

**xtintreg** — Random-effects interval-data regression models

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## Description

`xtintreg` fits a random-effects regression model in which the dependent variable may be measured as point data, interval data, left-censored data, or right-censored data. The dependent variable must be specified using two *depvars* that indicate how the dependent variable was measured. The user can request that a likelihood-ratio test comparing the panel interval regression model with the pooled model be conducted at estimation time.

## Quick start

Regression on `x` of an interval-measured dependent variable with lower endpoint `y_lower` and upper endpoint `y_upper` using `xtset` data

```
xtintreg y_lower y_upper x
```

Add [indicators](#) for levels of categorical variable `a` as covariates

```
xtintreg y_lower y_upper x i.a
```

Perform likelihood-ratio test against pooled model

```
xtintreg y_lower y_upper x i.a, intreg
```

## Menu

Statistics > Longitudinal/panel data > Censored outcomes > Interval regression (RE)

## Syntax

```
xtintreg depvarlower depvarupper [indepvars] [if] [in] [weight] [, options]
```

The values in *depvar*<sub>lower</sub> and *depvar*<sub>upper</sub> should have the following form:

Type of data		<i>depvar</i> <sub>lower</sub>	<i>depvar</i> <sub>upper</sub>
point data	$a = [a, a]$	$a$	$a$
interval data	$[a, b]$	$a$	$b$
left-censored data	$(-\infty, b]$	.	$b$
right-censored data	$[a, +\infty)$	$a$	.
missing		.	.

<i>options</i>	Description
<b>Model</b>	
<code>noconstant</code>	suppress constant term
<code>offset(<i>varname</i>)</code>	include <i>varname</i> in model with coefficient constrained to 1
<code>constraints(<i>constraints</i>)</code>	apply specified linear constraints
<code>collinear</code>	keep collinear variables
<b>SE</b>	
<code>vce(<i>vcetype</i>)</code>	<i>vcetype</i> may be <code>oim</code> , <code>bootstrap</code> , or <code>jackknife</code>
<b>Reporting</b>	
<code>level(#)</code>	set confidence level; default is <code>level(95)</code>
<code>noskip</code>	perform overall model test as a likelihood-ratio test
<code>intreg</code>	perform likelihood-ratio test against pooled model
<code>nocnsreport</code>	do not display constraints
<code>display_options</code>	control columns and column formats, row spacing, line width, display of omitted variables and base and empty cells, and factor-variable labeling
<b>Integration</b>	
<code>intmethod(<i>intmethod</i>)</code>	integration method; <i>intmethod</i> may be <code>mvaghermite</code> (the default) or <code>ghermite</code>
<code>intpoints(#)</code>	use # quadrature points; default is <code>intpoints(12)</code>
<b>Maximization</b>	
<code>maximize_options</code>	control the maximization process; see [R] <a href="#">maximize</a>
<code>coeflegend</code>	display legend instead of statistics

A panel variable must be specified; use `xtset`; see [XT] [xtset](#).

*indepvars* may contain factor variables; see [U] [11.4.3 Factor variables](#).

*depvar*<sub>lower</sub>, *depvar*<sub>upper</sub>, and *indepvars* may contain time-series operators; see [U] [11.4.4 Time-series varlists](#). `by` and `statsby` are allowed; see [U] [11.1.10 Prefix commands](#).

*iweights* are allowed; see [U] [11.1.6 weight](#). Weights must be constant within panel.

`coeflegend` does not appear in the dialog box.

See [U] [20 Estimation and postestimation commands](#) for more capabilities of estimation commands.

## Options

### Model

`noconstant`, `offset(varname)`, `constraints(constraints)`, `collinear`; see [R] [estimation options](#).

### SE

`vce(vcetype)` specifies the type of standard error reported, which includes types that are derived from asymptotic theory (`oim`) and that use bootstrap or jackknife methods (`bootstrap`, `jackknife`); see [XT] [vce\\_options](#).

