

# Linear Regression Models with Interaction/Moderation

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## 1 Introduction

### 1.1 Goals

#### Goals

- Learn how to use factor variable notation when fitting models involving
    - ◇ Categorical variables
    - ◇ Interactions
    - ◇ Polynomial terms
  
  - Learn how to use postestimation tools to interpret interactions
    - ◇ Tests for group differences
    - ◇ Tests of slopes
    - ◇ Graphs
-

## A Linear Model

- We'll use data from the National Health and Nutrition Examination Survey (NHANES) for our examples

```
. webuse nhanes2
```

- We'll start with a basic a model for bmi using age and sex (female).

- Before we fit the model, let's investigate the variables using codebook

```
. codebook bmi age female
```

```
-----  
bmi                                                    Body Mass Index (BMI)  
-----
```

```
      type: numeric (float)  
  
      range: [12.385596,61.129696]      units: 1.000e-07  
unique values: 9,941                    missing .: 0/10,351  
  
      mean: 25.5376  
      std. dev: 4.91497  
  
percentiles:      10%      25%      50%      75%      90%  
                20.1037  22.142  24.8181  28.0267  31.7259
```

```
-----  
age                                                    age in years  
-----
```

```
      type: numeric (byte)  
  
      range: [20,74]                    units: 1  
unique values: 55                       missing .: 0/10,351  
  
      mean: 47.5797  
      std. dev: 17.2148  
  
percentiles:      10%      25%      50%      75%      90%  
                24       31       49       63       69
```

```
-----  
female                                                1=female, 0=male  
-----
```

```
      type: numeric (byte)  
  
      range: [0,1]                      units: 1  
unique values: 2                         missing .: 0/10,351  
  
tabulation: Freq. Value  
            4,915 0  
            5,436 1
```

- Now we can fit the model

```
. regress bmi age female
```

Source	SS	df	MS	Number of obs	=	10,351
Model	7330.98402	2	3665.49201	F(2, 10348)	=	156.29
Residual	242693.178	10,348	23.4531483	Prob > F	=	0.0000
				R-squared	=	0.0293
				Adj R-squared	=	0.0291
Total	250024.162	10,350	24.1569239	Root MSE	=	4.8428

bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
age	.0488667	.0027653	17.67	0.000	.0434462	.0542872
female	.0380616	.0953249	0.40	0.690	-.1487936	.2249168
_cons	23.19255	.1482223	156.47	0.000	22.90201	23.48309

## 2 Estimation

### 2.1 Including Categorical Variables

#### Working with Categorical Variables

- We would now like to include region in the model, let's take a look at this variable

```
. codebook region
```

```
-----
region                                                    1=NE, 2=MW, 3=S, 4=W
-----

      type:  numeric (byte)
      label:  region

      range:  [1,4]
unique values: 4

      units:  1
missing .:  0/10,351

      tabulation:  Freq.  Numeric  Label
                   2,096      1  NE
                   2,774      2  MW
                   2,853      3   S
                   2,628      4   W
```

- It cannot simply be added to the list of covariates because it has 4 categories
- To include a categorical variable, put an `i.` in front of its name—this declares the variable to be a categorical variable, or in Stataese, a *factor variable*

- For example, to add region to our model we use

```
. regress bmi age i.female i.region
```

Source	SS	df	MS	Number of obs	=	10,351
Model	7390.19781	5	1478.03956	F(5, 10345)	=	63.02
Residual	242633.964	10,345	23.4542256	Prob > F	=	0.0000
				R-squared	=	0.0296
				Adj R-squared	=	0.0291
Total	250024.162	10,350	24.1569239	Root MSE	=	4.843

bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
age	.0488851	.0027674	17.66	0.000	.0434605 .0543097
female					
0	0 (base)				
1	.0372717	.0953357	0.39	0.696	-.1496047 .2241481
region					
NE	0 (base)				
MW	.0064779	.1402121	0.05	0.963	-.268365 .2813207
S	.0387957	.1393383	0.28	0.781	-.2343342 .3119256
W	-.1537648	.1418286	-1.08	0.278	-.4317762 .1242466
_cons	23.2187	.1760452	131.89	0.000	22.87362 23.56378

## Niceities

- Value labels associated with factor variables are displayed in the regression table
- We can tell Stata to show the base categories for our factor variables

```
. set showbaselevels on
```

## Factor Notation as Operators

- The `i.` operator can be applied to many variables at once:

```
. regress bmi age i.(female region)
```

Source	SS	df	MS	Number of obs	=	10,351
Model	7390.19781	5	1478.03956	F(5, 10345)	=	63.02
Residual	242633.964	10,345	23.4542256	Prob > F	=	0.0000
				R-squared	=	0.0296
				Adj R-squared	=	0.0291
Total	250024.162	10,350	24.1569239	Root MSE	=	4.843

bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
age	.0488851	.0027674	17.66	0.000	.0434605	.0543097
female						
0	0	(base)				
1	.0372717	.0953357	0.39	0.696	-.1496047	.2241481
region						
NE	0	(base)				
MW	.0064779	.1402121	0.05	0.963	-.268365	.2813207
S	.0387957	.1393383	0.28	0.781	-.2343342	.3119256
W	-.1537648	.1418286	-1.08	0.278	-.4317762	.1242466
_cons	23.2187	.1760452	131.89	0.000	22.87362	23.56378

- In other words, it understands the distributive property
  - ◇ This is useful when using variable ranges, for example
- For the curious, factor variable notation works with wildcards
  - ◇ If there were many variables starting with `u`, then `i.u*` would include them all as factor variables

---

## Using Different Base Categories

- By default, the smallest-valued category is the base category
- This can be overridden within commands
  - ◇ `b#`. specifies the value `#` as the base
  - ◇ `b(##)`. specifies the `#`'th largest value as the base
  - ◇ `b(first)`. specifies the smallest value as the base
  - ◇ `b(last)`. specifies the largest value as the base
  - ◇ `b(freq)`. specifies the most prevalent value as the base
  - ◇ `bn`. specifies there should be no base

## Playing with the Base

- We can use `region=3` as the base class on the fly:  

```
. regress bmi age i.female b3.region
```
  - We can use the most prevalent category as the base  

```
. regress bmi age i.female b(freq).region
```
  - Factor variables can be distributed across many variables  

```
. regress bmi age b(freq).(female region)
```
  - The base category can be omitted (with some care here)  

```
. regress bmi age i.female bn.region, noconstant
```
  - We can also include a term for `region=4` only  

```
. regress bmi age i.female 4.region
```
- 

## 2.2 Including Interactions

### Specifying Interactions

- Factor variables are also used for specifying interactions
    - ◇ This is where they really shine
  - To include both main effects and interaction terms in a model, put `##` between the variables
  - To include only the interaction terms, put `#` between the terms
-

## Categorical by Categorical Interactions

- For example, to fit a model that includes main effects for age, female, and region, as well as the interaction of female, and region

```
. regress bmi age female##region
```

Source	SS	df	MS	Number of obs	=	10,351
Model	7559.19099	8	944.898874	F(8, 10342)	=	40.30
Residual	242464.971	10,342	23.4446888	Prob > F	=	0.0000
				R-squared	=	0.0302
				Adj R-squared	=	0.0295
Total	250024.162	10,350	24.1569239	Root MSE	=	4.842

bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
age	.0488087	.0027671	17.64	0.000	.0433846 .0542328
female					
0	0	(base)			
1	-.2939562	.2116093	-1.39	0.165	-.7087514 .1208389
region					
NE	0	(base)			
MW	-.1420836	.2023593	-0.70	0.483	-.538747 .2545798
S	-.3347762	.2015721	-1.66	0.097	-.7298965 .0603441
W	-.2694841	.204234	-1.32	0.187	-.6698222 .1308541
female#region					
1#MW	.2897474	.280525	1.03	0.302	-.2601358 .8396306
1#S	.7124639	.2789251	2.55	0.011	.1657169 1.259211
1#W	.2266557	.2837887	0.80	0.424	-.3296251 .7829365
_cons	23.39271	.2013939	116.15	0.000	22.99793 23.78748

- Variables involved in interactions are assumed to be categorical, so no `i.` is needed

- To see all the omitted terms we can add the `allbaselevels` option

```
. regress bmi age female##region, allbaselevels
```

