

[Description](#)[Remarks and examples](#)[Also see](#)

## Description

This entry comprises an example from start to finish.

You may also be interested in introductions to other aspects of Sp. Below, we provide links to those other introductions.

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<a href="#">Intro 1</a>	A brief introduction to SAR models
<a href="#">Intro 2</a>	The <b>W</b> matrix
<a href="#">Intro 3</a>	Preparing data for analysis
<a href="#">Intro 4</a>	Preparing data: Data with shapefiles
<a href="#">Intro 5</a>	Preparing data: Data containing locations (no shapefiles)
<a href="#">Intro 6</a>	Preparing data: Data without shapefiles or locations
<a href="#">Intro 8</a>	The Sp estimation commands

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## Remarks and examples

Remarks are presented under the following headings:

[Research plan](#)  
[Finding and preparing data](#)  
    [Finding a shapefile for Texas counties](#)  
    [Creating the Stata-format shapefile](#)  
    [Merging our data with the Stata-format shapefile](#)  
[Analyzing texas\\_ue.dta](#)  
    [Testing whether ordinary regression is adequate](#)  
    [spregress can reproduce regress results](#)  
    [Fitting models with a spatial lag of the dependent variable](#)  
    [Interpreting models with a spatial lag of the dependent variable](#)  
    [Fitting models with a spatial lag of independent variables](#)  
    [Interpreting models with a spatial lag of the independent variables](#)  
    [Fitting models with spatially autoregressive errors](#)  
    [Models can have all three kinds of spatial lag terms](#)

## Research plan

We are going to analyze unemployment in counties of Texas. We are going to use `texas_ue.dta`. The data contain unemployment rates and college graduation rates for Texas counties, but they do not include the locations of the counties or a map. The data can be used to fit models with `regress`, but they do not contain the information necessary to fit models with `spregress` that could account for spillover effects.

We will

1. find and download a US counties shapefile,
2. translate the downloaded file to Stata format,
3. merge the translated file with our existing data, and
4. analyze the merged data.

Please keep in mind that this is just an example in a computer software manual. We will model the unemployment rate as a function of college graduation rate only, though we ought to include other explanatory variables. We analyze data for Texas only, though we should use the entire United States. We will draw conclusions that are unjustified, and we will not qualify them appropriately. We will, however, show you how to use `spregress` and interpret its output.

## Finding and preparing data

We first find and download an appropriate shapefile from the web. Then, we will prepare it as described in [SP] Intro 4.

### Finding a shapefile for Texas counties

We looked for a county shapefile for Texas but could not find one. We did find shapefiles for the entire United States, however. We used our browser to search for “shapefile US counties census”. From the results, we selected *TIGER/Line Shapefile, 2016, nation, US, Current County and Equivalent National Shapefile*. On the resulting page, we clicked to download the **Shapefile Zip File** from the **Downloads & Resources** section. File `tl_2016_us_county.zip` was downloaded to the Downloads directory on our computer.

### Creating the Stata-format shapefile

We found a standard-format shapefile, `tl_2016_us_county.zip`. We now follow the instructions in [SP] Intro 4 to create a Stata-format shapefile. Here is the result:

```
. // -----
. // [SP] intro 4, step 2: Translate the shapefile
.
. copy ~/Downloads/tl_2016_us_county.zip .
. unzipfile tl_2016_us_county.zip
  inflating: tl_2016_us_county.cpg
  inflating: tl_2016_us_county.dbf
  inflating: tl_2016_us_county.prj
  inflating: tl_2016_us_county.shp
  inflating: tl_2016_us_county.shp.ea.iso.xml
  inflating: tl_2016_us_county.shp.iso.xml
  inflating: tl_2016_us_county.shp.xml
  inflating: tl_2016_us_county.shx
successfully unzipped tl_2016_us_county.zip to current directory
total processed: 8
      skipped: 0
    extracted: 8
```



2.

