Motivation	Two-stage estimation	Stata syntax	Example	Conclusion
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xtseqreg: Sequential (two-stage) estimation of linear panel data models and some pitfalls in the estimation of dynamic panel models

Sebastian Kripfganz

University of Exeter Business School, Department of Economics, Exeter, UK

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net install xtseqreg, from(http://www.kripfganz.de/stata/) Or ssc install xtseqreg



- In many applications, important determinants of the outcome variable can be time invariant.
 - Education, gender, nationality, ethnic and religious background, and other individual-specific characteristics play important roles in the determination of labor market or health outcomes.
 - Institutional, socio-economic, and geographic factors matter in convergence models of economic growth, and they are key variables in gravity models of international trade and investment flows.
- A researcher might be particularly interested in their effects. Yet, traditional "fixed-effects" procedures (xtreg, fe) wipe out all time-invariant variables from the model.



- To identify the coefficients of time-invariant regressors, the assumption that a sufficient number of regressors (or excluded instrumental variables) is uncorrelated with the unit-specific error component cannot be avoided.
- Identification strategies for static panel models include:
 - Classical "random-effects" model: xtreg, re,
 - "Correlated random-effects" (Mundlak, 1978; Chamberlain, 1982) or "hybrid" models (Allison, 2009; Schunck, 2013): xthybrid (Schunck and Perales, 2017),
 - Hausman and Taylor (1981) model: xthtaylor,
 - Other instrumental variables strategies: xtivreg.



- In the context of dynamic panel models, generalized method of moments (GMM) estimators (Arellano and Bover, 1995; Blundell and Bond, 1998) are frequently employed: xtdpd, xtdpdsys, and xtabond2 (Roodman, 2009).
- Incorrect assumptions about the exogeneity of some variables may cause inconsistency of all coefficient estimates.
- A sequential procedure can provide partial robustness to such misspecification. In a first stage, only the coefficients of time-varying regressors are estimated. In a second stage, the coefficients of time-invariant regressors are recovered.
- \Rightarrow New Stata command: xtseqreg



• Linear panel data model with time-invariant regressors and error-components structure:

$$y_{it} = \mathbf{x}'_{it}\boldsymbol{\beta} + \mathbf{f}'_i\boldsymbol{\gamma} + u_i + e_{it}$$

• Sequential estimation procedure:

Estimation of the coefficients of time-varying regressors:

$$y_{it} = \mathbf{x}'_{it}\boldsymbol{\beta} + \tilde{u}_i + e_{it}, \quad \tilde{u}_i = \mathbf{f}'_i\boldsymbol{\gamma} + u_i$$

Stimation of the coefficients of time-invariant regressors:

$$y_{it} - \mathbf{x}'_{it}\hat{\boldsymbol{\beta}} = \mathbf{f}'_i \boldsymbol{\gamma} + u_i + \tilde{e}_{it}, \quad \tilde{e}_{it} = e_{it} - \mathbf{x}'_{it}(\hat{\boldsymbol{\beta}} - \boldsymbol{\beta})$$

- Conventional standard errors at the second stage are incorrect and often far too small.
- \Rightarrow xtseqreg computes proper standard errors with the analytical correction term derived by Kripfganz and Schwarz (2015).

Motivation 000	Two-stage estimation O	Example 000000000000	Conclusion
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Stata syntax of the xtseqreg command

Syntax

xtseqreg depvar [(indepvars1)] [indepvars2] [if] [in] [, options]

options	Description
Model	
<pre>first(first_spec)</pre>	specify first-stage estimation results
both	estimate both stages
nocommonesample	do not restrict estimation samples to be the same
iv(iv_spec)	standard instruments; can be specified more than once
<pre>gmmiv(gmmiv_spec)</pre>	GMM-type instruments; can be specified more than once
wmatrix(wmat spec)	specify initial weighting matrix
twostep	compute two-step instead of one-step estimator
teffects	add time effects to the model
noconstant	suppress constant term
SE/Robust	
vce (vcetype)	vcetype may be conventional, ec, or robust
Reporting	
combine	combine the estimation results for both equations
level(#)	set confidence level; default is level(95)
noheader	suppress output header
notable	suppress coefficient table
noomitted	suppress omitted variables

Stata syntax of xtseqreg postestimation commands

Syntax for predict

predict [type] newvar [if] [in] [, xb stdp ue xbu u e equation(eqno)]

predict [type] {stub* newvar1 ... newvarq} [if] [in] , scores

Syntax for estat

Arellano-Bond test for autocorrelated residuals

estat serial [, ar(numlist)]

Hansen's J-test of overidentifying restrictions

estat overid

Difference-in-Hansen test of overidentifying restrictions

estat overid name

Generalized Hausman test for model misspecification

estat hausman name [(varlist)] [, df(#) nonested]

where *name* is a name under which estimation results were stored via estimates store.



