

## Postestimation commands

The following postestimation commands are of special interest after `arfima`:

| Command                   | Description                                   |
|---------------------------|---|
| <code>estat acplot</code> | estimate autocorrelations and autocovariances |
| <code>irf</code>          | create and analyze IRFs                       |
| <code>psdensity</code>    | estimate the spectral density                 |

The following standard postestimation commands are also available:

| Command                      | Description   |
|------------------------------|---|
| <code>contrast</code>        | contrasts and ANOVA-style joint tests of parameters   |
| * <code>estat ic</code>      | Akaike's, consistent Akaike's, corrected Akaike's, and Schwarz's Bayesian information criteria (AIC, CAIC, AICc, and BIC, respectively) |
| <code>estat summarize</code> | summary statistics for the estimation sample  |
| <code>estat vce</code>       | variance–covariance matrix of the estimators (VCE)  |
| <code>estimates</code>       | cataloging estimation results   |
| <code>etable</code>          | table of estimation results   |
| <code>forecast</code>        | dynamic forecasts and simulations   |
| <code>lincom</code>          | point estimates, standard errors, testing, and inference for linear combinations of parameters  |
| <code>lrtest</code>          | likelihood-ratio test   |
| * <code>margins</code>       | marginal means, predictive margins, marginal effects, and average marginal effects  |
| * <code>marginsplot</code>   | graph the results from margins (profile plots, interaction plots, etc.)   |
| * <code>nlcom</code>         | point estimates, standard errors, testing, and inference for nonlinear combinations of parameters                                       |
| <code>predict</code>         | linear predictions, innovations, standardized innovations, etc.   |
| * <code>predictnl</code>     | point estimates, standard errors, testing, and inference for generalized predictions  |
| <code>pwcompare</code>       | pairwise comparisons of parameters  |
| <code>test</code>            | Wald tests of simple and composite linear hypotheses  |
| <code>testnl</code>          | Wald tests of nonlinear hypotheses  |

\*`estat ic`, `margins`, `marginsplot`, `nlcom`, and `predictnl` are not appropriate after `arfima`, `mpl`.

# predict

## Description for predict

predict creates a new variable containing predictions such as expected values, fractionally differenced series, and innovations. All predictions are available as static one-step-ahead predictions, and the dependent variable is also available as a dynamic multistep prediction.

## Menu for predict

Statistics > Postestimation

## Syntax for predict

```
predict [type] newvar [if] [in] [ , statistic options ]
```

| statistic   | Description                     |
|-------------|---------------------------------|
| Main        |                                 |
| xb          | predicted values; the default   |
| residuals   | predicted innovations           |
| rstandard   | standardized innovations        |
| fdifference | fractionally differenced series |

These statistics are available both in and out of sample; type predict ... if e(sample) ... if wanted only for the estimation sample.

| options             | Description  |
|---------------------|--|
| Options             |  |
| rmse([type] newvar) | put the estimated root mean squared error of the predicted statistic in a new variable; only permitted with options xb and residuals |
| dynamic(datetime)   | forecast the time series starting at datetime; only permitted with option xb   |

datetime is a # or a time literal, such as td(1jan1995) or tq(1995q1); see [D] Datetime.

## Options for predict

Main

- xb, the default, calculates the predictions for the level of depvar.
- residuals calculates the predicted innovations.
- rstandard calculates the standardized innovations.
- fdifference calculates the fractionally differenced predictions of depvar.

Options

`rmse([type] newvar)` puts the root mean squared errors of the predicted statistics into the specified new variables. The root mean squared errors measure the variances due to the disturbances but do not account for estimation error. `rmse()` is only permitted with the `xb` and `residuals` options.

`dynamic(datetime)` specifies when `predict` starts producing dynamic forecasts. The specified *datetime* must be in the scale of the time variable specified in `tsset`, and the *datetime* must be inside a sample for which observations on the dependent variables are available. For example, `dynamic(tq(2008q4))` causes dynamic predictions to begin in the fourth quarter of 2008, assuming that your time variable is quarterly; see [D] **Datetime**. If the model contains exogenous variables, they must be present for the whole predicted sample. `dynamic()` may only be specified with `xb`.

margins

Description for margins

`margins` estimates margins of response for expected values.

Menu for margins

Statistics > Postestimation

Syntax for margins

```
margins [marginlist] [ , options ]
margins [marginlist] , predict(statistic ...) [options]
```

| statistic                | Description                   |
|--------------------------|-------------------------------|
| <code>xb</code>          | predicted values; the default |
| <code>residuals</code>   | not allowed with margins      |
| <code>rstandard</code>   | not allowed with margins      |
| <code>fdifference</code> | not allowed with margins      |

Statistics not allowed with margins are functions of stochastic quantities other than  $e(b)$ .

