### cloglog — Complementary log-log regression

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## Description

cloglog fits a complementary log-log model for a binary dependent variable, typically with one of the outcomes rare relative to the other. It can also be used to fit a gompit model. cloglog can compute robust and cluster-robust standard errors and adjust results for complex survey designs.

# **Quick start**

Complementary log-log model of y on x1 and x2 cloglog y x1 x2

With robust standard errors

cloglog y x1 x2, vce(robust)

Adjust for complex survey design using svyset data

svy: cloglog y x1 x2

### Menu

Statistics > Binary outcomes > Complementary log-log regression

## Syntax

cloglog depvar [indepvars] [if] [in] [weight] [, options]

options	Description
Model	
<u>nocons</u> tant <u>off</u> set( <i>varname</i> ) asis <u>const</u> raints( <i>constraints</i> )	suppress constant term include <i>varname</i> in model with coefficient constrained to 1 retain perfect predictor variables apply specified linear constraints
SE/Robust	
vce( <i>vcetype</i> )	<pre>vcetype may be oim, robust, cluster clustvar, opg, bootstrap,</pre>
Reporting	
<u>l</u> evel(#)	set confidence level; default is level(95)
<u>ef</u> orm	report exponentiated coefficients
<u>nocnsr</u> eport	do not display constraints
display_options	control columns and column formats, row spacing, line width, display of omitted variables and base and empty cells, and factor-variable labeling
Maximization	
maximize_options	control the maximization process; seldom used
<u>col</u> linear <u>coefl</u> egend	keep collinear variables display legend instead of statistics

indepvars may contain factor variables; see [U] 11.4.3 Factor variables.

depvar and indepvars may contain time-series operators; see [U] 11.4.4 Time-series varlists.

bayes, bayesboot, bootstrap, by, collect, fmm, fp, jackknife, mi estimate, nestreg, rolling, statsby, stepwise, and svy are allowed; see [U] 11.1.10 Prefix commands. For more details, see [BAYES] bayes: cloglog and [FMM] fmm: cloglog.

vce(bootstrap) and vce(jackknife) are not allowed with the mi estimate prefix; see [MI] mi estimate.

Weights are not allowed with the bootstrap prefix; see [R] bootstrap.

vce() and weights are not allowed with the svy prefix; see [SVY] svy.

fweights, iweights, and pweights are allowed; see [U] 11.1.6 weight.

collinear and coeflegend do not appear in the dialog box.

See [U] 20 Estimation and postestimation commands for more capabilities of estimation commands.

## Options

Model

noconstant, offset(varname); see [R] Estimation options.

asis forces retention of perfect predictor variables and their associated perfectly predicted observations and may produce instabilities in maximization; see [R] probit.

constraints (*constraints*); see [R] Estimation options.

```
      SE/Robust

      vce(vcetype) specifies the type of standard error reported, which includes types that are derived from asymptotic theory (oim, opg), that are robust to some kinds of misspecification (robust), that allow for intragroup correlation (cluster clustvar), and that use bootstrap or jackknife methods (bootstrap, jackknife); see [R] vce_option.

      Reporting

      level(#); see [R] Estimation options.

      eform displays the exponentiated coefficients and corresponding standard errors and confidence intervals.

      nocnsreport; see [R] Estimation options.

      display_options: noci, nopvalues, noomitted, vsquish, noemptycells, baselevels, allbaselevels, nofvlabel, fvwrap(#), fvwrapon(style), cformat(%fmt), pformat(%fmt), sformat(%fmt), and nolstretch; see [R] Estimation options.
```

Maximization

maximize\_options: difficult, technique(algorithm\_spec), iterate(#), [no]log, trace, gradient, showstep, hessian, showtolerance, tolerance(#), ltolerance(#), nrtolerance(#), nonrtolerance, and from(init\_specs); see [R] Maximize. These options are seldom used.

Setting the optimization type to technique(bhhh) resets the default vcetype to vce(opg).

The following options are available with cloglog but are not shown in the dialog box:

collinear, coeflegend; see [R] Estimation options.

