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Description

`irt 3pl` fits three-parameter logistic (3PL) models to binary items. In the 3PL model, items vary in their difficulty and discrimination and the possibility of guessing is allowed.

Quick start

3PL model for binary items `b1` to `b10`

```
irt 3pl b1-b10
```

Group estimates by parameter type and sort items by difficulty

```
estat report, byparm sort(b)
```

Plot ICCs for all items

```
irtgraph icc
```

Menu

Statistics > IRT (item response theory)

Syntax

```
irt 3pl varlist [if] [in] [weight] [, options]
```

<i>options</i>	Description
<code>group(<i>varname</i>)</code>	fit model for different groups
Model	
<code>cns(<i>spec</i>)</code>	apply specified parameter constraints
<code>listwise</code>	drop observations with any missing items
<code>sepguessing</code>	estimate a separate pseudoguessing parameter for each item
<code>gsepguessing</code>	estimate separate pseudoguessing parameters for each group
SE/Robust	
<code>vce(<i>vcetype</i>)</code>	<i>vcetype</i> may be <code>oim</code> , <code>robust</code> , <code>cluster <i>clustvar</i></code> , <code>bootstrap</code> , or <code>jackknife</code>
Reporting	
<code>level(#)</code>	set confidence level; default is <code>level(95)</code>
<code>notable</code>	suppress coefficient table
<code>noheader</code>	suppress output header
<code>display_options</code>	control columns and column formats
Integration	
<code>intmethod(<i>intmethod</i>)</code>	integration method
<code>intpoints(#)</code>	set the number of integration points; default is <code>intpoints(7)</code>
Maximization	
<code>maximize_options</code>	control the maximization process; seldom used
<code>startvalues(<i>svmethod</i>)</code>	method for obtaining starting values
<code>noestimate</code>	do not fit the model; show starting values instead
<code>estmetric</code>	show parameter estimates in the estimation metric
<code>dnumerical</code>	use numerical derivative techniques
<code>coeflegend</code>	display legend instead of statistics
<i>intmethod</i>	Description
<code>mvaghermite</code>	mean–variance adaptive Gauss–Hermite quadrature; the default
<code>mcaghermite</code>	mode-curvature adaptive Gauss–Hermite quadrature
<code>ghermite</code>	nonadaptive Gauss–Hermite quadrature

`bootstrap`, `by`, `collect`, `jackknife`, `statsby`, and `svy` are allowed; see [U] 11.1.10 Prefix commands.

Weights are not allowed with the `bootstrap` prefix; see [R] `bootstrap`.

`vce()` and weights are not allowed with the `svy` prefix; see [SVY] `svy`.

`fweights`, `iweights`, and `pweights` are allowed; see [U] 11.1.6 `weight`.

`startvalues()`, `noestimate`, `estmetric`, `dnumerical`, and `coeflegend` do not appear in the dialog box.

See [U] 20 Estimation and postestimation commands for more capabilities of estimation commands.

Options

`group(varname)` specifies that the model be fit separately for the different values of *varname*; see [IRT] [irt, group\(\)](#) for details.

Model

`cns(spec)` constrains item parameters to a fixed value or constrains two or more parameters to be equal; see [IRT] [irt constraints](#) for details.

`listwise` handles missing values through listwise deletion, which means that the entire observation is omitted from the estimation sample if any of the items are missing for that observation. By default, all nonmissing items in an observation are included in the likelihood calculation; only missing items are excluded.

`sepguessing` specifies that a separate pseudoguessing parameter be estimated for each item. This is a seldom used option; see the [technical note](#) below.

`gsepguessing` specifies that separate pseudoguessing parameters be estimated for each group. This option is allowed only when fitting a group model.

SE/Robust

`vce(vcetype)` specifies the type of standard error reported, which includes types that are derived from asymptotic theory (`oim`), that are robust to some kinds of misspecification (`robust`), that allow for intragroup correlation (`cluster clustvar`), and that use bootstrap or jackknife methods (`bootstrap`, `jackknife`); see [R] [vce_option](#).