Source	SS	df	MS	Number of obs	=	10,351
Model	7559.19099	8	944.898874	F(8, 10342)	=	40.30
Residual	242464.971	10,342	23.4446888	Prob > F	=	0.0000
				R-squared	=	0.0302
				Adj R-squared	=	0.0295
Total	250024.162	10,350	24.1569239	Root MSE	=	4.842

bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
age	.0488087	.0027671	17.64	0.000	.0433846 .0542328
female					
0	0 (base)				
1	-.2939562	.2116093	-1.39	0.165	-.7087514 .1208389
region					
NE	0 (base)				
MW	-.1420836	.2023593	-0.70	0.483	-.538747 .2545798
S	-.3347762	.2015721	-1.66	0.097	-.7298965 .0603441
W	-.2694841	.204234	-1.32	0.187	-.6698222 .1308541
female#region					
0#NE	0 (base)				
0#MW	0 (base)				
0#S	0 (base)				
0#W	0 (base)				
1#NE	0 (base)				
1#MW	.2897474	.280525	1.03	0.302	-.2601358 .8396306
1#S	.7124639	.2789251	2.55	0.011	.1657169 1.259211
1#W	.2266557	.2837887	0.80	0.424	-.3296251 .7829365
_cons	23.39271	.2013939	116.15	0.000	22.99793 23.78748



## Categorical by Continuous Interactions

- To include continuous variables in interactions use `c.` to specify that a variable is continuous
  - ◊ Otherwise it will be assumed to be categorical
- Here is our model with an interaction between age and region

```
. regress bmi c.age##region i.female
```

Source	SS	df	MS	Number of obs	=	10,351
Model	7568.54189	8	946.067737	F(8, 10342)	=	40.35
Residual	242455.62	10,342	23.4437846	Prob > F	=	0.0000
				R-squared	=	0.0303
				Adj R-squared	=	0.0295
Total	250024.162	10,350	24.1569239	Root MSE	=	4.8419

bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
age	.0607829	.0062164	9.78	0.000	.0485975 .0729683
region					
NE	0 (base)				
MW	.3951518	.4106204	0.96	0.336	-.4097436 1.200047
S	1.051668	.4181868	2.51	0.012	.2319407 1.871395
W	.5921285	.4181932	1.42	0.157	-.227611 1.411868
region#c.age					
MW	-.0080245	.0081638	-0.98	0.326	-.0240272 .0079782
S	-.0211109	.008219	-2.57	0.010	-.0372217 -.0050002
W	-.0155977	.0082261	-1.90	0.058	-.0317225 .000527
female					
0	0 (base)				
1	.038259	.0953259	0.40	0.688	-.1485982 .2251161
_cons	22.64929	.3193208	70.93	0.000	22.02336 23.27522

## Continuous by Continuous Interactions

- Prefix both variables in the interaction with `c.` to fit models with continuous by continuous variable interactions
- For example, we can interact age with serum vitamin c levels (`vitaminc`)

```
. regress bmi c.age##c.vitaminc i.female i.region
```

Source	SS	df	MS	Number of obs	=	9,973
-----+-----				F(7, 9965)	=	63.61
Model	10298.9223	7	1471.27461	Prob > F	=	0.0000
Residual	230479.207	9,965	23.1288718	R-squared	=	0.0428
-----+-----				Adj R-squared	=	0.0421
Total	240778.13	9,972	24.1454201	Root MSE	=	4.8092

bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
-----+-----						
age	.0220407	.0059366	3.71	0.000	.0104038	.0336777
vitaminc	-2.331426	.2717928	-8.58	0.000	-2.864194	-1.798657
c.age#c.vitaminc	.029107	.0050017	5.82	0.000	.0193026	.0389115
female						
0	0 (base)					
1	.1858965	.0982311	1.89	0.058	-.0066564	.3784494
region						
NE	0 (base)					
MW	-.0936871	.1412331	-0.66	0.507	-.3705326	.1831584
S	-.2137082	.1431247	-1.49	0.135	-.4942615	.0668451
W	-.1626738	.1430181	-1.14	0.255	-.4430182	.1176706
_cons	25.45695	.3293507	77.29	0.000	24.81136	26.10255
-----+-----						

- To include polynomial terms, interact a variable with itself
- For example, a model that includes both age and age<sup>2</sup>

```
. regress bmi c.age##c.age i.female i.region
```

Source	SS	df	MS	Number of obs	=	10,351
Model	10269.3919	6	1711.56532	F(6, 10344)	=	73.84
Residual	239754.77	10,344	23.1781487	Prob > F	=	0.0000
-----				R-squared	=	0.0411
-----				Adj R-squared	=	0.0405
Total	250024.162	10,350	24.1569239	Root MSE	=	4.8144

bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
age	.2731368	.0203077	13.45	0.000	.2333297	.3129439
c.age#c.age	-.0024099	.0002162	-11.15	0.000	-.0028337	-.0019861
female						
0	0 (base)					
1	.0462855	.0947764	0.49	0.625	-.1394945	.2320656
region						
NE	0 (base)					
MW	.0322091	.1394036	0.23	0.817	-.2410489	.3054671
S	.0289346	.1385186	0.21	0.835	-.2425886	.3004579
W	-.1105093	.1410448	-0.78	0.433	-.3869844	.1659657
_cons	18.6987	.4416971	42.33	0.000	17.83289	19.56451

◇ The coefficient for age-squared is next to c.age#c.age

## Higher Order Interactions

- Factor variable syntax can be used to specify higher order interactions
- If the interactions are specified using ## all lower order terms are included

- For example, here we fit a model for bmi using a model that includes the three-way interaction of continuous variables age and vitaminc and categorical variable female

```
. regress bmi c.age##c.vitaminc##female
```

Source	SS	df	MS	Number of obs	=	9,973
Model	12294.4386	7	1756.34837	F(7, 9965)	=	76.60
Residual	228483.691	9,965	22.9286193	Prob > F	=	0.0000
				R-squared	=	0.0511
				Adj R-squared	=	0.0504
Total	240778.13	9,972	24.1454201	Root MSE	=	4.7884

	bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
age		-.0038595	.0084263	-0.46	0.647	-.0203767	.0126578
vitaminc		-2.008713	.4231851	-4.75	0.000	-2.838241	-1.179185
c.age#c.vitaminc		.0313728	.0078481	4.00	0.000	.0159889	.0467566
female							
0		0	(base)				
1		-2.098183	.6208318	-3.38	0.001	-3.315138	-.8812268
female#c.age							
1		.0646392	.0119517	5.41	0.000	.0412115	.0880668
female#c.vitaminc							
1		.0314475	.5539279	0.06	0.955	-1.054363	1.117258
female#c.age#c.vitaminc							
1		-.0166002	.0102645	-1.62	0.106	-.0367206	.0035203
_cons		26.16464	.4416624	59.24	0.000	25.29889	27.03039

### Some Factor Variable Notes

- If you plan to look at marginal effects of any kind, it is best to
  - Explicitly mark all categorical variables with *i*.
  - Specify all interactions using # or ##
  - Specify powers of a variable as interactions of the variable with itself
- There can be up to 8 categorical and 8 continuous interactions in one expression
  - Have fun with the interpretation

## 3 Postestimation

### 3.1 About Postestimation

#### Introduction to Postestimation

- In Stata jargon, postestimation commands are commands that can be run after a model is fit, for example

- ◇ Predictions
  - ◇ Additional hypothesis tests
  - ◇ Checks of assumptions
- We'll explore postestimation tools that can be used to help interpret the results of models that include interactions
  - The usefulness of specific tools will depend on the types of hypotheses you wish to examine

## 3.2 Investigating Categorical by Categorical Interactions

### Estimating a Model

- Lets begin by running a model with main effects for age, female and region, and the interaction of female and region

```
. regress bmi age female##region
```

Source	SS	df	MS	Number of obs	=	10,351
Model	7559.19099	8	944.898874	F(8, 10342)	=	40.30
Residual	242464.971	10,342	23.4446888	Prob > F	=	0.0000
				R-squared	=	0.0302
				Adj R-squared	=	0.0295
Total	250024.162	10,350	24.1569239	Root MSE	=	4.842

bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
age	.0488087	.0027671	17.64	0.000	.0433846 .0542328
female					
0	0	(base)			
1	-.2939562	.2116093	-1.39	0.165	-.7087514 .1208389
region					
NE	0	(base)			
MW	-.1420836	.2023593	-0.70	0.483	-.538747 .2545798
S	-.3347762	.2015721	-1.66	0.097	-.7298965 .0603441
W	-.2694841	.204234	-1.32	0.187	-.6698222 .1308541
female#region					
1#MW	.2897474	.280525	1.03	0.302	-.2601358 .8396306
1#S	.7124639	.2789251	2.55	0.011	.1657169 1.259211
1#W	.2266557	.2837887	0.80	0.424	-.3296251 .7829365
_cons	23.39271	.2013939	116.15	0.000	22.99793 23.78748