### Reporting

`level(#)`, `noskip`; see [R] [estimation options](#).

`intreg` specifies that a likelihood-ratio test comparing the random-effects model with the pooled (`intreg`) model be included in the output.

`nocnsreport`; see [R] [estimation options](#).

`display_options`: `nocl`, `nopvalues`, `noomitted`, `vsquish`, `noemptycells`, `baselevels`, `allbaselevels`, `nofvlabel`, `fvwrap(#)`, `fvwrapon(style)`, `cformat(%fmt)`, `pformat(%fmt)`, `sformat(%fmt)`, and `nolstretch`; see [R] [estimation options](#).

### Integration

`intmethod(intmethod)`, `intpoints(#)`; see [R] [estimation options](#).

### Maximization

`maximize_options`: `difficult`, `technique(algorithm_spec)`, `iterate(#)`, `[no]log`, `trace`, `gradient`, `showstep`, `hessian`, `showtolerance`, `tolerance(#)`, `ltolerance(#)`, `nrtolerance(#)`, `nonrtolerance`, and `from(init_specs)`; see [R] [maximize](#). These options are seldom used.

The following option is available with `xtintreg` but is not shown in the dialog box:

`coeflegend`; see [R] [estimation options](#).

## Remarks and examples

[stata.com](http://www.stata.com)

`xtintreg` may be used to fit a random-effects interval regression model. Consider the linear regression model with panel-level random effects

$$y_{it} = \mathbf{x}_{it}\boldsymbol{\beta} + \nu_i + \epsilon_{it}$$

for  $i = 1, \dots, n$  panels, where  $t = 1, \dots, n_i$ . The random effects,  $\nu_i$ , are i.i.d.,  $N(0, \sigma_\nu^2)$ , and  $\epsilon_{it}$  are i.i.d.,  $N(0, \sigma_\epsilon^2)$  independently of  $\nu_i$ . The observed data consist of the couples,  $(y_{1it}, y_{2it})$ , such that all that is known is that  $y_{1it} \leq y_{it} \leq y_{2it}$ , where  $y_{1it}$  is possibly  $-\infty$  and  $y_{2it}$  is possibly  $+\infty$ .

### ► Example 1: Random-effects interval regression

We begin with the `nlswork` dataset described in [XT] `xt` and create two fictional dependent variables, where the wages are instead reported sometimes as ranges. The wages have been adjusted to 1988 dollars and have further been recoded such that some of the observations are known exactly, some are left-censored, some are right-censored, and some are known only in an interval.

We wish to fit a random-effects interval regression model of adjusted (log) wages. We specify the `intreg` option to test our random-effects model against our pooled estimator.

```
. use http://www.stata-press.com/data/r14/nlswork5
(National Longitudinal Survey. Young Women 14-26 years of age in 1968)
. xtintreg ln_wage1 ln_wage2 i.union age grade not_smsa south#c.year, intreg
(output omitted)
Random-effects interval regression           Number of obs   =   19,224
Group variable: idcode                     Number of groups =    4,148
Random effects u_i ~ Gaussian              Obs per group:
                                           min =           1
                                           avg =           4.6
                                           max =           12
Integration method: mvaghermite            Integration pts. =           12
Wald chi2(7) = 2461.69
Prob > chi2 = 0.0000
Log likelihood = -23260.672
```

	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
1.union	.1229681	.0092943	13.23	0.000	.1047516	.1411846
age	.0096333	.0019	5.07	0.000	.0059094	.0133572
grade	.0756045	.0023828	31.73	0.000	.0709343	.0802747
not_smsa	-.1481304	.011433	-12.96	0.000	-.1705387	-.1257221
1.south	-.3586443	.0977512	-3.67	0.000	-.5502331	-.1670555
year	.0029219	.0020353	1.44	0.151	-.0010671	.0069109
south#c.year						
1	.0032699	.0012076	2.71	0.007	.000903	.0056368
_cons	.2747391	.1141328	2.41	0.016	.0510429	.4984352
/sigma_u	.3044775	.0052644	57.84	0.000	.2941594	.3147956
/sigma_e	.3516248	.00307	114.54	0.000	.3456078	.3576418
rho	.4285095	.0101261			.4087613	.4484385

```
LR test of sigma_u=0: chibar2(01) = 2683.77           Prob >= chibar2 = 0.000
4,781 left-censored observations
4,810 uncensored observations
4,848 right-censored observations
4,785 interval observations
```