_ID 2	_CX -123.43347	_CY 46.291134	STATEFP 53	COUNTYFP 069	COUNTYNS 01513275	GEOID 53069	
NAME Wahkiakum	NAMELSAD Wahkiakum County		LSAD 06	CLASSFP H1	MTFCC G4020	CSAFP	CBSAFP
METDIVFP	FUNCSTAT A		ALAND 680956787		AWATER 61588406		INTPTLAT +46.2946377
INTPTLON -123.4244583							

```
.
. // -----
. // [SP] intro 4, step 4: Create standard ID variable
.
. generate long fips = real(STATEFP + COUNTYFP)
. bysort fips: assert _N==1
. assert fips != .
.
. // -----
. // [SP] intro 4, step 5: Tell Sp to use standard ID variable
.
. spset fips, modify replace
  (_shp.dta file saved)
  (data in memory saved)
      Sp dataset: tl_2016_us_county.dta
Linked shapefile: tl_2016_us_county_shp.dta
      Data: Cross sectional
      Spatial-unit ID: _ID (equal to fips)
      Coordinates: _CX, _CY (planar)
.
. // -----
. // [SP] intro 4, step 6: Set coordinate units
.
. spset, modify coordsys(latlong, miles)
      Sp dataset: tl_2016_us_county.dta
Linked shapefile: tl_2016_us_county_shp.dta
      Data: Cross sectional
      Spatial-unit ID: _ID (equal to fips)
      Coordinates: _CY, _CX (latitude-and-longitude, miles)
. save, replace
file tl_2016_us_county.dta saved
. // -----
```

## Merging our data with the Stata-format shapefile

Recall that we are going to use `texas_ue.dta` containing unemployment rates and college graduation rates for Texas counties. We follow the instructions in [\[SP\] Intro 4, Step 7a](#) to merge our existing data with the Stata-format shapefile.

```
. copy https://www.stata-press.com/data/r19/texas_ue.dta .
. use texas_ue, clear
. describe
```

Contains data from texas\_ue.dta

Observations: 254

Variables: 4

10 Feb 2023 12:36

(\_dta has notes)

Variable name	Storage type	Display format	Value label	Variable label
fips	float	%9.0g		FIPS
college	float	%9.0g		* Percent college degree
income	long	%12.0g		Median household income
unemployment	float	%9.0g		Unemployment rate
* indicated variables have notes				

Sorted by: fips

```
. merge 1:1 fips using tl_2016_us_county
```

(variable **fips** was **float**, now **double** to accommodate using data's values)

Result	Number of obs	
Not matched	2,979	
from master	0	(_merge==1)
from using	2,979	(_merge==2)
Matched	254	(_merge==3)

```
. keep if _merge==3
```

(2,979 observations deleted)

```
. drop _merge
```

At this point, we type describe again and discover that `texas_ue.dta` has lots of unnecessary, leftover variables from `tl_2016_us_county.dta`, so we drop them. There is another variable that we rather like—the names of the counties—and we rename it.

```
. rename NAME countyname
. drop STATEFP COUNTYFP COUNTYNS GEOID
. drop NAMELSAD LSAD CLASSFP MTFCC CSAFP
. drop CBSAFP METDIVFP FUNCSTAT
. drop ALAND AWATER INTPTLAT INTPTLON
. save, replace
file texas_ue.dta saved
```

## Analyzing texas\_ue.dta

File `texas_ue.dta` is our updated analysis dataset that can be used with `Sp` commands.

```
. describe
Contains data from texas_ue.dta
Observations:      254
Variables:         8                               27 Mar 2025 21:06
                                                    (_dta has notes)
```

Variable name	Storage type	Display format	Value label	Variable label
fips	double	%9.0g		FIPS
college	float	%9.0g		* Percent college degree
income	long	%12.0g		Median household income
unemployment	float	%9.0g		Unemployment rate
_ID	long	%12.0g		Spatial-unit ID
_CX	double	%10.0g		x-coordinate of area centroid
_CY	double	%10.0g		y-coordinate of area centroid
countyname	str21	%21s		NAME
				* indicated variables have notes

Sorted by:

