stata.com

Title

mdsmat — Multidimensional scaling of proximity data in a matrix								
Syntax Remarks and exar Also see		Menu Stored results	Options References					
yntax								
mdsmat matname	, options	7]						
options	Descrip	tion						
Model								
<pre>method(method)</pre>		for performing	MDS					
loss(<i>loss</i>)	loss fun							
<u>trans</u> form(<i>tfunction</i>)		permitted transformations of dissimilarities						
<u>norm</u> alize(<i>norm</i>)		normalization method; default is normalize(principal)						
<u>nam</u> es(<i>namelist</i>)		e names corresp red with all but	onding to row and column shape(full)	n names of the matrix;				
<u>sh</u> ape(<u>f</u> ull)	matnam	e is a square sy	ymmetric matrix; the defa	ult				
$\underline{sh}ape(\underline{l}ower)$	matname is a vector with the rowwise lower triangle (with diagona							
<u>sh</u> ape(<u>ll</u> ower)			ith the rowwise strictly lo	- · ·				
<u>sh</u> ape(<u>u</u> pper)			ith the rowwise upper tria					
<u>sh</u> ape(<u>uu</u> pper)	<i>matname</i> is a vector with the rowwise strictly upper triangle (no dia							
s2d(<u>st</u> andard)	s2d(<u>st</u> andard) convert similarity to dissimilarity: $d_{ij} = \sqrt{s_{ii} + s_{jj} - 2s_{ij}}$							
s2d(<u>one</u> minus) convert similarity to dissimilarity: $d_{ij} = 1 - s_{ij}$								
Model 2								
<pre>dimension(#)</pre>	-		ns; default is dimension	(2)				
force	-	-	ity information					
addconstant		-	positive semidefinite (class	•				
<pre>weight(matname)</pre>	specifie	s a weight matr	rix with the same shape as	s the proximity matrix				
Reporting								
<pre>neigen(#)</pre>		im number of e sical MDS only)	igenvalues to display; def	ault is neigen(10)				
<u>con</u> fig display table with configuration coordinates								
<u>con</u> fig noplot	display	• ·	iguration coordinates					

Minimization	
<u>init</u> ialize(<i>initopt</i>)	start with configuration given in <i>initopt</i>
<u>tol</u> erance(#)	tolerance for configuration matrix; default is tolerance(1e-4)
<pre>ltolerance(#)</pre>	tolerance for loss criterion; default is ltolerance(1e-8)
<u>iter</u> ate(#)	perform maximum # of iterations; default is iterate(1000)
<pre>protect(#)</pre>	perform # optimizations and report best solution; default is protect(1)
nolog	suppress the iteration log
<u>tr</u> ace	display current configuration in iteration log
$\underline{\texttt{grad}}$ ient	display current gradient matrix in iteration log
<pre>sdprotect(#)</pre>	advanced; see Options below

sdprotect(#) does not appear in the dialog box.

See [U] 20 Estimation and postestimation commands for more capabilities of estimation commands.

method Description						
<u>c</u> lassical	<pre>classical MDS; default if neither loss() nor transform() is specified</pre>					
modern	<pre>modern MDS; default if loss() or transform() is specified; except when loss(stress) and transform(monotonic) an specified</pre>					
<u>n</u> onmetric	<pre>nonmetric (modern) MDS; default when loss(stress) and transform(monotonic) are specified</pre>					
loss	Description					
stress	stress criterion, normalized by distances; the default					
<u>nstr</u> ess	stress criterion, normalized by disparities					
<u>sstr</u> ess	squared stress criterion, normalized by distances					
<u>nsst</u> ress squared stress criterion, normalized by disparities						
<u>stra</u> in	strain criterion (with transform(identity) is equivalent to classical MDS)					
a ommon	Sammon mapping					
sammon						
tfunction	Description					
identity	no transformation; disparity = dissimilarity; the default					
power	power α : disparity = dissimilarity ^{α}					
monotonic	<pre>weakly monotonic increasing functions (nonmetric scaling); only with loss(stress)</pre>					
norm	Description					
principal	principal orientation; location $= 0$; the default					
_ classical	Procrustes rotation toward classical solution					
<pre>target(matname)[, copy] Procrustes rotation toward matname; ignore naming con if copy is specified</pre>						

initopt	Description
<u>c</u> lassical	start with classical solution; the default
<u>r</u> andom[(#)]	start at random configuration, setting seed to #
<u>f</u> rom(<i>matname</i>)[, copy]	start from <i>matname</i> ; ignore naming conflicts if copy is specified

Menu

Statistics > Multivariate analysis > Multidimensional scaling (MDS) > MDS of proximity matrix

Description

mdsmat performs multidimensional scaling (MDS) for two-way proximity data with an explicit measure of similarity or dissimilarity between objects, where the proximities are found in matrix *matname*. mdsmat performs classical metric MDS (Torgerson 1952) as well as modern metric and nonmetric MDS; see the loss() and transform() options.

If your proximities are stored as variables in long format, see [MV] **mdslong**. If you are looking for MDS on a dataset on the basis of dissimilarities between observations over variables, see [MV] **mds**.

Computing the classical solution is straightforward, but with modern MDS the minimization of the loss criteria over configurations is a high-dimensional problem that is easily beset by convergence to local minimums. mds, mdsmat, and mdslong provide options to control the minimization process 1) by allowing the user to select the starting configuration and 2) by selecting the best solution among multiple minimization runs from random starting configurations.

