

Description

To demonstrate a MIMIC model, we use the following summary statistics data:

```
. use https://www.stata-press.com/data/r19/sem_mimic1
(Multiple indicators and multiple causes)
. ssd describe
Summary statistics data from
https://www.stata-press.com/data/r19/sem_mimic1.dta
Observations:      432      Multiple indicators and multip..
Variables:          5      31 Oct 2024 08:29
                        (_dta has notes)
```

Variable name	Variable label
occpres	Occupational prestige, two-digit Dunca..
income	Total family income in units of \$2000,..
s_occupres	Subjective occupational prestige
s_income	Subjective income
s_socstat	Subjective overall social status

```
. notes
_dta:
1. Summary statistics data from Kluegel, J. R., R. Singleton, Jr., and C. E.
   Starnes. 1977. Subjective class identification: A multiple indicator
   approach. American Sociological Review 42: 599-611.
   https://doi.org/10.2307/2094558.
2. Data are also analyzed in Bollen, K. A. 1989. Structural Equations with
   Latent Variables. New York: Wiley.
3. The summary statistics represent 432 white adults included in the sample
   for the 1969 Gary Area Project for the Institute of Social Research at
   Indiana University.
4. The three subjective variables are measures of socioeconomic status based
   on an individuals perception of their own income, occupational prestige,
   and social status.
5. The income and occpres variables are objective measures of income and
   occupational prestige, respectively.
```

See *Structural models 10: MIMIC models* in [SEM] **Intro 5** for background.

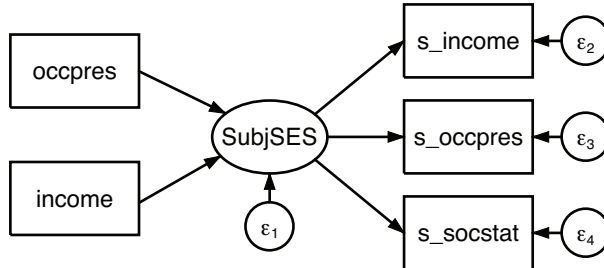
Remarks and examples

Remarks are presented under the following headings:

- [Fitting the MIMIC model](#)
- [Fitting the MIMIC model with the Builder](#)
- [Evaluating the residuals with estat residuals](#)
- [Performing likelihood-ratio tests with lrtest](#)

## Fitting the MIMIC model

Based on the data referenced above, [Bollen \(1989, 397–399\)](#) fits a MIMIC model, the path diagram of which is



Bollen includes paths that he constrains and we do not show. Our model is nonetheless equivalent to the one he shows. In his textbook, Bollen illustrates various ways the same model can be written.

```

. sem (SubjSES -> s_income s_occpres s_socstat) (SubjSES <- income occpres)
Endogenous variables
  Measurement: s_income s_occpres s_socstat
  Latent:      SubjSES
Exogenous variables
  Observed: income occpres
Fitting target model:
Iteration 0: Log likelihood = -4252.1834 (not concave)
Iteration 1: Log likelihood = -4022.8854 (not concave)
Iteration 2: Log likelihood = -3994.3867
Iteration 3: Log likelihood = -3978.7751 (not concave)
Iteration 4: Log likelihood = -3974.6636
Iteration 5: Log likelihood = -3972.0269
Iteration 6: Log likelihood = -3971.9238
Iteration 7: Log likelihood = -3971.9236
Structural equation model                                Number of obs = 432
Estimation method: ml
Log likelihood = -3971.9236
( 1) [s_income]SubjSES = 1

```

	OIM					
	Coefficient	std. err.	z	P> z	[95% conf. interval]	
Structural						
SubjSES						
income	.082732	.0138498	5.97	0.000	.0555869	.1098772
occpres	.0046275	.0012464	3.71	0.000	.0021847	.0070704
Measurement						
s_income						
SubjSES	1 (constrained)					
_cons	.9612091	.0794151	12.10	0.000	.8055583	1.11686
s_occpres						
SubjSES	.7301352	.0832915	8.77	0.000	.5668869	.8933835
_cons	1.114563	.0656195	16.99	0.000	.9859512	1.243175
s_socstat						
SubjSES	.9405161	.0934855	10.06	0.000	.7572878	1.123744
_cons	1.002113	.0706576	14.18	0.000	.863627	1.1406
var(e.s_in~e)	.2087546	.0254098			.1644474	.2649996
var(e.s_oc~s)	.2811852	.0228914			.2397153	.3298291
var(e.s_so~t)	.1807129	.0218405			.1425987	.2290146
var(e.Subj~S)	.1860097	.0270476			.1398822	.2473481

LR test of model vs. saturated: chi2(4) = 26.65      Prob > chi2 = 0.0000

#### Notes:

1. In this model, there are three observed variables that record the respondents' ideas of their perceived socioeconomic status (SES). One is the respondent's general idea of his or her SES (s\_socstat); another is based on the respondent's income (s\_income); and the last is based on the respondent's occupational prestige (s\_occpres). Those three variables form the latent variable SubjSES.
2. The other two observed variables are the respondent's income (income) and occupation, the latter measured by the two-digit Duncan SEI scores for occupations (occpres). These two variables are treated as predictors of SubjSES.