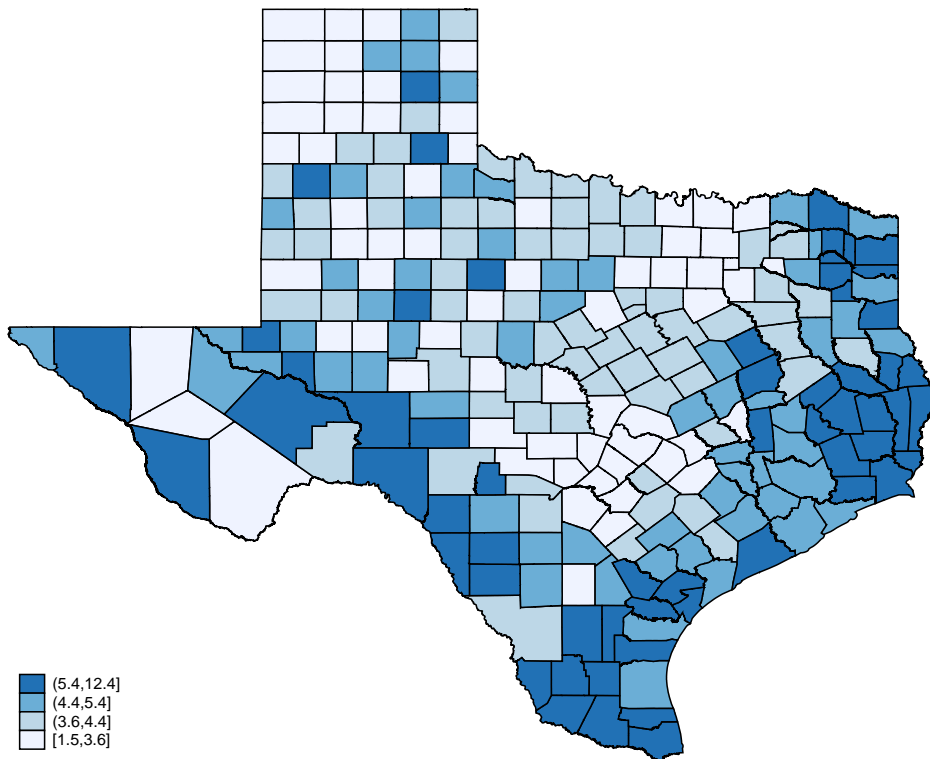
Our example uses the unemployment rate. It varies between 1.5% and 12.4% across the counties of Texas:

```
. summarize unemployment
```

Variable	Obs	Mean	Std. dev.	Min	Max
unemployment	254	4.731102	1.716514	1.5	12.4

Because `texas_ue.dta` has been `spset` and has a shapefile, we can draw choropleth maps, such as this one of the unemployment rate:

```
. grmap unemployment
```



Unemployment appears to be clustered, which suggests that there are spillover effects between counties.

### Testing whether ordinary regression is adequate

These data are suitable for both spatial and nonspatial analysis. (Spatial data always are.) We will fit a linear regression of the unemployment rate on the college graduation rate, mostly for illustrative purposes. After fitting the linear regression, we will use an `Sp` command to determine whether the residuals of the model are spatially correlated, and we find that they are.

Here is the regression:

```
. regress unemployment college
```

Source	SS	df	MS	Number of obs	=	254
Model	139.314746	1	139.314746	F(1, 252)	=	57.92
Residual	606.129539	252	2.40527595	Prob > F	=	0.0000
				R-squared	=	0.1869
				Adj R-squared	=	0.1837
Total	745.444285	253	2.9464201	Root MSE	=	1.5509

unemployment	Coefficient	Std. err.	t	P> t	[95% conf. interval]	
college	-.1008791	.0132552	-7.61	0.000	-.1269842	-.0747741
_cons	6.542796	.2571722	25.44	0.000	6.036316	7.049277

The results of this oversimplified model indicate that the college graduation rate reduces unemployment markedly.

Are we done? If the residuals show no signs of being spatially clustered, then we are. We can perform a statistical test.

