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## Description

`mi impute poisson` fills in missing values of a count variable using a Poisson regression imputation method. You can perform separate imputations on different subsets of the data by specifying the `by()` option. You can also account for frequency, importance, and sampling weights.

## Menu

Statistics > Multiple imputation

## Syntax

`mi impute poisson ivar [indepvars] [if] [weight] [, impute_options options]`

*impute\_options*

Description

Main	
* <code>add(#)</code>	specify number of imputations to add; required when no imputations exist
* <code>replace</code>	replace imputed values in existing imputations
<code>rseed(#)</code>	specify random-number seed
<code>double</code>	store imputed values in double precision; the default is to store them as float
<code>by(<i>varlist</i> [, <i>byopts</i>])</code>	impute separately on each group formed by <i>varlist</i>
Reporting	
<code>dots</code>	display dots as imputations are performed
<code>noisily</code>	display intermediate output
<code>nolegend</code>	suppress all table legends
Advanced	
<code>force</code>	proceed with imputation, even when missing imputed values are encountered
<code>noupdate</code>	do not perform mi update; see <a href="#">[MI] noupdate option</a>

\*`add(#)` is required when no imputations exist; `add(#)` or `replace` is required if imputations exist.  
`noupdate` does not appear in the dialog box.

<i>options</i>	Description
<b>Main</b>	
<code>noconstant</code>	suppress constant term
<code>exposure(<i>varname<sub>e</sub></i>)</code>	include $\ln(\text{varname}_e)$ in model with coefficient constrained to 1
<code>offset(<i>varname<sub>o</sub></i>)</code>	include $\text{varname}_o$ in model with coefficient constrained to 1
<code>conditional(<i>if</i>)</code>	perform conditional imputation
<code>bootstrap</code>	estimate model parameters using sampling with replacement
<b>Maximization</b>	
<code>maximize_options</code>	control the maximization process; seldom used
<p>You must <code>mi set</code> your data before using <code>mi impute poisson</code>; see <a href="#">[MI] mi set</a>.</p> <p>You must <code>mi register ivar</code> as imputed before using <code>mi impute poisson</code>; see <a href="#">[MI] mi set</a>.</p> <p><code>indepvars</code> may contain factor variables; see <a href="#">[U] 11.4.3 Factor variables</a>.</p> <p><code>collect</code> is allowed; see <a href="#">[U] 11.1.10 Prefix commands</a>.</p> <p><code>fweights</code>, <code>iweights</code>, and <code>pweights</code> are allowed; see <a href="#">[U] 11.1.6 weight</a>.</p>	

## Options

### Main

`noconstant`; see [\[R\] Estimation options](#).

`add()`, `replace`, `rseed()`, `double`, `by()`; see [\[MI\] mi impute](#).

`exposure(varnamee)`, `offset(varnameo)`; see [\[R\] Estimation options](#).

`conditional(if)` specifies that the imputation variable be imputed conditionally on observations satisfying *exp*; see [\[U\] 11.1.3 if exp](#). That is, missing values in a conditional sample, the sample identified by the *exp* expression, are imputed based only on data in that conditional sample. Missing values outside the conditional sample are replaced with a conditional constant, the value of the imputation variable in observations outside the conditional sample. As such, the imputation variable is required to be constant outside the conditional sample. Also, if any conditioning variables (variables involved in the conditional specification *if exp*) contain soft missing values (`.`), their missing values must be nested within missing values of the imputation variables. See [Conditional imputation](#) under *Remarks and examples* in [\[MI\] mi impute](#).

`bootstrap` specifies that posterior estimates of model parameters be obtained using sampling with replacement; that is, posterior estimates are estimated from a bootstrap sample. The default is to sample the estimates from the posterior distribution of model parameters or from the large-sample normal approximation of the posterior distribution. This option is useful when asymptotic normality of parameter estimates is suspect.

