055919 _Smeanbp3 -
> .2569899 _Smeanbp4 = 0
```

	hospdead	Odds Ratio	Std. Err.	z	P> z	[95% Conf. Interval]
(1)		3.65455	1.044734	4.53	0.000	2.086887 6.399835

Mortal odds ratio for patients with **meanbp = 60** vs. **meanbp = 90**.

```
. lincom (5.531072 + 120*_Smeanbp1 + 56.40007*_Smeanbp2 + 22.30039*_Smeanbp3 + 10.11355*_Smeanbp4)
>
>
> - (5.531072 + 90*_Smeanbp1 + 11.82436*_Smeanbp2 + 2.055919*_Smeanbp3 + .2569899*_Smeanbp4)
>
```

```
( 1) 30 _Smeanbp1 + 44.57571 _Smeanbp2 + 20.24447 _Smeanbp3 + 9.85656
> _Smeanbp4 = 0
```

	hospdead	Odds Ratio	Std. Err.	z	P> z	[95% Conf. Interval]
(1)		2.283625	.5871892	3.21	0.001	1.379606 3.780023

Mortal odds ratio for patients with **meanbp = 120** vs. **meanbp = 90**.

Stone CJ, Koo CY: Additive splines in statistics *Proceedings of the Statistical Computing Section ASA*. Washington D.C.: American Statistical Association, 1985:45-8.

### Stata Software

Goldstein, R: `srd15`, Restricted cubic spline functions. 1992; *STB*-10: 29-32. `spline.ado`

Sasieni, P: `snp7.1`, Natural cubic splines. 1995; *STB*-24. `spline.ado`

Dupont WD, Plummer WD: `rc_spline` from SSC-IDEAS <http://fmwww.bc.edu/RePEc/bocode/r>

### General Reference

Harrell FE: *Regression Modeling Strategies: With Applications to Linear Models, Logistic Regression, and Survival Analysis*. New York: Springer, 2001.



### Cubic B-Splines

de Boor, C: *A Practical Guide to Splines*.  
New York: Springer-Verlag 1978

- ❖ Similar to restricted cubic splines
- ❖ More complex
- ❖ More numerically stable
- ❖ Does not perform as well outside of the knots

### Software

Newson, R: sg151, B-splines & splines parameterized  
by values at ref. points on x-axis. 2000; *STB-57*: 20-27.  
*bspline.ado*

### nl – Nonlinear least-squares regression

- ❖ Effective when you know the correct form of the non-linear relationship between the dependent and independent variable.
- ❖ Has fewer post-estimation commands and **predict** options than **regress**.

## Conclusions

- ❖ Restricted cubic splines can be used with any regression program that uses a linear predictor – e.g. **regress**, **logistic**, **glm**, **stcox** etc.
- ❖ Can greatly increase the power of these methods to model non-linear relationships.
- ❖ Simple technique that is easy to use and easy to explain.
- ❖ Can be used to test the linearity assumption of generalized linear regression models.
- ❖ Allows users to take advantage of the very mature post-estimation commands associated with generalized linear regression programs to produce sophisticated graphics and residual analyses.