Sp provides the Moran test for determining whether the residuals of a model fit by `regress` are correlated with nearby residuals. To use it, we must define “nearby”. We do that by defining a spatial weighting matrix, which is created by the `spmatrix` command. We will define a contiguity matrix.

```
. spmatrix create contiguity W
```

This contiguity matrix sets “nearby” to mean “shares a border”.

`spmatrix` can create other types of weighting matrices. It even allows you to create custom matrices or to import matrices. See [\[SP\] spmatrix](#).

We can now run the Moran test.

```
. estat moran, errorlag(W)
Moran test for spatial dependence
H0: Error terms are i.i.d.
Errorlags: W
chi2(1)      =    94.06
Prob > chi2   =    0.0000
```

The test reports that we can reject that the residuals from the model above are independent and identically distributed (i.i.d.). In particular, the test considered the alternative hypothesis that residuals are correlated with nearby residuals as defined by **W**.

## spregress can reproduce regress results

`spregress` is the spatial autoregression command. `spregress` fits models in which the observations are not independent, as defined by the **W** weighting matrix.

Above, we fit a model under the assumption that the counties are independent. We used `regress`, Stata’s ordinary linear regression command. We typed

```
. regress unemployment college
```



We could have fit the same model and obtained the same results by using `spregress`. We would have typed

```
. sprepress unemployment college, gs2s1s
```

or

```
. sprepress unemployment college, ml
```

`spregress` is seldom used for fitting models without spatial lags or autocorrelated errors, but when it is, it reports the same linear regression results that `regress` reports, although there are some differences. Standard errors are slightly different, and `spregress` reports  $Z$  and  $\chi^2$  statistics instead of  $t$  and  $F$  statistics. `spregress` does not include the finite-sample adjustments that `regress` does because it does not expect to be used in situations where those adjustments would be appropriate.

### Fitting models with a spatial lag of the dependent variable

We will use `spregress` to fit the same model we fit using `regress` but with the addition of a spatial lag of unemployment. The model we fit will be

$$y_{uc} = \beta_0 + \beta_1 x_{cr} + \beta_2 W y_{uc} + \epsilon$$

$y_{uc}$  is the unemployment rate corresponding to variable `unemployment` in our data.  $x_{cr}$  is the college graduation rate corresponding to variable `college`.

The model we fit will include the term  $\beta_2 W y_{uc}$ , meaning that we will assume the unemployment rate spills over from nearby counties. There is a real logic to such a model. One would expect workers in high unemployment counties to seek employment nearby.

`spregress` provides two ways of fitting models: generalized spatial two-stage least squares (`gs2s1s`) and maximum likelihood (`ml`). To fit the above model, we could type

```
. sprepress unemployment college, gs2s1s dvarlag(W)
```

or

```
. sprepress unemployment college, ml dvarlag(W)
```

`spregress, ml` is statistically more efficient than `gs2s1s` when the errors are normally distributed. Efficiency is desirable, so we should use `ml`, right? That same property said differently is that `gs2s1s` is robust to violations of normality. Robustness is desirable, too. So now the choice between them hinges on whether we believe the normality assumption. That said, `ml` will provide standard errors that are also robust to violations of normality if we specify its `vce(robust)` option. Finally, `ml` takes longer to run, and that computation time increases as the number of observations increases. We will use `gs2s1s`.

```
. spreghress unemployment college, gs2sls dvarlag(W)
(254 observations)
(254 observations (places) used)
(weighting matrix defines 254 places)
```

Spatial autoregressive model  
GS2SLS estimates

Number of obs = 254  
Wald chi2(2) = 67.66  
Prob > chi2 = 0.0000  
Pseudo R2 = 0.1453

unemployment	Coefficient	Std. err.	z	P> z	[95% conf. interval]	
unemployment						
college	-.0939834	.0131033	-7.17	0.000	-.1196653	-.0683015
_cons	5.607379	.5033813	11.14	0.000	4.620769	6.593988
W						
unemployment	.2007728	.0942205	2.13	0.033	.016104	.3854415

Wald test of spatial terms: chi2(1) = 4.54 Prob > chi2 = 0.0331

Results for  $\beta_0$  and  $\beta_1$  are similar to those reported by `regress`, but that is a fluke of this example. Usually, when spillover effects are significant, other parameters change. Meanwhile, we find that  $\beta_2$  (which multiplies  $\mathbf{W}\mathbf{y}_{uc}$ ) is significant, but it is not sharply estimated. The 95% confidence interval places  $\beta_2$  in the range  $[0.02, 0.39]$ .