Options

Model

method(method) specifies the method for MDS.

method(classical) specifies classical metric scaling, also known as "principal coordinates analysis" when used with Euclidean proximities. Classical MDS obtains equivalent results to modern MDS with loss(strain) and transform(identity) without weights. The calculations for classical MDS are fast; consequently, classical MDS is generally used to obtain starting values for modern MDS. If the options loss() and transform() are not specified, mds computes the classical solution, likewise if method(classical) is specified loss() and transform() are not allowed.

method(modern) specifies modern scaling. If method(modern) is specified but not loss() or transform(), then loss(stress) and transform(identity) are assumed. All values of loss() and transform() are valid with method(modern).

method(nonmetric) specifies nonmetric scaling, which is a type of modern scaling. If method(nonmetric) is specified, loss(stress) and transform(monotonic) are assumed. Other values of loss() and transform() are not allowed.

loss(loss) specifies the loss criterion.

loss(stress) specifies that the stress loss function be used, normalized by the squared Euclidean distances. This criterion is often called Kruskal's stress-1. Optimal configurations for loss(stress) and for loss(nstress) are equivalent up to a scale factor, but the iteration paths may differ. loss(stress) is the default.

loss(nstress) specifies that the stress loss function be used, normalized by the squared disparities, that is, transformed dissimilarities. Optimal configurations for loss(stress) and for loss(nstress) are equivalent up to a scale factor, but the iteration paths may differ.

loss(sstress) specifies that the squared stress loss function be used, normalized by the fourth power of the Euclidean distances.

loss(nsstress) specifies that the squared stress criterion, normalized by the fourth power of the disparities (transformed dissimilarities) be used.

loss(strain) specifies the strain loss criterion. Classical scaling is equivalent to loss(strain) and transform(identity) but is computed by a faster noniterative algorithm. Specifying loss(strain) still allows transformations.

loss(sammon) specifies the Sammon (1969) loss criterion.

transform(*tfunction*) specifies the class of allowed transformations of the dissimilarities; transformed dissimilarities are called disparities.

transform(identity) specifies that the only allowed transformation is the identity; that is, disparities are equal to dissimilarities. transform(identity) is the default.

transform(power) specifies that disparities are related to the dissimilarities by a power function,

disparity = dissimilarity^{α}, $\alpha > 0$

transform(monotonic) specifies that the disparities are a weakly monotonic function of the dissimilarities. This is also known as nonmetric MDS. Tied dissimilarities are handled by the primary method; that is, ties may be broken but are not necessarily broken. transform(monotonic) is valid only with loss(stress).

normalize(*norm*) specifies a normalization method for the configuration. Recall that the location and orientation of an MDS configuration is not defined ("identified"); an isometric transformation (that is, translation, reflection, or orthonormal rotation) of a configuration preserves interpoint Euclidean distances.

normalize(principal) performs a principal normalization, in which the configuration columns have zero mean and correspond to the principal components, with positive coefficient for the observation with lowest value of id(). normalize(principal) is the default.

normalize(classical) normalizes by a distance-preserving Procrustean transformation of the configuration toward the classical configuration in principal normalization; see [MV] procrustes. normalize(classical) is not valid if method(classical) is specified.

normalize(target(*matname*) [, copy]) normalizes by a distance-preserving Procrustean transformation toward *matname*; see [MV] **procrustes**. *matname* should be an $n \times p$ matrix, where n is the number of observations and p is the number of dimensions, and the rows of *matname* should be ordered with respect to id(). The rownames of *matname* should be set correctly but will be ignored if copy is also specified.

Note on normalize(classical) and normalize(target()): the Procrustes transformation comprises any combination of translation, reflection, and orthonormal rotation—these transformations preserve distance. Dilation (uniform scaling) would stretch distances and is not applied. However, the output reports the dilation factor, and the reported Procrustes statistic is for the dilated configuration.

names(namelist) is required with all but shape(full). The number of names should equal the number of rows (and columns) of the full similarity or dissimilarity matrix and should not contain duplicates. shape(shape) specifies the storage mode of the existing similarity or dissimilarity matrix matname. The following storage modes are allowed:

full specifies that *matname* is a symmetric $n \times n$ matrix.

lower specifies that *matname* is a row or column vector of length n(n+1)/2, with the rowwise lower triangle of the similarity or dissimilarity matrix including the diagonal.

$$D_{11} D_{21} D_{22} D_{31} D_{32} D_{33} \dots D_{n1} D_{n2} \dots D_{nn}$$

llower specifies that *matname* is a row or column vector of length n(n-1)/2, with the rowwise lower triangle of the similarity or dissimilarity matrix excluding the diagonal.

 $D_{21} D_{31} D_{32} D_{41} D_{42} D_{43} \dots D_{n1} D_{n2} \dots D_{n,n-1}$

upper specifies that *matname* is a row or column vector of length n(n+1)/2, with the rowwise upper triangle of the similarity or dissimilarity matrix including the diagonal.

 $D_{11} D_{12} \dots D_{1n} D_{22} D_{23} \dots D_{2n} D_{33} D_{34} \dots D_{3n} \dots D_{nn}$

uupper specifies that *matname* is a row or column vector of length n(n-1)/2, with the rowwise upper triangle of the similarity or dissimilarity matrix excluding the diagonal.

 $D_{12} D_{13} \dots D_{1n} D_{23} D_{24} \dots D_{2n} D_{34} D_{35} \dots D_{3n} \dots D_{n-1,n}$

s2d(standard | oneminus) specifies how similarities are converted into dissimilarities. By default, the command dissimilarity data. Specifying s2d() indicates that your proximity data are similarities.

Dissimilarity data should have zeros on the diagonal (that is, an object is identical to itself) and nonnegative off-diagonal values. Dissimilarities need not satisfy the triangular inequality, $D(i, j)^2 \leq D(i, h)^2 + D(h, j)^2$. Similarity data should have ones on the diagonal (that is, an object is identical to itself) and have off-diagonal values between zero and one. In either case, proximities should be symmetric. See option force if your data violate these assumptions.

