

rreg — Robust regression[Description](#)[Quick start](#)[Menu](#)[Syntax](#)[Options](#)[Remarks and examples](#)[Stored results](#)[Methods and formulas](#)[Acknowledgment](#)[References](#)[Also see](#)

Description

`rreg` performs one version of robust regression of *depvar* on *indepvars*.

Also see [Robust standard errors](#) in [R] [regress](#) for standard regression with robust variance estimates and [R] [qreg](#) for quantile (including median or least-absolute-residual) regression.

Quick start

Robust regression of *y* on *x*

```
rreg y x
```

Add a categorical covariate *a* using factor-variable syntax

```
rreg y x i.a
```

As above, but set the tuning parameter to 8

```
rreg y x i.a, tune(8)
```

Generate a new variable *wvar* containing the final weight for each observation

```
rreg y x i.a, genwt(wvar)
```

Set confidence level to 99%

```
rreg y x i.a, level(99)
```

Menu

Statistics > Linear models and related > Other > Robust regression

Syntax

```
rreg depvar [indepvars] [if] [in] [, options]
```

<i>options</i>	Description
Model	
<u>tune</u> (#)	use # as the biweight tuning constant; default is <code>tune(7)</code>
Reporting	
<u>level</u> (#)	set confidence level; default is <code>level(95)</code>
<u>genwt</u> (<i>newvar</i>)	create <i>newvar</i> containing the weights assigned to each observation
<u>display_options</u>	control columns and column formats, row spacing, line width, display of omitted variables and base and empty cells, and factor-variable labeling
Optimization	
<u>optimization_options</u>	control the optimization process; seldom used
<u>graph</u>	graph weights during convergence
<u>coeflegend</u>	display legend instead of statistics

indepvars may contain factor variables; see [U] 11.4.3 [Factor variables](#).

depvar and *indepvars* may contain time-series operators; see [U] 11.4.4 [Time-series varlists](#).

`by`, `mfp`, `mi estimate`, `rolling`, and `statsby` are allowed; see [U] 11.1.10 [Prefix commands](#).

`coeflegend` does not appear in the dialog box.

See [U] 20 [Estimation and postestimation commands](#) for more capabilities of estimation commands.

Options

Model

`tune(#)` is the biweight tuning constant. The default is 7, meaning seven times the median absolute deviation (MAD) from the median residual; see [Methods and formulas](#). Lower tuning constants downweight outliers rapidly but may lead to unstable estimates (less than 6 is not recommended). Higher tuning constants produce milder downweighting.

Reporting

`level(#)`; see [R] [estimation options](#).

`genwt(newvar)` creates the new variable *newvar* containing the weights assigned to each observation.

`display_options`: `nocl`, `nopvalues`, `noomitted`, `vsquish`, `noemptycells`, `baselevels`, `allbaselevels`, `nofvlabel`, `fvwrap(#)`, `fvwrapon(style)`, `cformat(%fmt)`, `pformat(%fmt)`, `sformat(%fmt)`, and `nolstretch`; see [R] [estimation options](#).

Optimization

`optimization_options`: `iterate(#)`, `tolerance(#)`, `[no]log`. `iterate()` specifies the maximum number of iterations; iterations stop when the maximum change in weights drops below `tolerance()`; and `log/nolog` specifies whether to show the iteration log. These options are seldom used.

graph allows you to graphically watch the convergence of the iterative technique. The weights obtained from the most recent round of estimation are graphed against the weights obtained from the previous round.

The following option is available with rreg but is not shown in the dialog box:

coeflegend; see [R] estimation options.

Remarks and examples

stata.com

rreg first performs an initial screening based on Cook's distance > 1 to eliminate gross outliers before calculating starting values and then performs Huber iterations followed by biweight iterations, as suggested by Li (1985).