### Interpreting models with a spatial lag of the dependent variable

You might be tempted to think of  $\beta_1$  as the direct effect of education and  $\beta_2$  as the spillover effect, but they are not. They are ingredients into a recursive calculation of those effects. The model we fit is

$$\mathbf{y}_{uc} = \beta_0 + \beta_1 \mathbf{x}_{cr} + \beta_2 \mathbf{W}\mathbf{y}_{uc} + \epsilon$$

If  $\mathbf{x}_{cr}$  increases, that reduces  $\mathbf{y}_{uc}$  by  $\beta_1$ , and that reduction in  $\mathbf{y}_{uc}$  spills over to produce a further reduction in  $\mathbf{y}_{uc}$  of  $\beta_2 \mathbf{W}$ , and that reduction spills over to produce yet another reduction in  $\mathbf{y}_{uc}$ , and so on.

`estat impact` reports the average effects from the recursive process.

```
. estat impact
progress :100%
Average impacts
```

Number of obs = 254

	Delta-Method					
	dy/dx	std. err.	z	P> z	[95% conf. interval]	
direct						
college	-.0945245	.0130576	-7.24	0.000	-.120117	-.0689321
indirect						
college	-.0195459	.010691	-1.83	0.068	-.0405	.0014081
total						
college	-.1140705	.0171995	-6.63	0.000	-.1477808	-.0803602

In these data, both the unemployment and the graduation rates are measured in percentage points. A change of 1 is a change of 1 percentage point. The table above reports derivatives, but we can be forgiven for interpreting the results as if they were for a one-unit change. Everybody does it, and sometimes it is even justifiable, for example, if the model is linear in the variables as this one is. Even if the model were nonlinear, it would be a tolerable approximation to the truth as long as a one-unit change were small.

The table reports average changes for a 1-percentage-point increase in the college graduation rate. The direct effect is the effect of the change within the county, ignoring spillover effects. The own-county direct effect is to reduce the unemployment rate by 0.09 percentage points.

The indirect effect is the spillover effect. A 1-percentage-point increase in the college graduation rate reduces unemployment, and that reduction spills over to further reduce unemployment. The result is a 0.02 reduction in unemployment.

The total effect is the sum of the direct and indirect effects, which is  $-0.09 + -0.02 = -0.11$ .

You must use `estat impact` to interpret effects. Do not try to judge them from the coefficients that `spregress` reports because they can mislead you. For instance, if we multiplied variable `unemployment` by 100, that would not substantively change anything about the model, yet the effect on the coefficients that `spregress` estimates is surprising.

#### Summary of `spregress` results

##### Regression of unemployment and 100\*unemployment on college and W\*unemployment

	unemployment	100*unemployment
college	-0.094	-9.4
W*unemployment	0.201	0.201

Notes: Column 1 from `spregress` output above.

Column 2 from:

generate ue100 = 100\*unemployment

spregress unemployment college, gs2sls dvarlag(W)

The effect of the change in units is to multiply the coefficient on `college` ( $\beta_1$ ) by 100 just as you would expect. Yet  $\beta_2$ , the coefficient on  $\mathbf{W}y_{uc}$ , is unchanged! Comparing these two models, you might mislead yourself into thinking that the ratio of the indirect-to-direct effects is smaller in the second model, but it is not. `estat impact` continues to report the same results as it did previously, multiplied by 100:

```
. estat impact
progress   :100%
Average impacts                                     Number of obs   =       254
```