3. In the model, **item 1** is viewed as subjective and **item 2** is viewed as objective.
4. All variables are statistically significant at the 5% level, but the model versus saturated test suggests that we are not modeling the covariances well.

## Fitting the MIMIC model with the Builder

Use the diagram above for reference.

1. Open the dataset.

In the Command window, type

```
. use https://www.stata-press.com/data/r19/sem_mimic1
```

2. Open a new Builder diagram.

Select menu item **Statistics > SEM (structural equation modeling) > Model building and estimation**.

3. Create the measurement component for subject socioeconomic status.



Select the Add measurement component tool, , and then click near the center of the diagram.

In the resulting dialog box,



- a. change the *Latent variable name* to SubjSES;
- b. select `s_income`, `s_occpres`, and `s_socstat` by using the *Measurement variables control*;
- c. select **Right** in the *Measurement direction control*;
- d. click on **OK**.

If you wish, move this component by clicking on any variable and dragging it.



4. Create the two variables for the formative indicators of socioeconomic status.

- a. Select the Add observed variable tool, , and then click in the diagram to add the new variable slightly above and to the left of SubjSES. After adding it, you can click inside the rectangle and move the variable if you wish.
- b. In the Contextual Toolbar, select `occpres` by using the *Variable control*.
- c. Continuing with the  tool, click in the diagram to add another new variable below the `occpres` variable.
- d. In the Contextual Toolbar, select `income` by using the *Variable control*.


5. Create the paths for the formative indicators of socioeconomic status.

- a. Select the Add path tool, .
- b. Click in the lower-right quadrant of the `occpres` rectangle (it will highlight when you hover over it), and drag a path to the upper-left quadrant of the SubjSES oval (it will highlight when you can release to connect the path).
- c. Continuing with the  tool, click in the upper-right quadrant of the `income` rectangle, and drag a path to the lower-left quadrant of the SubjSES oval.

## 6. Clean up the direction of the errors.

The error on SubjSES is likely to have been created atop one of the existing paths. Choose the Select tool, , and then click in the SubjSES oval. Click on one of the **Error rotation** buttons, , in the Contextual Toolbar until the error is where you want it.

## 7. Estimate.

Click on the **Estimate** button, , in the Standard Toolbar, and then click on **OK** in the resulting *SEM estimation options* dialog box.

Tip: See the [tips](#) of [\[SEM\] Example 9](#) to make creating paths somewhat easier than described above.

You can open a completed diagram in the Builder by typing

```
. webgetsem sem_mimic1
```

## Evaluating the residuals with estat residuals

Remember that SEM fits covariances and means. Residuals in the SEM sense thus refer to covariances and means. If we are not fitting well, we can examine the residuals.

```
. estat residuals, normalized
```

Residuals of observed variables

Mean residuals

	s_income	s_occpres	s_socstat	income	occpres
raw	0.000	0.000	0.000	0.000	0.000
normalized	0.000	0.000	0.000	0.000	0.000

Covariance residuals

	s_income	s_occpres	s_socstat	income	occpres
s_income	0.000				
s_occpres	-0.009	0.000			
s_socstat	0.000	0.008	0.000		
income	0.101	-0.079	-0.053	0.000	
occpres	-0.856	1.482	0.049	0.000	0.000

Normalized covariance residuals

	s_income	s_occpres	s_socstat	income	occpres
s_income	0.000				
s_occpres	-0.425	0.000			
s_socstat	0.008	0.401	0.000		
income	1.362	-1.137	-0.771	0.000	
occpres	-1.221	2.234	0.074	0.000	0.000

Notes:

1. The residuals can be partitioned into two subsets: mean residuals and covariance residuals.
2. The normalized option caused the normalized residuals to be displayed.
3. Concerning mean residuals, the raw residuals and the normalized residuals are shown on a separate line of the first table.

4. Concerning covariance residuals, the raw residuals and the normalized residuals are shown in separate tables.
5. Distinguish between normalized residuals and standardized residuals. Both are available from `estat residuals`; if we wanted standardized residuals, we would have specified the `standardized` option instead of or along with `normalized`.
6. Both normalized and standardized residuals attempt to adjust the residuals in the same way. The normalized residuals are always valid, but they do not follow a standard normal distribution. The standardized residuals do follow a standard normal distribution if they can be calculated; otherwise, they will equal missing values. When both can be calculated (equivalent to both being appropriate), the normalized residuals will be a little smaller than the standardized residuals.
7. The normalized covariance residuals between `income` and `s_income` and between `occpres` and `s_occpres` are large.

