Title

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heckoprobit — Ordered probit model with sample selection

Description

`heckoprobit` fits maximum-likelihood ordered probit models with sample selection.

Quick start

Ordered probit model of `y` on `x` with selection indicated by binary variable `selected` and predicted by `v`

```
heckoprobit y x, select(selected = v x)
```

Add indicator variables for each level of categorical variable `a`

```
heckoprobit y x i.a, select(selected = v x i.a)
```

Account for complex sampling design using `svyset` data

```
svy: heckoprobit y x, select(selected = v x)
```

Menu

Statistics > Sample-selection models > Ordered probit model with selection

Syntax

```
heckoprobit depvar indepvars [if] [in] [weight],
  select([depvar_s =] varlist_s [, noconstant offset(varname_o)]) [options]
```

options	Description
Model	
* `select()`	specify selection equation: dependent and independent variables; whether to have constant term and offset variable
`offset(`*varname*`)`	include *varname* in model with coefficient constrained to 1
`constraints(`*constraints*`)`	apply specified linear constraints
SE/Robust	
`vce(`*vcetype*`)`	*vcetype* may be `oim`, `robust`, `cluster` *clustvar*, `opg`, `bootstrap`, or `jackknife`
Reporting	
`level(#)`	set confidence level; default is `level(95)`
`first`	report first-step probit estimates
`noheader`	do not display header above parameter table
`nofootnote`	do not display footnotes below parameter table
`nocnsreport`	do not display constraints
display_options	control columns and column formats, row spacing, line width, display of omitted variables and base and empty cells, and factor-variable labeling
Maximization	
maximize_options	control the maximization process; seldom used
`collinear`	keep collinear variables
`coeflegend`	display legend instead of statistics

* `select()` is required.

The full specification is `select(`[*depvar*$_s$ `=`] *varlist*$_s$ [`, noconstant offset(`*varname*$_o$`)`]`)`.

indepvars and *varlist*$_s$ may contain factor variables; see **[U] 11.4.3 Factor variables**.

depvar, *indepvars*, *depvar*$_s$, and *varlist*$_s$ may contain time-series operators; see **[U] 11.4.4 Time-series varlists**.

`bayes`, `bootstrap`, `by`, `jackknife`, `rolling`, `statsby`, and `svy` are allowed; see **[U] 11.1.10 Prefix commands**. For more details, see [BAYES] **bayes: heckoprobit**.

Weights are not allowed with the `bootstrap` prefix; see [R] **bootstrap**.

`vce()`, `first`, and weights are not allowed with the `svy` prefix; see [SVY] **svy**.

`pweights`, `fweights`, and `iweights` are allowed; see **[U] 11.1.6 weight**.

`collinear` and `coeflegend` do not appear in the dialog box.

See **[U] 20 Estimation and postestimation commands** for more capabilities of estimation commands.

Options

Model

`select(`[*depvar*$_s$ `=`] *varlist*$_s$ [`, noconstant offset(`*varname*$_o$`)`]`)` specifies the variables and options for the selection equation. It is an integral part of specifying a selection model and is

required. The selection equation should contain at least one variable that is not in the outcome equation.

If $depvar_s$ is specified, it should be coded as 0 or 1, 0 indicating an observation not selected and 1 indicating a selected observation. If $depvar_s$ is not specified, observations for which *depvar* is not missing are assumed selected, and those for which *depvar* is missing are assumed not selected.

`noconstant` suppresses the selection constant term (intercept).

`offset(`$varname_o$`)` specifies that selection offset $varname_o$ be included in the model with the coefficient constrained to be 1.

`offset(`*varname*`)`, `constraints(`*constraints*`)`; see [R] **Estimation options**.

SE/Robust

`vce(`*vcetype*`)` specifies the type of standard error reported, which includes types that are derived from asymptotic theory (`oim`, `opg`), that are robust to some kinds of misspecification (`robust`), that allow for intragroup correlation (`cluster` *clustvar*), and that use bootstrap or jackknife methods (`bootstrap`, `jackknife`); see [R] ***vce_option***.

Reporting

`level(#)`; see [R] **Estimation options**.

`first` specifies that the first-step probit estimates of the selection equation be displayed before estimation.

`noheader` suppresses the header above the parameter table.

`nofootnote` suppresses the footnotes displayed below the parameter table.

`nocnsreport`; see [R] **Estimation options**.

display_options: `noci`, `nopvalues`, `noomitted`, `vsquish`, `noemptycells`, `baselevels`, `allbaselevels`, `nofvlabel`, `fvwrap(#)`, `fvwrapon(`*style*`)`, `cformat(%`*fmt*`)`, `pformat(%`*fmt*`)`, `sformat(%`*fmt*`)`, and `nolstretch`; see [R] **Estimation options**.