The available s2d() options, standard and oneminus, are defined as follows:

standard dissim_{ij} = $\sqrt{\sin_{ii} + \sin_{jj} - 2\sin_{ij}} = \sqrt{2(1 - \sin_{ij})}$ oneminus dissim_{ij} = $1 - \sin_{ij}$

s2d(standard) is the default.

s2d() should be specified only with measures in similarity form.

Model 2

dimension(#) specifies the dimension of the approximating configuration. # defaults to 2 and should not exceed the number of positive eigenvalues of the centered distance matrix.

force corrects problems with the supplied proximity information. force specifies that the dissimilarity matrix be symmetrized; the mean of D_{ij} and D_{ji} is used. Also, problems on the diagonal (similarities: $D_{ii} \neq 1$; dissimilarities: $D_{ii} \neq 0$) are fixed. force does not fix missing values or out-of-range values (that is, $D_{ij} < 0$ or similarities with $D_{ij} > 1$). Analogously, force symmetrizes the weight matrix.

addconstant specifies that if the double-centered distance matrix is not positive semidefinite (psd), a constant should be added to the squared distances to make it psd and, hence, Euclidean.

weight(matname) specifies a symmetric weight matrix with the same shape as the proximity matrix; that is, if shape(lower) is specified, the weight matrix must have this shape. Weights should be nonnegative. Missing weights are assumed to be 0. Weights must also be irreducible; that is, it is not possible to split the objects into disjointed groups with all intergroup weights 0. weight() is not allowed with method(classical), but see loss(strain).

Reporting

neigen(#) specifies the number of eigenvalues to be included in the table. The default is neigen(10). Specifying neigen(0) suppresses the table. This option is allowed with classical MDS only.

- config displays the table with the coordinates of the approximating configuration. This table may also be displayed using the postestimation command estat config; see [MV] mds postestimation.
- noplot suppresses the graph of the approximating configuration. The graph can still be produced later via mdsconfig, which also allows the standard graphics options for fine-tuning the plot; see [MV] mds postestimation plots.

Minimization

These options are available only with method(modern) or method(nonmetric):

initialize(initopt) specifies the initial values of the criterion minimization process.

initialize(classical), the default, uses the solution from classical metric scaling as initial values. With protect(), all but the first run start from random perturbations from the classical solution. These random perturbations are independent and normally distributed with standard error equal to the product of sdprotect(#) and the standard deviation of the dissimilarities. initialize(classical) is the default.

initialize(random) starts an optimization process from a random starting configuration. These random configurations are generated from independent normal distributions with standard error equal to the product of sdprotect(#) and the standard deviation of the dissimilarities. The means of the configuration are irrelevant in MDS.

initialize(from(matname) [, copy]) sets the initial value to matname. matname should be an $n \times p$ matrix, where n is the number of observations and p is the number of dimensions, and the rows of matname should be ordered with respect to id(). The rownames of matname should be set correctly but will be ignored if copy is specified. With protect(), the second-to-last runs start from random perturbations from matname. These random perturbations are independent normal distributed with standard error equal to the product of sdprotect(#) and the standard deviation of the dissimilarities.

- tolerance(#) specifies the tolerance for the configuration matrix. When the relative change in the configuration from one iteration to the next is less than or equal to tolerance(), the tolerance() convergence criterion is satisfied. The default is tolerance(1e-4).
- ltolerance(#) specifies the tolerance for the fit criterion. When the relative change in the fit criterion from one iteration to the next is less than or equal to ltolerance(), the ltolerance() convergence is satisfied. The default is ltolerance(1e-8).

Both the tolerance() and ltolerance() criteria must be satisfied for convergence.

iterate(#) specifies the maximum number of iterations. The default is iterate(1000).

protect(#) requests that # optimizations be performed and that the best of the solutions be reported. The default is protect(1). See option initialize() on starting values of the runs. The output contains a table of the return code, the criterion value reached, and the seed of the random number used to generate the starting value. Specifying a large number, such as protect(50), provides reasonable insight whether the solution found is a global minimum and not just a local minimum.

If any of the options log, trace, or gradient is also specified, iteration reports will be printed for each optimization run. Beware: this option will produce a lot of output.

- nolog suppresses the iteration log, showing the progress of the minimization process.
- trace displays the configuration matrices in the iteration report. Beware: this option may produce a lot of output.
- gradient displays the gradient matrices of the fit criterion in the iteration report. Beware: this option may produce a lot of output.
- The following option is available with mdsmat but is not shown in the dialog box:
- sdprotect(#) sets a proportionality constant for the standard deviations of random configurations
 (init(random)) or random perturbations of given starting configurations (init(classical) or
 init(from())). The default is sdprotect(1).

Remarks and examples

stata.com

Remarks are presented under the following headings:

Introduction Proximity data in a Stata matrix Modern MDS and local minimums

Introduction

Multidimensional scaling (MDS) is a dimension-reduction and visualization technique. Dissimilarities (for instance, Euclidean distances) between observations in a high-dimensional space are represented in a lower-dimensional space (typically two dimensions) so that the Euclidean distance in the lower-dimensional space approximates the dissimilarities in the higher-dimensional space. See Kruskal and Wish (1978) for a brief nontechnical introduction to MDS. Young and Hamer (1987) and Borg and Groenen (2005) are more advanced textbook-sized treatments.

mdsmat performs MDS on a similarity or dissimilarity matrix *matname*. You may enter the matrix as a symmetric square matrix or as a vector (matrix with one row or column) with only the upper or lower triangle; see option shape() for details. *matname* should not contain missing values. The diagonal elements should be 0 (dissimilarities) or 1 (similarities). If you provide a square matrix (that is, shape(full)), names of the objects are obtained from the matrix row and column names. The row names should all be distinct, and the column names should equal the row names. Equation names, if any, are ignored. In any of the vectorized shapes, names are specified with option names(), and the matrix row and column names are ignored.

