Extended multivariate generalised linear and non-linear mixed effects models

Stata UK Meeting Cass Business School 7th September 2017

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Outline

- Motivation for this work
- Extended multivariate generalised linear and non-linear mixed effects models
- megenreg
- Methods development using megenreg
- Future directions

- More data \rightarrow more questions

need for appropriate statistical modelling techniques, and implementations $% \label{eq:constraint}$

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• Growth in access to EHR

biomarkers < patients < GP practice area < geographical regions...

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- More challenges
 - time-dependent effects, non-linear covariate effects

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• More challenges

time-dependent effects, non-linear covariate effects

We need modelling frameworks that can accommodate a lot of different things

Joint longitudinal-survival models



Linking via - current value, gradient, AUC, random effects...

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Joint longitudinal-survival models - extensions

- Competing risks [1]
- Different types of outcomes [2]
- Multiple continuous outcomes [3]
- Delayed entry [4]
- Recurrent events and a terminal event [5]
- Prediction [6]
- Many others...

Joint longitudinal-survival models - software

- stjm in Stata [7]
- gsem in Stata, see Yulia's talk from last year
- frailtypack in R [8]
- joineR in R [9]
- JM and JMBayes in R [10, 11]
- Many others...

(My) Methods development - software

- stjm joint longitudinal-survival models
- stmixed multilevel survival models
- stgenreg general parametric survival models

• ...

(My) Methods development - software

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Each new project brings a new code base to maintain...could I make my life easier?

A general framework for the analysis of data of all types

- Multiple outcomes of varying types
- Measurement schedule can vary across outcomes
- Any number of levels and random effects
- Sharing and linking random effects between outcomes
- Sharing functions of the expected value of other outcomes
- A reliable estimation engine
- Easily extendable by the user

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I think I made my life more difficult!

Extended multivariate generalised linear and non-linear mixed effects models

megenreg

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Extended multivariate generalised linear and non-linear mixed effects models

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- Much of what megenreg can do, cannot be done with gsem

A general level likelihood

Straight from the Stata manual...for a one-level model with \boldsymbol{n} response variables:

$$p(\boldsymbol{y}|\boldsymbol{x}, \boldsymbol{b}, \boldsymbol{\beta}) = \prod_{i=1}^{n} p_i(y_i|\boldsymbol{x}, \boldsymbol{b}, \boldsymbol{\beta})$$

For a two-level model:

$$p(\boldsymbol{y}|\boldsymbol{x}, \boldsymbol{b}, \boldsymbol{\beta}) = \prod_{i=1}^{n} \prod_{j=1}^{t} p_i(y_{ij}|\boldsymbol{x}, \boldsymbol{b}, \boldsymbol{\beta})$$

A general level likelihood

The log likelihood is obtained by integrating out the unobserved random effects

$$ll(\boldsymbol{\beta}) = \log \int_{\mathcal{R}^r} p(\boldsymbol{y}|\boldsymbol{x}, \boldsymbol{b}, \boldsymbol{\beta}) \phi(\boldsymbol{b}|\boldsymbol{\Sigma}_{\boldsymbol{b}}) \ \mathsf{d}\boldsymbol{b}$$

we assume $\phi()$ is the multivariate normal density for \boldsymbol{b} , with mean vector $\boldsymbol{0}$ and variance-covariance matrix $\Sigma_{\boldsymbol{b}}$. We have $\Sigma_{\boldsymbol{b}}$ becoming block diagonal with further levels, with a block for each level

A general level likelihood

Alternatively, exploiting conditional independence amongst level l-1 units, given the random effects at higher levels,

$$ll(\boldsymbol{\beta}) = \log \int \phi(\boldsymbol{b}^{(L)} | \boldsymbol{\Sigma}^{(L)}) \prod p^{(L-1)}(\boldsymbol{y} | \boldsymbol{x}, \boldsymbol{b}^{L}, \boldsymbol{\beta}) \, d\boldsymbol{b}^{(L)}$$

where, for $l=2,\ldots,L$

$$p^{(l)}(\boldsymbol{y}|\boldsymbol{x}, \boldsymbol{B}^{l+1}, \boldsymbol{\beta}) = \int \phi(\boldsymbol{b}^{(l)}|\boldsymbol{\Sigma}^{(l)}) \prod p^{(l-1)}(\boldsymbol{y}|\boldsymbol{x}, \boldsymbol{B}^{l}, \boldsymbol{\beta}) \, \mathrm{d}\boldsymbol{b}^{(l)}$$