### Reporting

`dots`, `noisily`, `nolegend`; see [\[MI\] mi impute](#). `noisily` specifies that the output from the Poisson regression fit to the observed data be displayed. `nolegend` suppresses all legends that appear before the imputation table. Such legends include a legend about conditional imputation that appears when the `conditional()` option is specified and group legends that may appear when the `by()` option is specified.

## Maximization

*maximize\_options*; see [R] [poisson](#). These options are seldom used.

## Advanced

force; see [MI] [mi impute](#).

The following option is available with `mi impute` but is not shown in the dialog box:

noupdate; see [MI] [noupdate option](#).

## Remarks and examples

Remarks are presented under the following headings:

*Univariate imputation using Poisson regression*  
*Using mi impute poisson*

See [MI] [mi impute](#) for a general description and details about options common to all imputation methods, *impute\_options*. Also see [MI] [Workflow](#) for general advice on working with `mi`.

## Univariate imputation using Poisson regression

The Poisson regression imputation method can be used to fill in missing values of a count variable (for example, [Raghuathan et al. \[2001\]](#) and [van Buuren \[2007\]](#)). It is a parametric method that assumes an underlying Poisson model for the imputed variable (given other predictors). For imputation of overdispersed count variables, see [MI] [mi impute nbreg](#). The Poisson method is based on the asymptotic approximation of the posterior predictive distribution of the missing data.

## Using mi impute poisson

To illustrate the use of `mi impute poisson`, we continue with our heart attack data analysis [example](#) in [MI] [Intro substantive](#) and consider an additional predictor, `npreg`, which records the number of pregnancies:

```
. use https://www.stata-press.com/data/r19/mheartpois
(Fictional heart attack data; npreg missing)
. misstable summarize
```

Variable	Obs=.	Obs>.	Obs<.	Obs<.		
				Unique values	Min	Max
npreg	10		144	6	0	5

```
. tab female if npreg==.
```

Gender	Freq.	Percent	Cum.
Male	7	70.00	70.00
Female	3	30.00	100.00
Total	10	100.00	

According to `misstable summarize`, `npreg` is the only variable containing missing values, and it has 10 out of 154 observations missing. The tabulation of missing values of `npreg` by gender reveals that most missing values (7) correspond to males.

In this example, we could replace missing `npreg` for males with 0 and proceed with complete-data analysis, disregarding the remaining three missing observations. Instead, as an illustration, we use `mi impute poisson` to impute missing values of `npreg`. Our dataset is not declared yet, so we use `mi set` to declare it. We also use `mi register` to register `npreg` as the imputed variable before using `mi impute poisson`:

```
. mi set mlong
. mi register imputed npreg
(10 m=0 obs now marked as incomplete)

. mi impute poisson npreg attack smokes age bmi hsgrad, add(20)
> conditional(if female==1)

Univariate imputation                      Imputations =      20
Poisson regression                          added =      20
Imputed: m=1 through m=20                  updated =       0

Conditional imputation:
  npreg: incomplete out-of-sample obs replaced with value 0
```

Variable	Observations per <i>m</i>			
	Complete	Incomplete	Imputed	Total
npreg	144	10	10	154

(Complete + Incomplete = Total; Imputed is the minimum across *m* of the number of filled-in observations.)

The `npreg` variable is relevant to females only, so we used the `conditional()` option to restrict imputation to observations with `female==1`; see [Conditional imputation](#) in [\[MI\] mi impute](#).

We can analyze these multiply imputed data using logistic regression with `mi estimate`:

```
. mi estimate: logit attack smokes age bmi female hsgrad npreg
```

## Stored results

`mi impute poisson` stores the following in `r()`:

Scalars

- `r(M)` total number of imputations
- `r(M_add)` number of added imputations
- `r(M_update)` number of updated imputations
- `r(k_ivars)` number of imputed variables (always 1)
- `r(N_g)` number of imputed groups (1 if `by()` is not specified)

Macros

- `r(method)` name of imputation method (`poisson`)
- `r(ivars)` names of imputation variables
- `r(rngstate)` random-number state used
- `r(by)` names of variables specified within `by()`