Empirical example: distance and FDI

- Estimation of a gravity model for U.S. outward FDI.
- Annual data, 1989–1999, for 341 bilateral industry-level relationships, compiled by Egger and Pfaffermayr (2004).

describe

obs: vars: size:	2,767 13 118,981			Egger and Pfaffermayr (2004, JAE) 8 Aug 2003 03:39
	storage	display	value	
variable name	type	format	label	variable label
ind	byte	%9.0g		industry identifier
codeim	int	%8.0g		country identifier
year	int	%9.0g		year
lrfdi	float	%9.0g		log real outward foreign direct investment
lgdt	float	%9.0g		log bilateral gross domestic product
lsimi	float	%9.0g		log similarity in country size
lrk	float	%9.0g		log relative physical capital endowment
lrh	float	%9.0g		log relative human capital endowment
lrl	float	%9.0g		log relative labor endowment
ldist	float	%9.0g		log geographical distance
lkgdt	float	%9.0g		= lgdt * abs(lrk)
lkldist	float	%9.0g		= ldist * (lrk - lrl)
id	int	%9.0a		group (codeim ind)

Sorted by: id year

Motivation	Two-stage estimation	Stata syntax	Example	Conclusion
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First-stage system GMM estimation

. xtseqreg L(0/1).lrfdi lkldist lgdt lkgdt lsimi lrk lrh lrl, twostep vce(robust) ///

- > gmmiv(L.lrfdi, lag(1 5) collapse model(difference)) ///
- > gmmiv(lkldist lgdt lkgdt lsimi lrk lrh lrl, lag(0 4) collapse model(difference)) ///
- > iv(L.lrfdi, difference model(level)) ///

> iv(lkldist lgdt lkgdt lsimi lrk lrh lrl, difference model(level))

Group variable: id	Number of obs	=	2198
Time variable: year	Number of groups		337
	Obs per group:	min = avg = max =	1 6.522255 10

Number of instruments = 49

(Std. Err. adjusted for clustering on id)

lrfdi	Coef.	WC-Robust Std. Err.	z	₽> z	[95% Conf.	Interval]
lrfdi						
L1.	.8956164	.063313	14.15	0.000	.7715252	1.019708
lkldist	0978499	.1490779	-0.66	0.512	3900371	.1943374
lgdt	1502013	.2320426	-0.65	0.517	6049964	. 3045939
lkgdt	.0072154	.0053281	1.35	0.176	0032276	.0176584
lsimi	.3100215	.2370884	1.31	0.191	1546632	.7747062
lrk	.7471581	1.291878	0.58	0.563	-1.784877	3.279193
lrh	0897363	.1311771	-0.68	0.494	3468386	.1673661
lrl	8973519	1.30242	-0.69	0.491	-3.450048	1.655344
cons	4.926161	5.971464	0.82	0.409	-6.777694	16.63002

. estimates store gmm1

Motivation	Two-stage estimation	Stata syntax	Example	Conclusion
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First-stage	system GMM e	stimation		

. estat serial, ar(1/3)

Arellano-Bond test for au	tocorrelation	of the first-di	fferenced residual	ls
H0: no autocorrelation of	order 1:	z = -7.3012	Prob > z =	0.0000
H0: no autocorrelation of	f order 2:	z = -0.0535	Prob > z =	0.9573
H0: no autocorrelation of	f order 3:	z = -0.3725	Prob > z =	0.7095

. estat overid

Hansen's J-test	chi2(40)	=	45.7042
H0: overidentifying restrictions are valid	Prob > chi2	=	0.2471