For the full syntax, see [R] **margins**.

Remarks and examples

Remarks are presented under the following headings:

Forecasting after ARFIMA  
IRF results for ARFIMA

## Forecasting after ARFIMA

We assume that you have already read [TS] [arfima](#). In this section, we illustrate some of the features of `predict` after fitting an ARFIMA model using `arfima`.

### ► Example 1

We have monthly data on the one-year Treasury bill secondary market rate imported from the Federal Reserve Bank (FRED) database using `import fred`; see [D] [import fred](#). Below we fit an ARFIMA model with two autoregressive terms and one moving-average term to the data.

```
. use https://www.stata-press.com/data/r19/tb1yr
(FRED, 1-year Treasury bill; secondary market rate, monthly 1959–2001)

. arfima tb1yr, ar(1/2) ma(1)
Iteration 0: Log likelihood = -235.31856
Iteration 1: Log likelihood = -235.26104 (backed up)
Iteration 2: Log likelihood = -235.25974 (backed up)
Iteration 3: Log likelihood = -235.2544 (backed up)
Iteration 4: Log likelihood = -235.13355
Iteration 5: Log likelihood = -235.13064
Iteration 6: Log likelihood = -235.12108
Iteration 7: Log likelihood = -235.11917
Iteration 8: Log likelihood = -235.11869
Iteration 9: Log likelihood = -235.11868
Refining estimates:
Iteration 0: Log likelihood = -235.11868
Iteration 1: Log likelihood = -235.11868

ARFIMA regression
Sample: 1959m7 thru 2001m8
Log likelihood = -235.11868
```

```
Number of obs =      506
Wald chi2(4)   = 1864.15
Prob > chi2    = 0.0000
```

| tb1yr   | Coefficient | Std. err. | z     | P> z  | [95% conf. interval] |          |
|---------|-------------|-----------|-------|-------|----------------------|----------|
| tb1yr   |             |           |       |       |                      |          |
| _cons   | 5.496708    | 2.920375  | 1.88  | 0.060 | -.2271229            | 11.22054 |
| ARFIMA  |             |           |       |       |                      |          |
| ar      |             |           |       |       |                      |          |
| L1.     | .2326102    | .1136814  | 2.05  | 0.041 | .0097987             | .4554217 |
| L2.     | .388521     | .0835709  | 4.65  | 0.000 | .2247251             | .552317  |
| ma      |             |           |       |       |                      |          |
| L1.     | .7755849    | .0669629  | 11.58 | 0.000 | .6443401             | .9068297 |
| d       | .4606494    | .064656   | 7.12  | 0.000 | .3339259             | .5873728 |
| /sigma2 | .1466495    | .009232   | 15.88 | 0.000 | .1285551             | .1647439 |

Note: The test of the variance against zero is one sided, and the two-sided confidence interval is truncated at zero.

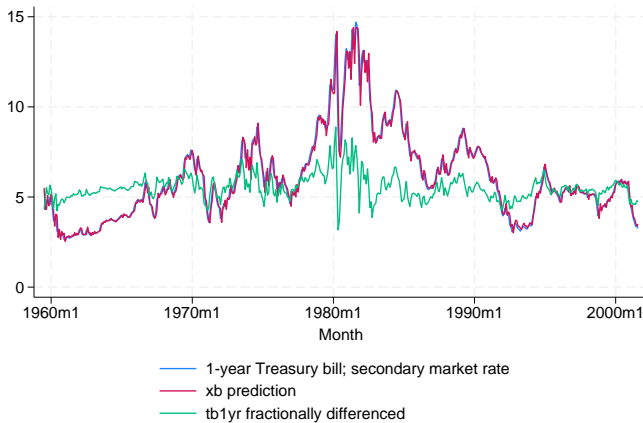
All the parameters are statistically significant at the 5% level, and they indicate a high degree of dependence in the series. In fact, the confidence interval for the fractional-difference parameter  $d$  indicates that the series may be nonstationary. We will proceed as if the series is stationary and suppose that it is fractionally integrated of order 0.46.

We begin our postestimation analysis by predicting the series in sample:

```
. predict ptb
(option xb assumed)
```

We continue by using the estimated fractional-difference parameter to fractionally difference the original series and by plotting the original series, the predicted series, and the fractionally differenced series. See [TS] [arfima](#) for a definition of the fractional-difference operator.

```
. predict fdtb, fdifference
. tsline tb1yr ptb fdtb, legend(cols(1))
```



The above graph shows that the in-sample predictions appear to track the original series well and that the fractionally differenced series looks much more like a stationary series than does the original.