See option force if your matrix violates these assumptions.

In some applications, the similarity or dissimilarity of objects is defined by the researcher in terms of variables (attributes) measured on the objects. If you need to do MDS of this form, you should continue by reading [MV] mds.

8 mdsmat — Multidimensional scaling of proximity data in a matrix

Often, however, proximities—that is, similarities or dissimilarities—are measured directly. For instance, psychologists studying the similarities or dissimilarities in a set of stimuli—smells, sounds, faces, concepts, etc.—may have subjects rate the dissimilarity of pairs of stimuli. Linguists have subjects rate the similarity or dissimilarity of pairs of dialects. Political scientists have subjects rate the similarity or dissimilarity of political parties or candidates for political office. In other fields, relational data are studied that may be interpreted as proximities in a more abstract sense. For instance, sociologists study interpresonal contact frequencies in groups ("social networks"); these measures are sometimes interpreted in terms of similarities.

A wide variety of MDS methods have been proposed. mdsmat performs classical and modern scaling. Classical scaling has its roots in Young and Householder (1938) and Torgerson (1952). MDS requires complete and symmetric dissimilarity interval-level data. To explore modern scaling, see Borg and Groenen (2005). Classical scaling results in an eigen decomposition, whereas modern scaling is accomplished by the minimization of a loss function. Consequently, eigenvalues are not available after modern MDS.

Proximity data in a Stata matrix

To perform MDS of relational data, you must enter the data in a suitable format. One convenient format is a Stata matrix. You may want to use this format for analyzing data that you obtain from a printed source.

Example 1

Many texts on multidimensional scaling illustrate how locations can be inferred from a table of geographic distances. We will do this too, using an example of distances in miles between 14 locations in Texas, representing both manufactured and natural treasures:

Big Bend	0	523	551	243	322	412	263	596	181	313	553
Corpus Christi	523	0	396	280	705	232	619	226	342	234	30
Dallas	551	396	0	432	643	230	532	243	494	317	426
Del Rio	243	280	432	0	427	209	339	353	62	70	310
El Paso	322	705	643	427	0	528	110	763	365	525	735
Enchanted Rock	412	232	230	209	528	0	398	260	271	69	262
Guadalupe Mnt	263	619	532	339	110	398	0	674	277	280	646
Houston	596	226	243	353	763	260	674	0	415	292	256
Langtry	181	342	494	62	365	271	277	415	0	132	372
Lost Maples	313	234	317	70	525	69	280	292	132	0	264
Padre Island	553	30	426	310	735	262	646	256	372	264	0
Pedernales Falls	434	216	235	231	550	40	420	202	293	115	246
San Antonio	397	141	274	154	564	91	475	199	216	93	171
StataCorp	426	205	151	287	606	148	512	83	318	202	316
Big Bend	434	397	426								
Corpus Christi	216	141	205								
Dallas	235	274	151								
Del Rio	231	154	287								
El Paso	550	564	606								
Enchanted Rock	40	91	148								
Guadalupe Mnt	420	475	512								

202 199

293 216

93 202

0 154

171

115

246

0 75 116

75

116 154

Houston Langtry

Lost Maples

Padre Island

San Antonio

StataCorp

Pedernales Falls

83

318

316

0

Note the inclusion of StataCorp, which is located in the twin cities of Bryan/College Station (BCS). To get the data into Stata, we will enter only the strictly upper triangle as a Stata one-dimensional matrix and collect the names in a global macro for later use. We are using the strictly upper triangle (that is, omitting the diagonal) because the diagonal of a dissimilarity matrix contains all zeros—there is no need to enter them.

	matrix	input	D =	(
>	523	551	243	322	412	263	596	181	313	553	434	397	426	
>		396	280	705	232	619	226	342	234	30	216	141	205	
>			432	643	230	532	243	494	317	426	235	274	151	
>				427	209	339	353	62	70	310	231	154	287	
>					528	110	763	365	525	735	550	564	606	
>						398	260	271	69	262	40	91	148	
>							674	277	280	646	420	475	512	
>								415	292	256	202	199	83	
>									132	372	293	216	318	
>										264	115	93	202	
>											246	171	316	
>												75	116	
>													154)	
	global	names												
>	Big_	Bend		Corpus	_Chri	sti	Dalla	s		Del_R	io			
>	E1_H	Paso		Enchan	ted_R	ock	Guada	lupe_	Mnt	Houst	on			
>	Lang	gtry		Lost_M	aples		Padre	_Isla	nd	Peder	nales	_Fall	s	
>	San	Anton	io	StataC	orp									

The triangular data entry is just typographical and is useful for catching data entry errors. As far as Stata is concerned, we could have typed all the numbers in one long row. We use matrix input D = rather than matrix define D = or just matrix D = so that we do not have to separate entries with commas.

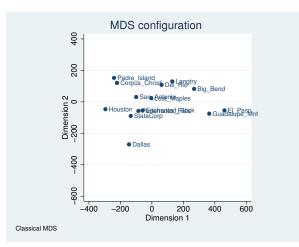
With the data now in Stata, we may use mdsmat to infer the locations in Texas and produce a map:

.

