spearman — Spearman's and Kendall's correlations

Description Options for spearman Methods and formulas Quick start Options for ktau Acknowledgment Menu Remarks and examples References Syntax Stored results Also see

Description

spearman displays Spearman's rank correlation coefficients for all pairs of variables in *varlist* or, if *varlist* is not specified, for all the variables in the dataset. When there are two variables, an exact *p*-value can be calculated optionally using permutations.

ktau displays Kendall's rank correlation coefficients between the variables in *varlist* or, if *varlist* is not specified, for all the variables in the dataset. ktau is intended for use on small- and moderate-sized datasets; it requires considerable computation time for larger datasets.

Quick start

Spearman's rank correlation coefficient with approximate p-value for v1 and v2

spearman v1 v2

Same as above, but report an exact p-value calculated using Monte Carlo permutations

spearman v1 v2, exact

Same as above, but perform 100,000 Monte Carlo permutations rather than the default of 10,000, and set the random-number seed for reproducibility

```
spearman v1 v2, exact(montecarlo, reps(100000) rseed(1234))
```

Display Spearman's rank correlation coefficients in a matrix for all pairs of v1, v2, and v3 spearman v1 v2 v3

Display *p*-values as well as correlation coefficients spearman v1 v2 v3, stats(rho p)

Same as above, but perform Bonferroni's adjustment to *p*-values spearman v1 v2 v3, stats(rho p) bonferroni

Kendall's rank correlation coefficients, scores, and standard errors of the scores for pairs of v1, v2, and v3

ktau v1 v2 v3, stats(taua taub score se)

Same as above, but use pairwise instead of casewise deletion ktau v1 v2 v3, stats(taua taub score se) pw

Menu

spearman

 $Statistics > Nonparametric \ analysis > Tests \ of \ hypotheses > Spearman's \ rank \ correlation$

ktau

 $Statistics > Nonparametric \ analysis > Tests \ of \ hypotheses > Kendall's \ rank \ correlation$

Syntax

Spearman's rank correlation coefficients

Spearman's rank correlation coefficients				
<pre>spearman [varlist] [if] [in] [, spearman_options]</pre>				
Kendall's rank correlation	coefficients			
ktau [varlist] [if] [in	[, ktau_options]			
spearman_options	Description			
Main				
<pre>stats(spearman_list)</pre>	list of statistics; select up to three statistics; default is stats(rho)			
$print(\hat{\#})$	<i>p</i> -value cutoff for displaying coefficients			
<u></u>				
<u>b</u> onferroni				
<u>sid</u> ak	idak report Šidák-adjusted p-values			
рพ	calculate each pairwise correlation coefficient using all available data			
<u>mat</u> rix	display output in matrix form			
<pre>exact[(exact_specs)]</pre>	report an exact <i>p</i> -value (available only when <i>varlist</i> is two variables)			
ktau_options	Description			
Main				
<pre>stats(ktau_list)</pre>	list of statistics; select up to six statistics; default is stats(taua)			
<pre>print(#)</pre>	<i>p</i> -value cutoff for displaying coefficients			
<u>b</u> onferroni				
<u>sid</u> ak	report Šidák-adjusted p-values			
Ъм	calculate each pairwise correlation coefficient using all available data			
<u>mat</u> rix	display output in matrix form			

by and collect are allowed with spearman and ktau; see [U] 11.1.10 Prefix commands.

where the elements of spearman_list may be

rho	correlation coefficient
obs	number of observations
р	<i>p</i> -value

and the elements of *ktau_list* may be

taua	correlation coefficient τ_a
taub	correlation coefficient τ_b
score	score
se	standard error of score
obs	number of observations
р	<i>p</i> -value