	Delta-Method dy/dx	std. err.	z	P> z	[95% conf. interval]	
direct college	-9.452455	1.30576	-7.24	0.000	-12.0117	-6.893213
indirect college	-1.954593	1.069105	-1.83	0.068	-4.05	.1408134
total college	-11.40705	1.719946	-6.63	0.000	-14.77808	-8.036016

## Fitting models with a spatial lag of independent variables

We fit a model above with a spatial lag of the dependent variable:

$$\mathbf{y}_{uc} = \beta_0 + \beta_1 \mathbf{x}_{cr} + \beta_2 \mathbf{W} \mathbf{y}_{uc} + \epsilon$$

We could instead fit a model with a spatial lag of the independent variable:

$$\mathbf{y}_{uc} = \beta_0 + \beta_1 \mathbf{x}_{cr} + \beta_2 \mathbf{W} \mathbf{x}_{cr} + \epsilon$$

We do that by typing

```
. spregress unemployment college, gs2sls ivarlag(W:college)
(254 observations)
(254 observations (places) used)
(weighting matrix defines 254 places)
```

```
Spatial autoregressive model      Number of obs =    254
GS2SLS estimates                  Wald chi2(2)   =   81.13
                                  Prob > chi2    = 0.0000
                                  Pseudo R2      = 0.2421
```

unemployment	Coefficient	Std. err.	z	P> z	[95% conf. interval]	
unemployment						
college	-.077997	.0138127	-5.65	0.000	-.1050695	-.0509245
_cons	7.424453	.3212299	23.11	0.000	6.794854	8.054053
W						
college	-.0823959	.0191586	-4.30	0.000	-.1199461	-.0448458
Wald test of spatial terms:			chi2(1) = 18.50	Prob > chi2 = 0.0000		

## Interpreting models with a spatial lag of the independent variables

Just as with lags of the dependent variable, the easy way to obtain the direct and indirect effects of independent variables is to use `estat impact`.

```
. estat impact
progress   :100%
Average impacts                                     Number of obs   =    254
```

	Delta-Method					
	dy/dx	std. err.	z	P> z	[95% conf. interval]	
direct						
college	-.077997	.0138127	-5.65	0.000	-.1050695	-.0509245
indirect						
college	-.0715273	.0166314	-4.30	0.000	-.1041243	-.0389303
total						
college	-.1495243	.0170417	-8.77	0.000	-.1829255	-.1161231

The table reports that the own-county direct effect of a 1-percentage-point increase in the college graduation rate is to reduce unemployment by 0.078 percentage points.

The across-county spillover effect of a 1-percentage-point increase in the college graduation rate is to reduce unemployment by 0.072 percentage points on average.

For those curious how the results were calculated, here are the details.

- The direct effect of college graduation rate is  $\beta_1 \mathbf{x}_{\text{cr}}$ .
- The indirect effect of college graduation rate is  $\beta_2 \mathbf{W} \mathbf{x}_{\text{cr}}$ .
- The direct effect of increasing  $\mathbf{x}_{\text{cr}}$  by 1 in all counties is

$$\Delta \mathbf{y}_{\text{uc}} = \beta_1 (\mathbf{x}_{\text{cr}} + \mathbf{1}) - \beta_1 \mathbf{x}_{\text{cr}} = \beta_1 \mathbf{1}$$

where  $\mathbf{1}$  is an  $N \times 1$  vector of 1s.

- The direct effect is that  $\mathbf{y}_{\text{uc}}$  increases by  $\beta_1$  in each county.
- The indirect effect follows the same logic:

$$\Delta \mathbf{y}_{\text{uc}} = \beta_2 \mathbf{W} (\mathbf{x}_{\text{cr}} + \mathbf{1}) - \beta_2 \mathbf{W} \mathbf{x}_{\text{cr}} = \beta_2 \mathbf{W} \mathbf{1}$$

This result states that  $\mathbf{y}_{\text{uc}}$  increases by  $(\beta_2 \mathbf{W} \mathbf{1})_i$  in county  $i$ . For different counties, there are different effects because each county is affected by its own neighbors. The average effect across counties is the average of  $\beta_2 \mathbf{W} \mathbf{1}$ .

