

**heckpoisson postestimation** — Postestimation tools for heckpoisson

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## Postestimation commands

The following postestimation commands are available after `heckpoisson`:

Command	Description
<code>contrast</code>	contrasts and ANOVA-style joint tests of estimates
<code>estat ic</code>	Akaike's and Schwarz's Bayesian information criteria (AIC and BIC)
<code>estat summarize</code>	summary statistics for the estimation sample
<code>estat vce</code>	variance–covariance matrix of the estimators (VCE)
<code>estat (svy)</code>	postestimation statistics for survey data
<code>estimates</code>	cataloging estimation results
* <code>forecast</code>	dynamic forecasts and simulations
* <code>hausman</code>	Hausman's specification test
<code>lincom</code>	point estimates, standard errors, testing, and inference for linear combinations of coefficients
* <code>lrtest</code>	likelihood-ratio test
<code>margins</code>	marginal means, predictive margins, marginal effects, and average marginal effects
<code>marginsplot</code>	graph the results from margins (profile plots, interaction plots, etc.)
<code>nlcom</code>	point estimates, standard errors, testing, and inference for nonlinear combinations of coefficients
<code>predict</code>	predictions, residuals, influence statistics, and other diagnostic measures
<code>predictnl</code>	point estimates, standard errors, testing, and inference for generalized predictions
<code>pwcompare</code>	pairwise comparisons of estimates
<code>suest</code>	seemingly unrelated estimation
<code>test</code>	Wald tests of simple and composite linear hypotheses
<code>testnl</code>	Wald tests of nonlinear hypotheses

\* `forecast`, `hausman`, and `lrtest` are not appropriate with `svy` estimation results.

# predict

## Description for predict

`predict` creates new variables containing predictions such as number of events, incidence rates, conditional predicted number of events, probabilities, linear predictions, and equation-level scores.

## Menu for predict

Statistics > Postestimation

## Syntax for predict

```
predict [type] newvar [if] [in] [, statistic nooffset]
```

```
predict [type] { stub* | newvarlist } [if] [in] , scores
```

<i>statistic</i>	Description
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Main

<code>n</code>	number of events; the default
<code>ir</code>	incidence rate
<code>ncond</code>	predicted number of events conditional on $y_j$ being observed
<code>pr(<math>n</math>)</code>	$\Pr(y_j = n)$
<code>pr(<math>a, b</math>)</code>	$\Pr(a \leq y_j \leq b)$
<code>pselect</code>	$\Pr(y_j \text{ observed})$
<code>xb</code>	linear prediction
<code>xbselect</code>	linear prediction for selection equation

These statistics are available both in and out of sample; type `predict ... if e(sample) ...` if wanted only for the estimation sample.

## Options for predict

Main

`n`, the default, calculates the predicted number of events, which is  $\exp(\mathbf{x}_j\beta + \sigma^2/2)$  if neither `offset()` nor `exposure()` was specified when the model was fit; is  $\exp(\mathbf{x}_j\beta + \sigma^2/2 + \text{offset}_j)$  if `offset()` was specified; or is  $\exp(\mathbf{x}_j\beta + \sigma^2/2) \times \text{exposure}_i$  if `exposure()` was specified.

`ir` calculates the incidence rate  $\exp(\mathbf{x}_j\beta + \sigma^2/2)$ , which is the predicted number of events when exposure is 1. Specifying `ir` is equivalent to specifying `n` when neither `offset()` nor `exposure()` was specified when the model was fit.

`ncond` calculates the predicted number of events conditional on  $y_j$  being observed, which is  $\exp(\mathbf{x}_j\beta + \sigma^2/2)\Phi(\mathbf{w}_j\gamma + \rho\sigma)/\Phi(\mathbf{w}_j\gamma)$ .

`pr( $n$ )` calculates the probability  $\Pr(y_j = n)$ , where  $n$  is a nonnegative integer that may be specified as a number or a variable.

`pr(a,b)` calculates the probability  $\Pr(a \leq y_j \leq b)$ , where  $a$  and  $b$  are nonnegative integers that may be specified as numbers or variables;

$b$  missing ( $b \geq .$ ) means  $+\infty$ ;

`pr(20,.)` calculates  $\Pr(y_j \geq 20)$ ;

`pr(20,b)` calculates  $\Pr(y_j \geq 20)$  in observations for which  $b \geq .$  and calculates

$\Pr(20 \leq y_j \leq b)$  elsewhere.