Reporting

`level(#)`; see [R] [Estimation options](#).

`notable` suppresses the estimation table, either at estimation or upon replay.

`noheader` suppresses the output header, either at estimation or upon replay.

`display_options:` `noci`, `nopvalues`, `cformat(%fmt)`, `pformat(%fmt)`, `sformat(%fmt)`, and `nolstretch`; see [R] [Estimation options](#).

Integration

`intmethod(intmethod)` specifies the integration method to be used for computing the log likelihood. `mvaghermite` performs mean and variance adaptive Gauss–Hermite quadrature; `mcaghermite` performs mode and curvature adaptive Gauss–Hermite quadrature; and `ghermite` performs nonadaptive Gauss–Hermite quadrature.

The default integration method is `mvaghermite`.

`intpoints(#)` sets the number of integration points for quadrature. The default is `intpoints(7)`, which means that seven quadrature points are used to compute the log likelihood.

The more integration points, the more accurate the approximation to the log likelihood. However, computation time increases with the number of integration points.

Maximization

maximize_options: difficult, technique(*algorithm_spec*), iterate(#), [no]log, trace, gradient, showstep, hessian, showtolerance, tolerance(#), ltolerance(#), nrtolerance(#), nonrtolerance, and from(*init_specs*); see [R] **Maximize**. Those that require special mention for *irt* are listed below.

from() accepts a properly labeled vector of initial values or a list of coefficient names with values. A list of values is not allowed.

The following options are available with *irt* but are not shown in the dialog box:

startvalues() specifies how starting values are to be computed. Starting values specified in *from*() override the computed starting values.

startvalues(zero) specifies that all starting values be set to 0. This option is typically useful only when specified with the *from*() option.

startvalues(constantonly) builds on *startvalues*(zero) by fitting a constant-only model for each response to obtain estimates of intercept and cutpoint parameters.

startvalues(fixedonly) builds on *startvalues*(constantonly) by fitting a full fixed-effects model for each response variable to obtain estimates of coefficients along with intercept and cutpoint parameters. You can also add suboption *iterate*(#) to limit the number of iterations *irt* allows for fitting the fixed-effects model.

startvalues(ivloadings) builds on *startvalues*(fixedonly) by using instrumental-variable methods with the generalized residuals from the fixed-effects models to compute starting values for latent-variable loadings. This is the default behavior.

noestimate specifies that the model is not to be fit. Instead, starting values are to be shown (as modified by the above options if modifications were made), and they are to be shown using the *coeflegend* style of output. An important use of this option is before you have modified starting values at all; you can type the following:

```
. irt ..., ... noestimate
. matrix b = e(b)
. ... (modify elements of b) ...
. irt ..., ... from(b)
```

estmetric displays parameter estimates in the slope-intercept metric that is used for estimation.

dnnumerical specifies that during optimization, the gradient vector and Hessian matrix be computed using numerical techniques instead of analytical formulas. By default, *irt* uses analytical formulas for computing the gradient and Hessian for all integration methods.

coeflegend; see [R] **Estimation options**.

Remarks and examples

Remarks are presented under the following headings:

[Overview](#)

[Video example](#)

Overview

The following discussion is about how to use `irt` to fit (3PL) models to binary items. If you are new to the IRT features in Stata, we encourage you to read [\[IRT\] irt](#) first.

In the 3PL model, item responses are typically of the form yes or no, correct or incorrect, agree or disagree, etc. Items are assumed to vary in discrimination and difficulty, and the model accommodates the possibility of guessing on a test. The probability of person j providing a positive answer to item i is given by

$$\Pr(y_{ij} = 1|\theta_j) = c_i + (1 - c_i) \frac{\exp\{a_i(\theta_j - b_i)\}}{1 + \exp\{a_i(\theta_j - b_i)\}} \quad \theta_j \sim N(0, 1) \quad (1)$$

where a_i represents the discrimination of item i , b_i represents the difficulty of item i , c_i represents the pseudoguessing parameter, and θ_j is the latent trait of person j . By default, the c_i are constrained to be the same across all items; see the [technical note](#) below.