Options for spearman

Main

- stats(spearman_list) specifies the statistics to be displayed in the matrix of output. stats(rho) is the default. Up to three statistics may be specified; stats(rho obs p) would display the correlation coefficient, number of observations, and p-value. If varlist contains only two variables, all statistics are shown in tabular form, and stats(), print(), and star() have no effect unless the matrix option is specified.
- print(#) specifies the p-value cutoff for correlation coefficients to be printed. Correlation coefficients with larger p-values are left blank in the matrix. Typing spearman, print(.10) would display only those correlation coefficients that have p-values less than or equal to 0.10.
- star(#) specifies the p-value cutoff for correlation coefficients to be marked with a star. Typing spearman, star(.05) would star all correlation coefficients that have p-values less than or equal to 0.05.
- bonferroni makes the Bonferroni adjustment to p-values. This adjustment affects displayed p-values and the print() and star() options. Thus, spearman, print(.05) bonferroni prints coefficients with Bonferroni-adjusted p-values of 0.05 or less.
- sidak makes the Šidák adjustment to p-values. This adjustment affects displayed p-values and the print() and star() options. Thus, spearman, print(.05) sidak prints coefficients with Šidákadjusted p-values of 0.05 or less.
- pw specifies that correlations be calculated using pairwise deletion of observations with missing values. By default, spearman uses casewise deletion, where observations are ignored if any of the variables in *varlist* are missing.
- matrix forces spearman to display the statistics as a matrix, even if *varlist* contains only two variables. matrix is implied if more than two variables are specified.
- exact and exact(*exact_specs*) specify that an exact *p*-value be reported. This option is available only when *varlist* contains only two variables.
 - exact specifies that an exact *p*-value from a Monte Carlo permutation test be reported. exact is a synonym for exact (montecarlo).
 - exact(montecarlo[, options] | enumerate[, options]) specifies that an exact p-value be reported in addition to the approximate p-value. Specifying exact(montecarlo) does a Monte Carlo permutation test. Specifying exact(enumerate) does an enumeration of all possible permutations. Because the number of all possible permutations is typically extremely large, enumeration is feasible only for very small datasets. The number of permutations will be displayed, and you can click on Break to stop the computation. The exact p-value is computed by permute.
 - exact(montecarlo[, options]) allows options show, reps(#), rseed(#),
 - saving(filename[, sav_options]), level(#), dots(#), nodots, and eps(#). The show option specifies that the table produced by permute also be displayed. By default, 10,000 Monte Carlo permutations are done. That is, the default is the same as specifying exact(montecarlo, reps(10000)). The default for dots() is dots(100) when reps() is \geq 10,000; otherwise, it is dots(1). See *Options* in [R] permute.
 - exact(enumerate[, options]) allows options show, saving(filename[, sav_options]), dots(#), nodots, and eps(#). The show option specifies that the table produced by permute also be displayed. The default for dots() is dots(100). See Options in [R] permute.

Options for ktau

Main

- stats(*ktau_list*) specifies the statistics to be displayed in the matrix of output. stats(taua) is the default. Up to six statistics may be specified; stats(taua taub score se obs p) would display the correlation coefficients τ_a and τ_b , score, standard error of score, number of observations, and *p*-value. If *varlist* contains only two variables, all statistics are shown in tabular form and stats(), print(), and star() have no effect unless the matrix option is specified.
- print(#) specifies the p-value cutoff for correlation coefficients to be printed. Correlation coefficients with larger p-values are left blank in the matrix. Typing ktau, print(.10) would display only those correlation coefficients that have p-values less than or equal to 0.10.
- star(#) specifies the p-value cutoff for correlation coefficients to be marked with a star. Typing ktau, star(.05) would star all correlation coefficients that have p-values less than or equal to 0.05.
- bonferroni makes the Bonferroni adjustment to *p*-values. This adjustment affects displayed *p*-values and the print() and star() options. Thus, ktau, print(.05) bonferroni prints coefficients with Bonferroni-adjusted *p*-values of 0.05 or less.
- sidak makes the Šidák adjustment to p-values. This adjustment affects displayed p-values and the print() and star() options. Thus, ktau, print(.05) sidak prints coefficients with Šidákadjusted p-values of 0.05 or less.
- pw specifies that correlations be calculated using pairwise deletion of observations with missing values. By default, ktau uses casewise deletion, where observations are ignored if any of the variables in *varlist* are missing.
- matrix forces ktau to display the statistics as a matrix, even if *varlist* contains only two variables. matrix is implied if more than two variables are specified.