- How might we begin?
  - ◇ Perform joint tests of coefficients
  - ◇ Estimate and test hypotheses about group differences

## Finding the Coefficient Names

- Some postestimation commands require that you know the names used to store the coefficients
- To see these names we can replay the model and showing the *coefficient legend*

```
. regress, coeflegend
```

Source	SS	df	MS	Number of obs	=	10,351
Model	7559.19099	8	944.898874	F(8, 10342)	=	40.30
Residual	242464.971	10,342	23.4446888	Prob > F	=	0.0000
				R-squared	=	0.0302
				Adj R-squared	=	0.0295
Total	250024.162	10,350	24.1569239	Root MSE	=	4.842

bmi	Coef.	Legend
age	.0488087	_b[age]
female		
0	0	_b[0b.female]
1	-.2939562	_b[1.female]
region		
NE	0	_b[1b.region]
MW	-.1420836	_b[2.region]
S	-.3347762	_b[3.region]
W	-.2694841	_b[4.region]
female#region		
1#MW	.2897474	_b[1.female#2.region]
1#S	.7124639	_b[1.female#3.region]
1#W	.2266557	_b[1.female#4.region]
_cons	23.39271	_b[_cons]

- From here, we can see the full specification of the factor levels:

\_b[2.region] corresponds to region=2 which is "MW" or midwest

\_b[3.region] corresponds to region=3 which is "S" or south

- We can also see the terms for the interaction:

\_b[1.female#2.region] corresponds to the term for the interaction of region=2 and female=1

\_b[1.female#3.region] corresponds to the term for the interaction of region=3 and female=1

---

## Joint Tests

- The test command performs a Wald test of the specified null hypothesis
  - ◇ The default test is that the listed terms are equal to 0
- test takes a list of terms, which may be variable names, but can also be terms associated with factor variables

- To perform a joint test of the null hypothesis that the coefficients for the levels of `region` are all equal to 0

```
. test 2.region 3.region 4.region
```

```
( 1) 2.region = 0  
( 2) 3.region = 0  
( 3) 4.region = 0
```

```
F( 3, 10342) = 1.07  
Prob > F = 0.3600
```

- ◊ Since the model contains an interaction, this is a test of the effect of `region` when `female=0`

---

## Testing Sets of Coefficients

- To test that all of the coefficients associated with the interaction of `female` and `region` we would need to give the full name of all the coefficients

```
. test 1.female#2.region 1.female#3.region 1.female#4.region
```

- `testparm` also performs Wald tests, but it accepts lists of variables, rather than coefficients in the model
- So we can perform joint tests with less typing, for example

```
. testparm i.region#i.female
```

```
( 1) 1.female#2.region = 0  
( 2) 1.female#3.region = 0  
( 3) 1.female#4.region = 0
```

```
F( 3, 10342) = 2.40  
Prob > F = 0.0656
```

---

## An Alternative Test

- Likelihood ratio tests provide an alternative method of testing sets of coefficients
- To test the coefficients associated with the interaction of `female` and `region` we need to store our model results. The name is arbitrary, we'll call them `m1`

```
. estimates store m1
```

- Now we can rerun our model without region

```
. regress bmi age i.female i.region
```

Source	SS	df	MS	Number of obs	=	10,351
Model	7390.19781	5	1478.03956	F(5, 10345)	=	63.02
Residual	242633.964	10,345	23.4542256	Prob > F	=	0.0000
				R-squared	=	0.0296
				Adj R-squared	=	0.0291
Total	250024.162	10,350	24.1569239	Root MSE	=	4.843

bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
age	.0488851	.0027674	17.66	0.000	.0434605 .0543097
female					
0	0 (base)				
1	.0372717	.0953357	0.39	0.696	-.1496047 .2241481
region					
NE	0 (base)				
MW	.0064779	.1402121	0.05	0.963	-.268365 .2813207
S	.0387957	.1393383	0.28	0.781	-.2343342 .3119256
W	-.1537648	.1418286	-1.08	0.278	-.4317762 .1242466
_cons	23.2187	.1760452	131.89	0.000	22.87362 23.56378

- If we were removing one of these variables entirely, we would want to add `if e(sample)` to make sure the same sample, what Stata calls the *estimation sample*, is used for both models

### Likelihood Ratio Tests (Continued)

- Now we store the second set of estimates

```
. estimates store m2
```

- And use the `lrtest` command to perform the likelihood ratio test

```
. lrtest m1 m2
```

```
Likelihood-ratio test                LR chi2(3) =      7.21
(Assumption: m2 nested in m1)       Prob > chi2 =    0.0654
```

- We'll restore the results from `m1`

```
. estimates restore m1
```

```
(results m1 are active now)
```

- Now it's as if we just ran the model stored in `m1`



## Tests of Differences

- test can also be used to the equality of coefficients

```
. test 3.region#1.female = 4.region#1.female

(1) 1.female#3.region - 1.female#4.region = 0

      F( 1, 10342) =    3.43
      Prob > F =    0.0640
```

- A likelihood ratio test can also be used; see help constraint for information on setting the necessary constraints
- The `lincom` command can be used to calculate linear combinations of coefficients, along with standard errors, hypothesis tests, and confidence intervals
- For example, to obtain the difference in coefficients

```
. lincom 3.region#1.female - 4.region#1.female

(1) 1.female#3.region - 1.female#4.region = 0
```

bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
(1)	.4858082	.2622654	1.85	0.064	-.0282827 .9998991

## Contrasts

- The `contrast` command allows us to test a wide variety of comparisons across groups
- For example comparing regions separately for men and women

```
. contrast region@female, effects
```

Contrasts of marginal linear predictions

Margins : asbalanced

	df	F	P>F
region@female			
0	3	1.07	0.3600
1	3	2.17	0.0890
Joint	6	1.62	0.1364
Denominator	10342		

	Contrast	Std. Err.	t	P> t	[95% Conf. Interval]
region@female					
(MW vs base) 0	-.1420836	.2023593	-0.70	0.483	-.538747 .2545798
(MW vs base) 1	.1476637	.1943419	0.76	0.447	-.2332839 .5286114
(S vs base) 0	-.3347762	.2015721	-1.66	0.097	-.7298965 .0603441
(S vs base) 1	.3776878	.1927872	1.96	0.050	-.0002125 .755588
(W vs base) 0	-.2694841	.204234	-1.32	0.187	-.6698222 .1308541
(W vs base) 1	-.0428284	.1970381	-0.22	0.828	-.4290612 .3434044

- ◇ The @ symbol requests comparisons of the levels of `region` at each value of `female`
- ◇ The `effects` option requests that individual contrasts be displayed along with their standard errors, hypothesis tests, and confidence intervals

### Adjusting for Multiple Comparisons

- Use of `contrast` can result in a large number of hypothesis tests
- The `mcompare()` option can be used to adjust p-values and confidence intervals for multiple comparisons within factor variable terms
- The available methods are
  - ◇ `noadjust`
  - ◇ `bonferroni`
  - ◇ `sidak`
  - ◇ `scheffe`
- To apply Bonferroni's adjustment to our previous contrast

```
. contrast region@female, effects mcompare(bonferroni)
```

Contrasts of marginal linear predictions

Margins : asbalanced

	df	F	P>F
region@female			
0	3	1.07	0.3600
1	3	2.17	0.0890
Joint	6	1.62	0.1364
Denominator	10342		

Note: Bonferroni-adjusted p-values are reported for tests on individual contrasts only.