Although (1) is not in logistic form, the model is commonly referred to as a three-parameter logistic model.

The 3PL model was proposed by [Birnbaum \(1968\)](#). An earlier three-parameter model with a probit link was developed by [Finney \(1952\)](#).

□ Technical note

By default, `irt 3pl` constrains the pseudoguessing parameter to be the same across all items. You can use the advanced option `sepguessing` to request a separate pseudoguessing parameter for each item. We do not recommend this option because this version of the 3PL model is plagued with identification problems; see, for example, [Samejima \(1973\)](#), [Holland \(1990\)](#), [Yen, Burket, and Sykes \(1991\)](#), [Maris \(2002\)](#), and [San Martín, Rolin, and Castro \(2013\)](#).

The `sepguessing` option can be useful in the context of hybrid IRT models, where separate pseudoguessing parameters can be estimated for a subset of items; see [example 2](#) in [\[IRT\] irt hybrid](#).

See [Balov \(2016\)](#) for an example of Bayesian estimation of a 3PL model with separate pseudoguessing parameters.



► Example 1: Fitting a 3PL model

To illustrate the 3PL model, we use an abridged version of the mathematics and science data from [De Boeck and Wilson \(2004\)](#). Student responses to test items are coded 1 for correct and 0 for incorrect. Here we list the first five observations.

```
. use https://www.stata-press.com/data/r19/masc1
(Data from De Boeck & Wilson (2004))

. list in 1/5
```

	q1	q2	q3	q4	q5	q6	q7	q8	q9
1.	1	1	1	0	0	0	0	1	0
2.	0	0	1	0	0	0	0	1	1
3.	0	0	0	1	0	0	1	0	0
4.	0	0	1	0	0	0	0	0	1
5.	0	1	1	0	0	0	0	1	0

Looking across the rows, we see that the first student correctly answered items q1, q2, q3, and q8, the second student correctly answered items q3, q8, and q9, and so on.

We fit a 3PL model to binary items q1–q9 as follows:

```
. irt 3pl q1-q9
Fitting fixed-effects model:
Iteration 0:  Log likelihood = -5322.8824
Iteration 1:  Log likelihood = -4317.9868
Iteration 2:  Log likelihood = -4273.6659
Iteration 3:  Log likelihood = -4269.7862
Iteration 4:  Log likelihood = -4269.7825
Iteration 5:  Log likelihood = -4269.7825
Fitting full model:
Iteration 0:  Log likelihood = -4226.5553 (not concave)
Iteration 1:  Log likelihood = -4126.9014 (not concave)
Iteration 2:  Log likelihood = -4120.7233
Iteration 3:  Log likelihood = -4117.2619
Iteration 4:  Log likelihood = -4116.3931
Iteration 5:  Log likelihood = -4116.3434
Iteration 6:  Log likelihood = -4116.3404
Iteration 7:  Log likelihood = -4116.3404
```