. .

```
. mdsmat D, names($names) shape(uupper)
Classical metric multidimensional scaling
    dissimilarity matrix: D
```

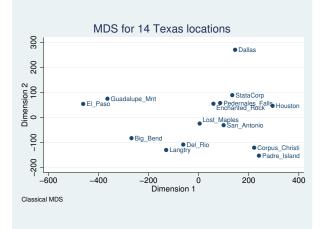
			Number of	obs =	14
Eigenvalues >	0 =	8	Mardia fit	: measure 1 =	0.7828
Retained dime	nsions =	2	Mardia fit	measure 2 =	0.9823
	_				
		abs(eige	envalue)	(eigenv	alue)^2
Dimension	Eigenvalue	Percent	Cumul.	Percent	Cumul.
1	691969.62	62.63	62.63	92.45	92.45
2	172983.05	15.66	78.28	5.78	98.23
3	57771.995	5.23	83.51	0.64	98.87
4	38678.916	3.50	87.01	0.29	99.16
5	19262.579	1.74	88.76	0.07	99.23
6	9230.7695	0.84	89.59	0.02	99.25
7	839.70996	0.08	89.67	0.00	99.25
8	44.989372	0.00	89.67	0.00	99.25



The representation of the distances in two dimensions provides a reasonable, but not great, fit; the percentage of eigenvalues accounted for is 78%.

By default, mdsmat produces a configuration plot. Enhancements to the configuration plot are possible using the mdsconfig postestimation graphics command; see [MV] mds postestimation plots. We present the configuration plot with the autoaspect option to obtain better use of the available space while preserving the equivalence of distance in the x and y axes. We negate the direction of the x axis with the xnegate option to flip the configuration horizontally and flip the direction of the y axis with the ynegate option. We also change the default title and control the placement of labels.

```
. set obs 14
obs was 0, now 14
. generate pos = 3
. replace pos = 4 in 6
(1 real change made)
. replace pos = 2 in 10
(1 real change made)
. mdsconfig, autoaspect xnegate ynegate mlabvpos(pos)
> title(MDS for 14 Texas locations)
```



Look at the graph produced by mdsconfig after mdsmat. You will probably recognize a twisted (and slightly distorted) map of Texas. The vertical orientation of the map is not correctly north-south; you would probably want to turn the map some 20 degrees clockwise. Why didn't mdsmat get it right? It could not have concluded the correct rotation from the available distance information. Any orthogonal rotation of the map would produce the same distances. The orientation of the map is *not identified*. Finally, the "location" of the map cannot be inferred from the distances. Translating the coordinates does not change the distances. As far as mdsmat is concerned, Texas could be part of China.

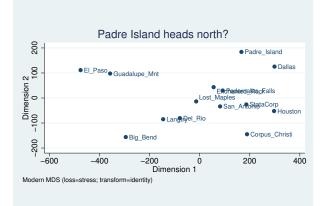
Modern MDS and local minimums

Modern MDS can converge to a local rather than a global minimum. We give an example where this happens and show how the protect() option can guard against this. protect(#) performs multiple minimizations and reports the best one. The output is explained in [MV] mds.

Example 2

Continuing from where we left off, we perform modern MDS, using an initial random configuration with the init(random(512308)) option. The number 512,308 sets the seed so that this run may be replicated.

```
. mdsmat D, names($names) shape(uupper) meth(modern) init(random(512308)) nolog
> noplot
(loss(stress) assumed)
(transform(identity) assumed)
Modern multidimensional scaling
    dissimilarity matrix: D
    Loss criterion: stress = raw_stress/norm(distances)
    Transformation: identity (no transformation)
                                                   Number of obs
                                                                              14
                                                                    =
                                                   Dimensions
                                                                               2
                                                                     =
                                                                          0.1639
    Normalization: principal
                                                   Loss criterion
                                                                    =
. mdsconfig, autoaspect xnegate ynegate mlabvpos(pos)
> title(Padre Island heads north?)
```



4

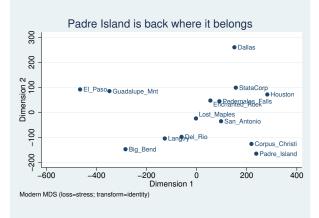
This graph has some resemblance to the one we saw before, but any Texan can assure you that Padre Island should not end up north of Dallas.

We check this result by rerunning with protect(10). This will repeat the minimization and report the best result. Larger values of protect() give us more assurance that we have found the global minimum, but here protect(10) is sufficient to tell us that our original mdsmat found a local, not a global, minimum.

```
. mdsmat D, names($names) shape(uupper) meth(modern) init(random(512308)) nolog
> protect(10) noplot
(loss(stress) assumed)
(transform(identity) assumed)
          #iter
run
     mrc
                     lossval
                               seed random configuration
       0
  1
             54
                   .06180059
                               X4f2d0d0cc0f1b3343a20359ad2ef90f100013c36
  2
       0
             50
                   .06180059
                               X75445b1482f7cbeca4262f9391b6f5e100013232
  3
       0
             42
                               X7fa3d0c0ff14ace95d0e18ed6b30fed800013c05
                    .0618006
  4
       0
             54
                    .0618006
                               Xbca058982f163fb5735186b9c7f29262000133b9
  5
       0
             46
                    .0618006
                               X7ce217b44e3a967be8b4ef2ca63b100300012ba8
  6
             47
       0
                    .0618006
                               Xcaf38d70eba11618b2169652cd28c9a6000117ca
  7
                               Xa3aacf488b86c16fde811ec95b22fdf700012b7e
       0
             83
                   .08581202
  8
       0
            111
                   .08581202
                               Xaaa0cebc256ebbd6da55f0075594212700013851
  9
                   .08581202
                               X448b7f64f30e23ca382ef3eeeaec6da4000118b5
       0
             74
 10
       0
            100
                    .1639279
                               X6c6495e8c43f462544a474abacbdd93d00014486
```

```
Modern multidimensional scaling
dissimilarity matrix: D
Loss criterion: stress = raw_stress/norm(distances)
Transformation: identity (no transformation)
Number of obs = 14
Dimensions = 2
Normalization: principal Loss criterion = 0.0618
. mdsconfig, autoaspect xnegate ynegate mlabvpos(pos)
```