	Number of Comparisons
region@female	6

	Contrast	Std. Err.	Bonferroni t	Bonferroni P> t	Bonferroni [95% Conf. Interval]
region@female					
(MW vs base) 0	-.1420836	.2023593	-0.70	1.000	-.6760623 .391895
(MW vs base) 1	.1476637	.1943419	0.76	1.000	-.3651588 .6604863
(S vs base) 0	-.3347762	.2015721	-1.66	0.581	-.8666776 .1971252
(S vs base) 1	.3776878	.1927872	1.96	0.301	-.1310325 .886408
(W vs base) 0	-.2694841	.204234	-1.32	1.000	-.8084096 .2694415
(W vs base) 1	-.0428284	.1970381	-0.22	1.000	-.5627657 .4771089

## Average Predicted Values

- We might want to explore predictions based on our model and data
- Predictions for individual observations can be made using the `predict` command, see `help predict`
- To find out about our model more generally, we may be more interested in average predicted values
  - ◊ Also known as predictive margins or recycled predictions
- To obtain the average predicted value of `bmi`

```
. margins
```

```
Predictive margins          Number of obs   =    10,351
Model VCE      : OLS
```

```
Expression   : Linear prediction, predict()
```

---

		Delta-method				[95% Conf. Interval]	
	Margin	Std. Err.	t	P> t			
-----+-----							
_cons	25.5376	.0475917	536.60	0.000	25.44431	25.63089	
-----+-----							

---

## Predictions at Specified Values of Factor Variables

- Stata calls the list of variables that follow the `margins` command the *marginslist*
  - ◊ To appear in the *marginslist* a variable must have been specified as factor variable in the model
- To obtain the average predicted value of `bmi` at different values of `region`

```
. margins region
```

```
Predictive margins          Number of obs   =    10,351
Model VCE      : OLS
```

```
Expression   : Linear prediction, predict()
```

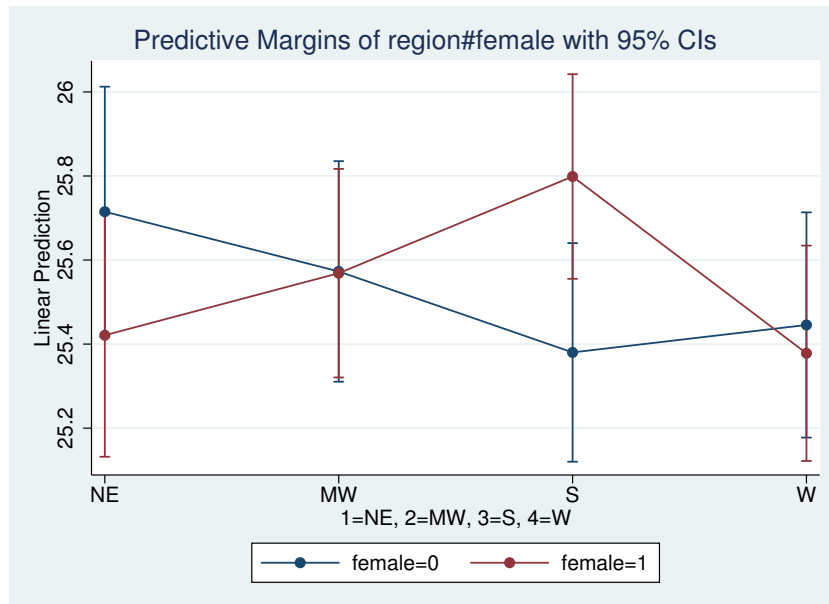
---

		Delta-method				[95% Conf. Interval]	
	Margin	Std. Err.	t	P> t			
-----+-----							
region							
NE	25.56063	.1057882	241.62	0.000	25.35327	25.768	
MW	25.57071	.09198	278.00	0.000	25.39042	25.75101	
S	25.60002	.0906777	282.32	0.000	25.42227	25.77776	
W	25.41018	.0944557	269.02	0.000	25.22503	25.59533	
-----+-----							

---

- How were these values generated?
  1. Calculate the predicted value of `bmi` setting `region=1` and using each case's observed values of `female` and `age`
  2. Find the mean of the predicted values
  3. Repeat steps 1 and 2 for each value of `region`





◇ If our model did not include a region by female interaction, the lines would be parallel

### Predicted Values for Specific Groups

- When we specify the variables in the *marginlist* Stata calculates predicted values treating each case as though it belonged to each group
- The `over()` option allows us to obtain predictions separately for each group, for example

```
. margins, over(female)
```

```
Predictive margins                                Number of obs    =    10,351
Model VCE      : OLS
```

```
Expression   : Linear prediction, predict()
over         : female
```

		Delta-method				
		Margin	Std. Err.	t	P> t	[95% Conf. Interval]
female						
	0	25.50999	.0690654	369.36	0.000	25.37461 25.64538
	1	25.56256	.0656723	389.24	0.000	25.43383 25.69129

- This time the table shows
  - ◇ The average predicted value of `bmi` for cases where `female=0` using each case's observed values of `age` and `region`
  - ◇ The average predicted value of `bmi` for cases where `female=1` using each case's observed values of `age` and `region`
- This can be useful when we want to compare groups

### 3.3 Investigating Categorical by Continuous Interactions

#### A Categorical by Continuous Interaction

- For this set of examples, we'll fit a model that includes an interaction between the continuous variable age and the categorical variable region

```
. regress bmi c.age##region i.female
```

Source	SS	df	MS	Number of obs	=	10,351
Model	7568.54189	8	946.067737	F(8, 10342)	=	40.35
Residual	242455.62	10,342	23.4437846	Prob > F	=	0.0000
Total	250024.162	10,350	24.1569239	R-squared	=	0.0303
				Adj R-squared	=	0.0295
				Root MSE	=	4.8419

bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
age	.0607829	.0062164	9.78	0.000	.0485975 .0729683
region					
NE	0 (base)				
MW	.3951518	.4106204	0.96	0.336	-.4097436 1.200047
S	1.051668	.4181868	2.51	0.012	.2319407 1.871395
W	.5921285	.4181932	1.42	0.157	-.227611 1.411868
region#c.age					
MW	-.0080245	.0081638	-0.98	0.326	-.0240272 .0079782
S	-.0211109	.008219	-2.57	0.010	-.0372217 -.0050002
W	-.0155977	.0082261	-1.90	0.058	-.0317225 .000527
female					
0	0 (base)				
1	.038259	.0953259	0.40	0.688	-.1485982 .2251161
_cons	22.64929	.3193208	70.93	0.000	22.02336 23.27522

- Let's take a look at how the coefficients are stored

```
. regress, coeflegend
```

Source	SS	df	MS	Number of obs	=	10,351
Model	7568.54189	8	946.067737	F(8, 10342)	=	40.35
Residual	242455.62	10,342	23.4437846	Prob > F	=	0.0000
				R-squared	=	0.0303
				Adj R-squared	=	0.0295
Total	250024.162	10,350	24.1569239	Root MSE	=	4.8419

bmi	Coef.	Legend
age	.0607829	_b[age]
region		
NE	0	_b[1b.region]
MW	.3951518	_b[2.region]
S	1.051668	_b[3.region]
W	.5921285	_b[4.region]
region#c.age		
MW	-.0080245	_b[2.region#c.age]
S	-.0211109	_b[3.region#c.age]
W	-.0155977	_b[4.region#c.age]
female		
0	0	_b[0b.female]
1	.038259	_b[1.female]
_cons	22.64929	_b[_cons]

### test and testparm

- As before, we can test the null hypothesis that all of the coefficients associated with the interaction of age and region are equal to 0 using testparm

```
. testparm c.age#i.region
```

```
( 1) 2.region#c.age = 0
( 2) 3.region#c.age = 0
( 3) 4.region#c.age = 0
```

```
F( 3, 10342) = 2.54
Prob > F = 0.0549
```

- We could also use lrtest
- We can test specific hypotheses about the slopes
- For example we might want to test whether the slope of age is significantly different in the south (region=3) versus the west (region=4)

```
. test 3.region#c.age = 4.region#c.age
```

```
( 1) 3.region#c.age - 4.region#c.age = 0
```

```
F( 1, 10342) = 0.52
Prob > F = 0.4689
```

## Estimated Slopes

- We can use `lincom` to estimate the slope of age for the south (`region=3`)

```
. lincom c.age + 3.region#c.age
```

```
( 1) age + 3.region#c.age = 0
```

	bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
(1)		.0396719	.0053765	7.38	0.000	.0291329 .0502109

- We can also use `margins` with the `dydx()` option to calculate the slope of age for each region

```
. margins region, dydx(age)
```

```
Average marginal effects          Number of obs   =   10,351
Model VCE      : OLS
```

```
Expression      : Linear prediction, predict()
dy/dx w.r.t.    : age
```