Replication with xtabond2:

```
. xtabond2 L(0/1).lrfdi lkldist lqdt lkqdt lsimi lrk lrh lrh, twostep robust ar(3) ///
> gmm(lrfdi, lag(2 6) collapse equation(diff)) ///
> gmm(lkldist lgdt lkgdt lsimi lrk lrh lrl, lag(0 4) collapse equation(diff)) ///
> iv(LD.lrfdi, equation(level) mz) ///
> iv(D.lkldist D.lgdt D.lkgdt D.lsimi D.lrk D.lrh D.lrl, equation(level) mz)
```

```
Arellano-Bond test for AR(1) in first differences: z = -6.69 Pr > z = 0.000 Arellano-Bond test for AR(2) in first differences: z = -0.05 Pr > z = 0.957 Arellano-Bond test for AR(3) in first differences: z = -0.37 Pr > z = 0.792 Pr
```

```
Sargan test of overid. restrictions: chi2(40) = 80.12 Prob > chi2 = 0.000
(Not robust, but not weakened by many instruments.)
Hansen test of overid. restrictions: chi2(40) = 45.70 Prob > chi2 = 0.247
(Robust, but weakened by many instruments.)
```



• The first two specifications yield identical estimation results. The results from the last specification differ (but should not):

```
. xtseqreg 1(0/1).lrfdi lkldist lgdt lsimi lrk lrh lrl, twostep vce(robust) ///
> gmmiv(lkldist lgdt lkgdt lsimi lrk lrh lrl, lag(0 4) collapse model(difference)) ///
> iv(L.lrfdi, difference model(level)) ///
> iv(lkldist lgdt lkgdt lsimi lrk lrh lrl, difference model(level))

. xtabond2 L(0/1).lrfdi lkldist lgdt lkgdt lsimi lrk lrh lrl, twostep robust ar(3) ///
> gmm(lrfdi, lag(2 6) collapse equation(diff)) ///
> gmm(lrfdi,st lgdt lsimi lrk lrh lrl, lag(0 4) collapse equation(diff)) ///
> iv(D.lrfdi, equation(level) mz) ///
> iv(D.lkldist D.lgdt D.lkgdt D.lsimi D.lrk D.lrh D.lrl, equation(level) mz)
```

- . xtabond2 L(0/1).lrfdi lkldist lgdt lkgdt lsimi lrk lrh lrl, twostep robust ar(3) ///
- > gmm(L.lrfdi, lag(1 5) collapse equation(diff)) ///
- > gmm(lkldist lgdt lkgdt lsimi lrk lrh lrl, lag(0 4) collapse equation(diff)) ///
- > iv(LD.lrfdi, equation(level) mz) ///
- > iv(D.lkldist D.lgdt D.lkgdt D.lsimi D.lrk D.lrh D.lrl, equation(level) mz)

Motivation	Two-stage estimation	Stata syntax	Example	Conclusion
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Second-stage 2SLS estimation

. xtseqreg lrfdi (L.lrfdi lkldist lgdt lkgdt lsimi lrk lrh lrl) ldist, vce(robust) ///
> first(gmm1. nocons) iv(lsimi lrh)

Group variable: id Time variable: year		Number of obs Number of groups	=	2198 337	
 Equation _first			Equation _second		
Number of obs	=	2198	Number of obs	=	2198
Number of groups	=	337	Number of groups	=	337
Obs per group:	min =	1	Obs per group: 1	nin =	1
	avg =	6.522255		vg =	6.522255
	max =	10	I	nax =	10
Number of instrum	ments =	49	Number of instrumer	nts =	3

(Std. Err. adjusted for clustering on id)

lrfdi	Coef.	Robust Std. Err.	z	₽> z	[95% Conf.	Interval]
first						
lrfdi						
L1.	.8956164	.063313	14.15	0.000	.7715252	1.019708
lkldist	0978499	.1490779	-0.66	0.512	3900371	.1943374
lgdt	1502013	.2320426	-0.65	0.517	6049964	. 3045939
lkgdt	.0072154	.0053281	1.35	0.176	0032276	.0176584
lsimi	.3100215	.2370884	1.31	0.191	1546632	.7747062
lrk	.7471581	1.291878	0.58	0.563	-1.784877	3.279193
lrh	0897363	.1311771	-0.68	0.494	3468386	.1673661
lrl	8973519	1.30242	-0.69	0.491	-3.450048	1.655344
second						
- ldist	1213967	.5854263	-0.21	0.836	-1.268811	1.026018
_cons	5.966496	8.5777	0.70	0.487	-10.84549	22.77848

Motivation 000	Two-stage estimation 0	Stata syntax 00	Example ○○○○○○○○○○○○○	Conclusion 00
Second-	stage 2SLS estin	nation		
	. estat overid			
	Hansen's J-test for equation H0: overidentifying restricti		chi2(40) = 45.7042 Prob > chi2 = 0.2471	