> title(Padre Island is back where it belongs)



The original run had a loss criterion of 0.1639, but after using the protect() option the loss criterion was much lower—0.0618. We also see that Padre Island is back down south where it belongs. It is clear that the original run converged to a local minimum. You can see the original results appear as the final output line of the first table in the output after using protect(10). The seed in the

table is a hexadecimal representation of how the seed is stored internally. The number 512,308 in init(random(512308)) is convenient shorthand for specifying the seed; the two are equivalent. If we wish, we could repeat the command with a larger value of protect() to assure ourselves that 0.0618 is indeed the global minimum.

4

After mdsmat, all MDS postestimation tools are available. For instance, you may analyze residuals with estat quantile, you may produce a Shepard diagram, etc.; see [MV] mds postestimation and [MV] mds postestimation plots.

Stored results

mdsmat stores the following in e():

Sca	lars	
	e(N)	number of rows or columns (i.e., number of observations)
	e(p)	number of dimensions in the approximating configuration
	e(np)	number of strictly positive eigenvalues
	e(addcons)	constant added to squared dissimilarities to force positive semidefiniteness
	e(mardia1)	Mardia measure 1
	e(mardia2)	Mardia measure 2
	e(critval)	loss criterion value
	e(wsum)	sum of weights
	e(alpha)	parameter of transform(power)
	e(ic)	iteration count
	e(rc)	return code
	e(converged)	1 if converged, 0 otherwise
Mad	cros	
	e(cmd)	mdsmat
	e(cmdline)	command as typed
	e(method)	classical or modern MDS method
	e(method2)	nonmetric, if method(nonmetric)
	e(loss)	loss criterion
	e(losstitle)	description loss criterion
	e(dmatrix)	name of analyzed matrix
	e(tfunction)	identity, power, or monotonic, transformation function
	e(transftitle)	description of transformation
	e(mxlen)	maximum length of category labels
	e(dtype)	similarity or dissimilarity; type of proximity data
	e(s2d)	standard or oneminus (when e(dtype) is similarity)
	e(unique)	1 if eigenvalues are distinct, 0 otherwise
	e(init)	initialization method
	e(iseed)	seed for init(random)
	e(seed)	seed for solution
	e(norm)	normalization method
	e(targetmatrix)	name of target matrix for normalize(target)
	e(properties)	nob noV for modern or nonmetric MDS; nob noV eigen for classical MDS
	e(estat_cmd)	program used to implement estat
	e(predict)	program used to implement predict
	e(marginsnotok)	predictions disallowed by margins
Mat	rices	
	e(D)	dissimilarity matrix
	e(Disparities)	disparity matrix for nonmetric MDS
	e(Y)	approximating configuration coordinates
	e(Ev)	eigenvalues
	e(W)	weight matrix
	e(norm_stats)	normalization statistics
	e(linearf)	two element vector defining the linear transformation; distance
		equals first element plus second element times dissimilarity

Methods and formulas

Methods and formulas are presented under the following headings:

Classical multidimensional scaling Modern multidimensional scaling Conversion of similarities to dissimilarities

Classical multidimensional scaling

Let D be an $n \times n$ dissimilarity matrix. The matrix D is said to be *Euclidean* if there are coordinates Y so that

$$D_{ij}^2 = (\mathbf{Y}_i - \mathbf{Y}_j)(\mathbf{Y}_i - \mathbf{Y}_j)'$$

Here \mathbf{Y}_i and \mathbf{Y}_j are the *i*th and *j*th column vectors extracted from \mathbf{Y} . Let $\mathbf{A} = -(1/2)\mathbf{D} \odot \mathbf{D}$, with \odot being the Hadamard or elementwise matrix product, and define \mathbf{B} as the double-centered distance matrix

$$\mathbf{B} = \mathbf{H}\mathbf{A}\mathbf{H}$$
 with $\mathbf{H} = \mathbf{I} - \frac{1}{n}\mathbf{1}\mathbf{1}^{\prime}$

D is Euclidean if and only if **B** is positive semidefinite. Assume for now that **D** is indeed Euclidean. The spectral or eigen decomposition of **B** is written as $\mathbf{B} = \mathbf{U}\mathbf{\Lambda}\mathbf{U}'$, with **U** the orthonormal matrix of eigenvectors normed to 1, and $\mathbf{\Lambda}$ a diagonal matrix with nonnegative values (the eigenvalues of **B**) in decreasing order. The coordinates **Y** are defined in terms of the spectral decomposition $\mathbf{Y} = \mathbf{U}\mathbf{\Lambda}^{1/2}$. These coordinates are centered $\mathbf{Y}'\mathbf{1} = \mathbf{0}$.

The spectral decomposition can also be used to obtain a low-dimensional configuration $\hat{\mathbf{Y}}$, $n \times p$, so that the interrow distances of $\tilde{\mathbf{Y}}$ approximate **D**. Mardia, Kent, and Bibby (1979, sec. 14.4) discuss some characterizations under which the leading p columns of \mathbf{Y} are an optimal choice of $\tilde{\mathbf{Y}}$. These characterizations also apply to the case when **B** is not positive semidefinite, so some of the λ 's are negative; we require that $\lambda_p > 0$.