The results from an interval regression can be interpreted as we would those from a linear regression. Because the dependent variable is log transformed, the coefficients can be interpreted in terms of a percentage change. We see, for example, that on average, union members make 12.3% more than nonunion members.

The output also includes the overall and panel-level variance components (labeled `sigma_e` and `sigma_u`, respectively) together with  $\rho$  (labeled `rho`),

$$\rho = \frac{\sigma_v^2}{\sigma_\epsilon^2 + \sigma_v^2}$$

which is the proportion of the total variance contributed by the panel-level variance component.

When `rho` is zero, the panel-level variance component is unimportant, and the panel estimator is not different from the pooled estimator. A likelihood-ratio test of this is included at the bottom of the output. This test formally compares the pooled estimator (interval regression) with the panel estimator. In this case, we reject the null hypothesis that there are no panel-level effects.

◀

## □ Technical note

The random-effects model is calculated using quadrature, which is an approximation whose accuracy depends partially on the number of integration points used. We can use the `quadchk` command to see if changing the number of integration points affects the results. If the results change, the quadrature approximation is not accurate given the number of integration points. Try increasing the number of integration points using the `intpoints()` option and run `quadchk` again. Do not attempt to interpret the results of estimates when the coefficients reported by `quadchk` differ substantially. See [\[XT\] quadchk](#) for details and [\[XT\] xtprobit](#) for an [example](#).

Because the `xtintreg` likelihood function is calculated by Gauss–Hermite quadrature, on large problems the computations can be slow. Computation time is roughly proportional to the number of points used for the quadrature.

□

## Stored results

`xtintreg` stores the following in `e()`:

Scalars

<code>e(N)</code>	number of observations
<code>e(N_g)</code>	number of groups
<code>e(N_unc)</code>	number of uncensored observations
<code>e(N_lc)</code>	number of left-censored observations
<code>e(N_rc)</code>	number of right-censored observations
<code>e(N_int)</code>	number of interval observations
<code>e(N_cd)</code>	number of completely determined observations
<code>e(k)</code>	number of parameters
<code>e(k_aux)</code>	number of auxiliary parameters
<code>e(k_eq)</code>	number of equations in <code>e(b)</code>
<code>e(k_eq_model)</code>	number of equations in overall model test
<code>e(k_dv)</code>	number of dependent variables
<code>e(df_m)</code>	model degrees of freedom
<code>e(ll)</code>	log likelihood
<code>e(ll_0)</code>	log likelihood, constant-only model
<code>e(chi2)</code>	$\chi^2$
<code>e(chi2_c)</code>	$\chi^2$ for comparison test
<code>e(rho)</code>	$\rho$
<code>e(sigma_u)</code>	panel-level standard deviation
<code>e(sigma_e)</code>	standard deviation of $\epsilon_{it}$
<code>e(n_quad)</code>	number of quadrature points
<code>e(g_min)</code>	smallest group size