Estimation challenges

- At each level, we need to integrate out our normally distributed random effects
- Generally this is done using Gauss-Hermite numerical quadrature

intmethod(mvaghermite | ghermite)

- Issue with GH quadrature is it doesn't scale up well:
 - 7-point quadrature; for 1 random effect we evaluate our function at 7-points
 - 7-point quadrature; for 6 random effects, we evaluate it at $7^6 = 117,649$ points

Estimation challenges - alternatives

- An alternative is Monte Carlo integration
- Also known for its use in maximum simulated likelihood see the special issue in the Stata Journal Vol 6 No 2
- This is a rather brute force approach, but it's usefulness is in it's simplicity

$$L(\boldsymbol{\theta}) = \int f(\boldsymbol{y}|\boldsymbol{\theta}, \boldsymbol{b}) \phi(\boldsymbol{b}) \partial \boldsymbol{b} = \frac{1}{m} \sum_{u=1}^{m} f(y|\boldsymbol{\theta}, \boldsymbol{b}_{u})$$

The important thing to note is m doesn't have to change when extra random effects are added.

Estimation challenges - alternatives

Monte Carlo integration can be improved by:

- antithetic sampling [12]
- Halton sequences [13]
- an adaptive procedure just like adaptive GH quadrature, resulting in an importance sampling approximation

Extensions - level-specific random effect distributions

$$ll(\boldsymbol{\theta}) = \log \int \phi_{\boldsymbol{L}}(\boldsymbol{b}^{(L)} | \boldsymbol{\Sigma}^{(L)}) \prod p^{(L-1)}(\boldsymbol{y} | \boldsymbol{x}, \boldsymbol{b}^{L}, \boldsymbol{\beta}) \,\, \mathrm{d}\boldsymbol{b}^{(L)}$$

where, for $l=2,\ldots,L$

$$p^{(l)}(\boldsymbol{y}|\boldsymbol{x}, \boldsymbol{B}^{l+1}, \boldsymbol{\beta}) = \int \phi_{\boldsymbol{l}}(\boldsymbol{b}^{(l)}) |\boldsymbol{\Sigma}^{(l)} \prod p^{(l-1)}(\boldsymbol{y}|\boldsymbol{x}, \boldsymbol{B}^{l}, \boldsymbol{\beta}) \; \mathrm{d}\boldsymbol{b}^{(l)}$$

Extensions - level-specific random effect distributions and integration techniques

- This formulation now allows us to specify different distributions at each level
- Assess robustness using the *t*-distribution
- Issue of which integration techniques to apply at each level
 - e.g. one random effect at level 1, many at level 2, then use AGHQ at level 3, and MCI at level 2 intmethod(mvaghermite mcarlo) redistribution(normal t) df(3)

Standard linear predictor

The standard linear predictor for a general level model can be written as follows,

$$\eta = oldsymbol{X}oldsymbol{eta} + \sum_{l=2}^L oldsymbol{X}^l oldsymbol{b}^l$$

where subscripts are omitted. We have X our vector of covariates, which could vary at any level, with associated fixed effect coefficient vector β , and X^l the vector of covariates with random effects b^l at level l.

Extended linear predictor

$$\eta_i = g_i(E[y_i|\boldsymbol{X}, \boldsymbol{b}]) = \sum_{r=1}^{R_i} \prod_{s=1}^{S_{ir}} \psi_{irs}$$

where $g_i()$ is the link function for the *i*th outcome. To maintain generality, $\psi_{irs}(t)$ can take many forms, including,

$$\begin{split} \psi_{irs}(t) &= X\\ \psi_{irs}(t) &= \beta\\ \psi_{irs}(t) &= b\\ \psi_{irs}(t) &= q(t)\\ \psi_{irs}(t) &= d_{rs}(E[y_j]), \end{split} \text{ where } j = 1, \dots, k, j \neq i \end{split}$$

Title

megenreg — Extended multivariate generalised linear and non-linear mixed effects models

Syntax

megenreg models [if] [in] [, options]

where models are the model specifications; see megenreg models.

options	Description
model_description_options	fully define, along with <i>models</i> , the model to be fit
estimation_options	method used to obtain estimation results
reporting_options	reporting of estimation results

fweights, iweights, and pweights are allowed; see weight.