		Delta-method				
		dy/dx	Std. Err.	t	P> t	[95% Conf. Interval]
age						
	region					
	NE	.0607829	.0062164	9.78	0.000	.0485975 .0729683
	MW	.0527584	.0052919	9.97	0.000	.0423853 .0631315
	S	.0396719	.0053765	7.38	0.000	.0291329 .0502109
	W	.0451852	.0053875	8.39	0.000	.0346246 .0557457

- The `dydx()` option calculates derivative of the predicted values with respect to the specified variable, also known as the marginal effect

---

## Predictions at Specified Values

- To obtain margins at set values of continuous variables use the `at()` option



- For example, the predicted value of bmi at each level of region setting age=20

```
. margins region, at(age=20) vsquish
```

```
Predictive margins                                Number of obs    =    10,351
Model VCE      : OLS
```

```
Expression   : Linear prediction, predict()
at           : age                =         20
```

```
-----
```

		Delta-method				[95% Conf. Interval]	
	Margin	Std. Err.	t	P> t			
region							
NE	23.88504	.2026955	117.84	0.000	23.48772	24.28236	
MW	24.1197	.1678019	143.74	0.000	23.79078	24.44862	
S	24.51449	.1766004	138.81	0.000	24.16832	24.86066	
W	24.16521	.1772397	136.34	0.000	23.81779	24.51264	

```
-----
```

◇ The vsquish option reduces the vertical space in the output

- The at() option accepts numlists so we aren't restricted to a single value of age

```
. margins region, at(age=(20(25)70)) vsquish
```

```
Predictive margins                                Number of obs    =    10,351
Model VCE      : OLS
```

```
Expression   : Linear prediction, predict()
1._at       : age                =         20
2._at       : age                =         45
3._at       : age                =         70
```

```
-----
```

		Delta-method				[95% Conf. Interval]	
	Margin	Std. Err.	t	P> t			
_at#region							
1#NE	23.88504	.2026955	117.84	0.000	23.48772	24.28236	
1#MW	24.1197	.1678019	143.74	0.000	23.79078	24.44862	
1#S	24.51449	.1766004	138.81	0.000	24.16832	24.86066	
1#W	24.16521	.1772397	136.34	0.000	23.81779	24.51264	
2#NE	25.40461	.1072029	236.98	0.000	25.19447	25.61475	
2#MW	25.43866	.0922856	275.65	0.000	25.25776	25.61956	
2#S	25.50629	.0922593	276.46	0.000	25.32544	25.68713	
2#W	25.29484	.0956797	264.37	0.000	25.10729	25.48239	
3#NE	26.92418	.1737943	154.92	0.000	26.58351	27.26485	
3#MW	26.75762	.1545335	173.15	0.000	26.4547	27.06054	
3#S	26.49809	.148221	178.77	0.000	26.20754	26.78863	
3#W	26.42447	.1522388	173.57	0.000	26.12605	26.72289	

```
-----
```

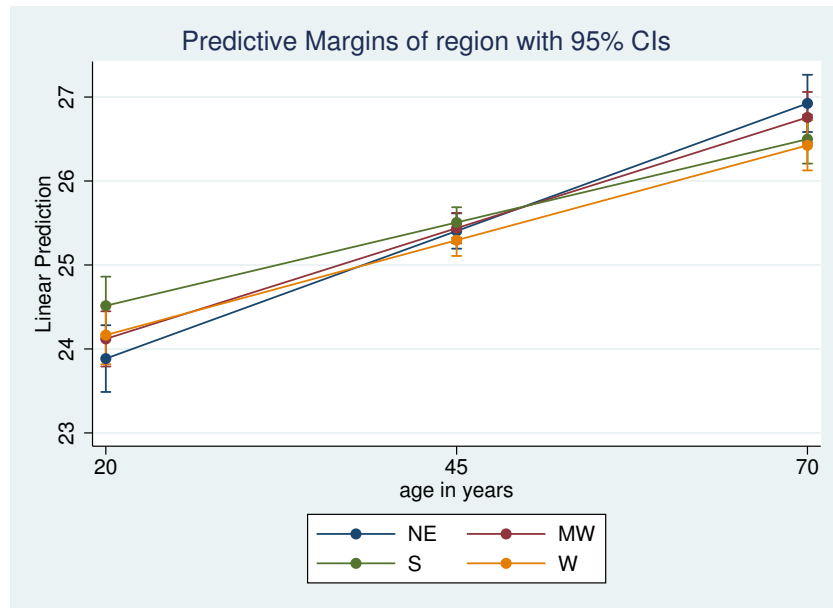
◇ The observed values of age are from 20 to 74

---

## Graphing Predicted Values

- And we can plot the results

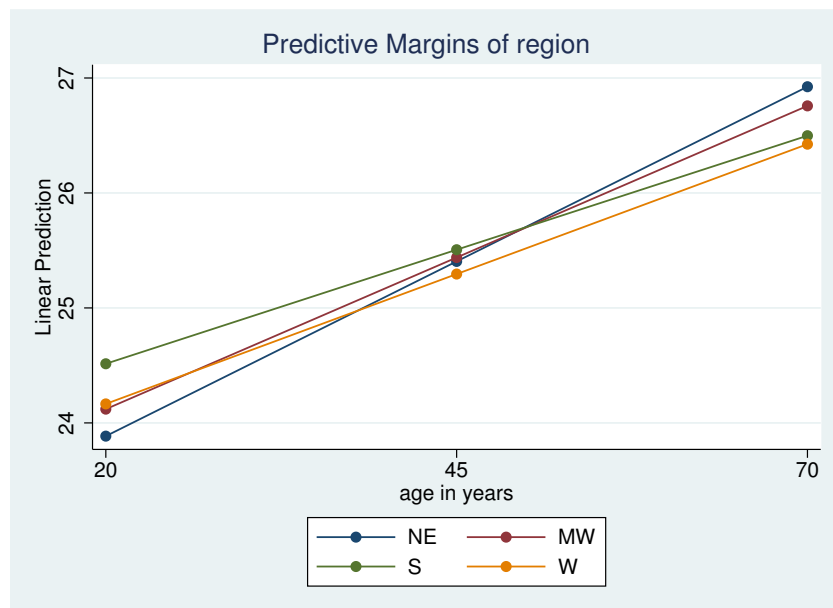
```
. marginsplot
```



### Suppressing Confidence Intervals

- The confidence intervals can make the graph appear messy; we can suppress them

```
. marginsplot, noci
```



- ◇ This is dangerous because it makes the predictions look more precise than they are

## Testing for Differences

- We might want to perform tests of differences at different levels of the continuous variable
- To obtain tests of differences between levels of region at each level of age

```
. margins region, at(age=(20(10)70)) vsquish contrast
```

```
Contrasts of predictive margins
Model VCE      : OLS
```

```
Expression    : Linear prediction, predict()
1._at        : age          =          20
2._at        : age          =          30
3._at        : age          =          40
4._at        : age          =          50
5._at        : age          =          60
6._at        : age          =          70
```

	df	F	P>F
region@_at			
1	3	1.94	0.1200
2	3	1.59	0.1884
3	3	1.06	0.3642
4	3	0.93	0.4251
5	3	1.56	0.1974
6	3	2.05	0.1041
Joint	6	1.69	0.1193
Denominator	10342		

## Predicted Values Over Groups

- As with *marginslist*, when we specify `at()` Stata calculates predicted values treating each case as though they belong to each group or combination of values
- As before, we can use the `over()` option after models with categorical by continuous interactions
- For example, to obtain predicted values for each region using the observed values of `female` and `age` in that region

```
. margins, over(region)
```

```
Predictive margins          Number of obs      =      10,351
Model VCE      : OLS
```

```
Expression    : Linear prediction, predict()
over          : region
```

	Margin	Delta-method Std. Err.	t	P> t	[95% Conf. Interval]
region					
NE	25.57535	.1057592	241.83	0.000	25.36804 25.78266
MW	25.51936	.0919307	277.59	0.000	25.33916 25.69956
S	25.63317	.090649	282.77	0.000	25.45548 25.81086
W	25.42299	.0944498	269.17	0.000	25.23785 25.60813

### 3.4 Investigating Continuous by Continuous Interactions

#### A Continuous by Continuous Interaction

- For this example we'll use a similar model for bmi but we'll add a main effect of serum vitamin c (vitaminc), and an interaction between age and vitaminc
- Before we fit the model, let's take a closer look at vitaminc

```
. summ vitaminc, detail
```

serum vitamin C (mg/dL)					
Percentiles		Smallest			
1%	.2	.1			
5%	.3	.1			
10%	.3	.1	Obs	9,973	
25%	.6	.1	Sum of Wgt.	9,973	
50%	1		Mean	1.034814	
		Largest	Std. Dev.	.5813791	
75%	1.4	8.3			
90%	1.7	9.4	Variance	.3380017	
95%	1.9	13.9	Skewness	4.539869	
99%	2.4	18.1	Kurtosis	108.2617	

◇ The distribution has a long tail, but most observations are between .2 and 2.