`pr(.,b)` produces a syntax error. A missing value in an observation of the variable  $a$  causes a missing value in that observation for `pr(a,b)`.

`pselect` calculates the probability of selection (or being observed):

$\Pr(y_j \text{ observed}) = \Pr(\mathbf{w}_j\gamma + \epsilon_{2j} > 0)$

`xb` calculates the linear prediction for the dependent count variable, which is  $\mathbf{x}_j\beta$  if neither `offset()` nor `exposure()` was specified;  $\mathbf{x}_j\beta + \text{offset}_j^\beta$  if `offset()` was specified; or  $\mathbf{x}_j\beta + \ln(\text{exposure}_j)$  if `exposure()` was specified.

`xbselect` calculates the linear prediction for the selection equation, which is  $\mathbf{w}_j\gamma$  if `offset()` was not specified in `select()` and is  $\mathbf{w}_j\gamma + \text{offset}_j^\gamma$  if `offset()` was specified in `select()`.

`nooffset` is relevant only if you specified `offset()` or `exposure()` when you fit the model. It modifies the calculations made by `predict` so that they ignore the offset or exposure variable; the linear prediction is treated as  $\mathbf{x}_j\beta$  rather than as  $\mathbf{x}_j\beta + \text{offset}_j$  or  $\mathbf{x}_j\beta + \ln(\text{exposure}_j)$ .

`scores` calculates equation-level score variables.

The first new variable will contain  $\partial \ln L / \partial (\mathbf{x}_j\beta)$ .

The second new variable will contain  $\partial \ln L / \partial (\mathbf{w}_j\gamma)$ .

The third new variable will contain  $\partial \ln L / \partial \text{atanh } \rho$ .

The fourth new variable will contain  $\partial \ln L / \partial \ln \sigma$ .

## margins

### Description for margins

`margins` estimates margins of response for number of events, incidence rates, conditional predicted number of events, probabilities, and linear predictions.

### Menu for margins

Statistics > Postestimation

### Syntax for margins

```
margins [marginlist] [, options]
```

```
margins [marginlist] , predict(statistic ...) [predict(statistic ...) ...] [options]
```

<i>statistic</i>	Description
<code>n</code>	number of events; the default
<code>ir</code>	incidence rate
<code>ncond</code>	predicted number of events conditional on $y_j$ being observed
<code>pr(<math>n</math>)</code>	$\Pr(y_j = n)$
<code>pr(<math>a, b</math>)</code>	$\Pr(a \leq y_j \leq b)$
<code>p<math>sel</math></code>	$\Pr(y_j \text{ observed})$
<code>xb</code>	linear prediction
<code>x<math>b</math>sel</code>	linear prediction for selection equation

Statistics not allowed with `margins` are functions of stochastic quantities other than `e(b)`.

For the full syntax, see [\[R\] margins](#).

## Remarks and examples

[stata.com](http://www.stata.com)

### ► Example 1: Obtaining margins for a count model with selection

In [example 1](#) of [\[R\] heckpoisson](#), we fit a model for the number of patents. In that example, we are interested in the effect of R&D expenditures on the number of patents received by a firm. We continue that example to determine the magnitude of the effect of R&D expenditures on the number of patents and compare this effect for IT and non-IT sectors.

After reading in the data and fitting the model, we use `margins` to estimate the effect of an increase of a million dollars in R&D expenditures (`expenditure`) on the number of patents (`npatents`) for firms in the IT and non-IT sectors (`tech`).

To do this, we use the `at()` option of `margins`. We use the observed values in our first scenario, so we tell `margins` to set `expenditure` equal to itself. For our second scenario, we tell `margins` to set `expenditure` equal to the observed value plus 1 because expenditures are measured in millions of dollars. We include the `post` option so that we can perform additional calculations later.

```

. use https://www.stata-press.com/data/r16/patent
(Fictional data on patents and R&D)
. quietly heckpoisson npatents expenditure i.tech,
> select(applied = expenditure size i.tech)
. margins i.tech, at(expenditure = generate(expenditure))
> at(expenditure = generate(expenditure+1)) post
Predictive margins                                Number of obs      =      10,000
Model VCE      : OIM
Expression     : Predicted number of events, predict()
1._at         : expenditure      = expenditure
2._at         : expenditure      = expenditure+1