## Performing likelihood-ratio tests with `lrtest`

Thus [Bollen \(1989, 397–399\)](#) suggests adding a direct path from the objective measures to the corresponding subjective measures. We are about to fit the model

```
(SubjSES -> s_income s_occpres s_socstat)   ///
(SubjSES <- income occpres)                 ///
(s_income <- income)                         /// <- new
(s_occpres <- occpres)                       //  <- new
```

For no other reason than we want to demonstrate the likelihood-ratio test, we will then use `lrtest` rather than `test` to test the joint significance of the new paths. `lrtest` compares the likelihood values of two fitted models. Thus we will use `lrtest` to compare this new model with the one above. To do that, we must plan ahead and store in memory the currently fit model:

```
. estimates store mimic1
```

Alternatively, we could skip that and calculate the joint significance of the two new paths by using a Wald test and the `test` command.

In any case, having stored the current estimates under the name `mimic1`, we can now fit our new model:

```
. sem (SubjSES -> s_income s_occpres s_socstat)
>   (SubjSES <- income occpres)
>   (s_income <- income)
>   (s_occpres <- occpres)
```

Endogenous variables

```
Observed:   s_income s_occpres
Measurement: s_socstat
Latent:      SubjSES
```

Exogenous variables

```
Observed: income occpres
```

Fitting target model:

```
Iteration 0: Log likelihood = -4267.0974 (not concave)
Iteration 1: Log likelihood = -4022.6745 (not concave)
Iteration 2: Log likelihood = -3977.0648
Iteration 3: Log likelihood = -3962.9229
Iteration 4: Log likelihood = -3962.1604
Iteration 5: Log likelihood = -3960.8404
Iteration 6: Log likelihood = -3960.7133
Iteration 7: Log likelihood = -3960.7111
Iteration 8: Log likelihood = -3960.7111
```

Structural equation model

Number of obs = 432

Estimation method: ml

Log likelihood = -3960.7111

( 1) [s\_income]SubjSES = 1

	OIM					
	Coefficient	std. err.	z	P> z	[95% conf. interval]	
Structural						
s_income						
SubjSES	1	(constrained)				
income	.0532425	.0142862	3.73	0.000	.025242	.081243
_cons	.8825314	.0781685	11.29	0.000	.7293239	1.035739
s_occpres						
SubjSES	.783781	.1011457	7.75	0.000	.585539	.982023
occpres	.0045201	.0013552	3.34	0.001	.0018641	.0071762
_cons	1.06586	.0696058	15.31	0.000	.9294353	1.202285
SubjSES						
income	.0538025	.0129158	4.17	0.000	.028488	.0791171
occpres	.0034324	.0011217	3.06	0.002	.0012339	.0056309
Measurement						
s_socstat						
SubjSES	1.195539	.1582735	7.55	0.000	.8853282	1.505749
_cons	1.07922	.078323	13.78	0.000	.9257097	1.23273
var(e.s_in~e)	.2292697	.0248905			.1853261	.2836329
var(e.s_oc~s)	.2773786	.0223972			.2367783	.3249407
var(e.s_so~t)	.1459009	.028228			.0998556	.2131785
var(e.Subj~S)	.1480275	.0278381			.1023918	.2140029

LR test of model vs. saturated:  $\chi^2(2) = 4.22$

Prob >  $\chi^2 = 0.1211$

Now we can perform the likelihood-ratio test:

```
. lrtest mimic1 .
Likelihood-ratio test
Assumption: mimic1 nested within .
LR chi2(2) = 22.42
Prob > chi2 = 0.0000
```

Notes:

1. The syntax of `lrtest` is `lrtest modelname1 modelname2`. We specified the first model name as `mimic1`, the model we previously stored. We specified the second model name as `period (.)`, meaning the model most recently fit. The order in which we specify the names is irrelevant.
2. We find the two added paths to be extremely significant.

## Reference

Bollen, K. A. 1989. *Structural Equations with Latent Variables*. New York: Wiley. <https://doi.org/10.1002/9781118619179>.

## Also see

[SEM] [Example 36g](#) — MIMIC model (generalized response)

[SEM] [sem](#) — Structural equation model estimation command

[SEM] [estat residuals](#) — Display mean and covariance residuals

[SEM] [lrtest](#) — Likelihood-ratio test of linear hypothesis

Stata, Stata Press, and Mata are registered trademarks of StataCorp LLC. Stata and Stata Press are registered trademarks with the World Intellectual Property Organization of the United Nations. StataNow and NetCourseNow are trademarks of StataCorp LLC. Other brand and product names are registered trademarks or trademarks of their respective companies. Copyright © 1985–2025 StataCorp LLC, College Station, TX, USA. All rights reserved.

For suggested citations, see the FAQ on [citing Stata documentation](#).