Maximization

maximize_options: `difficult`, `technique(`*algorithm_spec*`)`, `iterate(#)`, `[no]log`, `trace`, `gradient`, `showstep`, `hessian`, `showtolerance`, `tolerance(#)`, `ltolerance(#)`, `nrtolerance(#)`, `nonrtolerance`, and `from(`*init_specs*`)`; see [R] **Maximize**. These options are seldom used.

Setting the optimization type to `technique(bhhh)` resets the default *vcetype* to `vce(opg)`.

The following options are available with `heckoprobit` but are not shown in the dialog box:

`collinear`, `coeflegend`; see [R] **Estimation options**.

Remarks and examples

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`heckoprobit` estimates the parameters of a regression model for an ordered categorical outcome from a nonrandom sample known as a selected sample. Selected samples suffer from "selection on unobservables" because the errors that determine whether a case is missing are correlated with the errors that determine the outcome.

For ordered categorical regression from samples that do not suffer from selection on unobservables, see [R] **oprobit** or [R] **ologit**. For regression of a continuous outcome variable from a selected sample, see [R] **heckman**.

Even though we are interested in modeling a single ordinal outcome, there are two dependent variables in the ordered probit sample-selection model because we must also model the sample-selection process. First, there is the ordinal outcome y_j. Second, there is a binary variable that indicates whether each case in the sample is observed or unobserved. To handle the sample-selection problem, we model both dependent variables jointly. Both variables are categorical. Their categorical values are determined by the values of linear combinations of covariates and normally distributed error terms relative to certain cutpoints that partition the real line. The error terms used in the determination of selection and the ordinal outcome value may be correlated.

The probability that the ordinal outcome y_j is equal to the value v_h is given by the probability that $\mathbf{x}_j\boldsymbol{\beta} + u_{1j}$ falls within the cutpoints κ_{h-1} and κ_h,

$$\Pr(y_j = v_h) = \Pr(\kappa_{h-1} < \mathbf{x}_j\boldsymbol{\beta} + u_{1j} \leq \kappa_h)$$

where $\mathbf{x}_j$ is the outcome covariates, $\boldsymbol{\beta}$ is the coefficients, and u_{1j} is a random-error term. The observed outcome values $v_1, \ldots, v_H$ are integers such that $v_i < v_m$ for $i < m$. κ_0 is taken as $-\infty$ and κ_H is taken as $+\infty$.

We model the selection process for the outcome by

$$s_j = \mathbf{1}(\mathbf{z}_j\boldsymbol{\gamma} + u_{2j} > 0)$$

where $s_j = 1$ if we observed y_j and 0 otherwise, $\mathbf{z}_j$ is the covariates used to model the selection process, $\boldsymbol{\gamma}$ is the coefficients for the selection process, $\mathbf{1}(\cdot)$ denotes the indicator function, and u_{2j} is a random-error term.

(u_{1j}, u_{2j}) have bivariate normal distribution with mean zero and variance matrix

$$\begin{bmatrix} 1 & \rho \\ \rho & 1 \end{bmatrix}$$

When $\rho \neq 0$, standard ordered probit techniques applied to the outcome equation yield inconsistent results. `heckoprobit` provides consistent, asymptotically efficient estimates for all the parameters in such models.

De Luca and Perotti (2011) describe the maximum likelihood estimator used in `heckoprobit`.

▷ Example 1

We have a simulated dataset containing a sample of 5,000 women, 3,480 of whom work. The outcome of interest is a woman's job satisfaction, and we suspect that unobservables that determine job satisfaction and the unobservables that increase the likelihood of employment are correlated. Women may make a decision to work based on how satisfying their job would be. We estimate the parameters of an ordered probit sample-selection model for the outcome of job satisfaction (`satisfaction`) with selection on employment (`work`). Age (`age`) and years of education (`education`) are used as outcome covariates, and we also expect that they affect selection. Additional covariates for selection are marital status (`married`) and the number of children at home (`children`).