Remarks and examples

Example 1

We wish to calculate the correlation coefficients among marriage rate (mrgrate), divorce rate (divorce_rate), and median age (medage) in state data. We can calculate the standard Pearson correlation coefficients and *p*-values by typing

```
. use https://www.stata-press.com/data/r19/states2
(State data)
. pwcorr mrgrate divorce rate medage, sig
```

1 0	-	-	0 0
	mrgrate o	divorc~e	medage
mrgrate	1.0000		
divorce_rate	0.7895 0.0000	1.0000	
medage	0.0011 0.9941	-0.1526 0.2900	1.0000

We can calculate Spearman's rank correlation coefficients by typing

```
. spearman mrgrate divorce_rate medage, stats(rho p)
Number of observations = 50
```

Кеу			
rho p-value			
	mrgrate	divorc~e	medage
mrgrate	1.0000		
divorce_rate	0.6933 0.0000	1.0000	
medage	-0.4869 0.0004	-0.2455 0.0857	1.0000

The large difference in the results is caused by one observation. Nevada's marriage rate is almost 10 times higher than the state with the next-highest marriage rate. An important feature of the Spearman rank correlation is its reduced sensitivity to extreme values compared with the Pearson correlation.

We can calculate Kendall's rank correlations by typing

. ktau mrgrate divorce_rate medage, stats(taua taub p) Number of observations = 50

Key tau_a tau_b p-value			
	mrgrate	divorc~e	medage
mrgrate	0.9829 1.0000		
divorce_rate	0.5110 0.5206 0.0000	0.9804 1.0000	
medage	-0.3486 -0.3544 0.0004	-0.1698 -0.1728 0.0828	0.9845 1.0000

There are tied values for variables mrgrate, divorce_rate, and medage, so average ranks are used for the tied values. As a result, $\tau_a < 1$ on the diagonal (see *Methods and formulas* for the definition of τ_a).

According to Conover (1999, 323), "Spearman's ρ tends to be larger than Kendall's τ in absolute value. However, as a test of significance, there is no strong reason to prefer one over the other because both will produce nearly identical results in most cases."

Newson (2000a, 2000b, 2000c, 2001, 2003, 2005, 2006) introduces confidence intervals for Kendall's τ_a . The community-contributed somersd command provides these confidence intervals along with additional rank statistics such as Somers' D and Harrell's C and their corresponding confidence intervals.

See Seed (2001) for confidence intervals for Spearman's rank correlation.

Example 2

We illustrate spearman and ktau with the auto data, which contains some missing values.

Because we specified two variables, spearman displayed the sample size, correlation, and *p*-value in tabular form. To obtain just the correlation coefficient displayed in matrix form, we type

We can specify the pw option with spearman and ktau so that all nonmissing observations between a pair of variables when calculating their correlation coefficient are used. In the output below, some correlations are based on 74 observations, whereas others are based on 69 because 5 observations contain a missing value for rep78.

```
. spearman mpg price rep78, pw stats(rho obs p) star(0.01)
Number of observations:
                       min = 69
                       avg = 71
                       max = 74
   Key
    rho
    Number of obs
    p-value
                             price
                                      rep78
                     mpg
                  1.0000
         mpg
                      74
                        .
       price
                 -0.5419*
                           1.0000
                      74
                                74
                  0.0000
                                 .
                  0.3098*
                                     1.0000
       rep78
                           0.1028
                      69
                                69
                                         69
                  0.0098
                           0.4000
                                           .
```