Various other approaches have been proposed to deal with the case when the matrix **B** is not positive semidefinite, that is, when **B** has negative eigenvalues (see Cox and Cox 2001, 45–48). An easy solution is to add a constant to the off-diagonal elements of $\mathbf{D} \odot \mathbf{D}$ to make **B** positive semidefinite. The smallest such constant is $-2\lambda_n$, where λ_n is the smallest eigenvalue of **B** (Lingoes 1971). See Cailliez (1983) for a solution to the additive constant problem in terms of the dissimilarities instead of the squared dissimilarities.

Goodness-of-fit statistics for a configuration in p dimensions have been proposed by Mardia (1978) in characterizations of optimality properties of the classical solution

$$Mardia_1 = \frac{\sum_{i=1}^{p} |\lambda_i|}{\sum_{i=1}^{n} |\lambda_i|}$$

and

Mardia₂ =
$$\frac{\sum_{i=1}^{p} \lambda_i^2}{\sum_{i=1}^{n} \lambda_i^2}$$

Modern multidimensional scaling

Let **D** be a symmetric $n \times n$ matrix of observed dissimilarities. We assume that proximity data in the form of similarities have already been transformed into dissimilarities. Let **W** be an $n \times n$ matrix of nonnegative weights. With unweighted MDS, we define $W_{ij} = 1$. For a configuration of n points in k-dimensional space represented by the $n \times k$ matrix **Y**, let **B**(**Y**) be the $n \times n$ matrix of Euclidean distances between the rows of **Y**. We consider \mathcal{F} to be some class of permitted transformations from $n \times n$ real matrices to $n \times n$ real matrices.

Modern metric and nonmetric multidimensional scaling involves the minimization of a loss criterion

$$L\left\{f(\mathbf{D}), \mathbf{B}(\mathbf{Y}), \mathbf{W}\right\}$$

over the configurations \mathbf{Y} and transformations $f \in \mathcal{F}$. Whether a scaling method is labeled metric or nonmetric depends on the class \mathcal{F} . In nonmetric scaling, \mathcal{F} is taken to be the class of monotonic functions. If \mathcal{F} is a regular parametrized set of functions, one commonly labels the scaling as metric.

D is the matrix of proximities or dissimilarities, $\mathbf{B}(\mathbf{Y})$ is the matrix of distances, and the result of $f(\mathbf{D}) = \widehat{\mathbf{D}}$ is the matrix of disparities.

The mdsmat command supports the following loss criteria:

(1) **stress** specifies Kruskal's stress-1 criterion: the Euclidean norm of the difference between the distances and the disparities, normalized by the Euclidean norm of the distances.

$$\operatorname{stress}(\widehat{\mathbf{D}}, \mathbf{B}, \mathbf{W}) = \left\{ \frac{\sum_{ij} W_{ij} (B_{ij} - \widehat{D}_{ij})^2}{\sum_{ij} W_{ij} B_{ij}^2} \right\}^{1/2}$$

(2) **nstress** specifies the square root of the normalized stress criterion: the Euclidean norm of the difference between the distances and the disparities, normalized by the Euclidean norm of the disparities.

$$\operatorname{nstress}(\widehat{\mathbf{D}}, \mathbf{B}, \mathbf{W}) = \left\{ \frac{\sum_{ij} W_{ij} (B_{ij} - \widehat{D}_{ij})^2}{\sum_{ij} W_{ij} \widehat{D}_{ij}^2} \right\}^{1/2}$$

nstress normalizes with the disparities, stress with the distances.

(3) **sammon** specifies the Sammon mapping criterion (Sammon 1969; Neimann and Weiss 1979): the sum of the scaled, squared differences between the distances and the disparities, normalized by the sum of the disparities.

$$\mathbf{sammon}(\widehat{\mathbf{D}}, \mathbf{B}, \mathbf{W}) = \frac{\sum_{ij} W_{ij} (B_{ij} - \widehat{D}_{ij})^2 / \widehat{D}_{ij}}{\sum_{ij} W_{ij} \widehat{D}_{ij}}$$

(4) **sstress** specifies the squared stress criterion: the Euclidean norm of the difference between the squared distances and the squared disparities, normalized by the Euclidean norm of the squared distances.

$$\operatorname{sstress}(\widehat{\mathbf{D}}, \mathbf{B}, \mathbf{W}) = \left\{ \frac{\sum_{ij} W_{ij} (B_{ij}^2 - \widehat{D}_{ij}^2)^2}{\sum_{ij} W_{ij} B_{ij}^4} \right\}^{1/2}$$

(5) **nsstress** specifies the normalized squared stress criterion: the Euclidean norm of the difference between the squared distances and the squared disparities, normalized by the Euclidean norm of the squared disparities.

$$\operatorname{nsstress}(\widehat{\mathbf{D}}, \mathbf{B}, \mathbf{W}) = \left\{ \frac{\sum_{ij} W_{ij} (B_{ij}^2 - \widehat{D}_{ij}^2)^2}{\sum_{ij} W_{ij} \widehat{D}_{ij}^4} \right\}^{1/2}$$

nsstress normalizes with the disparities, sstress with the distances.

(6) strain specifies the strain criterion:

$$\operatorname{strain}(\widehat{\mathbf{D}}, \mathbf{B}, \mathbf{W}) = \frac{\sqrt{\operatorname{trace}(\mathbf{X}'\mathbf{X})}}{\sum_{ij} W_{ij}}$$

where

$$\mathbf{X} = \mathbf{W} \odot \left\{ \widehat{\mathbf{D}} - \mathbf{B}(\widetilde{\mathbf{Y}})
ight\}$$

where $\widetilde{\mathbf{Y}}$ is the centered configuration of \mathbf{Y} . Without weights, $W_{ij} = 1$, and without transformation, that is, $\widehat{\mathbf{D}} = \mathbf{D}$, minimization of the strain criterion is equivalent to classical metric scaling.