```

	Delta-method			z	P> z	[95% Conf. Interval]	
	Margin	Std. Err.					
_at#tech							
1 #							
non-IT se..	1.276213	.0556644	22.93	0.000	1.167112	1.385313	
1#IT sector	2.287013	.080119	28.55	0.000	2.129983	2.444044	
2 #							
non-IT se..	2.099539	.131364	15.98	0.000	1.84207	2.357007	
2#IT sector	3.76244	.2226221	16.90	0.000	3.326109	4.198771	

The output indicates that the expected number of patents for non-IT firms is about 1.28 compared with 2.29 for firms in the IT sector.

The second scenario shows the expected number of patents after our hypothetical increase in R&D expenditures. In the non-IT sector, the expected number of patents received would be about 2.10 compared with 3.76 in the IT sector. It appears that increasing expenditures may have a larger effect for IT firms—the difference between the two scenarios is 1.47 for IT firms and only 0.82 for non-IT firms. We can test whether the effect of increasing expenditures is different for IT and non-IT firms. We use `lincom` to obtain an estimate of the difference in the differences between scenarios for the two sectors and a test of its significance. We ask for the differences by referring to the scenarios as 1.\_at and 2.\_at and by referring to the sector using the value that corresponds to the IT sector indicator, 1.tech for IT firms and 0.tech otherwise.

```

. lincom (_b[2._at#1.tech] - _b[1._at#1.tech]) -
> (_b[2._at#0.tech] - _b[1._at#0.tech])
(1) 1bn._at#0bn.tech - 1bn._at#1.tech - 2._at#0bn.tech + 2._at#1.tech = 0

```

	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
(1)	.6521006	.0917299	7.11	0.000	.4723133	.8318878

We find that the expected effect of increasing R&D expenditures by one million dollars is 0.65 patents larger for IT firms than for non-IT firms, and this difference is significantly different from 0.

◀

## Methods and formulas

Suppose that the count outcome  $y_j$  has covariates  $\mathbf{x}_j$  and the selection outcome  $s_j$  has covariates  $\mathbf{w}_j$ .  $y_j$  is assumed to have a Poisson distribution, conditional on  $\mathbf{x}_j$ , with conditional mean

$$E(y_j | \mathbf{x}_j, \epsilon_{1j}) = \mu_j = \exp(\mathbf{x}_j \boldsymbol{\beta} + \epsilon_{1j})$$

$s_j$  is a binary outcome from a latent-variable model:

$$s_j = \begin{cases} 1, & \text{if } \mathbf{w}_j\gamma + \epsilon_{2j} > 0 \\ 0, & \text{otherwise} \end{cases}$$

The expectation of  $y_j$  conditional on covariates  $\mathbf{x}_j$  for the whole population is

$$E(y_j|\mathbf{x}_j) = \exp(\mathbf{x}_j\boldsymbol{\beta} + \sigma^2/2)$$

Furthermore, if we want the expectation of  $y_j$  only if it was observed, then the formula is

$$E(y_j|\mathbf{x}_j, \mathbf{w}_j, s_j = 1) = \exp(\mathbf{x}_j\boldsymbol{\beta} + \sigma^2/2) \frac{\Phi(\mathbf{w}_j\gamma + \rho\sigma)}{\Phi(\mathbf{w}_j\gamma)}$$

We note that if  $\rho = 0$ , this expectation is the same as its population version.

We can also predict the probability of  $y_j$  conditional on  $\mathbf{x}_j$ . Note that although  $y_j$  is Poisson-distributed conditional on  $\epsilon_1$  and  $\mathbf{x}_j$ , the distribution of  $y_j$  is unknown unconditional on  $\epsilon_1$ .

$$\Pr(y_j = n|\mathbf{x}_j) = \int_{-\infty}^{\infty} \Pr(y_j = n|\mathbf{x}_j, \epsilon_1) \phi(\epsilon_1/\sigma) d\epsilon_1$$

As in the implementation of log likelihood, we approximate this integral by Gauss–Hermite quadrature.

## Also see

[R] **heckpoisson** — Poisson regression with sample selection

[U] **20 Estimation and postestimation commands**