

Usefulness and estimation of proportionality constraints

The `propcnstreg` package

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Outline

usefulness

- proportionality constraint
- a latent variable
- scale for a categorical variable

estimation

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a latent variable
scale for a categorical variable

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example

Hypothesis:

Effect of father's and mother's socioeconomic status on child's education can change over cohorts,

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example

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Effect of father's and mother's socioeconomic status on child's education can change over cohorts, but the relative contribution of the father and the mother remains constant.

$$ed = \beta_0 + \beta_1 coh + (1 + \lambda_1 0)(\gamma_1 pasei + \gamma_2 masei)$$

example

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Effect of father's and mother's socioeconomic status on child's education can change over cohorts, but the relative contribution of the father and the mother remains constant.

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example

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empirical example

- ▶ 7 surveys held between 1994 and 2006 in the USA from the General Social Survey (GSS) containing data on 2,500 white male.

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- ▶ Variable *degree*: educational attainment in pseudo years
- ▶ Variable *byr*: cohort centered in 1940 and measuring time in decades, ranges between 1929 and 1979.
- ▶ Variables *pasei* and *masei*: Father's and mother's occupational status, ranges between 0 and 1.

example output

```
. propcnsreg degree byr, lambda(byr) constrained(masei pasei) lcons
Constraint: [lambda]_cons = 1
```

| degree | Coef. | Std. Err. | z | P> z | [95% Conf. Interval] | |
|--|-----------|-----------|-------|-------|----------------------|----------|
| ----- | | | | | | |
| unconstrained | ----- | | | | | |
| byr | .0392033 | .1418648 | 0.28 | 0.782 | -.2388465 | .3172531 |
| _cons | 10.2406 | .2762536 | 37.07 | 0.000 | 9.699157 | 10.78205 |
| ----- | | | | | | |
| constrained | ----- | | | | | |
| masei | 3.363018 | .3688164 | 9.12 | 0.000 | 2.640152 | 4.085885 |
| pasei | 3.948723 | .3972388 | 9.94 | 0.000 | 3.170149 | 4.727296 |
| ----- | | | | | | |
| lambda | ----- | | | | | |
| byr | -.0323712 | .037854 | -0.86 | 0.392 | -.1065637 | .0418212 |
| _cons | 1 | . | . | . | . | . |
| ----- | | | | | | |
| ln_sigma | ----- | | | | | |
| _cons | .837853 | .014199 | 59.01 | 0.000 | .8100234 | .8656826 |
| ----- | | | | | | |
| LR test vs. unconstrained model: chi2(1) = | | | | 0.04 | Prob > chi2 = 0.849 | |

alternative way of looking

$$ed = \beta_0 + \beta_1 coh + (\lambda_0 + \lambda_1 coh) \underbrace{(\gamma_1 pasei + \gamma_2 masei)}_{\text{latent family sei}}$$

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$$ed = \beta_0 + \beta_1 coh + (\lambda_0 + \lambda_1 coh) \underbrace{(\gamma_1 palsei + \gamma_2 malsei)}_{\text{latent family sei}}$$

- ▶ Need to identify the latent variable by fixing the origin and the scale.

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- ▶ If the minimum value of *palsei* and *malsei* is 0 then the origin is fixed to when both variables are minimum.

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- ▶ If the minimum value of *palsei* and *malsei* is 0 then the origin is fixed to when both variables are minimum.
- ▶ If the maximum value of *palsei* and *malsei* is 1, and

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- ▶ If the maximum value of *palsei* and *malsei* is 1, and their parameters are constrained to sum to 1, then the unit is fixed to the distance between both variables at minimum and both variables at maximum.

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$$\text{latent family sei} = \gamma_1 palsei + \gamma_2 malsei$$

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$$ed = \beta_0 + \beta_1 coh + (\lambda_0 + \lambda_1 coh) \underbrace{(\gamma_1 palsei + \gamma_2 malsei)}_{\text{latent family sei}}$$