Three-parameter logistic model
Log likelihood = -4116.3404

Number of obs = 800

		Coefficient	Std. err.	z	P> z	[95% conf. interval]	
q1	Discrim	1.911892	.3633509	5.26	0.000	1.199737	2.624047
	Diff	-.3040607	.0970816	-3.13	0.002	-.4943372	-.1137842
q2	Discrim	.750889	.1414085	5.31	0.000	.4737334	1.028045
	Diff	.1506376	.1667842	0.90	0.366	-.1762535	.4775287
q3	Discrim	.9674965	.1682051	5.75	0.000	.6378205	1.297172
	Diff	-1.508912	.2358768	-6.40	0.000	-1.971222	-1.046602
q4	Discrim	.9846883	.1860968	5.29	0.000	.6199454	1.349431
	Diff	.5726226	.1491574	3.84	0.000	.2802794	.8649659
q5	Discrim	1.439631	.4426227	3.25	0.001	.5721063	2.307156
	Diff	1.605677	.2144335	7.49	0.000	1.185395	2.025959
q6	Discrim	1.369119	.3249596	4.21	0.000	.7322101	2.006028
	Diff	.7818615	.1236333	6.32	0.000	.5395447	1.024178
q7	Discrim	.4823135	.1727569	2.79	0.005	.1437162	.8209108
	Diff	3.010922	.8924986	3.37	0.001	1.261656	4.760187
q8	Discrim	1.436069	.2482751	5.78	0.000	.9494586	1.922679
	Diff	-1.594747	.1918747	-8.31	0.000	-1.970815	-1.21868
q9	Discrim	.6772551	.1314525	5.15	0.000	.419613	.9348971
	Diff	-1.213933	.2661804	-4.56	0.000	-1.735637	-.6922291
	Guess	.0904473	.0359669			.0199534	.1609412

In the 3PL model, each test item has its own parameter estimates for discrimination and difficulty. The estimated common pseudoguessing parameter is reported at the end of the table.

In the following, we use `estat report` to replay the table of estimated IRT parameters and control how the output is reported. We include the `byparm` option, which arranges the output by parameter rather than by item, and the `sort(b)` option, which displays the items in an ascending order of difficulty. This makes it easy to see that item q8 is least difficult and item q7 is most difficult.

```
. estat report, byparm sort(b)
```

```
Three-parameter logistic model  
Log likelihood = -4116.3404
```

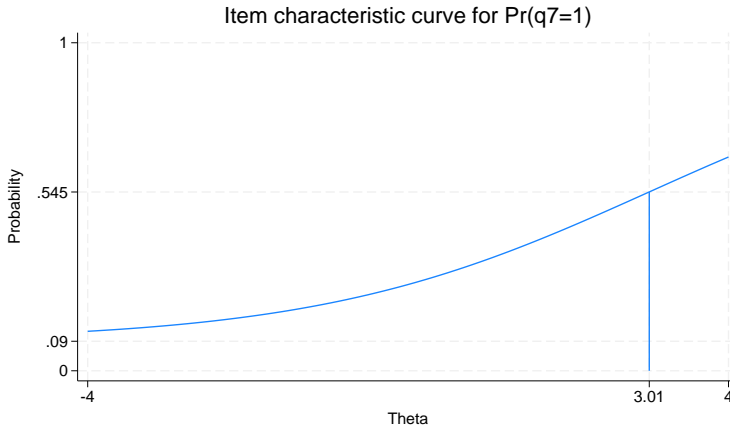
Number of obs = 800

		Coefficient	Std. err.	z	P> z	[95% conf. interval]	
Discrim	q8	1.436069	.2482751	5.78	0.000	.9494586	1.922679
	q3	.9674965	.1682051	5.75	0.000	.6378205	1.297172
	q9	.6772551	.1314525	5.15	0.000	.419613	.9348971
	q1	1.911892	.3633509	5.26	0.000	1.199737	2.624047
	q2	.750889	.1414085	5.31	0.000	.4737334	1.028045
	q4	.9846883	.1860968	5.29	0.000	.6199454	1.349431
	q6	1.369119	.3249596	4.21	0.000	.7322101	2.006028
	q5	1.439631	.4426227	3.25	0.001	.5721063	2.307156
	q7	.4823135	.1727569	2.79	0.005	.1437162	.8209108
Diff	q8	-1.594747	.1918747	-8.31	0.000	-1.970815	-1.21868
	q3	-1.508912	.2358768	-6.40	0.000	-1.971222	-1.046602
	q9	-1.213933	.2661804	-4.56	0.000	-1.735637	-.6922291
	q1	-.3040607	.0970816	-3.13	0.002	-.4943372	-.1137842
	q2	.1506376	.1667842	0.90	0.366	-.1762535	.4775287
	q4	.5726226	.1491574	3.84	0.000	.2802794	.8649659
	q6	.7818615	.1236333	6.32	0.000	.5395447	1.024178
	q5	1.605677	.2144335	7.49	0.000	1.185395	2.025959
	q7	3.010922	.8924986	3.37	0.001	1.261656	4.760187
Guess		.0904473	.0359669			.0199534	.1609412