Also see megenreg postestimation for features available after estimation.

megenreg in Stata

- Everything I've talked about will be available in the megenreg package in Stata
- It is a simplified/modified version of Stata's official gsem
- megenreg will have many extensions, such as
 - Alternative models, such as spline based survival models
 - Extending sharing between outcomes, motivated by joint modelling
 - User-defined likelihood functions
 - Other things...

Distributional choices

- Gaussian, Poisson, binomial, beta, negative binomial
- exponential, Weibull, Gompertz, log-normal, log-logistic, gamma, Royston-Parmar
- Non-linear outcome models
- User-defined hazard functions
- More to add...



The Royston-Parmar survival model uses restricted cubic splines of log time, on the log cumulative hazard scale, i.e.,

$$\log H(y_i) = s(\log(y_i)|\beta_k) + \eta_i$$

. list patient time infect age female in 1/4, noobs

patient	time	infect	age	female
1	8	1	28	0
1	16	1	28	0
2	13	0	48	1
2	23	1	48	1

. megenreg (time age female M1[patient], ///
> family(rp, failure(infect) scale(h) df(3)))

Relax the normally dist. random effects assumption;

. megenreg (time age female M1[patient], family(rp, failure(infect) scale(h) df(3)))
> , redistribution(t) df(3)

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Higher levels of clustering;

. megenreg (time trt M1[trial] M2[trial>patient], ...)

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Random coefficients;

. megenreg (time trt M1[trial] trt#M1[trial] M2[trial>patient], ...)

Time-dependent effects;

. megenreg (stime trt trt#{log(&t)} M1[id1] M2[id1>id2], ... timevar(stime))

Relax the normally dist. random effects assumption;

. megenreg (time age female M1[patient], family(rp, failure(infect) scale(h) df(3)))
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Random coefficients;

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Time-dependent effects;

. megenreg (stime trt trt#{log(&t)} M1[id1] M2[id1>id2], ... timevar(stime))

Non-linear covariate effects

```
. gen age2 = age<sup>2</sup>
. megenreg (stime trt trt#{log(&t)} age age2 M1[id1] M2[id1>id2], ... )
```

2. A general level relative survival model

Relative survival models are used widely, particularly in population based cancer epidemiology [14]. They model the excess mortality in a population with a particular disease, compared to a reference population.

$$h(y) = h^*(y) + \lambda(y)$$

where $h^*(y)$ is the expected mortality in the reference population. Any of the previous models can be turned into a relative survival model;

```
. megenreg (stime trt trt#log(&t) M1[id1] M2[id1>id2], ///
> family(rp, failure(died) df(3) scale(h) bhazard(bhaz)))
```

- 3. General level joint frailty survival models
 - An area of intense research in recent years is in the field of joint frailty survival models, for the analysis of joint recurrent event and terminal event data
 - Here I focus on the two most popular approaches, proposed by Liu et al. (2004) [15] and Mazroui et al. (2012) [16]
 - In both, we have a survival model for the recurrent event process, and a survival model for the terminal event process, linked through shared random effects

3. General level joint frailty survival models

$$h_{ij}(y) = h_0(y) \exp(\mathbf{X}_{1ij}\boldsymbol{\beta}_1 + b_i)$$

$$\lambda_i(y) = \lambda_0(y) \exp(\mathbf{X}_{1i}\boldsymbol{\beta}_2 + \alpha b_i)$$

where $h_{ij}(y)$ is the hazard function for the *j*th event of the *i*th patient, $\lambda_i(y)$ is the hazard function for the terminal event, and $b_i \sim N(0, \sigma^2)$. We can fit such a model with megenreg, adjusting for treatment in each outcome model,