The bonferroni and sidak options provide adjusted *p*-values:

. ktau mpg price rep78, stats(taua taub score se p) bonferroni Number of observations = 69

Кеу	
tau_a tau_b score se of score p-value	

	mpg	price	rep78
mpg	0.9471		
	1.0000		
	2222.0000		
	191.8600		
price	-0.3973	1.0000	
-	-0.4082	1.0000	
	-932.0000	2346.0000	
	192.4561	193.0682	
	0.0000	•	
rep78	0.2076	0.0648	0.7136
1	0.2525	0.0767	1.0000
	487.0000	152.0000	1674.0000
	181.7024	182.2233	172.2161
	0.0224	1.0000	

Example 3

We continue with the auto data and show an example of spearman with the exact option.

By default, exact does a Monte Carlo permutation test with 10,000 permutations. Because it is a Monte Carlo test, we set the random-number generator seed before running spearman for reproducibility.

The exact *p*-value from the Monte Carlo permutation test is 0.0214, which is close to the approximate *p*-value of 0.0222. Note that the approximate *p*-value is not based on the normal distribution or the *t* distribution. It is calculated using a beta distribution fit to the first four moments of the null distribution of Spearman's rank correlation, and these four moments are calculated exactly for any value of *N*, the number of observations. See *Methods and formulas* below.

Let's increase the number of Monte Carlo permutations to 1,000,000 to see how close the approximate *p*-value is to the exact *p*-value. We specify dots(10000) to see a dot every 10,000th permutation, rather than the default of every 1,000th permutation. This time we set the random-number generator seed using a suboption. The exact *p*-value is computed by **permute**, and we can see **permute**'s output by specifying the show suboption.

```
. spearman mpg gear_ratio if foreign == 1,
> exact(montecarlo, reps(1000000) dots(10000) rseed(1234) show)
Permutations (1,000,000): ......100,000......200,000......300,000......
> ....400,000.......500,000........600,000.......700,000......800,000.....
> .....900,000.......1,000,000 done
Monte Carlo permutation results Number of observations = 22
Permutation variable: mpg Number of permutations = 1,000,000
```

						Monte	Carlo	error
Т	T(obs)	Test	с	n	р	SE(p)	[95%	CI(p)]
rho	.4880642			1000000 1000000	.9885 .0115 .0230	.0001	.9883 .0113 .0227	

```
Notes: For lower one-sided test, c = #{T <= T(obs)} and p = p_lower = c/n.
For upper one-sided test, c = #{T >= T(obs)} and p = p_upper = c/n.
For two-sided test, p = 2*min(p_lower, p_upper); SE and CI approximate.
Number of observations = 22
Spearman's rho = 0.4881
Test of H0: mpg and gear_ratio are independent
Prob = 0.0222
Exact prob = 0.0230 (1,000,000 Monte Carlo permutations)
```

With this increase in the number of permutations, the exact *p*-value is now calculated as 0.0230. From the table produced by permute, we see that this *p*-value has a confidence interval of [0.0227, 0.0233], accounting for the Monte Carlo error. The approximate *p*-value of 0.0222 falls outside this confidence interval, but it is still very close. Not bad for N = 22.

Example 4

For very small sample sizes, an exact *p*-value can be computed by enumerating the full permutation distribution. Here is an example using the auto data with sample size N = 11.

The exact *p*-value is 0.0747. The approximate *p*-value is 0.0722, which is quite close to the exact *p*-value in this case, even with only 11 observations.

For N = 11, the permutation distribution consists of 11! = 39,916,800 permutations. However, the output above says that there are only 831,600 permutations. This is because the values of mpg in this sample are not all unique.