Here we estimate the parameters of the model with `heckoprobit`. We use the factorial interaction of `married` and `children` in `select()`. This specifies that the number of children and marital status affect selection, and it allows the effect of the number of children to differ among married and nonmarried women. The factorial interaction is specified using factor-variable notation, which is described in [U] **11.4.3 Factor variables**.

```
. use https://www.stata-press.com/data/r16/womensat
(Job satisfaction, female)
. heckoprobit satisfaction education age,
> select(work=education age i.married##c.children)
Fitting oprobit model:
Iteration 0:   log likelihood = -3934.1474
Iteration 1:   log likelihood =  -3571.886
Iteration 2:   log likelihood = -3570.2616
Iteration 3:   log likelihood = -3570.2616
Fitting selection model:
Iteration 0:   log likelihood = -3071.0775
Iteration 1:   log likelihood = -2565.5092
Iteration 2:   log likelihood = -2556.8369
Iteration 3:   log likelihood = -2556.8237
Iteration 4:   log likelihood = -2556.8237
Comparison:    log likelihood = -6127.0853
Fitting full model:
Iteration 0:   log likelihood = -6127.0853
Iteration 1:   log likelihood = -6093.8868
Iteration 2:   log likelihood =  -6083.215
Iteration 3:   log likelihood = -6083.0376
Iteration 4:   log likelihood = -6083.0372

Ordered probit model with sample selection      Number of obs     =      5,000
                                                      Selected    =      3,480
                                                      Nonselected =      1,520
                                                Wald chi2(2)      =     842.42
Log likelihood = -6083.037                      Prob > chi2       =     0.0000
```

	Coef.	Std. Err.	z	P>|z|	[95% Conf.	Interval]
satisfaction						
education	.1536381	.0068266	22.51	0.000	.1402583	.1670179
age	.0334463	.0024049	13.91	0.000	.0287329	.0381598
work						
education	.0512494	.0068095	7.53	0.000	.037903	.0645958
age	.0288084	.0026528	10.86	0.000	.023609	.0340078
1.married	.6120876	.0700055	8.74	0.000	.4748794	.7492958
children	.5140995	.0288529	17.82	0.000	.4575489	.5706501
married# c.children						
1	-.1337573	.035126	-3.81	0.000	-.202603	-.0649117
_cons	-2.203036	.125772	-17.52	0.000	-2.449545	-1.956528
/cut1	1.728757	.1232063	14.03	0.000	1.487277	1.970237
/cut2	2.64357	.116586	22.67	0.000	2.415066	2.872075
/cut3	3.642911	.1178174	30.92	0.000	3.411993	3.873829
/athrho	.7430919	.0780998	9.51	0.000	.5900191	.8961646
rho	.6310096	.0470026			.5299093	.7144252

```
LR test of indep. eqns. (rho = 0):   chi2(1) =    88.10   Prob > chi2 = 0.0000
```

The output shows several iteration logs. The first iteration log corresponds to running the ordered probit model for those observations in the sample where we have observed the outcome. The second iteration log corresponds to running the selection probit model, which models whether we observe

our outcome of interest. If $\rho = 0$, the sum of the log likelihoods from these two models will equal the log likelihood of the ordered probit sample-selection model; this sum is printed in the iteration log as the comparison log likelihood. The final iteration log is for fitting the full ordered probit sample-selection model.

The Wald test in the header is highly significant, indicating a good model fit. All the covariates are statistically significant. The likelihood-ratio test in the footer indicates that we can reject the null hypothesis that the errors for outcome and selection are uncorrelated. This means that we should use the ordered probit sample-selection model instead of the simple ordered probit model.

The positive estimate of 0.63 for ρ indicates that unobservables that increase job satisfaction tend to occur with unobservables that increase the chance of having a job.

◁

Stored results

`heckoprobit` stores the following in `e()`:

Scalars

`e(N)`	number of observations
`e(N_selected)`	number of selected observations
`e(N_nonselected)`	number of nonselected observations
`e(N_cd)`	number of completely determined observations
`e(k_cat)`	number of categories
`e(k)`	number of parameters
`e(k_eq)`	number of equations in `e(b)`
`e(k_eq_model)`	number of equations in overall model test
`e(k_aux)`	number of auxiliary parameters
`e(k_dv)`	number of dependent variables
`e(df_m)`	model degrees of freedom
`e(ll)`	log likelihood
`e(ll_c)`	log likelihood, comparison model
`e(N_clust)`	number of clusters
`e(chi2)`	χ^2
`e(chi2_c)`	χ^2 for comparison test
`e(p)`	p-value for model test
`e(p_c)`	p-value for comparison test
`e(rho)`	ρ
`e(rank)`	rank of `e(V)`
`e(ic)`	number of iterations
`e(rc)`	return code
`e(converged)`	`1` if converged, `0` otherwise