▶ Example 1

We wish to examine the relationship between mileage rating, weight, and location of manufacture for the 74 cars in the auto.dta. As a point of comparison, we begin by fitting an ordinary regression:

```
. use http://www.stata-press.com/data/r15/auto
(1978 Automobile Data)
. regress mpg weight foreign
```

Source	SS	df	MS	Number of obs	=	74
Model	1619.2877	2	809.643849	F(2, 71)	=	69.75
Residual	824.171761	71	11.608053	Prob > F	=	0.0000
Total	2443.45946	73	33.4720474	R-squared	=	0.6627
				Adj R-squared	=	0.6532
				Root MSE	=	3.4071

mpg	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
weight	-.0065879	.0006371	-10.34	0.000	-.0078583 - .0053175
foreign	-1.650029	1.075994	-1.53	0.130	-3.7955 .4954422
_cons	41.6797	2.165547	19.25	0.000	37.36172 45.99768

We now compare this with the results from rreg:

```
. rreg mpg weight foreign
  Huber iteration 1: maximum difference in weights = .80280176
  Huber iteration 2: maximum difference in weights = .2915438
  Huber iteration 3: maximum difference in weights = .08911171
  Huber iteration 4: maximum difference in weights = .02697328
  Biweight iteration 5: maximum difference in weights = .29186818
  Biweight iteration 6: maximum difference in weights = .11988101
  Biweight iteration 7: maximum difference in weights = .03315872
  Biweight iteration 8: maximum difference in weights = .00721325
Robust regression
```

mpg	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
weight	-.0063976	.0003718	-17.21	0.000	-.007139 - .0056562
foreign	-3.182639	.627964	-5.07	0.000	-4.434763 -1.930514
_cons	40.64022	1.263841	32.16	0.000	38.1202 43.16025

Note the large change in the foreign coefficient.

□ Technical note

It would have been better if we had fit the previous robust regression by typing `rreg mpg weight foreign, genwt(w)`. The new variable, `w`, would then contain the estimated weights. Let's pretend that we did this:

```
. rreg mpg weight foreign, genwt(w)
(output omitted)
. summarize w, detail
```

Robust Regression Weight			
Percentiles		Smallest	
1%	0	0	
5%	.0442957	0	
10%	.4674935	0	Obs 74
25%	.8894815	.0442957	Sum of Wgt. 74
50%	.9690193		Mean .8509966
		Largest	Std. Dev. .2746451
75%	.9949395	.9996715	
90%	.9989245	.9996953	Variance .0754299
95%	.9996715	.9997343	Skewness -2.287952
99%	.9998585	.9998585	Kurtosis 6.874605

We discover that 3 observations in our data were dropped altogether (they have weight 0). We could further explore our data:

```
. sort w
. list make mpg weight w if w <.467, sep(0)
```

	make	mpg	weight	w
1.	Subaru	35	2,050	0
2.	Datsun 210	35	2,020	0
3.	VW Diesel	41	2,040	0
4.	Plym. Arrow	28	3,260	.04429567
5.	Cad. Seville	21	4,290	.08241943
6.	Toyota Corolla	31	2,200	.10443129
7.	Olds 98	21	4,060	.28141296

Being familiar with the automobile data, we immediately spotted two things: the VW is the only diesel car in our data, and the weight recorded for the Plymouth Arrow is incorrect.

□

▶ Example 2

If we specify no explanatory variables, `rreg` produces a robust estimate of the mean:

```
. rreg mpg
      Huber iteration 1: maximum difference in weights = .64471879
      Huber iteration 2: maximum difference in weights = .05098336
      Huber iteration 3: maximum difference in weights = .0099887
      Biweight iteration 4: maximum difference in weights = .25197391
      Biweight iteration 5: maximum difference in weights = .00358606
Robust regression                                Number of obs    =          74
                                                F( 0,          73) =          0.00
                                                Prob > F          =          .
```

mpg	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
_cons	20.68825	.641813	32.23	0.000	19.40912 21.96738