-	og if foreign	== 1 & mpg	<= 24
Mileage (mpg)	Freq.	Percent	Cum.
14	1	9.09	9.09
17	2	18.18	27.27
18	2	18.18	45.45
21	2	18.18	63.64
23	3	27.27	90.91
24	1	9.09	100.00
Total	11	100.00	

The multiplicities in the values of mpg yield permutations that give identical results, and we need to enumerate only the permutations that are distinct. From the values of mpg, we see that each distinct permutation has a multiplicity of 2! 2! 2! 3! = 48, and 39,916,800/48 = 831,600, which reduces considerably the number of permutations that need to be computed.

permute, which computes the permutations, permutes the first variable of the two in the spearman *varlist* command. So to best exploit this, the variable that produces the most multiplicities in the permutations should be placed first in *varlist* when doing an enumeration.

4

Charles Edward Spearman (1863–1945) was a British psychologist who made contributions to correlation, factor analysis, test reliability, and psychometrics. After several years' military service, he obtained a PhD in experimental psychology at Leipzig and became a professor at University College London, where he sustained a long program of work on the interpretation of intelligence tests. Ironically, the rank correlation version bearing his name is not the formula he advocated.

Maurice George Kendall (1907–1983) was a British statistician who contributed to rank correlation, time series, multivariate analysis, among other topics, and wrote many statistical texts. Most notably, perhaps, his advanced survey of the theory of statistics went through several editions, later ones with Alan Stuart; the baton has since passed to others. Kendall was employed in turn as a government and business statistician, as a professor at the London School of Economics, as a consultant, and as director of the World Fertility Survey. He was knighted in 1974.

Stored results

spearman stores the following in r():

r(N)	number of observations (last variable pair)
r(rho)	ρ (last variable pair)
r(p)	two-sided <i>p</i> -value (last variable pair)
r(p_1)	lower one-sided p-value (last variable pair)
r(p_u)	upper one-sided p-value (last variable pair)
r(p_exact)	two-sided exact p-value
r(p_l_exact)	lower one-sided exact p-value
r(p_u_exact)	upper one-sided exact p-value
r(n_perm)	number of permutations performed
Macros	
r(exact)	"montecarlo" or "enumerate"
r(rngstate)	random-number state used for Monte Carlo permutations
Matrices	
r(Nobs)	number of observations
r(Rho)	ρ
r(P)	two-sided p-value

If exact(..., show) is specified, the stored results from permute are returned as well; see Stored results in [R] permute.

ktau stores the following in r():

Scalars

r(N) r(tau_a) r(tau_b) r(score) r(se_score)	number of observations (last variable pair) τ_a (last variable pair) τ_b (last variable pair) Kendall's score (last variable pair) standard error of score (last variable pair)
r(p)	two-sided <i>p</i> -value (last variable pair)
Matrices	
r(Nobs)	number of observations
r(Tau_a)	$ au_{ m a}$
r(Tau_b)	$ au_{ m b}$
r(Score)	Kendall's score
r(Se_Score)	standard error of score
r(P)	two-sided p-value

Methods and formulas

Methods and formulas are presented under the following headings:

Spearman's rank correlation Exact p-values Kendall's tau

Spearman's rank correlation

Spearman's (1904) rank correlation is calculated as Pearson's correlation computed on the ranks (averaged for ties) (Conover 1999, 314–315). Ranks are as calculated by egen; see [D] egen.

If x_i and y_i , where i = 1, 2, ..., n, are the ranks of one variable pair, and n is the number of observations, then Spearman's rank correlation is

$$\rho = \frac{\sum_i (x_i - \overline{x})(y_i - \overline{y})}{\sqrt{\sum_i (x_i - \overline{x})^2} \sqrt{\sum_i = (y_i - \overline{y})^2}}$$

where $\overline{x} = (\sum_{i} x_{i})/n$ is the mean of x and \overline{y} is defined similarly.