The mdsmat command supports three classes of permitted transformations, $f \in \mathcal{F}$: (1) the class of all weakly monotonic transformations, (2) the power class of functions where f is defined elementwise on **D** as $f(D_{ij}, \alpha) = D_{ij}^{\alpha}$ (Critchley 1978; Cox and Cox 2001), and (3) the trivial identity case of $f(\mathbf{D}) = \mathbf{D}$.

Minimization of a loss criterion with respect to the configuration \mathbf{Y} and the permitted transformation $f \in \mathcal{F}$ is performed with an alternating algorithm in which the configuration \mathbf{Y} is modified (the C-step) and the transformation f is adjusted (the T-step) to reduce loss. Obviously, no T-step is made with the identity transformation. The classical solution is the default starting configuration. Iteration continues until the C-step and T-step reduce loss by less than the tolerance for convergence or the maximum number of iterations is performed. The C-step is taken by steepest descent using analytical gradients and an optimal stepsize computed using Brent's bounded minimization (Brent 1973). The implementation of the T-step varies with the specified class of transformations. In the nonmetric case of monotonic transformations, we use isotonic regression (Kruskal 1964a, 1964b; Cox and Cox 2001), using the primary approach to ties (Borg and Groenen 2005, 40). For power transformations, we again apply Brent's minimization method.

Given enough iterations, convergence is usually not a problem. However, the alternating algorithm may not converge to a global minimum. mdsmat provides some protection by repeated runs from different initial configurations. However, as Euclidean distances B(Y) are invariant with respect to isometric transformations (rotations, translations) of Y, some caution is required to compare different runs and, similarly, to compare the configurations obtained from different scaling methods. mdsmat normalizes the optimal configuration by centering and via the orthogonal Procrustean rotation without dilation toward the classical or a user-specified solution; see [MV] procrustes.

Conversion of similarities to dissimilarities

If a similarity measure was selected, it is turned into a dissimilarity measure by using one of two methods. The *standard* conversion method is

$$\operatorname{dissim}_{ij} = \sqrt{\operatorname{sim}_{ii} + \operatorname{sim}_{jj} - 2\operatorname{sim}_{ij}}$$

With the similarity of an object to itself being 1, this is equivalent to

$$\operatorname{dissim}_{ij} = \sqrt{2(1 - \operatorname{sim}_{ij})}$$

This conversion method has the attractive property that it transforms a positive-semidefinite similarity matrix into a Euclidean distance matrix (see Mardia, Kent, and Bibby 1979, 402).

We also offer the one-minus method

$$\operatorname{dissim}_{ij} = 1 - \operatorname{sim}_{ij}$$

References

- Borg, I., and P. J. F. Groenen. 2005. Modern Multidimensional Scaling: Theory and Applications. 2nd ed. New York: Springer.
- Brent, R. P. 1973. Algorithms for Minimization without Derivatives. Englewood Cliffs, NJ: Prentice Hall. (Reprinted in paperback by Dover Publications, Mineola, NY, January 2002).
- Cailliez, F. 1983. The analytical solution of the additive constant problem. Psychometrika 48: 305-308.
- Cox, T. F., and M. A. A. Cox. 2001. Multidimensional Scaling. 2nd ed. Boca Raton, FL: Chapman & Hall/CRC.
- Critchley, F. 1978. Multidimensional scaling: A short critique and a new method. In COMPSTAT 1978: Proceedings in Computational Statistics, ed. L. C. A. Corsten and J. Hermans. Vienna: Physica.
- Kruskal, J. B. 1964a. Multidimensional scaling by optimizing goodness of fit to a nonmetric hypothesis. *Psychometrika* 29: 1–27.
- —. 1964b. Nonmetric multidimensional scaling: A numerical method. Psychometrika 29: 115–129.

Kruskal, J. B., and M. Wish. 1978. Multidimensional Scaling. Newbury Park, CA: Sage.

- Lingoes, J. C. 1971. Some boundary conditions for a monotone analysis of symmetric matrices. Psychometrika 36: 195–203.
- Mardia, K. V. 1978. Some properties of classical multidimensional scaling. Communications in Statistics, Theory and Methods 7: 1233–1241.
- Mardia, K. V., J. T. Kent, and J. M. Bibby. 1979. Multivariate Analysis. London: Academic Press.
- Neimann, H., and J. Weiss. 1979. A fast-converging algorithm for nonlinear mapping of high-dimensional data to a plane. *IEEE Transactions on Computers* 28: 142–147.
- Sammon, J. W., Jr. 1969. A nonlinear mapping for data structure analysis. *IEEE Transactions on Computers* 18: 401–409.
- Torgerson, W. S. 1952. Multidimensional scaling: I. Theory and method. Psychometrika 17: 401-419.
- Young, F. W., and R. M. Hamer. 1987. Multidimensional Scaling: History, Theory, and Applications. Hillsdale, NJ: Erlbaum Associates.
- Young, G., and A. S. Householder. 1938. Discussion of a set of points in terms of their mutual distances. *Psychometrika* 3: 19–22.

Also see

- [MV] mds postestimation Postestimation tools for mds, mdsmat, and mdslong
- [MV] mds postestimation plots Postestimation plots for mds, mdsmat, and mdslong
- [MV] **biplot** Biplots
- [MV] ca Simple correspondence analysis
- [MV] factor Factor analysis
- [MV] mds Multidimensional scaling for two-way data
- [MV] mdslong Multidimensional scaling of proximity data in long format
- [MV] pca Principal component analysis
- [U] 20 Estimation and postestimation commands