- Now lets fit the model

```
. regress bmi c.age##c.vitaminc i.female i.region
```

Source	SS	df	MS	Number of obs	=	9,973
Model	10298.9223	7	1471.27461	F(7, 9965)	=	63.61
Residual	230479.207	9,965	23.1288718	Prob > F	=	0.0000
				R-squared	=	0.0428
				Adj R-squared	=	0.0421
Total	240778.13	9,972	24.1454201	Root MSE	=	4.8092

bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
age	.0220407	.0059366	3.71	0.000	.0104038 .0336777
vitaminc	-2.331426	.2717928	-8.58	0.000	-2.864194 -1.798657
c.age#c.vitaminc	.029107	.0050017	5.82	0.000	.0193026 .0389115
female					
0	0	(base)			
1	.1858965	.0982311	1.89	0.058	-.0066564 .3784494
region					
NE	0	(base)			
MW	-.0936871	.1412331	-0.66	0.507	-.3705326 .1831584
S	-.2137082	.1431247	-1.49	0.135	-.4942615 .0668451
W	-.1626738	.1430181	-1.14	0.255	-.4430182 .1176706
_cons	25.45695	.3293507	77.29	0.000	24.81136 26.10255

- We can replay the model using `coeflegend`

```
. regress, coeflegend
```

### Estimating Slopes

- We can use `lincom` to calculate the slope for `vitaminc` when `age=49` (it's median)

```
. lincom vitaminc + c.vitaminc#c.age*49
```

```
( 1) vitaminc + 49*c.age#c.vitaminc = 0
```

	bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
(1)		-.9051806	.0870624	-10.40	0.000	-1.075841 - .7345206

- We could also calculate the slope of `age` when `vitaminc=1` (it's median)

```
. lincom age + c.vitaminc#c.age*1
```

```
( 1) age + c.age#c.vitaminc = 0
```

	bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
(1)		.0511478	.0028212	18.13	0.000	.0456176 .0566779

- `margins` can produce estimates of the slopes for a range of values

```
. margins, dydx(vitaminc) at(age=(20(10)70)) vsquish
```

```
Average marginal effects          Number of obs   =   9,973
Model VCE      : OLS
```

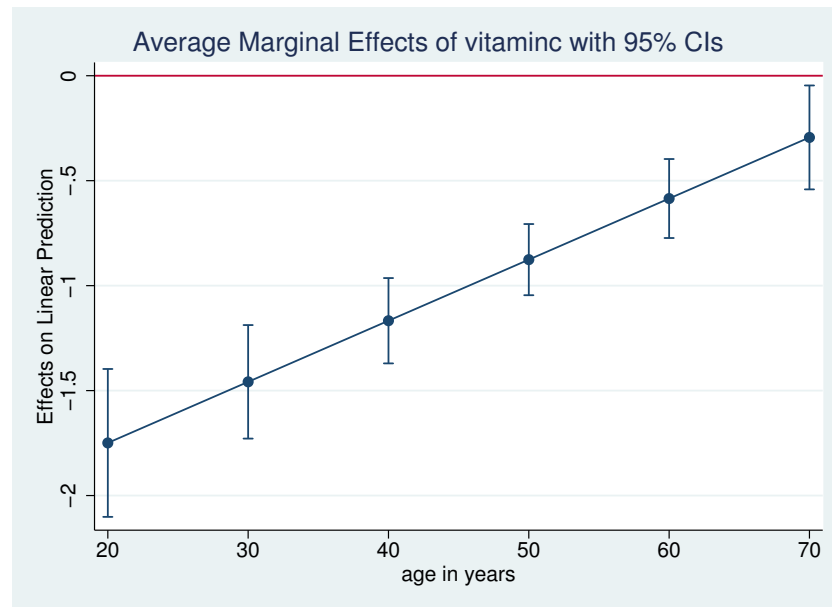
```
Expression      : Linear prediction, predict()
dy/dx w.r.t.    : vitaminc
1._at           : age                =       20
2._at           : age                =       30
3._at           : age                =       40
4._at           : age                =       50
5._at           : age                =       60
6._at           : age                =       70
```

		Delta-method				
		dy/dx	Std. Err.	t	P> t	[95% Conf. Interval]
vitaminc						
	_at					
	1	-1.749285	.1797317	-9.73	0.000	-2.101595 -1.396974
	2	-1.458214	.1379304	-10.57	0.000	-1.728586 -1.187843
	3	-1.167144	.1036802	-11.26	0.000	-1.370378 -.9639098
	4	-.8760735	.0864746	-10.13	0.000	-1.045581 -.7065659
	5	-.5850031	.0959667	-6.10	0.000	-.7731173 -.3968889
	6	-.2939327	.126273	-2.33	0.020	-.5414532 -.0464122

## Graphing Slopes

- We can graph the slopes of vitaminc across age

```
. marginsplot, yline(0)
```



## Predicted Values

- Specifying multiple variables in the `at()` option results in predictions at each combination of values

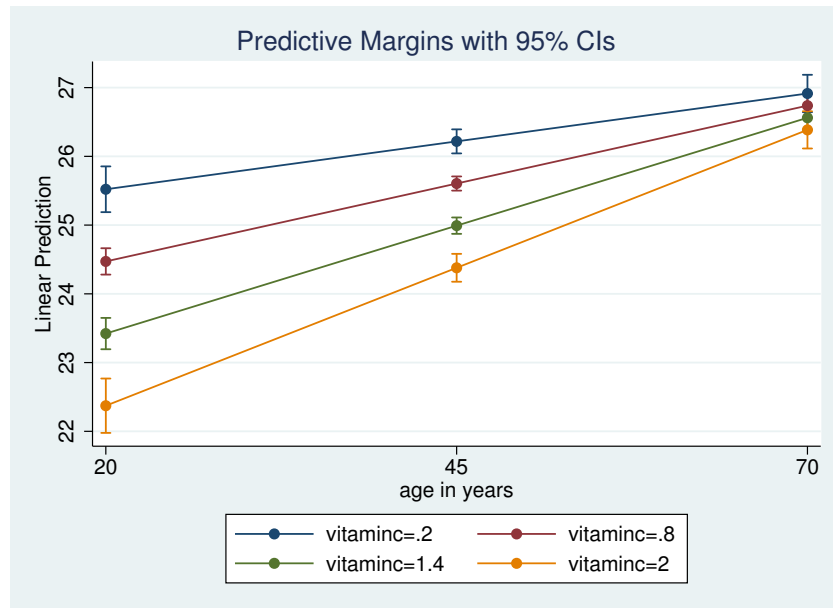
```
. margins , at(age=(20(25)70) vitaminc=(.2(.6)2)) vsquish
```

```
Predictive margins          Number of obs   =       9,973
Model VCE      : OLS
```

```
Expression   : Linear prediction, predict()
1._at       : age           =       20
              vitaminc      =       .2
2._at       : age           =       20
              vitaminc      =       .8
3._at       : age           =       20
              vitaminc      =      1.4
4._at       : age           =       20
              vitaminc      =       2
5._at       : age           =      45
              vitaminc      =       .2
6._at       : age           =      45
              vitaminc      =       .8
7._at       : age           =      45
              vitaminc      =      1.4
8._at       : age           =      45
              vitaminc      =       2
9._at       : age           =      70
              vitaminc      =       .2
10._at      : age           =      70
              vitaminc      =       .8
11._at      : age           =      70
              vitaminc      =      1.4
12._at      : age           =      70
              vitaminc      =       2
```