Under the null hypothesis of independence (or, more generally, exchangeability), the distribution of ρ is given by all the possible permutations of x_i with y_i fixed. For the permutation distribution, an equivalent statistic to ρ is

$$T = \sum_{i} x_{i} y_{i}$$

The moments of T for the permutation distribution can be computed exactly. Its mean is

$$E(T) = \frac{1}{n} \{\sum_i x_i\} \{\sum_i y_i\}$$

Assume now that the ranks, x_i and y_i , are adjusted so that their means are zero. With this assumption, we have (Stuart, Ord, and Arnold 1999, eqs. 27.42 and 27.43)

$$\begin{split} E(T^2) &= \frac{1}{n-1} \{\sum_i x_i^2\} \{\sum_i y_i^2\} \\ E(T^3) &= \frac{1}{n} \{\sum_i x_i^3\} \{\sum_i y_i^3\} \\ &+ \frac{3}{n(n-1)} \{\sum_{i \neq j} x_i^2 x_j\} \{\sum_{i \neq j} y_i^2 y_j\} \\ &+ \frac{36}{n(n-1)(n-2)} \{\sum_{i < j < k} x_i x_j x_k\} \{\sum_{i < j < k} y_i y_j y_k\} \end{split}$$

$$\begin{split} E(T^4) &= \frac{1}{n} \{ \sum_i x_i^4 \} \{ \sum_i y_i^4 \} \\ &+ \frac{4}{n(n-1)} \{ \sum_{i \neq j} x_i^3 x_j \} \{ \sum_{i \neq j} y_i^3 y_j \} \\ &+ \frac{12}{n(n-1)} \{ \sum_{i < j} x_i^2 x_j^2 \} \{ \sum_{i < j} y_i^2 y_j^2 \} \\ &+ \frac{24}{n(n-1)(n-2)} \{ \sum_{i \neq j,k; j < k} x_i^2 x_j x_k \} \{ \sum_{i \neq j,k; j < k} y_i^2 y_j y_k \} \\ &+ \frac{576}{n(n-1)(n-2)(n-3)} \{ \sum_{i < j < k < l} x_i x_j x_k x_l \} \{ \sum_{i < j < k < l} y_i y_j y_k y_l \} \end{split}$$

Note that Stuart, Ord, and Arnold (1999, eq. 27.43) express $E(T^3)$ and $E(T^4)$ in terms of k statistics, but that formulation is exactly equivalent to the equations given above.

An approximate *p*-value for Spearman's rank correlation is calculated by fitting a four-parameter beta distribution to the first four moments of *T*. (See Lord [1965] and Hanson [1991] for a description of the technique of fitting a four-parameter beta distribution to moments of another distribution.) The four-parameter beta distribution with domain [l, u] is

$$f(x) = \frac{(-l+x)^{\alpha-1}(u-x)^{\beta-1}}{(u-l)^{\alpha+\beta-1}B(\alpha,\beta)}$$

where $0 \le l < u \le 1$ and $B(\alpha, \beta)$ is the beta function with $\alpha > 0$ and $\beta > 0$.

The parameters for the beta distribution are calculated as follows. Let m be the mean of T, v its variance, g_3 its skewness, and g_4 its kurtosis (see Methods and formulas in [R] summarize). Define

$$r = \frac{6(g_4 - g_3^2 - 1)}{6 + 3g_3^2 - 2g_4}$$

and

$$d=1-\frac{24(r+1)}{(r+2)(r+3)g_4-3(r-6)(r+1)}$$

Let

$$a = \frac{r(1 - \sqrt{d})}{2}$$
$$b = \frac{r(1 + \sqrt{d})}{2}$$

If $g_3 > 0$, then $\alpha = a$ and $\beta = b$; otherwise, $\alpha = b$ and $\beta = a$. The domain boundaries are given by

$$l = m - \alpha \sqrt{\frac{v(\alpha + \beta + 1)}{\alpha \beta}}$$
$$u = m + \beta \sqrt{\frac{v(\alpha + \beta + 1)}{\alpha \beta}}$$