### **Remarks and examples**

Remarks are presented under the following headings:

Introduction to complementary log-log regression Robust standard errors

#### Introduction to complementary log-log regression

cloglog fits maximum likelihood models with dichotomous dependent variables coded as 0/1 (or, more precisely, coded as 0 and not 0).

```
Example 1
```

We have data on the make, weight, and mileage rating of 22 foreign and 52 domestic automobiles. We wish to fit a model explaining whether a car is foreign based on its weight and mileage. Here is an overview of our data:

```
. use https://www.stata-press.com/data/r19/auto
(1978 automobile data)
. keep make mpg weight foreign
. describe
Contains data from https://www.stata-press.com/data/r19/auto.dta
 Observations:
                           74
                                                 1978 automobile data
    Variables:
                            4
                                                 13 Apr 2024 17:45
                                                 ( dta has notes)
                                     Value
Variable
               Storage
                         Display
    name
                  type
                          format
                                     label
                                                 Variable label
make
                 str18
                         %-18s
                                                 Make and model
                         %8.0g
                                                 Mileage (mpg)
mpg
                 int
weight
                 int
                         %8.0gc
                                                 Weight (lbs.)
foreign
                 byte
                         %8.0g
                                     origin
                                                 Car origin
Sorted by: foreign
     Note: Dataset has changed since last saved.
. inspect foreign
foreign: Car origin
                                                   Number of observations
                                                Total
                                                           Integers
                                                                       Nonintegers
   #
                               Negative
   #
                               Zero
                                                   52
                                                                  52
   #
                                                                  22
                               Positive
                                                   22
   #
   #
       #
                               Total
                                                   74
                                                                  74
   #
       #
                               Missing
                                                    _
                                                   74
'n
                       1
   (2 unique values)
```

foreign is labeled and all values are documented in the label.

The variable foreign takes on two unique values, 0 and 1. The value 0 denotes a domestic car, and 1 denotes a foreign car.

The model that we wish to fit is

$$\Pr(\texttt{foreign} = 1) = F(\beta_0 + \beta_1 \texttt{weight} + \beta_2 \texttt{mpg})$$

where  $F(z) = 1 - \exp\{-\exp(z)\}.$ 

To fit this model, we type

. cloglog fore Iteration 0: Iteration 1: Iteration 2: Iteration 3:	eign weight mp Log likelihoo Log likelihoo Log likelihoo Log likelihoo	g d = -34.054 d = -27.869 d = -27.742 d = -27.742	593 915 997 769				
Iteration 4:	Log likelihoo	d = -27.7427	769				
Complementary log-log regression				Number Zero ou Nonzero	of obs tcomes outcomes	= = 5 =	74 52 22
Log likelihood	a = −27.742769			LR chi2 Prob >	(2) chi2	=	34.58 0.0000
foreign	Coefficient	Std. err.	Z	P> z	[95% 0	conf.	interval]
weight mpg _cons	0029153 1422911 10.09694	.0006974 .076387 3.351841	-4.18 -1.86 3.01	0.000 0.062 0.003	00428 29200 3.5274	323 069 148	0015483 .0074247 16.66642

We find that heavier cars are less likely to be foreign and that cars yielding better gas mileage are also less likely to be foreign, at least when holding the weight of the car constant.

See [R] Maximize for an explanation of the output.

#### Technical note

Stata interprets a value of 0 as a negative outcome (failure) and treats all other values (except missing) as positive outcomes (successes). Thus, if your dependent variable takes on the values 0 and 1, 0 is interpreted as failure and 1 as success. If your dependent variable takes on the values 0, 1, and 2, 0 is still interpreted as failure, but both 1 and 2 are treated as successes.