Hansen's J-test for equation _second	chi2(1) =	1.1989
H0: overidentifying restrictions are valid	Prob > chi2 =	0.2735

• Replication with ivregress (incorrect standard errors):

. quietly estimates restore gmm1

. quietly predict residuals, ue

. ivregress 2sls residuals (ldist = lsimi lrh), vce(cluster id)

Instrumental variables (2SLS) regression

Number of obs	=	2,198
Wald chi2(1)	=	2.15
Prob > chi2	=	0.1422
R-squared	=	0.0107
Root MSE	=	.46723

(Std. Err. adjusted for 337 clusters in id)

residuals	Coef.	Robust Std. Err.	z	P> z	[95% Conf.	Interval]
ldist _cons	1213967 1.040335	.0827107 .7110881	-1.47 1.46	0.142 0.143	2835066 3533725	.0407132 2.434042
Instrumented: Instruments:	ldist lsimi lrh					

Motivation	Two-stage estimation	Stata syntax	Example	Conclusion
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One-stage GMM estimation

. xtseqreg L(0/1).lrfdi lkldist lgdt lkgdt lsimi lrk lrh lrl ldist, twostep vce(robust) ///

- > gmmiv(L.lrfdi, lag(1 5) collapse model(difference)) ///
- > gmmiv(lkldist lgdt lkgdt lsimi lrk lrh lrl, lag(0 4) collapse model(difference)) ///
- > iv(L.lrfdi, difference model(level)) ///
- > iv(lkldist lgdt lkgdt lsimi lrk lrh lrl, difference model(level)) ///
- > iv(lsimi lrh)

Group variable: id	Number of obs	=	2198
Time variable: year	Number of groups		337
	Obs per group:	min = avg = max =	1 6.522255 10

Number of instruments = 51

(Std. Err. adjusted for clustering on id)

lrfdi	Coef.	WC-Robust Std. Err.	z	₽> z	[95% Conf.	Interval]
lrfdi						
L1.	.874835	.0658537	13.28	0.000	.7457641	1.003906
lkldist	0894573	.1552895	-0.58	0.565	3938191	.2149044
lgdt	100095	.2389068	-0.42	0.675	5683437	.3681537
lkgdt	.0103749	.0053781	1.93	0.054	000166	.0209159
lsimi	.3735686	.2467129	1.51	0.130	1099798	.857117
lrk	.6246915	1.349609	0.46	0.643	-2.020494	3.26987
lrh	0007819	.1125051	-0.01	0.994	2212878	.219724
lrl	7648876	1.37943	-0.55	0.579	-3.468521	1.938746
ldist	0825973	.1385583	-0.60	0.551	3541665	.1889719
_cons	4.320648	6.06585	0.71	0.476	-7.5682	16.2095

. estat hausman gmml (L.lrfdi lkldist lgdt lkgdt lsimi lrk lrh lrl)

Generalized Hausman test	chi2(1) =	4.4792
H0: coefficients do not systematically diffe	er Prob > chi2 =	0.0343



- Instruments for the first-differenced equation are uncorrelated with time-invariant variables by construction, first-differenced instruments for the level equation by assumption.
- ⇒ Difference-in-Hansen tests might be based on asymptotically incorrect (or at least debatable) degrees of freedom:

```
. xtabond2 1(0/1).lrfdi lkldist lqdt lkqdt lsimi lrk lrh lrh lrh ldist, twostep robust ///
> gmm(lrfdi, lag(2 6) collapse equation(diff)) ///
> gmm(lkldist lgdt lsimi lrk lrh lrl, lag(0 4) collapse equation(diff)) ///
> iv(D.lrfdi, equation(level) mz) ///
> iv(D.lrfdist D.lgdt D.lkgdt D.lsimi D.lrk D.lrh D.lrh, equation(level) mz) ///
> iv(lsimi lrh, equation(level) mz)
```

```
iv(lsimi lrh, mz eq(level))
Hansen test excluding group: chi2(39) = 45.95 Prob > chi2 = 0.206
Difference (null H = exogenous): chi2(2) = 2.44 Prob > chi2 = 0.295
```

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Motivation 000	Two-stage estimation 0	Stata syntax 00	Example 000000000000	Conclusion 00
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Alternative first-stage QML estimator

• First-stage QML estimator of Hsiao et al. (2002):