## ► Example 2

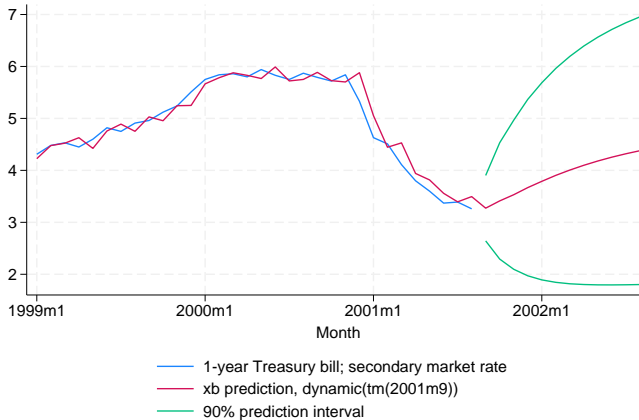
In this example, we use the above estimates to produce a dynamic forecast and a confidence interval for the forecast for the one-year treasury bill rate and plot them.

We begin by extending the dataset and using `predict` to put the dynamic forecast in the new `ftb` variable and the root mean squared error of the forecast in the new `rtb` variable. (As discussed in [Methods and formulas](#), the root mean squared error of the forecast accounts for the idiosyncratic error but not for the estimation error.)

```
. tsappend, add(12)
. predict ftb, xb dynamic(tm(2001m9)) rmse(rtb)
```

Now we compute a 90% confidence interval around the dynamic forecast and plot the original series, the in-sample forecast, the dynamic forecast, and the confidence interval of the dynamic forecast.

```
. scalar z = invnormal(0.95)
. generate lb = ftb - z*rtb if month>=tm(2001m9)
(506 missing values generated)
. generate ub = ftb + z*rtb if month>=tm(2001m9)
(506 missing values generated)
. tsline tbyr ftb if month>tm(1998m12) ||
> tsline lb ub if month>=tm(2001m9),
> legend(cols(1) label(3 "90% prediction interval"))
```



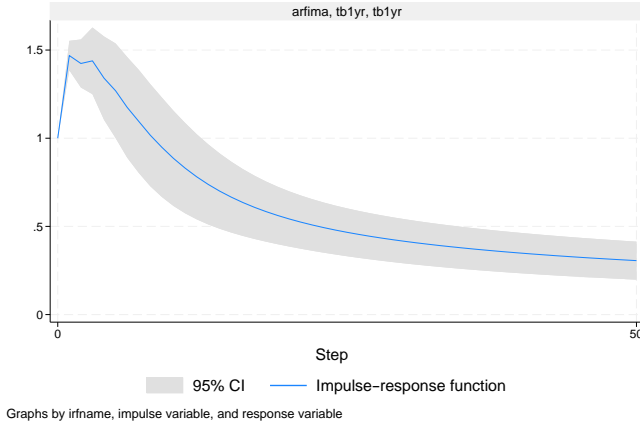
## IRF results for ARFIMA

We assume that you have already read [\[TS\] irf](#) and [\[TS\] irf create](#). In this section, we illustrate how to calculate the impulse–response function (IRF) of an ARFIMA model.

### ► Example 3

Here we use the estimates obtained in [example 1](#) to calculate the IRF of the ARFIMA model; see [\[TS\] irf](#) and [\[TS\] irf create](#) for more details about IRFs.

```
. irf create arfima, step(50) set(myirf)
(file myirf.irf created)
(file myirf.irf now active)
(file myirf.irf updated)
. irf graph irf
```



The figure shows that a shock to `tb1yr` causes an initial spike in `tb1yr`, after which the impact of the shock starts decaying slowly. This behavior is characteristic of long-memory processes.

◀

## Methods and formulas

Denote  $\gamma_h$ ,  $h = 1, \dots, t$ , to be the autocovariance function of the ARFIMA( $p, d, q$ ) process for two observations,  $y_t$  and  $y_{t-h}$ ,  $h$  time periods apart. The covariance matrix  $\mathbf{V}$  of the process of length  $T$  has a Toeplitz structure of

$$\mathbf{V} = \begin{pmatrix} \gamma_0 & \gamma_1 & \gamma_2 & \cdots & \gamma_{T-1} \\ \gamma_1 & \gamma_0 & \gamma_1 & \cdots & \gamma_{T-2} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ \gamma_{T-1} & \gamma_{T-2} & \gamma_{T-3} & \cdots & \gamma_0 \end{pmatrix}$$

where the process variance is  $\gamma_0 = \text{Var}(y_t)$ . We factor  $\mathbf{V} = \mathbf{L}\mathbf{D}\mathbf{L}'$ , where  $\mathbf{L}$  is lower triangular and  $\mathbf{D} = \text{Diag}(\nu_t)$ . The structure of  $\mathbf{L}^{-1}$  is of importance.