The estimate of the pseudoguessing parameter is 0.09, which suggests a modest degree of guessing on the test. The pseudoguessing parameter represents the smallest probability of a correct response. Thus, according to this model, even the least able student has, at minimum, a 9% chance of responding correctly on any given item.

After `irt 3pl`, we can use `irtgraph icc` to plot the ICCs using the estimated 3PL parameters; see [IRT] [irtgraph icc](#). To focus on the most difficult item, as reported by `estat` report, we restrict the plot to item q7. We use option `blocation` to add a vertical line at the estimated difficulty and option `ylabel()` to change the default labeling of the y axis to include the lower asymptote and the midpoint probability, where θ equals the estimated difficulty for q7.

```
. irtgraph icc q7, blocation ylabel(0 0.09 0.545 1)
```



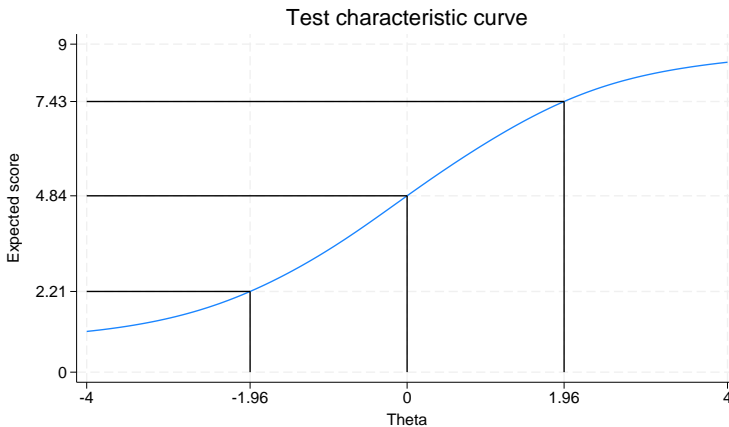
Notice that the estimate of the pseudoguessing parameter is now a lower asymptote for the plotted ICC. Also, because of the pseudoguessing parameter, the midpoint probability, where θ equals the estimated difficulty for q7, is

$$\hat{c} + (1 - \hat{c}) \times \frac{1}{2} = 0.09 + 0.91 \times \frac{1}{2} = 0.545$$

instead of 0.5, as in the case of 1PL and 2PL models.

The TCC plots the expected score as a function of θ , using the estimated 3PL parameters. We use `irtgraph tcc` to plot the TCC. For 9 binary items, it is clear that the total score ranges from 0 to 9; however, because of the pseudoguessing parameter, the minimum expected score is $\hat{c} \times 9 = 0.09 \times 9 = 0.81$. The `thetalines()` option plots the expected scores at the specified values for θ .

```
. irtgraph tcc, thetalines(-1.96 0 1.96)
```



This plot tells us what kind of scores we can expect from individuals with different levels of the latent trait. For example, we can expect above-average individuals to score 4.84 or above. Actually, no one is expected to score 4.84 on a 9-item test, so a more realistic statement is that we expect above-average individuals to score above 4.

Using the 95% critical values from the standard normal distribution (-1.96 and 1.96), this plot also tells us that we can expect 95% of randomly selected people to score between 2.21 and 7.43. A more realistic statement is that we expect about 95% of randomly selected people to score from 2 to 7.