. quietly xtdpdqml lrfdi lkldist lgdt lkgdt lsimi lrk lrh lrl, fe mlparam vce(robust)

. xtseqreg lrfdi (L.1rfdi lkldist lgdt lkgdt lsimi lrk lrh lrl) ldist, vce(robust) /// > first(, nocons) iv(lsimi lrh) noheader

note: first-stage variable names do not match with coefficient list from xtdpdgml note: dependent variable D.lrfdi from xtdpdgml does not match with lrfdi

lrfdi	Coef.	Robust Std. Err.	z	₽> z	[95% Conf.	Interval
first						
lrfdi						
L1.	.8000757	.0539962	14.82	0.000	. 6942451	. 905906
lkldist	7160072	.5053811	-1.42	0.157	-1.706536	.274521
lqdt	.4346637	.1907476	2.28	0.023	.0608052	.808522
lkqdt	.0028906	.0068807	0.42	0.674	0105954	.016376
lsimi	.3172032	.3605734	0.88	0.379	3895076	1.02391
lrk	6.152142	4.400668	1.40	0.162	-2.473009	14.7772
lrh	.0758457	.0869135	0.87	0.383	0945017	.246193
lrl	-5.60704	4.175718	-1.34	0.179	-13.7913	2.57721
second						
ldist	2.41061	2.285819	1.05	0.292	-2.069514	6.89073
_cons	-31.43894	21.15977	-1.49	0.137	-72.91133	10.0334

(Std. Err. adjusted for clustering on id)

. estat overid

Hansen's J-test for	equation _second	chi2(1) =	0.8358
H0: overidentifying	restrictions are valid	Prob > chi2 =	0.3606

Motivation 000	Two-stage estimation O	Stata syntax 00	Example ○○○○○○○○○●○○○	Conclusion

Alternative first-stage GMM estimator

• First-stage GMM estimator of Ahn and Schmidt (1995):

- . quietly xtdpdgmm L(0/1).lrfdi lkldist lgdt lkgdt lsimi lrk lrh lrl, twostep noserial ///
- > vce(robust) aux gmmiv(L.lrfdi, lag(1 5) collapse model(difference)) ///

> gmmiv(lkldist lgdt lkgdt lsimi lrk lrh lrl, lag(0 4) collapse model(difference))

- . xtseqreg lrfdi (L.lrfdi lkldist lgdt lkgdt lsimi lrk lrh lrl) ldist, vce(robust) ///
- > first(, copy) iv(lsimi lrh) noheader

note: first-stage standard errors may not be robust

(Std. Err. adjusted for clustering on id)

lrfdi	Coef.	Robust Std. Err.	z	P> z	[95% Conf.	Interval
first						
lrfdi						
L1.	.8017069	.1204806	6.65	0.000	.5655692	1.03784
lkldist	2290635	.7040092	-0.33	0.745	-1.608896	1.15076
lgdt	0748559	.2905325	-0.26	0.797	6442891	.494577
lkgdt	0186638	.0112666	-1.66	0.098	0407459	.003418
lsimi	.0212282	.3722118	0.06	0.955	7082936	.7507
lrk	1.784527	6.101738	0.29	0.770	-10.17466	13.7437
lrh	.0299533	.1551918	0.19	0.847	2742171	.334123
1r1	-1.580551	6.123368	-0.26	0.796	-13.58213	10.4210
_cons	3.642671	7.335562	0.50	0.619	-10.73477	18.0201
second						
ldist	.3209373	1.580573	0.20	0.839	-2.776928	3.41880
_cons	-2.761592	13.56865	-0.20	0.839	-29.35565	23.8324

. estat overid

Hansen's	J-test for	equation _se	cond	
H0: overi	dentifying	restrictions	are	valid

chi2(1) = 2.7079 Prob > chi2 = 0.0999

Motivation	Two-stage estimation	Stata syntax	Example	Conclusion
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Time of	facto			

l ime effects

. xtsegreg L(0/1).lrfdi, teffects twostep vce(robust) ///

> gmmiv(L.lrfdi, lag(1 5) collapse model(difference)) iv(L.lrfdi, difference model(level))

Group variable: id	Number of obs	=	2198
Time variable: year	Number of groups		337
	Obs per group:	min = avg = max =	1 6.522255 10