$$\mathbf{L}^{-1} = \begin{pmatrix} 1 & 0 & 0 & \cdots & 0 & 0 \\ -\tau_{1,1} & 1 & 0 & \cdots & 0 & 0 \\ -\tau_{2,2} & -\tau_{2,1} & 1 & \cdots & 0 & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots & \vdots \\ -\tau_{T-1,T-1} & -\tau_{T-1,T-2} & -\tau_{T-1,T-2} & \cdots & -\tau_{T-1,1} & 1 \end{pmatrix}$$

Let  $z_t = y_t - \mathbf{x}_t\beta$ . The best linear predictor of  $z_{t+1}$  based on  $z_1, z_2, \dots, z_t$  is  $\hat{z}_{t+1} = \sum_{k=1}^t \tau_{t,k} z_{t-k+1}$ . Define  $-\tau_t = (-\tau_{t,t}, -\tau_{t,t-1}, \dots, -\tau_{t-1,1})$  to be the  $t$ th row of  $\mathbf{L}^{-1}$  up to, but not including, the diagonal. Then  $\tau_t = \mathbf{V}_t^{-1}\gamma_t$ , where  $\mathbf{V}_t$  is the  $t \times t$  upper left submatrix of  $\mathbf{V}$  and  $\gamma_t = (\gamma_1, \gamma_2, \dots, \gamma_t)'$ . Hence, the best linear predictor of the innovations is computed as  $\hat{\epsilon} = \mathbf{L}^{-1}\mathbf{z}$ , and the one-step predictions are  $\hat{\mathbf{y}} = \hat{\epsilon} + \mathbf{X}\hat{\beta}$ . In practice, the computation is

$$\hat{\mathbf{y}} = \hat{\mathbf{L}}^{-1} (\mathbf{y} - \mathbf{X}\hat{\boldsymbol{\beta}}) + \mathbf{X}\hat{\boldsymbol{\beta}}$$

where  $\hat{\mathbf{L}}$  and  $\hat{\mathbf{V}}$  are computed from the maximum likelihood estimates. We use the Durbin–Levinson algorithm (Palma 2007; Golub and Van Loan 2013) to factor  $\hat{\mathbf{V}}$ , invert  $\hat{\mathbf{L}}$ , and scale  $\mathbf{y} - \mathbf{X}\hat{\boldsymbol{\beta}}$  using only the vector of estimated autocovariances  $\hat{\boldsymbol{\gamma}}$ .

The prediction error variances of the one-step predictions are computed recursively in the Durbin–Levinson algorithm. They are the  $\nu_t$  elements in the diagonal matrix  $\mathbf{D}$  computed from the Cholesky factorization of  $\mathbf{V}$ . The recursive formula is  $\nu_0 = \gamma_0$ , and  $\nu_t = \nu_{t-1}(1 - \tau_{t,t}^2)$ .

Forecasting is carried out as described by Beran (1994, sec. 8.7),  $\hat{\mathbf{z}}_{T+k} = \tilde{\boldsymbol{\gamma}}_k' \hat{\mathbf{V}}^{-1} \hat{\mathbf{z}}$ , where  $\tilde{\boldsymbol{\gamma}}_k' = (\hat{\gamma}_{T+k-1}, \hat{\gamma}_{T+k-2}, \dots, \hat{\gamma}_k)$ . The forecast mean squared error is computed as  $\text{MSE}(\hat{\mathbf{z}}_{T+k}) = \hat{\gamma}_0 - \tilde{\boldsymbol{\gamma}}_k' \hat{\mathbf{V}}^{-1} \tilde{\boldsymbol{\gamma}}_k$ . Computation of  $\hat{\mathbf{V}}^{-1} \tilde{\boldsymbol{\gamma}}_k$  is carried out efficiently using algorithm 4.7.2 of Golub and Van Loan (2013).

## References

- Beran, J. 1994. *Statistics for Long-Memory Processes*. Boca Raton, FL: Chapman and Hall/CRC.
- Golub, G. H., and C. F. Van Loan. 2013. *Matrix Computations*. 4th ed. Baltimore: Johns Hopkins University Press. <https://doi.org/10.56021/9781421407944>.
- Palma, W. 2007. *Long-Memory Time Series: Theory and Methods*. Hoboken, NJ: Wiley.

## Also see

- [TS] **arfima** — Autoregressive fractionally integrated moving-average models
- [TS] **estat acplot** — Plot parametric autocorrelation and autocovariance functions
- [TS] **irf** — Create and analyze IRFs, dynamic-multiplier functions, and FEVDs
- [TS] **psdensity** — Parametric spectral density estimation after arima, arfima, and ucm
- [U] **20 Estimation and postestimation commands**

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