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- ▶ If the minimum value of *palsei* and *malsei* is 0 then the origin is fixed to when both variables are minimum.
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$$\text{latent family sei} = \gamma_1 1 + \gamma_2 1 = 1$$

example output

```
. propcnsgreg degree byr, lambda(byr) constrained(masei pasei) unit(masei pasei)
Constraint: [constrained]masei + [constrained]pasei = 1
```

| degree | Coef. | Std. Err. | z | P> z | [95% Conf. Interval] | |
|--|-----------|-----------|-------|-------|----------------------|----------|
| ----- | | | | | | |
| unconstrained | ----- | | | | | |
| byr | .0392033 | .1418647 | 0.28 | 0.782 | -.2388464 | .3172529 |
| _cons | 10.2406 | .2762534 | 37.07 | 0.000 | 9.699158 | 10.78205 |
| ----- | | | | | | |
| constrained | ----- | | | | | |
| masei | .4599477 | .0323745 | 14.21 | 0.000 | .3964949 | .5234005 |
| pasei | .5400523 | .0323745 | 16.68 | 0.000 | .4765995 | .6035051 |
| ----- | | | | | | |
| lambda | ----- | | | | | |
| byr | -.2366899 | .2935214 | -0.81 | 0.420 | -.8119814 | .3386015 |
| _cons | 7.311741 | .601956 | 12.15 | 0.000 | 6.131929 | 8.491553 |
| ----- | | | | | | |
| ln_sigma | ----- | | | | | |
| _cons | .837853 | .014199 | 59.01 | 0.000 | .8100234 | .8656826 |
| ----- | | | | | | |
| LR test vs. unconstrained model: chi2(1) = | | | | 0.04 | Prob > chi2 = 0.849 | |

scale for a categorical variable

Example

- ▶ Differences in the effect of education in 5 dummies on occupational status between white and black US men:
 - ▶ < highschool (reference)
 - ▶ highschool (*hs*)
 - ▶ some college (*sc*)
 - ▶ college (*c*)
 - ▶ graduate (*g*)

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$$isei = \beta_0 + (\lambda_0 + \lambda_1 \mathit{black})(\gamma_1 \mathit{hs} + \gamma_2 \mathit{sc} + \gamma_3 \mathit{c} + \gamma_4 \mathit{g})$$

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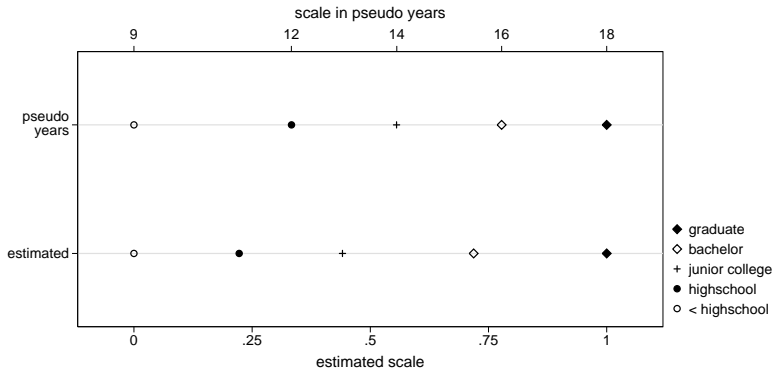
γ_1 , γ_2 , and γ_3 now measure the position of highschool, some college, and college education, relative to less than highschool (0) and graduate (1).

example output

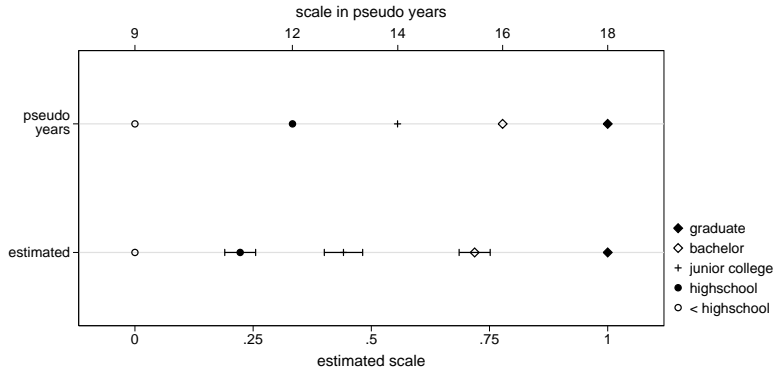
```
. propcnsreg sei black, lambda(black) constrained(hs sc c g) unit(g)
Constraint: [constrained]g = 1
```

