Local Maxima in the Estimation of the ZINB and Sample Selection models

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23rd London Stata Users Group Meeting 7 September 2017

- Maximum likelihood (ML) estimators have many desirable properties.
- However, ML estimators also have problems:
 - 1 The ML estimator may not exist;
 - 2 The likelihood function may have multiple maxima.
- Stata makes available many ML estimators to users that may not be aware of these potential problems.

- Non-existence issues are reasonably well understood and solutions are available.
- For example:
 - 1 Stata deals well with non-existence issues in the logit/probit;
 - The user-written ppml command deals with non-existence issues in Poisson regression;
 - 3 A similar issue exists with other estimators (e.g., Tobit) and ppml can be used to address some of these.

- The existence of multiple optima has received less attention.
- This is perhaps because the issue does not arise in some leading cases (Poisson, logit, probit, Tobit).
- However, the existence of multiple (local) maxima is a problem for many frequently used estimators.
- In this presentation I'll focus on two important examples, but there may be many others.

- This is one of the most used (abused?) estimators in applied economics.
- Olsen (1982) shows that the log-likelihood function of the sample selection estimator has a unique maximum for fixed values of ρ .
- However, when ρ has to be estimated, the log-likelihood is not globally concave and multiple maxima may exist.
- Olsen (1982) suggested that estimation should start with a grid search over ρ ; I believe Stata does not do that.

• Consider the following DGP:

$$egin{aligned} y &= 15 + x_1 - x_2 + (\kappa u_1 + u_2) \ / \ ig(1 + \kappa^2ig)^{0.5} \ y \ ext{is observed if} \ ig(1 + x_1 - x_2 + u_1ig) > 0 \ x_1 &\sim U\left(0,1
ight), \quad x_2 &\sim B\left(1,0.3
ight), \quad u_i &\sim N\left(0,1
ight) \end{aligned}$$

- The parameter κ controls the correlation between the errors of the two equations: $\rho = \kappa / \sqrt{(1 + \kappa^2)}$.
- I performed some simulations for different sample sizes and for different values of κ .
- Estimation was performed either using the default method or using as the starting values the ML estimates with the sign of ρ switched.

Table 1: Simulation results for the heckman command

n	250			1000		
κ	-2	0	2	-2	0	2
Both converged	959	999	951	1000	1000	1000
Alternative is better	151	37	125	58	9	58
Default is better	325	123	290	456	75	449

NB: results are considered different if the log-likelihoods differ by more than 0.1.

- Results based on 1000 replicas.
- None of the methods dominates the other.
- The existence of multiple maxima is an issue, especially with small samples.
- The differences between the results can be substantial.

3. The zinb command

- The zero-inflated negative binomial estimator is also very popular.
- Part of the reason for its popularity is due to misconceptions about overdispersion and to results of Vuong's test reported by Stata.
- Unfortunately:
 - zinb often converges to local maxima of the likelihood function.
 - Vuong's test as reported by Stata is not valid in this context.
- Next I use a small simulation to illustrate the existence of multiple maxima in the zinb.

• Consider the following DGP:

$$\begin{array}{rcl} y^{*} & \sim & \mathrm{Poisson}\left(\mu\right) \\ \mu & = & \exp\left(1 + x_{1} - x_{2}\right)\eta \\ y & = & y^{*} \times I\left(u > \frac{\exp\left(\kappa + x_{1} - x_{2}\right)}{1 + \exp\left(\kappa + x_{1} - x_{2}\right)}\right) \\ x_{1} & \sim & U\left(0, 1\right), \quad x_{2} \sim B\left(1, 0.3\right), \\ \eta & \sim & \Gamma\left(1, 1\right), \quad u \sim U\left(0, 1\right) \end{array}$$

- So, *y* is generated by a ZINB and the probability of zero inflation increases with *κ*.
- I performed 1000 simulations for κ ∈ {-∞, -2, -1}; these correspond to zero-inflation probabilities of about 0, 0.15, and 0.32.

- Estimation is performed using two different approaches:
 - The default (start by estimating a model where µ is constant and then estimate the full model);
 - **2** Estimate the ZINB starting form the nbreg estimates.

n	250			1000		
κ	$-\infty$	-2	-1	$-\infty$	-2	-1
Both converged	747	871	957	764	924	990
Alternative is better	103	179	50	133	271	9
Default is better	46	17	3	49	0	0

Table 2: Simulation results for the zinb command

NB: results are considered different if the log-likelihoods differ by more than 0.1.

- Like before, no method dominates and the existence of multiple maxima is an issue.
- Again, in some cases the differences are substantial.

4. Vuong's test

- Vuong (1989) presents model selection tests that can be applied to nested, non-nested, and overlapping models.
 - For nested models, Vuong's test coincides with the classical LR test.
 - For overlapping models, Vuong's test is based on a statistic that under the null is distributed as a weighted sum of χ^2 random variables.
 - For strictly non-nested models, Vuong's test is directional and is based on a statistic that under the null has a normal distribution.
- For non-nested models, Vuong's test is very different from the tests for non-nested hypotheses inspired by Cox (1961).

- Stata implements Vuong's test for non-nested model to test for zero-inflation (ZINB vs NB and ZIP vs Poisson).
- However, the competing models are overlapping, not non-nested.
- This problem has been noted by Santos Silva, Tenreyro, and Windmeijer (2015) and Wilson (2015).
- The results of the test can be very misleading.
- For example, if the data is generated by a NB process, the test of the Poisson vs the ZIP will never favour the Poisson model and generally favours the ZIP.

- Multiple maxima in ML can be a serious problem.
- It would be great if Stata could do more to deal with this.
- At least tnbr is also affected by this problem.
- The current vuong option should be removed from zip and zinb.

References

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