<code>e(g_avg)</code>	average group size
<code>e(g_max)</code>	largest group size
<code>e(p)</code>	significance
<code>e(rank)</code>	rank of $e(V)$
<code>e(rank0)</code>	rank of $e(V)$ for constant-only model
<code>e(ic)</code>	number of iterations
<code>e(rc)</code>	return code
<code>e(converged)</code>	1 if converged, 0 otherwise
<b>Macros</b>	
<code>e(cmd)</code>	<code>xtintreg</code>
<code>e(cmdline)</code>	command as typed
<code>e(depvar)</code>	names of dependent variables
<code>e(ivar)</code>	variable denoting groups
<code>e(wtype)</code>	weight type
<code>e(wexp)</code>	weight expression
<code>e(title)</code>	title in estimation output
<code>e(offset1)</code>	offset
<code>e(chi2type)</code>	Wald or LR; type of model $\chi^2$ test
<code>e(chi2_ct)</code>	Wald or LR; type of model $\chi^2$ test corresponding to <code>e(chi2_c)</code>
<code>e(vce)</code>	<code>vcetype</code> specified in <code>vce()</code>
<code>e(vcetype)</code>	title used to label Std. Err.
<code>e(intmethod)</code>	integration method
<code>e(distrib)</code>	Gaussian; the distribution of the random effect
<code>e(opt)</code>	type of optimization
<code>e(which)</code>	max or min; whether optimizer is to perform maximization or minimization
<code>e(ml_method)</code>	type of ml method
<code>e(user)</code>	name of likelihood-evaluator program
<code>e(technique)</code>	maximization technique
<code>e(properties)</code>	<code>b V</code>
<code>e(predict)</code>	program used to implement <code>predict</code>
<code>e(marginsok)</code>	predictions allowed by <code>margins</code>
<code>e(asbalanced)</code>	factor variables <code>fvset</code> as <code>asbalanced</code>
<code>e(asobserved)</code>	factor variables <code>fvset</code> as <code>asobserved</code>
<b>Matrices</b>	
<code>e(b)</code>	coefficient vector
<code>e(Cns)</code>	constraints matrix
<code>e(ilog)</code>	iteration log
<code>e(gradient)</code>	gradient vector
<code>e(V)</code>	variance-covariance matrix of the estimators
<b>Functions</b>	
<code>e(sample)</code>	marks estimation sample

## Methods and formulas

Assuming a normal distribution,  $N(0, \sigma_\nu^2)$ , for the random effects  $\nu_i$ , we have the joint (unconditional of  $\nu_i$ ) density of the observed data for the  $i$ th panel

$$f \{ (y_{1i1}, y_{2i1}), \dots, (y_{1in_i}, y_{2in_i}) | \mathbf{x}_{1i}, \dots, \mathbf{x}_{in_i} \} = \int_{-\infty}^{\infty} \frac{e^{-\nu_i^2/2\sigma_\nu^2}}{\sqrt{2\pi}\sigma_\nu} \left\{ \prod_{t=1}^{n_i} F(y_{1it}, y_{2it}, \mathbf{x}_{it}\boldsymbol{\beta} + \nu_i) \right\} d\nu_i$$

where

$$F(y_{1it}, y_{2it}, \Delta_{it}) = \begin{cases} (\sqrt{2\pi}\sigma_\epsilon)^{-1} e^{-(y_{1it}-\Delta_{it})^2/(2\sigma_\epsilon^2)} & \text{if } (y_{1it}, y_{2it}) \in C \\ \Phi\left(\frac{y_{2it}-\Delta_{it}}{\sigma_\epsilon}\right) & \text{if } (y_{1it}, y_{2it}) \in L \\ 1 - \Phi\left(\frac{y_{1it}-\Delta_{it}}{\sigma_\epsilon}\right) & \text{if } (y_{1it}, y_{2it}) \in R \\ \Phi\left(\frac{y_{2it}-\Delta_{it}}{\sigma_\epsilon}\right) - \Phi\left(\frac{y_{1it}-\Delta_{it}}{\sigma_\epsilon}\right) & \text{if } (y_{1it}, y_{2it}) \in I \end{cases}$$

where  $C$  is the set of noncensored observations ( $y_{1it} = y_{2it}$  and both nonmissing),  $L$  is the set of left-censored observations ( $y_{1it}$  missing and  $y_{2it}$  nonmissing),  $R$  is the set of right-censored observations ( $y_{1it}$  nonmissing and  $y_{2it}$  missing),  $I$  is the set of interval observations ( $y_{1it} < y_{2it}$  and both nonmissing), and  $\Phi(\cdot)$  is the cumulative normal distribution.