Video example

[Item response theory using Stata: Three-parameter logistic \(3PL\) models](#)

Stored results

`irt 3pl` stores the following in `e()`:

Scalars

<code>e(N)</code>	number of observations
<code>e(k)</code>	number of parameters
<code>e(k_eq)</code>	number of equations in <code>e(b)</code>
<code>e(k_dv)</code>	number of dependent variables
<code>e(k_rc)</code>	number of covariances
<code>e(k_rs)</code>	number of variances
<code>e(irt_k_eq)</code>	number of IRT equations
<code>e(k_items1)</code>	number of items in first IRT equation
<code>e(sepguess1)</code>	1 if model contains a separate pseudoguessing parameter
<code>e(ll)</code>	log likelihood
<code>e(N_clust)</code>	number of clusters
<code>e(N_groups)</code>	number of groups

<code>e(n_quad)</code>	number of integration points
<code>e(rank)</code>	rank of $e(V)$
<code>e(ic)</code>	number of iterations
<code>e(rc)</code>	return code
<code>e(converged)</code>	1 if target model converged, 0 otherwise
Macros	
<code>e(cmd)</code>	<code>gsem</code>
<code>e(cmd2)</code>	<code>irt</code>
<code>e(cmdline)</code>	command as typed
<code>e(model1)</code>	<code>3pl</code>
<code>e(items1)</code>	names of items in first IRT equation
<code>e(depvar)</code>	names of all item variables
<code>e(wtype)</code>	weight type
<code>e(wexp)</code>	weight expression
<code>e(title)</code>	title in estimation output
<code>e(clustvar)</code>	name of cluster variable
<code>e(groupvar)</code>	name of group variable
<code>e(family#)</code>	family for the <i>#th item</i>
<code>e(link#)</code>	link for the <i>#th item</i>
<code>e(intmethod)</code>	integration method
<code>e(vce)</code>	<i>vcetype</i> specified in <code>vce()</code>
<code>e(vcetype)</code>	title used to label Std. err.
<code>e(opt)</code>	type of optimization
<code>e(which)</code>	max or min; whether optimizer is to perform maximization or minimization
<code>e(method)</code>	estimation method: <code>ml</code>
<code>e(ml_method)</code>	type of <code>ml</code> method
<code>e(user)</code>	name of likelihood-evaluator program
<code>e(technique)</code>	maximization technique
<code>e(datasignature)</code>	the checksum
<code>e(datasignaturevars)</code>	variables used in calculation of checksum
<code>e(properties)</code>	<code>b V</code>
<code>e(estat_cmd)</code>	program used to implement <code>estat</code>
<code>e(predict)</code>	program used to implement <code>predict</code>
<code>e(covariates)</code>	list of covariates
<code>e(footnote)</code>	program used to implement the footnote display
Matrices	
<code>e(_N)</code>	sample size for each item
<code>e(b)</code>	coefficient vector, slope-intercept parameterization
<code>e(b_pclass)</code>	parameter class
<code>e(Cns)</code>	constraints matrix
<code>e(ilog)</code>	iteration log (up to 20 iterations)
<code>e(gradient)</code>	gradient vector
<code>e(V)</code>	variance–covariance matrix of the estimators
<code>e(V_modelbased)</code>	model-based variance
<code>e(groupvalue)</code>	vector of group values in <code>e(groupvar)</code>
<code>e(nobs)</code>	vector with number of observations per group
Functions	
<code>e(sample)</code>	marks estimation sample

In addition to the above, the following is stored in `r()`:

Matrices	
<code>r(table)</code>	matrix containing the coefficients with their standard errors, test statistics, <i>p</i> -values, and confidence intervals

Note that results stored in `r()` are updated when the command is replayed and will be replaced when any `r-class` command is run after the estimation command.