The estimate is given by the coefficient on `_cons`. The mean is 20.69 with an estimated standard error of 0.6418. The 95% confidence interval is [19.4, 22.0]. By comparison, `ci means` (see [\[R\] ci](#)) gives us the standard calculation:

```
. ci means mpg
      Variable |           Obs           Mean      Std. Err.      [95% Conf. Interval]
      -----+-----
      mpg      |           74      21.2973      .6725511      19.9569      22.63769
```

◀

Stored results

`rreg` stores the following in `e()`:

Scalars

```
e(N)           number of observations
e(mss)         model sum of squares
e(df_m)        model degrees of freedom
e(rss)         residual sum of squares
e(df_r)        residual degrees of freedom
e(r2)          R-squared
e(r2_a)        adjusted R-squared
e(F)           F statistic
e(rmse)        root mean squared error
e(rank)        rank of e(V)
```

Macros

```
e(cmd)         rreg
e(cmdline)     command as typed
e(depvar)      name of dependent variable
e(genwt)       variable containing the weights
e(title)       title in estimation output
e(model)       ols
e(properties)  b V
e(predict)     program used to implement predict
e(marginsok)   predictions allowed by margins
e(asbalanced)  factor variables fvset as asbalanced
e(asobserved)  factor variables fvset as asobserved
```

Matrices

```
e(b)           coefficient vector
e(V)           variance–covariance matrix of the estimators
```

Functions

```
e(sample)      marks estimation sample
```

Methods and formulas

See Berk (1990), Goodall (1983), and Rousseeuw and Leroy (1987) for a general description of the issues and methods. Hamilton (1991a, 1992) provides a more detailed description of `rreg` and some Monte Carlo evaluations.

`rreg` begins by fitting the regression (see [R] `regress`), calculating Cook's D (see [R] `predict` and [R] `regress postestimation`), and excluding any observation for which $D > 1$.

Thereafter `rreg` works iteratively: it performs a regression, calculates case weights from absolute residuals, and regresses again using those weights. Iterations stop when the maximum change in weights drops below `tolerance()`. Weights derive from one of two weight functions, Huber weights and biweights. Huber weights (Huber 1964) are used until convergence, and then, from that result, biweights are used until convergence. The biweight was proposed by Beaton and Tukey (1974, 151–152) after the Princeton robustness study (Andrews et al. 1972) had compared various estimators. Both weighting functions are used because Huber weights have problems dealing with severe outliers, whereas biweights sometimes fail to converge or have multiple solutions. The initial Huber weighting should improve the behavior of the biweight estimator.

In Huber weighting, cases with small residuals receive weights of 1; cases with larger residuals receive gradually smaller weights. Let $e_i = y_i - \mathbf{X}_i\mathbf{b}$ represent the i th-case residual. The i th scaled residual $u_i = e_i/s$ is calculated, where $s = M/0.6745$ is the residual scale estimate and $M = \text{med}(|e_i - \text{med}(e_i)|)$ is the median absolute deviation from the median residual. Huber estimation obtains case weights:

$$w_i = \begin{cases} 1 & \text{if } |u_i| \leq c_h \\ c_h/|u_i| & \text{otherwise} \end{cases}$$

`rreg` defines $c_h = 1.345$, so downweighting begins with cases whose absolute residual exceeds $(1.345/0.6745)M \approx 2M$.

With biweights, all cases with nonzero residuals receive some downweighting, according to the smoothly decreasing biweight function

$$w_i = \begin{cases} \{1 - (u_i/c_b)^2\}^2 & \text{if } |u_i| \leq c_b \\ 0 & \text{otherwise} \end{cases}$$

where $c_b = 4.685 \times \text{tune}()/7$. Thus when `tune() = 7`, cases with absolute residuals of $(4.685/0.6745)M \approx 7M$ or more are assigned 0 weight and thus are effectively dropped. Goodall (1983, 377) suggests using a value between 6 and 9, inclusive, for `tune()` in the biweight case and states that performance is good between 6 and 12, inclusive.