If you prefer a more formal mathematical statement, when you type cloglog y x, Stata fits the model

$$\Pr(y_j \neq 0 \mid \mathbf{x}_j) = 1 - \exp\left\{-\exp(\mathbf{x}_j \boldsymbol{\beta})\right\}$$

**Robust standard errors** 

If you specify the vce(robust) option, cloglog reports robust standard errors, as described in [U] **20.22 Obtaining robust variance estimates**. For the model of foreign on weight and mpg, the robust calculation increases the standard error of the coefficient on mpg by 44%:

4

. cloglog fo	oreign	weight m	pg, vce(r	obust)	)				
Iteration 0 Iteration 1 Iteration 2 Iteration 3 Iteration 4	: Log : Log : Log : Log : Log	pseudoli pseudoli pseudoli pseudoli pseudoli	kelihood kelihood kelihood kelihood kelihood	= -34 = -27 = -27 = -27 = -27	.054593 .869915 .742997 .742769 .742769				
Complementar	ry log-	-log regr	ession			Number	of obs	=	74
						Zero c	outcomes	=	52
						Nonzer	o outcome	es =	22
						Wald d	hi2(2)	=	29.74
Log pseudol:	ikeliho	ood = -27	.742769			Prob >	chi2	=	0.0000
foreig	n Coe	efficient	Robust std. er	r.	z	P> z	[95%	conf.	interval]
weight mpg _cons	t g s 1	0029153 1422911 .0.09694	.000748 .110246 4.31730	4 - 6 - 5	-3.90 -1.29 2.34	0.000 0.197 0.019	0043 3583 1.635	3822 3704 5174	0014484 .0737882 18.5587

Without vce(robust), the standard error for the coefficient on mpg was reported to be 0.076, with a resulting confidence interval of [-0.29, 0.01].

The vce(cluster *clustvar*) option can relax the independence assumption required by the complementary log-log estimator to being just independence between clusters. To demonstrate this ability, we will switch to a different dataset.

We are studying unionization of women in the United States by using the union dataset; see [XT] **xt**. We fit the following model, ignoring that women are observed an average of 5.9 times each in this dataset:

```
. use https://www.stata-press.com/data/r19/union, clear
(NLS Women 14-24 in 1968)
. cloglog union age grade not_smsa south##c.year
Iteration 0: Log likelihood = -13606.373
Iteration 1: Log likelihood = -13540.726
Iteration 2: Log likelihood = -13540.607
Iteration 3: Log likelihood = -13540.607
Complementary log-log regression
                                                Number of obs
                                                                        26,200
                                                                  =
                                                Zero outcomes
                                                                        20,389
                                                                  =
                                                Nonzero outcomes =
                                                                         5,811
                                                                  =
                                                                        647.24
                                                LR chi2(6)
Log likelihood = -13540.607
                                                Prob > chi2
                                                                        0.0000
                                                                  =
      union
               Coefficient Std. err.
                                                P>|z|
                                                          [95% conf. interval]
                                           7.
                                         4.25
                                                0.000
                                                           .009986
                                                                       .0270833
                 .0185346
                            .0043616
         age
                                         7.93
                                                0.000
                                                          .0340809
      grade
                 .0452772
                            .0057125
                                                                      .0564736
                -.1886592
                            .0317801
                                        -5.94
                                                0.000
                                                         -.2509471
                                                                     -.1263712
   not smsa
    1.south
               -1.422292
                            .3949381
                                        -3.60
                                                0.000
                                                         -2.196356
                                                                      -.648227
       year
                -.0133007
                            .0049576
                                        -2.68
                                                0.007
                                                         -.0230174
                                                                     -.0035839
south#c.year
          1
                 .0105659
                            .0049234
                                         2.15
                                                0.032
                                                           .0009161
                                                                       .0202157
       _cons
                -1.219801
                            .2952374
                                        -4.13
                                                0.000
                                                         -1.798455
                                                                     -.6411462
```

The reported standard errors in this model are probably meaningless. Women are observed repeatedly, and so the observations are not independent. Looking at the coefficients, we find a large southern effect against unionization and a different time trend for the south. The vce(cluster *clustvar*) option provides a way to fit this model and obtains correct standard errors:

