Title

> **rreg —** Robust regression

Syntax

> rreg *depvar* [*indepvars*] [*if*] [*in*] [, *options*]

options	Description
Model	
tune(*#*)	use *#* as the biweight tuning constant; default is tune(7)
Reporting	
level(*#*)	set confidence level; default is level(95)
genwt(*newvar*)	create *newvar* containing the weights assigned to each observation
display_options	control column formats, row spacing, line width, display of omitted variables and base and empty cells, and factor-variable labeling
Optimization	
optimization_options	control the optimization process; seldom used
graph	graph weights during convergence
coeflegend	display legend instead of statistics

indepvars may contain factor variables; see [U] **11.4.3 Factor variables**.

depvar and *indepvars* may contain time-series operators; see [U] **11.4.4 Time-series varlists**.

by, mfp, mi estimate, rolling, and statsby are allowed; see [U] **11.1.10 Prefix commands**.

coeflegend does not appear in the dialog box.

See [U] **20 Estimation and postestimation commands** for more capabilities of estimation commands.

Menu

Statistics > Linear models and related > Other > Robust regression

Description

rreg performs one version of robust regression of *depvar* on *indepvars*.

Also see *Robust standard errors* in [R] **regress** for standard regression with robust variance estimates and [R] **qreg** for quantile (including median or least-absolute-residual) regression.

Options

Model

tune(#) is the biweight tuning constant. The default is 7, meaning seven times the median absolute deviation (MAD) from the median residual; see *Methods and formulas*. Lower tuning constants downweight outliers rapidly but may lead to unstable estimates (less than 6 is not recommended). Higher tuning constants produce milder downweighting.

Reporting

level(#); see [R] **estimation options**.

genwt(*newvar*) creates the new variable *newvar* containing the weights assigned to each observation.

display_options: noomitted, vsquish, noemptycells, baselevels, allbaselevels, nofvlabel, fvwrap(#), fvwrapon(*style*), cformat(*%fmt*), pformat(*%fmt*), sformat(*%fmt*), and nolstretch; see [R] **estimation options**.

Optimization

optimization_options: iterate(#), tolerance(#), [no]log. iterate() specifies the maximum number of iterations; iterations stop when the maximum change in weights drops below tolerance(); and log/nolog specifies whether to show the iteration log. These options are seldom used.

graph allows you to graphically watch the convergence of the iterative technique. The weights obtained from the most recent round of estimation are graphed against the weights obtained from the previous round.

The following option is available with rreg but is not shown in the dialog box:

coeflegend; see [R] **estimation options**.

Remarks and examples   stata.com

rreg first performs an initial screening based on Cook's distance > 1 to eliminate gross outliers before calculating starting values and then performs Huber iterations followed by biweight iterations, as suggested by Li (1985).

▷ Example 1

We wish to examine the relationship between mileage rating, weight, and location of manufacture for the 74 cars in our automobile data. As a point of comparison, we begin by fitting an ordinary regression:

```
. use http://www.stata-press.com/data/r13/auto
(1978 Automobile Data)

. regress mpg weight foreign
```

Source	SS	df	MS		Number of obs =	74
				F(2, 71) =	69.75
Model	1619.2877	2	809.643849	Prob > F =	0.0000
Residual	824.171761	71	11.608053	R-squared =	0.6627
				Adj R-squared =	0.6532
Total	2443.45946	73	33.4720474	Root MSE =	3.4071

mpg	Coef.	Std. Err.	t	P>|t|	[95% Conf. Interval]	
weight	-.0065879	.0006371	-10.34	0.000	-.0078583	-.0053175
foreign	-1.650029	1.075994	-1.53	0.130	-3.7955	.4954422
_cons	41.6797	2.165547	19.25	0.000	37.36172	45.99768