### Fitting models with spatially autoregressive errors

We have fit models with a spatial lag of the dependent variable and with a spatial lag of the independent variable.

$$\begin{aligned} \mathbf{y}_{\text{uc}} &= \beta_0 + \beta_1 \mathbf{x}_{\text{cr}} + \beta_2 \mathbf{W} \mathbf{y}_{\text{uc}} + \boldsymbol{\epsilon} \\ \mathbf{y}_{\text{uc}} &= \beta_0 + \beta_1 \mathbf{x}_{\text{cr}} + \beta_2 \mathbf{W} \mathbf{x}_{\text{cr}} + \boldsymbol{\epsilon} \end{aligned}$$

We could instead fit a model with a spatial lag of the error:

$$\mathbf{y}_{\text{uc}} = \beta_0 + \beta_1 \mathbf{x}_{\text{cr}} + (\mathbf{I} - \rho \mathbf{W})^{-1} \boldsymbol{\epsilon}$$

We do that by typing

```
. spreghress unemployment college, gs2s1s errorlag(W)
(254 observations)
(254 observations (places) used)
(weighting matrix defines 254 places)
Estimating rho using 2SLS residuals:
Initial:      GMM criterion = .71251706
Alternative:  GMM criterion = .04381608
Rescale:      GMM criterion = .02453154
Iteration 0:  GMM criterion = .02453154
Iteration 1:  GMM criterion = .00420723
Iteration 2:  GMM criterion = .0002217
Iteration 3:  GMM criterion = .00021298
Iteration 4:  GMM criterion = .00021298
Estimating rho using GS2SLS residuals:
Iteration 0:  GMM criterion = .00566696
Iteration 1:  GMM criterion = .00486118
Iteration 2:  GMM criterion = .00486066
Iteration 3:  GMM criterion = .00486066
Spatial autoregressive model
GS2SLS estimates
Number of obs = 254
Wald chi2(1) = 37.76
Prob > chi2 = 0.0000
Pseudo R2 = 0.1869
```

unemployment	Coefficient	Std. err.	z	P> z	[95% conf. interval]	
unemployment						
college	-.0759125	.0123532	-6.15	0.000	-.1001243	-.0517008
_cons	6.292997	.2968272	21.20	0.000	5.711227	6.874768
W						
e.unemploy~t	.7697395	.0690499	11.15	0.000	.6344043	.9050748

Wald test of spatial terms:                      chi2(1) = 124.27              Prob > chi2 = 0.0000

The estimated value of the spatial autocorrelation parameter  $\rho$  is presented on the line above the Wald test:  $\hat{\rho} = 0.77$ . It is estimated to be large and significant.

$\rho$  is called the autocorrelation parameter because it is not a correlation coefficient, although it does share some characteristics with correlation coefficients. It is theoretically bounded by  $-1$  and  $1$ , and  $\rho = 0$  means that the autocorrelation is  $0$ .

estat impact does not report  $\rho$ :

```
. estat impact
progress   :100%
Average impacts                                     Number of obs   =      254
```

	Delta-Method					
	dy/dx	std. err.	z	P> z	[95% conf. interval]	
direct college	-.0759125	.0123532	-6.15	0.000	-.1001243	-.0517008
indirect college	0 (omitted)					
total college	-.0759125	.0123532	-6.15	0.000	-.1001243	-.0517008

The above output is an example of what `estat impact` produces when there are no lagged dependent or independent variables. There are no spillover effects. Spatially correlated errors do not induce spillover effects in the covariates.

Models can have all three kinds of spatial lag terms

We have shown models with each type of spatial lag term, but models can have more than one. Use `estat impact` to estimate the effects of covariates when you have lagged variables, whether dependent, independent, or both. If you include spatially correlated errors, check the size and significance of the estimated  $\rho$ .

Also see

- [SP] [Intro](#) — Introduction to spatial data and SAR models
- [SP] [spregress](#) — Spatial autoregressive models
- [SP] [spregress postestimation](#) — Postestimation tools for `spregress`
- [SP] [spset](#) — Declare data to be Sp spatial data