-----							
		Delta-method				[95% Conf. Interval]	
	Margin	Std. Err.	t	P> t			
-----							
_at							
1	25.52113	.1698638	150.24	0.000	25.18816	25.8541	
2	24.47156	.097744	250.36	0.000	24.27996	24.66316	
3	23.42199	.1162436	201.49	0.000	23.19413	23.64985	
4	22.37242	.2018162	110.86	0.000	21.97682	22.76802	
5	26.21768	.0891689	294.02	0.000	26.04289	26.39247	
6	25.60472	.0525344	487.39	0.000	25.50174	25.70769	
7	24.99175	.0606647	411.97	0.000	24.87284	25.11067	
8	24.37879	.1034993	235.55	0.000	24.17591	24.58167	
9	26.91423	.1388456	193.84	0.000	26.64207	27.1864	
10	26.73788	.0879343	304.07	0.000	26.56551	26.91024	
11	26.56152	.0875619	303.35	0.000	26.38988	26.73315	
12	26.38516	.1381377	191.01	0.000	26.11438	26.65593	
-----							

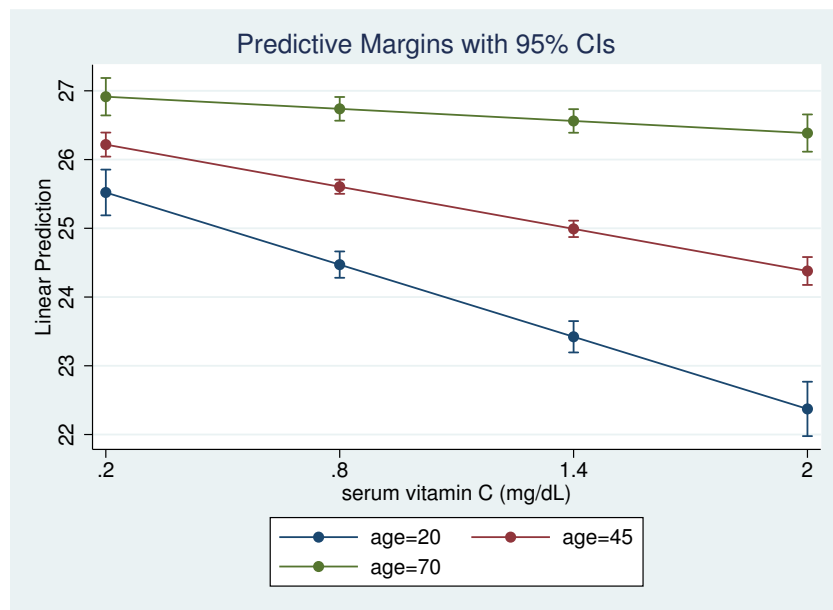
```
. marginsplot
```



### Changing the X-axis Variable

- We can select which variable appears on the x-axis using the `xdimension()` option

```
. marginsplot, xdimension(vitaminc)
```





## Models with Polynomial Terms

- We'll start by fitting a model that includes age and age<sup>2</sup>

```
. regress bmi c.age##c.age i.female i.region
```

Source	SS	df	MS	Number of obs	=	10,351
Model	10269.3919	6	1711.56532	F(6, 10344)	=	73.84
Residual	239754.77	10,344	23.1781487	Prob > F	=	0.0000
				R-squared	=	0.0411
				Adj R-squared	=	0.0405
Total	250024.162	10,350	24.1569239	Root MSE	=	4.8144

bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
age	.2731368	.0203077	13.45	0.000	.2333297 .3129439
c.age#c.age	-.0024099	.0002162	-11.15	0.000	-.0028337 -.0019861
female					
0	0	(base)			
1	.0462855	.0947764	0.49	0.625	-.1394945 .2320656
region					
NE	0	(base)			
MW	.0322091	.1394036	0.23	0.817	-.2410489 .3054671
S	.0289346	.1385186	0.21	0.835	-.2425886 .3004579
W	-.1105093	.1410448	-0.78	0.433	-.3869844 .1659657
_cons	18.6987	.4416971	42.33	0.000	17.83289 19.56451

- Graphs can be particularly useful in understanding models with polynomial terms
- Here we predict values of bmi at different values of age

```
. margins, at(age=(20(10)70)) vsquish
```

Predictive margins	Number of obs	=	10,351
Model VCE : OLS			
Expression : Linear prediction, predict()			
1._at : age = 20			
2._at : age = 30			
3._at : age = 40			
4._at : age = 50			
5._at : age = 60			
6._at : age = 70			

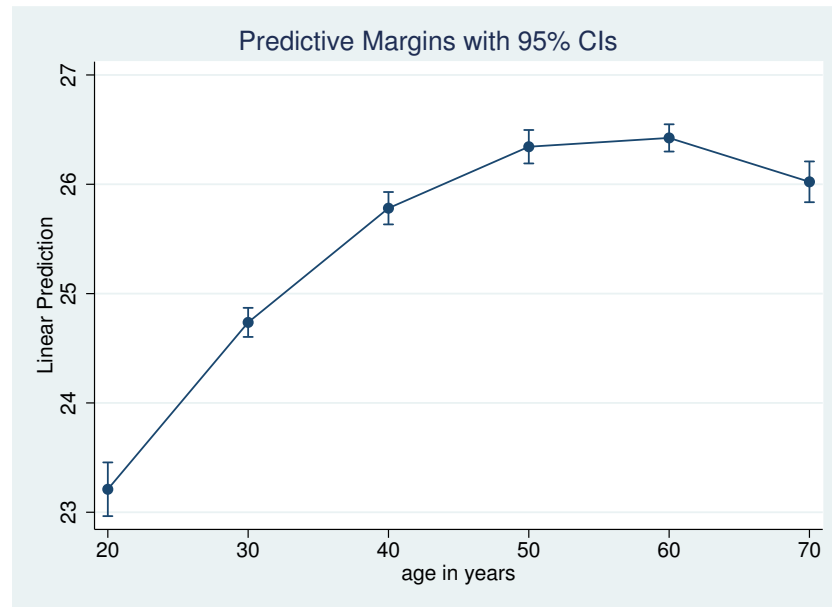
  

	Delta-method				
	Margin	Std. Err.	t	P> t	[95% Conf. Interval]
_at					
1	23.21033	.1253478	185.17	0.000	22.96462 23.45604
2	24.73675	.0678653	364.50	0.000	24.60372 24.86977
3	25.78118	.0755647	341.18	0.000	25.63306 25.9293
4	26.34363	.0780441	337.55	0.000	26.19065 26.49661
5	26.4241	.0635204	415.99	0.000	26.29959 26.54861
6	26.02259	.0951272	273.56	0.000	25.83612 26.20905

## Graphing Predicted Values

- And graph the predictions

```
. marginsplot
```



## Slopes

- We can also obtain estimates of the slope of age across its range
- To do so we'll include age in both the dyed() and at() options

```
. margins, dydx(age) at(age=(20(10)70)) vsquish
```

```
Average marginal effects          Number of obs    =    10,351
Model VCE      : OLS
```

```
Expression      : Linear prediction, predict()
dy/dx w.r.t.    : age
1._at          : age                =        20
2._at          : age                =        30
3._at          : age                =        40
4._at          : age                =        50
5._at          : age                =        60
6._at          : age                =        70
```

		Delta-method				
		dy/dx	Std. Err.	t	P> t	[95% Conf. Interval]
age						
	_at					
	1	.1767405	.0117968	14.98	0.000	.1536164 .1998646
	2	.1285424	.0076583	16.78	0.000	.1135307 .143554
	3	.0803442	.0039415	20.38	0.000	.0726181 .0880703
	4	.032146	.0031343	10.26	0.000	.0260022 .0382899
	5	-.0160521	.0064432	-2.49	0.013	-.028682 -.0034222
	6	-.0642503	.010517	-6.11	0.000	-.0848657 -.0436349

## Adding a Cubic Term

- The same process can be used with higher order polynomials, here we add a cubic term for age

```
. regress bmi c.age##c.age##c.age i.female i.region
```

Source	SS	df	MS	Number of obs	=	10,351
Model	10422.3157	7	1488.90224	F(7, 10343)	=	64.27
Residual	239601.846	10,343	23.1656044	Prob > F	=	0.0000
				R-squared	=	0.0417
				Adj R-squared	=	0.0410
Total	250024.162	10,350	24.1569239	Root MSE	=	4.8131