The panel-level likelihood  $l_i$  is given by

$$\begin{aligned} l_i &= \int_{-\infty}^{\infty} \frac{e^{-\nu_i^2/2\sigma_\nu^2}}{\sqrt{2\pi}\sigma_\nu} \left\{ \prod_{t=1}^{n_i} F(y_{1it}, y_{2it}, \mathbf{x}_{it}\beta + \nu_i) \right\} d\nu_i \\ &\equiv \int_{-\infty}^{\infty} g(y_{1it}, y_{2it}, x_{it}, \nu_i) d\nu_i \end{aligned}$$

This integral can be approximated with  $M$ -point Gauss–Hermite quadrature

$$\int_{-\infty}^{\infty} e^{-x^2} h(x) dx \approx \sum_{m=1}^M w_m^* h(a_m^*)$$

This is equivalent to

$$\int_{-\infty}^{\infty} f(x) dx \approx \sum_{m=1}^M w_m^* \exp\{(a_m^*)^2\} f(a_m^*)$$

where the  $w_m^*$  denote the quadrature weights and the  $a_m^*$  denote the quadrature abscissas. The log likelihood,  $L$ , is the sum of the logs of the panel-level likelihoods  $l_i$ .

The default approximation of the log likelihood is by adaptive Gauss–Hermite quadrature, which approximates the panel-level likelihood with

$$l_i \approx \sqrt{2\hat{\sigma}_i} \sum_{m=1}^M w_m^* \exp\{(a_m^*)^2\} g(y_{1it}, y_{2it}, x_{it}, \sqrt{2\hat{\sigma}_i} a_m^* + \hat{\mu}_i)$$

where  $\hat{\sigma}_i$  and  $\hat{\mu}_i$  are the adaptive parameters for panel  $i$ . Therefore, using the definition of  $g(y_{1it}, y_{2it}, x_{it}, \nu_i)$ , the total log likelihood is approximated by

$$L \approx \sum_{i=1}^n w_i \log \left[ \sqrt{2\hat{\sigma}_i} \sum_{m=1}^M w_m^* \exp\{(a_m^*)^2\} \frac{\exp\{-(\sqrt{2\hat{\sigma}_i}a_m^* + \hat{\mu}_i)^2/2\sigma_\nu^2\}}{\sqrt{2\pi}\sigma_\nu} \prod_{t=1}^{n_i} F(y_{1it}, y_{2it}, x_{it}\beta + \sqrt{2\hat{\sigma}_i}a_m^* + \hat{\mu}_i) \right] \quad (1)$$

where  $w_i$  is the user-specified weight for panel  $i$ ; if no weights are specified,  $w_i = 1$ .

The default method of adaptive Gauss–Hermite quadrature is to calculate the posterior mean and variance and use those parameters for  $\hat{\mu}_i$  and  $\hat{\sigma}_i$  by following the method of [Naylor and Smith \(1982\)](#), further discussed in [Skrondal and Rabe-Hesketh \(2004\)](#). We start with  $\hat{\sigma}_{i,0} = 1$  and  $\hat{\mu}_{i,0} = 0$ , and the posterior means and variances are updated in the  $k$ th iteration. That is, at the  $k$ th iteration of the optimization for  $l_i$  we use

$$l_{i,k} \approx \sum_{m=1}^M \sqrt{2\hat{\sigma}_{i,k-1}} w_m^* \exp\{(a_m^*)^2\} g(y_{1it}, y_{2it}, x_{it}, \sqrt{2\hat{\sigma}_{i,k-1}}a_m^* + \hat{\mu}_{i,k-1})$$