To calculate the p-values, we first scale the observed value of T to the domain of beta:

$$s = \frac{T_{\rm obs} - l}{u-l}$$

Then the lower and upper one-sided *p*-values are given by

$$\begin{split} p_{\text{lower}} &= \texttt{ibeta}(\alpha, \beta, s) \\ p_{\text{upper}} &= \texttt{ibetatail}(\alpha, \beta, s) \end{split}$$

where ibeta is Stata's two-parameter cumulative beta distribution and ibetatail is the function for its upper tail; see [FN] **Statistical functions**. The two-sided *p*-value is given by

$$p = \min\{1, 2\min(p_{\text{lower}}, p_{\text{upper}})\}$$

Exact p-values

Exact *p*-values for Spearman's rank correlation are computed by permute. For details on the permutation computation, see [R] permute.

Kendall's tau

Kendall's τ is calculated in the following manner. For any two pairs of ranks (x_i, y_i) and (x_j, y_j) , $1 \le i, j \le n$, define them as concordant if

$$(x_i - x_j)(y_i - y_j) > 0$$

and discordant if this product is less than zero.

Kendall's (1938; also see Kendall and Gibbons [1990] or Bland [2015], 187–188) score S is defined as C - D, where C(D) is the number of concordant (discordant) pairs. Let N = n(n-1)/2 be the total number of pairs, so τ_a is given by

$$\tau_{\rm a} = \frac{S}{N}$$

and $\tau_{\rm b}$ is given by

$$\tau_{\rm b} = \frac{S}{\sqrt{N - U}\sqrt{N - V}}$$

where

$$\begin{split} U &= \sum_{i=1}^{N_1} u_i (u_i - 1)/2 \\ V &= \sum_{j=1}^{N_2} v_j (v_j - 1)/2 \end{split}$$

and where N_1 is the number of sets of tied x values, u_i is the number of tied x values in the *i*th set, N_2 is the number of sets of tied y values, and v_i is the number of tied y values in the *j*th set.

Under the null hypothesis of independence, the variance of S is exactly (Kendall and Gibbons 1990, 66)

$$\begin{split} \mathrm{Var}(S) &= \frac{1}{18} \bigg\{ n(n-1)(2n+5) - \sum_{i=1}^{N_1} u_i(u_i-1)(2u_i+5) - \sum_{j=1}^{N_2} v_j(v_j-1)(2v_j+5) \bigg\} \\ &\quad + \frac{1}{9n(n-1)(n-2)} \bigg\{ \sum_{i=1}^{N_1} u_i(u_i-1)(u_i-2) \bigg\} \bigg\{ \sum_{j=1}^{N_2} v_j(v_j-1)(v_j-2) \bigg\} \\ &\quad + \frac{1}{2n(n-1)} \bigg\{ \sum_{i=1}^{N_1} u_i(u_i-1) \bigg\} \bigg\{ \sum_{j=1}^{N_2} v_j(v_j-1) \bigg\} \end{split}$$

Using a normal approximation with a continuity correction,

$$z = \frac{|S| - 1}{\sqrt{\operatorname{Var}(S)}}$$

For the hypothesis of independence, the statistics S, τ_a , and τ_b produce equivalent tests and give the same p-value.

For Kendall's τ , the normal approximation is surprisingly accurate for sample sizes as small as 8, at least for calculating *p*-values under the null hypothesis for continuous variables. See Kendall and Gibbons [1990, chap. 4], who also present some tables for calculating exact *p*-values for n < 10.

Let v be the number of variables specified so that k = v(v-1)/2 correlation coefficients are to be estimated. If bonferroni is specified, the adjusted p-value is $p' = \min\{1, 1-(1-p)^n\}$. See Methods and formulas in [R] **oneway** for a more complete description of the logic behind these adjustments.

Early work on rank correlation is surveyed by Kruskal (1958).

Acknowledgment

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Also see

- [R] correlate Correlations of variables
- [R] nptrend Tests for trend across ordered groups
- [R] permute Permutation tests

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