The tuning constants $c_h = 1.345$ and $c_b = 4.685$ (assuming `tune()` is set at the default 7) give `rreg` about 95% of the efficiency of OLS when applied to data with normally distributed errors (Hamilton 1991b). Lower tuning constants downweight outliers more drastically (but give up Gaussian efficiency); higher tuning constants make the estimator more like OLS.

Standard errors are calculated using the pseudovalues approach described in Street, Carroll, and Ruppert (1988).

Acknowledgment

The current version of `rreg` is due to the work of Lawrence Hamilton of the Department of Sociology at the University of New Hampshire.

References

- Andrews, D. F., P. J. Bickel, F. R. Hampel, P. J. Huber, W. H. Rogers, and J. W. Tukey. 1972. *Robust Estimates of Location: Survey and Advances*. Princeton, NJ: Princeton University Press.
- Beaton, A. E., and J. W. Tukey. 1974. The fitting of power series, meaning polynomials, illustrated on band-spectroscopic data. *Technometrics* 16: 147–185.
- Berk, R. A. 1990. A primer on robust regression. In *Modern Methods of Data Analysis*, ed. J. Fox and J. S. Long, 292–324. Newbury Park, CA: Sage.
- Goodall, C. 1983. M-estimators of location: An outline of the theory. In *Understanding Robust and Exploratory Data Analysis*, ed. D. C. Hoaglin, C. F. Mosteller, and J. W. Tukey, 339–431. New York: Wiley.
- Gould, W. W., and W. H. Rogers. 1994. Quantile regression as an alternative to robust regression. In *1994 Proceedings of the Statistical Computing Section*. Alexandria, VA: American Statistical Association.
- Hamilton, L. C. 1991a. `srd1`: How robust is robust regression? *Stata Technical Bulletin* 2: 21–26. Reprinted in *Stata Technical Bulletin Reprints*, vol. 1, pp. 169–175. College Station, TX: Stata Press.
- . 1991b. `ssi2`: Bootstrap programming. *Stata Technical Bulletin* 4: 18–27. Reprinted in *Stata Technical Bulletin Reprints*, vol. 1, pp. 208–220. College Station, TX: Stata Press.
- . 1992. *Regression with Graphics: A Second Course in Applied Statistics*. Belmont, CA: Duxbury.
- . 2013. *Statistics with Stata: Updated for Version 12*. 8th ed. Boston: Brooks/Cole.
- Huber, P. J. 1964. Robust estimation of a location parameter. *Annals of Mathematical Statistics* 35: 73–101.
- Li, G. 1985. Robust regression. In *Exploring Data Tables, Trends, and Shapes*, ed. D. C. Hoaglin, C. F. Mosteller, and J. W. Tukey, 281–340. New York: Wiley.
- Mosteller, C. F., and J. W. Tukey. 1977. *Data Analysis and Regression: A Second Course in Statistics*. Reading, MA: Addison-Wesley.
- Relles, D. A., and W. H. Rogers. 1977. Statisticians are fairly robust estimators of location. *Journal of the American Statistical Association* 72: 107–111.
- Rousseeuw, P. J., and A. M. Leroy. 1987. *Robust Regression and Outlier Detection*. New York: Wiley.
- Street, J. O., R. J. Carroll, and D. Ruppert. 1988. A note on computing robust regression estimates via iteratively reweighted least squares. *American Statistician* 42: 152–154.
- Verardi, V., and C. Croux. 2009. Robust regression in Stata. *Stata Journal* 9: 439–453.

Also see

- [R] [rreg postestimation](#) — Postestimation tools for rreg
- [R] [qreg](#) — Quantile regression
- [R] [regress](#) — Linear regression
- [MI] [estimation](#) — Estimation commands for use with mi estimate
- [U] [20 Estimation and postestimation commands](#)