Macros

e(cmd)	heckoprobit
e(cmdline)	command as typed
e(depvar)	names of dependent variables
e(wtype)	weight type
e(wexp)	weight expression
e(title)	title in estimation output
e(clustvar)	name of cluster variable
e(offset1)	offset for regression equation
e(offset2)	offset for selection equation
e(chi2type)	Wald or LR; type of model χ^2 test
e(chi2_ct)	type of comparison χ^2 test
e(vce)	*vcetype* specified in vce()
e(vcetype)	title used to label Std. Err.
e(opt)	type of optimization
e(which)	max or min; whether optimizer is to perform maximization or minimization
e(ml_method)	type of ml method
e(user)	name of likelihood-evaluator program
e(technique)	maximization technique
e(properties)	b V
e(predict)	program used to implement predict
e(marginsok)	predictions allowed by margins
e(marginsnotok)	predictions disallowed by margins
e(marginsdefault)	default predict() specification for margins
e(asbalanced)	factor variables fvset as asbalanced
e(asobserved)	factor variables fvset as asobserved

Matrices

e(b)	coefficient vector
e(Cns)	constraints matrix
e(ilog)	iteration log (up to 20 iterations)
e(gradient)	gradient vector
e(cat)	category values
e(V)	variance–covariance matrix of the estimators
e(V_modelbased)	model-based variance

Functions

e(sample)	marks estimation sample

In addition to the above, the following is stored in r():

Matrices

r(table)	matrix containing the coefficients with their standard errors, test statistics, p-values, and confidence intervals

Note that results stored in r() are updated when the command is replayed and will be replaced when any r-class command is run after the estimation command.

Methods and formulas

De Luca and Perotti (2011) provide an introduction to this model.

The ordinal outcome equation is

$$y_j = \sum_{h=1}^{H} v_h 1\left(\kappa_{h-1} < \mathbf{x}_j\boldsymbol{\beta} + u_{1j} \leq \kappa_h\right)$$

where $\mathbf{x}_j$ is the outcome covariates, $\boldsymbol{\beta}$ is the coefficients, and u_{1j} is a random-error term. The observed outcome values $v_1, \ldots, v_H$ are integers such that $v_i < v_m$ for $i < m$. $\kappa_1, \ldots, \kappa_{H-1}$ are real numbers such that $\kappa_i < \kappa_m$ for $i < m$. κ_0 is taken as $-\infty$ and κ_H is taken as $+\infty$.

The selection equation is

$$s_j = \mathbf{1}(\mathbf{z}_j\boldsymbol{\gamma} + u_{2j} > 0)$$

where $s_j = 1$ if we observed y_j and 0 otherwise, $\mathbf{z}_j$ is the covariates used to model the selection process, $\boldsymbol{\gamma}$ is the coefficients for the selection process, and u_{2j} is a random-error term.

(u_{1j}, u_{2j}) have bivariate normal distribution with mean zero and variance matrix

$$\begin{bmatrix} 1 & \rho \\ \rho & 1 \end{bmatrix}$$

Let $a_j = z_j\boldsymbol{\gamma} + \text{offset}_j^{\gamma}$ and $b_j = x_j\boldsymbol{\beta} + \text{offset}_j^{\beta}$. This yields the log likelihood

$$\ln L = \sum_{j \notin S} w_j \ln\{\Phi(-a_j)\} + \sum_{h=1}^{H} \sum_{\substack{j \in S \\ y_j = v_h}} w_j \ln\{\Phi_2(a_j, \kappa_h - b_j, -\rho) - \Phi_2(a_j, \kappa_{h-1} - b_j, -\rho)\}$$

where S is the set of observations for which y_j is observed, $\Phi_2(\cdot)$ is the cumulative bivariate normal distribution function (with mean $[\,0 \;\; 0\,]'$), $\Phi(\cdot)$ is the standard cumulative normal, and w_j is an optional weight for observation j.

In the maximum likelihood estimation, ρ is not directly estimated. Directly estimated is atanh ρ:

$$\text{atanh}\,\rho = \frac{1}{2}\ln\left(\frac{1+\rho}{1-\rho}\right)$$

From the form of the likelihood, it is clear that if $\rho = 0$, the log likelihood for the ordered probit sample-selection model is equal to the sum of the ordered probit model for the outcome y and the selection model. We can perform a likelihood-ratio test by comparing the log likelihood of the full model with the sum of the log likelihoods for the ordered probit and selection models.

References

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Also see

[R] **heckoprobit postestimation** — Postestimation tools for heckoprobit

[R] **heckman** — Heckman selection model

[R] **heckpoisson** — Poisson regression with sample selection

[R] **heckprobit** — Probit model with sample selection

[R] **oprobit** — Ordered probit regression

[R] **probit** — Probit regression

[R] **regress** — Linear regression

[R] **tobit** — Tobit regression

[BAYES] **bayes: heckoprobit** — Bayesian ordered probit model with sample selection

[ERM] **eoprobit** — Extended ordered probit regression

[SVY] **svy estimation** — Estimation commands for survey data

[U] 20 Estimation and postestimation commands