Matrices

- `r(N)` number of observations in imputation sample in each group
- `r(N_complete)` number of complete observations in imputation sample in each group
- `r(N_incomplete)` number of incomplete observations in imputation sample in each group
- `r(N_imputed)` number of imputed observations in imputation sample in each group

## Methods and formulas

Consider a univariate variable  $\mathbf{x} = (x_1, x_2, \dots, x_n)'$  that follows a Poisson model

$$\Pr(x_i = x | \mathbf{z}_i) = \frac{e^{-\lambda_i} \lambda_i^x}{x!}, \quad x = 0, 1, 2, \dots \quad (1)$$

where  $\lambda_i = \exp(\mathbf{z}_i' \boldsymbol{\beta} + \text{offset}_i)$  (see [R] **poisson**),  $\mathbf{z}_i = (z_{i1}, z_{i2}, \dots, z_{iq})'$  records values of predictors of  $\mathbf{x}$  for observation  $i$  and  $\boldsymbol{\beta}$  is the  $q \times 1$  vector of unknown regression coefficients. (When a constant is included in the model—the default— $z_{i1} = 1$ ,  $i = 1, \dots, n$ .)

$\mathbf{x}$  contains missing values that are to be filled in. Consider the partition of  $\mathbf{x} = (\mathbf{x}'_o, \mathbf{x}'_m)$  into  $n_o \times 1$  and  $n_1 \times 1$  vectors containing the complete and the incomplete observations. Consider a similar partition of  $\mathbf{Z} = (\mathbf{Z}_o, \mathbf{Z}_m)$  into  $n_o \times q$  and  $n_1 \times q$  submatrices.

`mi impute poisson` follows the steps below to fill in  $\mathbf{x}_m$ :

1. Fit a Poisson regression model (1) to the observed data  $(\mathbf{x}_o, \mathbf{Z}_o)$  to obtain the maximum likelihood estimates,  $\widehat{\boldsymbol{\beta}}$ , and their asymptotic sampling variance,  $\widehat{\mathbf{U}}$ .
2. Simulate new parameters,  $\boldsymbol{\beta}_*$ , from the large-sample normal approximation,  $N(\widehat{\boldsymbol{\beta}}, \widehat{\mathbf{U}})$ , to its posterior distribution assuming the noninformative prior  $\Pr(\boldsymbol{\beta}) \propto \text{const.}$
3. Obtain one set of imputed values,  $\mathbf{x}_m^1$ , by simulating from a Poisson distribution (1) with  $\lambda_i = \exp(\mathbf{z}_{i_m}' \boldsymbol{\beta}_* + \text{offset}_{i_m})$ .
4. Repeat steps 2 and 3 to obtain  $M$  sets of imputed values  $\mathbf{x}_m^1, \mathbf{x}_m^2, \dots, \mathbf{x}_m^M$ .

Steps 2 and 3 above correspond to only approximate draws from the posterior predictive distribution of the missing data,  $\Pr(\mathbf{x}_m | \mathbf{x}_o, \mathbf{Z}_o)$ , because  $\boldsymbol{\beta}_*$  is drawn from the asymptotic approximation to its posterior distribution.

If weights are specified, a weighted Poisson regression model is fit to the observed data in step 1 (see [R] **poisson** for details).

## References

- Ragunathan, T. E., J. M. Lepkowski, J. Van Hoewyck, and P. Solenberger. 2001. A multivariate technique for multiply imputing missing values using a sequence of regression models. *Survey Methodology* 27: 85–95.
- van Buuren, S. 2007. Multiple imputation of discrete and continuous data by fully conditional specification. *Statistical Methods in Medical Research* 16: 219–242. <https://doi.org/10.1177/0962280206074463>.

## Also see

- [MI] **mi impute** — Impute missing values
- [MI] **mi impute nbreg** — Impute using negative binomial regression
- [MI] **mi estimate** — Estimation using multiple imputations
- [MI] **Intro** — Introduction to mi

[MI] **Intro substantive** — Introduction to multiple-imputation analysis

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