Number of instruments = 16

(Std. Err. adjusted for clustering on id)

lrfdi	Coef.	WC-Robust Std. Err.	z	P> z	[95% Conf.	Interval]
lrfdi						
L1.	1.015676	.0727146	13.97	0.000	.8731579	1.158194
year						
1991	0975429	.0419594	-2.32	0.020	1797819	0153039
1992	0670002	.0476785	-1.41	0.160	1604484	.0264479
1993	0945048	.0457007	-2.07	0.039	1840766	0049331
1994	0644637	.0701426	-0.92	0.358	2019406	.0730132
1995	0513381	.0426408	-1.20	0.229	1349125	.0322363
1996	0605227	.0481965	-1.26	0.209	1549861	.0339408
1997	1211606	.0594696	-2.04	0.042	2377189	0046024
1998	1699316	.0552347	-3.08	0.002	2781895	0616736
1999	1261552	.0830178	-1.52	0.129	2888672	.0365568
_cons	.0937689	.3189754	0.29	0.769	5314114	.7189492

. estat overid

Hansen's J-test	chi2(5) =	13.2885
H0: overidentifying restrictions are valid	Prob > chi2 =	0.0208

Motivatio	on Two-stage estimation	Stata syntax	Example	Conclusion
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How (not) to do xtabond2: Beware of the dummy trap!

. xtabond2 L(0/1).lrfdi i.year, twostep robust ///

- > gmm(lrfdi, lag(2 6) collapse equation(diff)) iv(LD.lrfdi, equation(level) mz) ///
- > iv(i.year, equation(level))

lrfdi	Coef.	Corrected Std. Err.	Z	₽> z	[95% Conf	. Interval]
lrfdi						
L1.	1.015676	.0727146	13.97	0.000	.8731579	1.158194
year						
1989	0	(empty)				
1990	.0644637	.0701426	0.92	0.358	0730132	.2019406
1991	0330792	.0597255	-0.55	0.580	150139	.0839805
1992	0025366	.0513121	-0.05	0.961	1031064	.0980333
1993	0300412	.0579887	-0.52	0.604	1436969	.0836146
1994	0	(omitted)				
1995	.0131256	.0551362	0.24	0.812	0949394	.1211905
1996	.003941	.055217	0.07	0.943	1042823	.1121643
1997	056697	.0504278	-1.12	0.261	1555337	.0421398
1998	1054679	.04837	-2.18	0.029	2002714	0106643
1999	0616915	.0540627	-1.14	0.254	1676525	.0442694
_cons	.0293052	. 3703467	0.08	0.937	696561	.7551714

Hansen test of overid. restrictions: chi2(3) = 13.29 Prob > chi2 = 0.004
(Robust, but weakened by many instruments.)



- Instruments for the time dummies should only be included for the level equation. Asymptotically, the additional instruments for the first-differenced equation are redundant.
- \Rightarrow Hansen's J-test is based on incorrect degrees of freedom:

```
. xtabond2 L(0/1).lrfdi i.year, twostep robust ///
> gmm(lrfdi, lag(2 6) collapse equation(diff)) iv(LD.lrfdi, equation(level) mz) ///
> iv(i.year, equation(diff)) iv(i.year, equation(level))
Hansen test of overid. restrictions: chi2(12) = 14.82 Prob > chi2 = 0.252
(Robust, but weakened by many instruments.)
```

 Never use the iv() option without suboption equation()! It is not equivalent to the joint specification of iv(, equation(diff)) and iv(, equation(level)):

```
. xtabond2 L(0/1).lrfdi i.year, twostep robust ///
> gmm(lrfdi, lag(2 6) collapse equation(diff)) iv(LD.lrfdi, equation(level) mz) ///
> iv(i.year)
Hansen test of overid. restrictions: chi2(3) = 10.79 Prob > chi2 = 0.013
```

```
(Robust, but weakened by many instruments.)
```



- Sequential estimation can provide partial robustness to model misspecification.
- Is is important to compute corrected standard errors at the second stage that account for the first-stage estimation error.
- The new xtseqreg Stata command implements this standard error correction for two-stage linear panel data models.
- The two-stage procedure is particularly relevant in the presence of time-invariant regressors, but it can be easily applied to more general settings.

Kripfganz, S., and C. Schwarz (2015). Estimation of linear dynamic panel data models with time-invariant regressors. *ECB Working Paper 1838.* European Central Bank.

```
net install xtseqreg, from(http://www.kripfganz.de/stata/) Of ssc install xtseqreg
help xtseqreg
help xtseqreg postestimation
```

Motivation	Two-stage estimation	Stata syntax	Example	Conclusion
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22/22