. cloglog unio	on age grade n	ot_smsa sout	ch##c.yea	ar, vce(c	luster id) n	olog
Complementary	log-log regre	ssion		Number Zero ou Nonzero	of obs = tcomes = outcomes =	26,200 20,389 5,811
Log pseudolike	elihood = -135	40.607		Wald ch Prob >	i2(6) = chi2 =	160.76 0.0000
		(Std. err	. adjuste	ed for 4,	434 clusters	in idcode)
union	Coefficient	Robust std. err.	Z	P> z	[95% conf	. interval]
age grade not_smsa 1.south year	.0185346 .0452772 1886592 -1.422292 0133007	.0084873 .0125776 .0642068 .506517 .0090628	2.18 3.60 -2.94 -2.81 -1.47	0.029 0.000 0.003 0.005 0.142	.0018999 .0206255 3145021 -2.415047 0310633	.0351694 .069929 0628162 4295365 .004462
south#c.year 1 _cons	.0105659 -1.219801	.0063175 .5175129	1.67 -2.36	0.094 0.018	0018162 -2.234107	. 022948 2054942

These standard errors are larger than those reported by the inappropriate conventional calculation. By comparison, another way we could fit this model is with an equal-correlation population-averaged complementary log–log model:

. xtcloglog union age grade not_smsa south##c.year, pa nolog						
GEE population Group variable	el		N N	umber of obs umber of group:	= 26,200 s = 4,434	
Family: Binomi	ial			0	bs per group:	
Link: Comple	ementary log-l	og			mi	n = 1
Correlation: e	exchangeable				av	g = 5.9
					ma	x = 12
				W	ald chi2(6)	= 234.66
Scale paramete	er = 1			Р	rob > chi2	= 0.0000
union	Coefficient	Std. err.	z	P> z	[95% conf.	interval]
age	.0153737	.0081156	1.89	0.058	0005326	.03128
grade	.0549518	.0095093	5.78	0.000	.0363139	.0735897
not_smsa	1045232	.0431082	-2.42	0.015	1890138	0200326
1.south	-1.714868	.3384558	-5.07	0.000	-2.378229	-1.051507
year	0115881	.0084125	-1.38	0.168	0280763	.0049001
south#c.year						
1	.0149796	.0041687	3.59	0.000	.0068091	.0231501
_cons	-1.488278	.4468005	-3.33	0.001	-2.363991	6125652

The coefficient estimates are similar, but these standard errors are smaller than those produced by cloglog, vce(cluster *clustvar*). This finding is as we would expect. If the within-panel correlation assumptions are valid, the population-averaged estimator should be more efficient.

In addition to this estimator, we may use the xtgee command to fit a panel estimator (with complementary log-log link) and any number of assumptions on the within-idcode correlation.

cloglog, vce(cluster *clustvar*) is robust to assumptions about within-cluster correlation. That is, it inefficiently sums within cluster for the standard error calculation rather than attempting to exploit what might be assumed about the within-cluster correlation (as do the xtgee population-averaged models).

## **Stored results**

cloglog stores the following in e():

Sc	alars	
	e(N)	number of observations
	e(k)	number of parameters
	e(k_eq)	number of equations in e(b)
	e(k_eq_model)	number of equations in overall model test
	e(k_dv)	number of dependent variables
	e(N_f)	number of zero outcomes
	e(N_s)	number of nonzero outcomes
	e(df_m)	model degrees of freedom
	e(11)	log likelihood
	e(11_0)	log likelihood, constant-only model
	e(N_clust)	number of clusters
	e(chi2)	$\chi^2$
	e(p)	<i>p</i> -value for model test
	e(rank)	rank of e(V)
	e(ic)	number of iterations
	e(rc)	return code
	e(converged)	1 if converged, 0 otherwise
M	acros	
	e(cmd)	cloglog
	e(cmdline)	command as typed
	e(depvar)	name of dependent variable
	e(wtype)	weight type
	e(wexp)	weight expression
	e(title)	title in estimation output
	e(clustvar)	name of cluster variable
	e(offset)	linear offset variable
	e(chi2type)	Wald or LR; type of model $\chi^2$ test
	e(vce)	vcetype specified in vce()
	e(vcetype)	title used to label Std. err.
	e(opt)	type of optimization
	e(which)	max or min; whether optimizer is to perform maximization or minimization
	e(ml_method)	type of ml method
	e(user)	name of likelihood-evaluator program
	e(technique)	maximization technique
	e(properties)	b V
	e(predict)	program used to implement predict
	e(marginsok)	predictions allowed by margins
	e(marginsnotok)	predictions disallowed by margins
	e(asbalanced)	factor variables fvset as asbalanced
	e(asobserved)	factor variables fvset as asobserved