| | sei | Coef. | Std. Err. | z | P> z | [95% Conf. Interval] | |
|----------------------------------|-----------|-----------|-----------|---------|---------------|----------------------|-----------|
| ----- | | | | | | | |
| unconstrai~d | | | | | | | |
| | black | -.042371 | .009563 | -4.43 | 0.000 | -.0611141 | -.0236279 |
| | _cons | .3638307 | .0076114 | 47.80 | 0.000 | .3489126 | .3787488 |
| ----- | | | | | | | |
| constrained | | | | | | | |
| | hs | .2226429 | .016662 | 13.36 | 0.000 | .1899861 | .2552997 |
| | sc | .4411229 | .0206904 | 21.32 | 0.000 | .4005705 | .4816753 |
| | c | .7185653 | .01676 | 42.87 | 0.000 | .6857163 | .7514144 |
| | g | 1 | . | . | . | . | . |
| ----- | | | | | | | |
| lambda | | | | | | | |
| | black | .0458751 | .0227816 | 2.01 | 0.044 | .0012239 | .0905263 |
| | _cons | .38541 | .0099432 | 38.76 | 0.000 | .3659217 | .4048983 |
| ----- | | | | | | | |
| ln_sigma | | | | | | | |
| | _cons | -1.859163 | .0090043 | -206.48 | 0.000 | -1.876811 | -1.841515 |
| ----- | | | | | | | |
| LR test vs. unconstrained model: | chi2(3) = | | | 5.42 | Prob > chi2 = | | 0.144 |

Scaling of education



Scaling of education



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estimation

EM-algorithm for starting values

$$y = \beta_0 + \beta_1 x_1 + (\lambda_0 + \lambda_1 x_1)(\gamma_1 z_1 + \gamma_2 z_2)$$

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1. Given current estimates/starting value for γ , create a new variable containing the latent variable.

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1. Given current estimates/starting value for γ , create a new variable containing the latent variable. This simplifies the problem to: $y = \beta_0 + \beta_1 x_1 + \lambda_0 \textit{latent} + \lambda_1 x_1 \textit{latent}$

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2. Estimate β and λ using `regress`

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2. Estimate β and λ using `regress`
3. Given current estimates of λ , create a new variable containing the effect of the latent variable.

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1. Given current estimates/starting value for γ , create a new variable containing the latent variable. This simplifies the problem to: $y = \beta_0 + \beta_1 x_1 + \lambda_0 \textit{latent} + \lambda_1 x_1 \textit{latent}$
2. Estimate β and λ using `regress`
3. Given current estimates of λ , create a new variable containing the effect of the latent variable. This simplifies the problem to: $y = \beta_0 + \beta_1 x_1 + \gamma_1 \textit{effectz}_1 + \gamma_2 \textit{effectz}_2$

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2. Estimate β and λ using `regress`
3. Given current estimates of λ , create a new variable containing the effect of the latent variable. This simplifies the problem to: $y = \beta_0 + \beta_1 x_1 + \gamma_1 \textit{effectz}_1 + \gamma_2 \textit{effectz}_2$
4. Estimate β and γ using `cnsreg` imposing the constraint specified in the `unit` option.

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$$y = \beta_0 + \beta_1 x_1 + (\lambda_0 + \lambda_1 x_1)(\gamma_1 z_1 + \gamma_2 z_2)$$

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2. Estimate β and λ using `regress`
3. Given current estimates of λ , create a new variable containing the effect of the latent variable. This simplifies the problem to: $y = \beta_0 + \beta_1 x_1 + \gamma_1 \textit{effectz}_1 + \gamma_2 \textit{effectz}_2$
4. Estimate β and γ using `cnsreg` imposing the constraint specified in the `unit` option.
5. Repeat steps 1-4 till convergence.

speed and standard errors

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- ▶ To speed up convergence every 5th iteration will consist of two m_1 iterations for the complete model.
- ▶ Once the EM has converged, these estimates are fed into m_1 for the complete model to get the variance covariance matrix.

example iteration log

improving starting values

```
-----  
iteration  unconstrained  constrained  full model  
           part only      part only  
-----  
1           2712.7047       2716.1367  
2           2716.4376       2716.5608  
3           2716.6246       2716.6572  
4           2716.674        2716.6825  
-----  
5           two iterations from full model  
                                     2716.6914  
-----  
6           2716.6914       2716.6914  
-----
```

estimating full model

```
Iteration 0:  log likelihood = 2716.6899  
Iteration 1:  log likelihood = 2716.6914  
Iteration 2:  log likelihood = 2716.6914
```

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- ▶ It can also be interpreted in terms of a latent variable, e.g. father's and mother's status both measure family status.

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- ▶ This can be of interest in it's own right, e.g. the effect on child's education of father's and mother's status change, but the relative contribution of each parent can remain constant.
- ▶ It can also be interpreted in terms of a latent variable, e.g. father's and mother's status both measure family status.
- ▶ Standard `m1` can have a hard time converging, so starting values are created using a EM algorithm.