We now compare this with the results from `rreg`:

```
. rreg mpg weight foreign
   Huber iteration 1:  maximum difference in weights = .80280176
   Huber iteration 2:  maximum difference in weights = .2915438
   Huber iteration 3:  maximum difference in weights = .08911171
   Huber iteration 4:  maximum difference in weights = .02697328
Biweight iteration 5:  maximum difference in weights = .29186818
Biweight iteration 6:  maximum difference in weights = .11988101
Biweight iteration 7:  maximum difference in weights = .03315872
Biweight iteration 8:  maximum difference in weights = .00721325
```

Robust regression				Number of obs =	74	
				F(2, 71) =	168.32	
				Prob > F =	0.0000	

mpg	Coef.	Std. Err.	t	P>|t|	[95% Conf. Interval]	
weight	-.0063976	.0003718	-17.21	0.000	-.007139	-.0056562
foreign	-3.182639	.627964	-5.07	0.000	-4.434763	-1.930514
_cons	40.64022	1.263841	32.16	0.000	38.1202	43.16025

Note the large change in the `foreign` coefficient.

◁

❑ Technical note

It would have been better if we had fit the previous robust regression by typing `rreg mpg weight foreign, genwt(w)`. The new variable, `w`, would then contain the estimated weights. Let's pretend that we did this:

```
. rreg mpg weight foreign, genwt(w)
(output omitted )
. summarize w, detail
```

```
                      Robust Regression Weight

        Percentiles       Smallest
 1%            0                 0
 5%      .0442957                0
10%      .4674935                0          Obs                   74
25%      .8894815         .0442957          Sum of Wgt.           74

50%      .9690193                            Mean             .8509966
                          Largest           Std. Dev.        .2746451
75%      .9949395         .9996715
90%      .9989245         .9996953          Variance         .0754299
95%      .9996715         .9997343          Skewness        -2.287952
99%      .9998585         .9998585          Kurtosis         6.874605
```

We discover that 3 observations in our data were dropped altogether (they have weight 0). We could
further explore our data:

```
. sort w

. list make mpg weight w if w <.467, sep(0)
```

```
        make           mpg    weight          w

 1.     VW Diesel       41     2,040          0
 2.     Subaru          35     2,050          0
 3.     Datsun 210      35     2,020          0
 4.     Plym. Arrow     28     3,260    .04429567
 5.     Cad. Seville    21     4,290    .08241943
 6.     Toyota Corolla  31     2,200    .10443129
 7.     Olds 98         21     4,060    .28141296
```

Being familiar with the automobile data, we immediately spotted two things: the VW is the only
diesel car in our data, and the weight recorded for the Plymouth Arrow is incorrect.

❑

▷ Example 2

If we specify no explanatory variables, rreg produces a robust estimate of the mean:

```
. rreg mpg
    Huber iteration 1:  maximum difference in weights = .64471879
    Huber iteration 2:  maximum difference in weights = .05098336
    Huber iteration 3:  maximum difference in weights = .0099887
Biweight iteration 4:  maximum difference in weights = .25197391
Biweight iteration 5:  maximum difference in weights = .00358606

Robust regression                                   Number of obs =      74
                                                    F(  0,    73) =    0.00
                                                    Prob > F      =       .
```

mpg	Coef.	Std. Err.	t	P>|t|	[95% Conf. Interval]
_cons	20.68825	.641813	32.23	0.000	19.40912 21.96738

The estimate is given by the coefficient on _cons. The mean is 20.69 with an estimated standard error of 0.6418. The 95% confidence interval is $[19.4, 22.0]$. By comparison, ci (see [R] **ci**) gives us the standard calculation:

```
. ci mpg
     Variable │       Obs        Mean    Std. Err.      [95% Conf. Interval]
─────────────┼─────────────────────────────────────────────────────────────
          mpg │        74     21.2973    .6725511       19.9569      22.63769
```

◁

Stored results

rreg stores the following in e():

Scalars
  e(N)   number of observations
  e(mss)   model sum of squares
  e(df_m)   model degrees of freedom
  e(rss)   residual sum of squares
  e(df_r)   residual degrees of freedom
  e(r2) R-squared
  e(r2_a)   adjusted R-squared
  e(F) F statistic
  e(rmse)   root mean squared error
  e(rank)   rank of e(V)