. megenreg (rectime trt M1[id1] , family(rp, failure(recevent) scale(h) df(5))
> (stime trt M1[id1]@alpha , family(rp, failure(died) scale(h) df(3))

3. General level joint frailty survival models

$$h_{ij}(y) = h_0(y) \exp(\mathbf{X}_{1ij}\boldsymbol{\beta}_1 + b_{1i} + b_{2i})$$
$$\lambda_i(y) = \lambda_0(y) \exp(\mathbf{X}_{1i}\boldsymbol{\beta}_2 + b_{2i})$$

where $b_{1i} \sim N(0, \sigma_1^2)$ and $b_{2i} \sim N(0, \sigma_2^2)$. We give an example of how to fit this model with megenreg, this time illustrating how to use different distributions for the recurrent event and terminal event processes,

Multiple longitudinal biomarkers

$$Y_1 \sim Weib(\lambda, \gamma), \quad Y_2 \quad \sim N(\mu_2, \sigma_2^2), \quad Y_3 \sim N(\mu_3, \sigma_3^2)$$

The linear predictor of the survival outcome can be written as follows,

$$\eta_1(t) = \mathbf{X}\beta_0 + E[y_2(t)|\eta_2(t)]\beta_1 + E[y_3(t)|\eta_3(t)]\beta_2 + E[y_2(t)|\eta_2(t)] \times E[y_3(t)|\eta_3(t)]\beta_3$$

Competing risks

id	logb	logp	time	trt	stime	diedpbc	diedother
3	.3364722	2.484907	0	D-penicil	2.77078	1	0
3	.0953102	2.484907	.481875	D-penicil			
3	.4054651	2.484907	.996605	D-penicil			
3	.5877866	2.587764	2.03428	D-penicil			

. list id logb logp time trt stime diedpbc diedother if id==3, noobs

	megenreg	(stime	trt E	V[logb]@a1	EV[1	.ogp]@a2	,	family(weibull,	failure(diedpbc)))
>		(stime	trt E	V[logb]@a3	EV[1	.ogp]@a4	,	family(gompertz,	failure(diedother))
>		(logb	{&t}@	11 {&t}#M2	[id]	M1[id]	,	family(gaussian)	<pre>timevar(time))</pre>
>		(logp	{&t}@	12 {&t}#M4	[id]	M3[id]		familv(gaussian)	<pre>timevar(time))</pre>

Joint frailty - The extensive frailtypack in R has recently been extended to fit a joint model of a continuous biomarker, a recurrent event process, and a terminal event [5, 8]. We can use megenreg,

	megenreg	(cancti	ime trt	EV[logb]@a1	EV[logp]@a2	M5[id]	,)
>		(stime	trt	EV[logb]@a4	EV[logp]@a5	M5[id]@alpha	,	•)
>		(logb	{&t}@1	1 {&t}#M2[id]] M1[id]	_	,)
>		(logp	{&t}@1	2 {&t}#M4[id]] M3[id]		,)



	megenreg	(canctime	trt	EV[logh)@a1 !	EV []	logp]@a2	,	<pre>family(weibull, failure(canc)))</pre>	111
>		(stimenocanc	trt	EV[logh]@a4 !	EV []	logp]@a5	,		111
>				fami	ily(go	mpe	rtz, fail	Luı	re(diednocanc) ltrunc(canctime))	111
>		(stimecanc	trt	EV[logh]@a4 !	EV [:	logp]@a5	,	<pre>family(gompertz, failure(diedcanc)))</pre>	111
>		(logb	{&t	}@l1 {&́†	;}#M2[id]	M1[id]	,	family(gaussian) timevar(time))	111
>		(logp	{&t	}@12 {&†	;}#M4[:	id]	M3[id]	,	family(gaussian) timevar(time))	

5. A user-defined model - utility functions

A Gaussian response model

$$y \sim N(\eta, \sigma^2)$$

```
real matrix gauss_logl(transmorphic gml)
{
    y = gml_util_depvar(gml) //dep. var.
    linpred = gml_util_xzb(gml) //lin. pred.
    sdre = exp(gml_util_xb(gml,1)) //anc. param.
    return(lnnormalden(y,linpred,sdre)) //logl
}
```