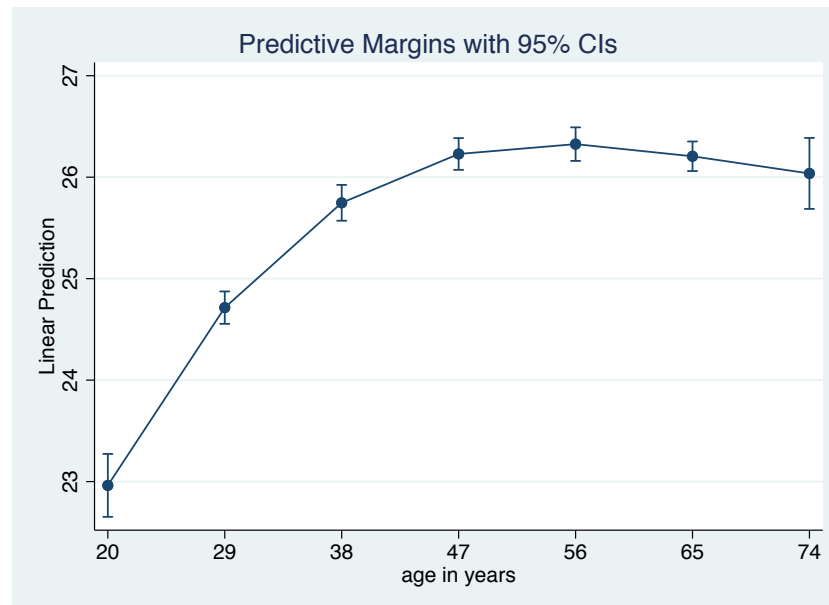
bmi	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
age	.5056311	.0927387	5.45	0.000	.3238453 .6874169
c.age#c.age	-.0077683	.0020967	-3.70	0.000	-.0118782 -.0036583
c.age#c.age#c.age	.0000383	.0000149	2.57	0.010	9.07e-06 .0000675
female					
0	0	(base)			
1	.0449127	.0947522	0.47	0.636	-.1408201 .2306454
region					
NE	0	(base)			
MW	.0274302	.1393783	0.20	0.844	-.2457782 .3006386
S	.025305	.1384883	0.18	0.855	-.2461589 .2967689
W	-.1172832	.1410312	-0.83	0.406	-.3937317 .1591653
_cons	15.6426	1.268785	12.33	0.000	13.15554 18.12967



## Graphing the Cubic Term

- And we can easily graph this as well

```
. marginsplot
```



---

## 4 Conclusion

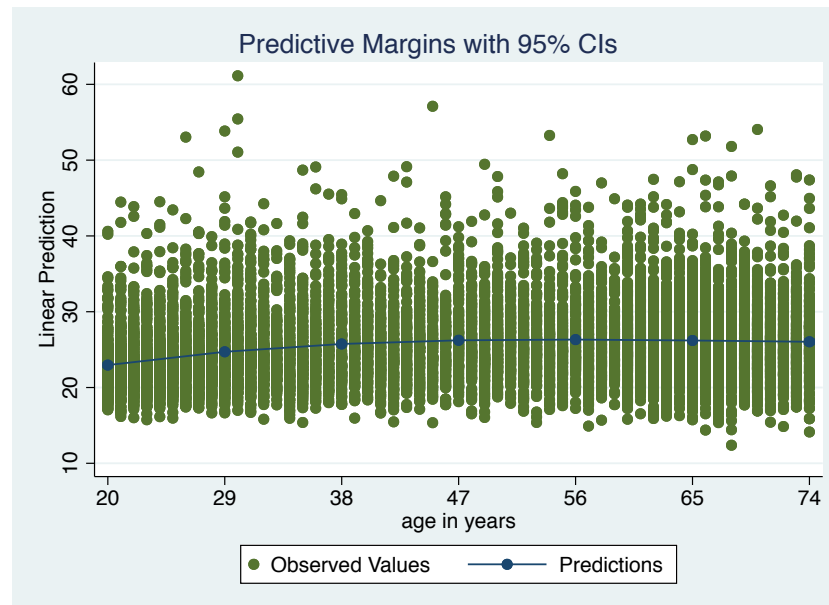
### 4.1 Graphing Extras

#### Adding Additional Plots

- We can add other types of twoway plots to the plots drawn by marginsplots
  - ◇ Continuing with our cubic example
- The addplot option allows us to add additional plots to our marginsplots
- We do want to be careful about the order in which graphs are drawn, we usually want the most dense graphs, for example individual data points, drawn first
  - ◇ Specifying `addplot(..., below)` draws the added plot below the marginsplot

## Adding Observed Data

```
. marginsplot, addplot(scatter bmi age, below ///  
    legend(order(3 "Observed Values" 2 "Predictions")) ///  
    xlabel(20(9)74))
```



- Note: The confidence intervals are in the plot, they're just small relative to the scale of the y-axis, so they're hard to see.

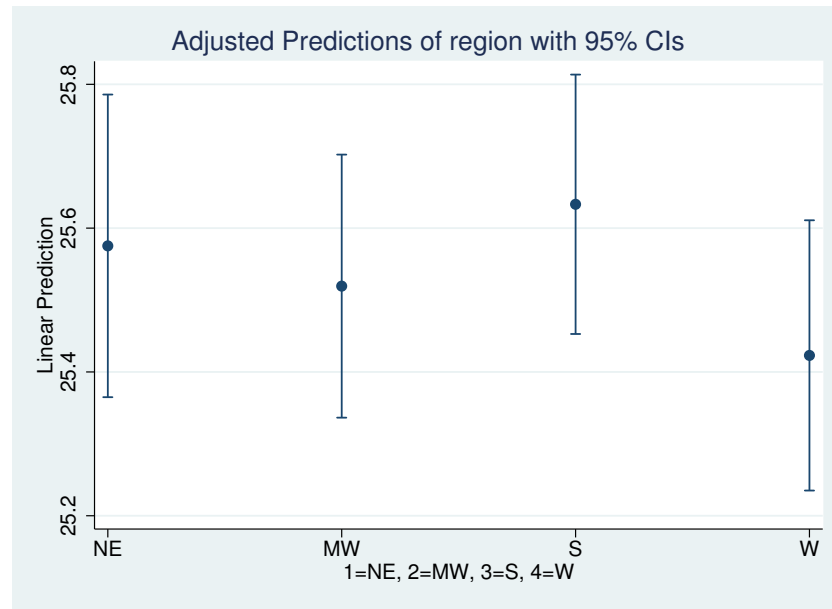
## Changing the Plot Type

- We can change the plots drawn by `marginsplot` to another twoway plot type
  - ◇ See `help twoway` for a list
- The `recast()` option changes the plot for the predictions
  - ◇ `recastci()` changes how the CIs are plotted



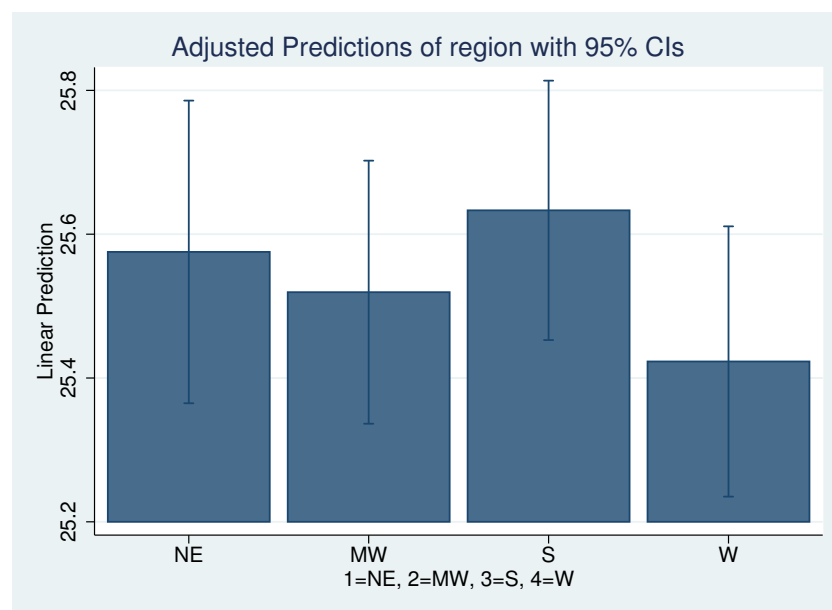
## Estimates as a Scatterplot

```
. marginsplot, recast(scatter)
```



## Estimates as a Bar plot

```
. marginsplot, recast(bar) plotopts(barwidth(.9))
```



- ◇ The `plotopts()` option allows you to specify options for the plots
- ◇ `barwidth()` specifies the width of the bars in units of the x variable



## 4.2 Conclusion

### Conclusion

- We've seen how to fit models that include interactions
  - We've learned how to use Stata's postestimation tools to explore the resulting models
  - We've learned how to graph predictions and how to modify those graphs
-

# Index