Letting

$$\tau_{i,m,k-1} = \sqrt{2\hat{\sigma}_{i,k-1}}a_m^* + \hat{\mu}_{i,k-1}$$

$$\hat{\mu}_{i,k} = \sum_{m=1}^M (\tau_{i,m,k-1}) \frac{\sqrt{2\hat{\sigma}_{i,k-1}} w_m^* \exp\{(a_m^*)^2\} g(y_{1it}, y_{2it}, x_{it}, \tau_{i,m,k-1})}{l_{i,k}}$$

and

$$\hat{\sigma}_{i,k} = \sum_{m=1}^M (\tau_{i,m,k-1})^2 \frac{\sqrt{2\hat{\sigma}_{i,k-1}} w_m^* \exp\{(a_m^*)^2\} g(y_{1it}, y_{2it}, x_{it}, \tau_{i,m,k-1})}{l_{i,k}} - (\hat{\mu}_{i,k})^2$$

and this is repeated until  $\hat{\mu}_{i,k}$  and  $\hat{\sigma}_{i,k}$  have converged for this iteration of the maximization algorithm. This adaptation is applied on every iteration until the log-likelihood change from the preceding iteration is less than a relative difference of  $1e-6$ ; after this, the quadrature parameters are fixed.

The log likelihood can also be calculated by nonadaptive Gauss–Hermite quadrature if the `intmethod(ghermite)` option is specified. For nonadaptive Gauss–Hermite quadrature, the following formula for the log likelihood is used in place of (1).

$$L = \sum_{i=1}^n w_i \log f \{ (y_{1i1}, y_{2i1}), \dots, (y_{1in_i}, y_{2in_i}) | \mathbf{x}_{1i}, \dots, \mathbf{x}_{in_i} \} \\ \approx \sum_{i=1}^n w_i \log \left\{ \frac{1}{\sqrt{\pi}} \sum_{m=1}^M w_m^* \prod_{t=1}^{n_i} F \left( y_{1it}, y_{2it}, \mathbf{x}_{it}\beta + \sqrt{2\sigma_\nu}a_m^* \right) \right\}$$



Both quadrature formulas require that the integrated function be well approximated by a polynomial of degree equal to the number of quadrature points. Panel size can affect whether

$$\prod_{t=1}^{n_i} F(y_{1it}, y_{2it}, \mathbf{x}_{it}\boldsymbol{\beta} + \nu_i)$$

is well approximated by a polynomial. As panel size and  $\rho$  increase, the quadrature approximation can become less accurate. For large  $\rho$ , the random-effects model can also become unidentified. Adaptive quadrature gives better results for correlated data and large panels than nonadaptive quadrature; however, we recommend that you use the `quadchk` command (see [XT] [quadchk](#)) to verify the quadrature approximation used in this command, whichever approximation you choose.

## References

- Naylor, J. C., and A. F. M. Smith. 1982. Applications of a method for the efficient computation of posterior distributions. *Journal of the Royal Statistical Society, Series C* 31: 214–225.
- Neuhaus, J. M. 1992. Statistical methods for longitudinal and clustered designs with binary responses. *Statistical Methods in Medical Research* 1: 249–273.
- Pendergast, J. F., S. J. Gange, M. A. Newton, M. J. Lindstrom, M. Palta, and M. R. Fisher. 1996. A survey of methods for analyzing clustered binary response data. *International Statistical Review* 64: 89–118.
- Skrondal, A., and S. Rabe-Hesketh. 2004. *Generalized Latent Variable Modeling: Multilevel, Longitudinal, and Structural Equation Models*. Boca Raton, FL: Chapman & Hall/CRC.

## Also see

- [XT] [xtintreg postestimation](#) — Postestimation tools for `xtintreg`
- [XT] [quadchk](#) — Check sensitivity of quadrature approximation
- [XT] [xtreg](#) — Fixed-, between-, and random-effects and population-averaged linear models
- [XT] [xtset](#) — Declare data to be panel data
- [XT] [xttobit](#) — Random-effects tobit models
- [R] [intreg](#) — Interval regression
- [R] [tobit](#) — Tobit regression
- [U] [20 Estimation and postestimation commands](#)