. megenreg (logb time time#M2[id] M1[id], family(user, loglf(gauss_logl)) np(1))

6. A NLME example with multiple linear predictors

Consider Murawska et al. (2012), they developed a Bayesian NL joint model, with Gaussian response variable, and multiple non-linear predictors each with fixed effects and a random intercept. The overall non-linear predictor is defined as,

$$f(t) = \beta_{1i} + \beta_{2i} \exp^{-\beta_{3i}t}$$

where each linear predictor was constrained to be positive,

$$\beta_{1i} = \exp(X_1\beta_1 + b_{1i}) \beta_{2i} = \exp(X_2\beta_2 + b_{2i}) \beta_{3i} = \exp(X_3\beta_3 + b_{3i})$$

and for the survival outcome

$$\lambda(t) = \lambda_0(t) \exp(\alpha_1 b_{1i} + \alpha_2 b_{2i} + \alpha_3 b_{3i})$$

6. A NLME example with multiple linear predictors

We can fit this, and extend it, easily with megenreg

```
real matrix nlme_logl(transmorphic gml, real matrix t)
ſ
               = gml_util_depvar(gml)
                                                            //dep.var.
     linpred1
               = exp(gml_util_xzb(gml))
                                                            //main lin. pred.
     linpred2
               = exp(gml_util_xzb_mod(gml,2))
                                                            //extra lin. preds
     linpred3
               = exp(gml_util_xzb_mod(gml,3))
               = exp(gml_util_xb(gml,1))
     sdre
                                                            //anc. param
     linpred
               = linpred1 :+ linpred2:*exp(-linpred3:*t)
                                                           //nonlin. pred
     return(lnnormalden(y,linpred,sdre))
                                                           //logl
```

}

7. Mixed effects for the level 1 variance function

A recent paper by Goldstein et al. (2017) [17] proposed a two-level model with complex level 1 variation, of the form,

$$y_{ij} = \mathbf{X}_{1ij} \boldsymbol{\beta}_1 + \mathbf{Z}_{1ij} \boldsymbol{b}_{1j} + \epsilon_{ij}$$

$$\epsilon_{ij} \sim N(0, \sigma_e^2)$$

$$\log(\sigma_e^2) = \mathbf{X}_{2ij} \boldsymbol{\beta}_2 + \mathbf{Z}_{2ij} \boldsymbol{b}_{2j}$$

$$\begin{pmatrix} \boldsymbol{b}_{1j} \\ \boldsymbol{b}_{2j} \end{pmatrix} \sim N \begin{bmatrix} \begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} \boldsymbol{\Sigma}_{b_1} \\ \boldsymbol{\Sigma}_{b_1 b_2} & \boldsymbol{\Sigma}_{b_2} \end{bmatrix}$$

7. Mixed effects for the level 1 variance function

We can fit this, and extend it, easily with megenreg

```
real matrix lev1_logl(transmorphic gml, real matrix t)
{
    y = gml_util_depvar(gml) //response
    linpred1 = gml_util_xzb(gml) //lin. pred.
    varresid = exp(gml_util_xzb_mod(gml,2)) //lev1 lin. pred
    return(lnnormalden(y,linpred,sqrt(varresid))) //logl
}
. megenreg (resp female age age#M2[id] M1[id], family(user, loglf(lev1_logl)))
    (age female M3[id], family(null))
    covariance(unstructured)
```

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Summary

- I've presented a very general, and hopefully usable, implementation which can fit a lot of different and new models
- Through the complex linear predictor, we allow seamless development of novel models, and crucially, a way of making them immediately available to researchers through an accessible implementation
 - Realised it can fit multivariate network IPD meta-analysis models this week
- I've incorporated level-specific random effect distributions, and integration techniques

Stuff I didn't show

- family(user, hazard(funcname) cumhazard(funcname))
- fp() and rcs() as elements
- dEV[], d2EV[], iEV[] as elements
- Shell files just like gsem

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- Crowther MJ. *Extended multivariate generalised linear and non-linear mixed effects models*. (Under review).
- Updates and tutorials here: www.mjcrowther.co.uk/software/megenreg

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