Matrices	
e(b)	coefficient vector
e(Cns)	constraints matrix
e(ilog)	iteration log (up to 20 iterations)
e(gradient)	gradient vector
e(V)	variance-covariance matrix of the estimators
e(V_modelbased)	model-based variance
Functions	
e(sample)	marks estimation sample

In addition to the above, the following is stored in r():

```
Matrices
```

r(table)

matrix containing the coefficients with their standard errors, test statistics, *p*-values, and confidence intervals

Note that results stored in r() are updated when the command is replayed and will be replaced when any r-class command is run after the estimation command.

### Methods and formulas

Complementary log-log analysis (related to the gompit model, so named because of its relationship to the Gompertz distribution) is an alternative to logit and probit analysis, but it is unlike these other estimators in that the transformation is not symmetric. Typically, this model is used when the positive (or negative) outcome is rare.

The log-likelihood function for complementary log-log is

$$\mathrm{ln} L = \sum_{j \in S} w_j \, \mathrm{ln} F(\mathbf{x}_j \mathbf{b}) + \sum_{j \notin S} w_j \, \mathrm{ln} \Big\{ 1 - F(\mathbf{x}_j \mathbf{b}) \Big\}$$

where S is the set of all observations j such that  $y_j \neq 0$ ,  $F(z) = 1 - \exp\{-\exp(z)\}$ , and  $w_j$  denotes the optional weights.  $\ln L$  is maximized as described in [R] Maximize.

We can fit a gompit model by reversing the success-failure sense of the dependent variable and using cloglog.

This command supports the Huber/White/sandwich estimator of the variance and its clustered version using vce(robust) and vce(cluster *clustvar*), respectively. See [P] **\_robust**, particularly *Maximum likelihood estimators* and *Methods and formulas*. The scores are calculated as  $\mathbf{u}_j = [\exp(\mathbf{x}_j \mathbf{b}) \exp\{-\exp(\mathbf{x}_j \mathbf{b})]\mathbf{x}_j$  for the positive outcomes and  $\{-\exp(\mathbf{x}_j \mathbf{b})\}\mathbf{x}_j$  for the negative outcomes.

cloglog also supports estimation with survey data. For details on VCEs with survey data, see [SVY] Variance estimation.

### Acknowledgment

We thank Joseph Hilbe (1944–2017), coauthor of the Stata Press book Generalized Linear Models and Extensions, for providing the inspiration for the cloglog command.

### References

Long, J. S. 1997. Regression Models for Categorical and Limited Dependent Variables. Thousand Oaks, CA: Sage.

Long, J. S., and J. Freese. 2014. Regression Models for Categorical Dependent Variables Using Stata. 3rd ed. College Station, TX: Stata Press.

Xu, J., and J. S. Long. 2005. Confidence intervals for predicted outcomes in regression models for categorical outcomes. Stata Journal 5: 537–559.

### Also see

- [R] cloglog postestimation Postestimation tools for cloglog
- [R] clogit Conditional (fixed-effects) logistic regression
- [R] glm Generalized linear models
- [R] logistic Logistic regression, reporting odds ratios
- [R] scobit Skewed logistic regression
- [BAYES] bayes: cloglog Bayesian complementary log-log regression
- [FMM] fmm: cloglog Finite mixtures of complementary log-log regression models
- [ME] mecloglog Multilevel mixed-effects complementary log-log regression
- [MI] Estimation Estimation commands for use with mi estimate
- [SVY] svy estimation Estimation commands for survey data
- [XT] xtcloglog Random-effects and population-averaged cloglog models
- [U] 20 Estimation and postestimation commands

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