Macros
  e(cmd)   rreg
  e(cmdline)   command as typed
  e(depvar)   name of dependent variable
  e(genwt)   variable containing the weights
  e(title)   title in estimation output
  e(model)   ols
  e(vce)   ols
  e(properties)   b V
  e(predict)   program used to implement predict
  e(marginsok)   predictions allowed by margins
  e(asbalanced)   factor variables fvset as asbalanced
  e(asobserved)   factor variables fvset as asobserved

Matrices
  e(b)   coefficient vector
  e(V)   variance–covariance matrix of the estimators

Functions
  e(sample)   marks estimation sample

Methods and formulas

See Berk (1990), Goodall (1983), and Rousseeuw and Leroy (1987) for a general description of the issues and methods. Hamilton (1991a, 1992) provides a more detailed description of rreg and some Monte Carlo evaluations.

rreg begins by fitting the regression (see [R] **regress**), calculating Cook's D (see [R] **predict** and [R] **regress postestimation**), and excluding any observation for which $D > 1$.

Thereafter rreg works iteratively: it performs a regression, calculates case weights from absolute residuals, and regresses again using those weights. Iterations stop when the maximum change in weights drops below tolerance(). Weights derive from one of two weight functions, Huber weights

and biweights. Huber weights (Huber 1964) are used until convergence, and then, from that result, biweights are used until convergence. The biweight was proposed by Beaton and Tukey (1974, 151–152) after the Princeton robustness study (Andrews et al. 1972) had compared various estimators. Both weighting functions are used because Huber weights have problems dealing with severe outliers, whereas biweights sometimes fail to converge or have multiple solutions. The initial Huber weighting should improve the behavior of the biweight estimator.

In Huber weighting, cases with small residuals receive weights of 1; cases with larger residuals receive gradually smaller weights. Let $e_i = y_i - \mathbf{X}_i\mathbf{b}$ represent the ith-case residual. The ith scaled residual $u_i = e_i/s$ is calculated, where $s = M/0.6745$ is the residual scale estimate and $M = \text{med}(|e_i - \text{med}(e_i)|)$ is the median absolute deviation from the median residual. Huber estimation obtains case weights:

$$w_i = \begin{cases} 1 & \text{if } |u_i| \leq c_h \\ c_h/|u_i| & \text{otherwise} \end{cases}$$

rreg defines $c_h = 1.345$, so downweighting begins with cases whose absolute residual exceeds $(1.345/0.6745)M \approx 2M$.

With biweights, all cases with nonzero residuals receive some downweighting, according to the smoothly decreasing biweight function

$$w_i = \begin{cases} \{1 - (u_i/c_b)^2\}^2 & \text{if } |u_i| \leq c_b \\ 0 & \text{otherwise} \end{cases}$$

where $c_b = 4.685 \times \texttt{tune()}/7$. Thus when $\texttt{tune()} = 7$, cases with absolute residuals of $(4.685/0.6745)M \approx 7M$ or more are assigned 0 weight and thus are effectively dropped. Goodall (1983, 377) suggests using a value between 6 and 9, inclusive, for tune() in the biweight case and states that performance is good between 6 and 12, inclusive.

The tuning constants $c_h = 1.345$ and $c_b = 4.685$ (assuming tune() is set at the default 7) give rreg about 95% of the efficiency of OLS when applied to data with normally distributed errors (Hamilton 1991b). Lower tuning constants downweight outliers more drastically (but give up Gaussian efficiency); higher tuning constants make the estimator more like OLS.

Standard errors are calculated using the pseudovalues approach described in Street, Carroll, and Ruppert (1988).

Acknowledgment

References

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Also see

[R] **rreg postestimation** — Postestimation tools for rreg

[R] **qreg** — Quantile regression

[R] **regress** — Linear regression

[MI] **estimation** — Estimation commands for use with mi estimate

[U] **20 Estimation and postestimation commands**