Methods and formulas

Let Y_{ij} represent the (yet to be observed) outcome for item i from person j , and let y_{ij} be the observed value of Y_{ij} . Without loss of generality, we will use the terms “correct” and “incorrect” in reference to the outcomes of Y_{ij} . Furthermore, we will refer to $y_{ij} = 1$ as correct and $y_{ij} = 0$ as incorrect.

Using the IRT parameterization, we see that the probability of person j with latent trait level θ_j (the latent trait) providing a correct response to item i is given by

$$\Pr(Y_{ij} = 1 | a_i, b_i, c_i, \theta_j) = c_i + (1 - c_i) \frac{\exp\{a_i(\theta_j - b_i)\}}{1 + \exp\{a_i(\theta_j - b_i)\}}$$

where a_i represents the discrimination of item i , b_i represents the difficulty of item i , and c_i represents the pseudoguessing parameter. `irt 3pl` fits the model using the slope-intercept form, so the probability for providing a correct answer is parameterized as

$$\Pr(Y_{ij} = 1 | \alpha_i, \beta_i, \gamma_i, \theta_j) = \frac{\exp(\gamma_i)}{1 + \exp(\gamma_i)} + \frac{1}{1 + \exp(\gamma_i)} \frac{\exp(\alpha_i \theta_j + \beta_i)}{1 + \exp(\alpha_i \theta_j + \beta_i)}$$

The transformation between these two parameterizations is

$$a_i = \alpha_i \quad b_i = -\frac{\beta_i}{\alpha_i} \quad c_i = \frac{\exp(\gamma_i)}{1 + \exp(\gamma_i)}$$

By default, the γ_i (and thus the c_i) are constrained to be the same across all items.

Let $p_{ij} \equiv \Pr(Y_{ij} = 1 | \alpha_i, \beta_i, \gamma_i, \theta_j)$ and $q_{ij} = 1 - p_{ij}$. Conditional on θ_j , the item responses are assumed to be independent, so the conditional density for person j is given by

$$f(\mathbf{y}_j | \mathbf{B}, \theta_j) = \prod_{i=1}^I p_{ij}^{y_{ij}} q_{ij}^{1-y_{ij}}$$

where $\mathbf{y}_j = (y_{1j}, \dots, y_{Ij})$, $\mathbf{B} = (\alpha_1, \dots, \alpha_I, \beta_1, \dots, \beta_I, \gamma_1, \dots, \gamma_I)$, and I is the number of items.

Missing items are skipped over in the above product by default. When the `listwise` option is specified, persons with any missing items are dropped from the estimation sample.

The likelihood for person j is computed by integrating out the latent variable from the joint density

$$L_j(\mathbf{B}) = \int_{-\infty}^{\infty} f(\mathbf{y}_j | \mathbf{B}, \theta_j) \phi(\theta_j) d\theta_j$$

where $\phi(\cdot)$ is the density function for the standard normal distribution. The log likelihood for the estimation sample is simply the sum of the log likelihoods from the N persons in the estimation sample.

$$\log L(\mathbf{B}) = \sum_{j=1}^N \log L_j(\mathbf{B})$$

The integral in the formula for $L_j(\mathbf{B})$ is generally not tractable, so we must use numerical methods.

Models for multiple groups, Gauss–Hermite quadrature, and adaptive quadrature are documented in *Methods and formulas* of [IRT] `irt hybrid`.

References

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Also see

- [IRT] **irt 3pl postestimation** — Postestimation tools for irt 3pl
- [IRT] **irt** — Introduction to IRT models
- [IRT] **irt 1pl** — One-parameter logistic model
- [IRT] **irt 2pl** — Two-parameter logistic model
- [IRT] **irt constraints** — Specifying constraints
- [SEM] **gsem** — Generalized structural equation model estimation command
- [SVY] **svy estimation** — Estimation commands for survey data
- [U] **20 Estimation and postestimation commands**

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