Title

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### mds - Multidimensional scaling for two-way data

Syntax Remarks and exar Also see	Menu nples Stored res	Description sults Methods and formulas	Options References					
Syntax								
mds varlist [if] [	in], id(varname)	[options]						
options	Description							
Model								
*id(varname)	identify observa	ations						
method(method)	method for per							
loss( <i>loss</i> )	loss function	0						
<u>trans</u> form( <i>tfunction</i> )	permitted transf	formations of dissimilarities						
<u>norm</u> alize( <i>norm</i> )	•	nethod; default is normalize	(principal)					
dimension(#)		imensions; default is dimensi						
addconstant	-	matrix positive semidefinite						
Model 2								
unit ( <i>varlist</i> <sub>2</sub> )	scale variables	to $\min = 0$ and $\max = 1$						
$std[(varlist_3)]$		to mean $= 0$ and sd $= 1$						
measure( <i>measure</i> )		ssimilarity measure; default is	I.2 (Euclidean)					
s2d( <u>st</u> andard)		ty to dissimilarity: $dissim_{ij} =$						
s2d( <u>one</u> minus)	convert similari	ty to dissimilarity: $dissim_{ij} =$	$1 - \sin_{ij}$					
Reporting		-	-					
neigen(#)	maximum num	ber of eigenvalues to display;	default is neigen(10)					
config	display table w	ith configuration coordinates						
noplot	suppress config	uration plot						
Minimization								
<u>init</u> ialize( <i>initopt</i> )	start with config	guration given in initopt						
<u>tol</u> erance(#)	tolerance for co	onfiguration matrix; default is	tolerance(1e-4)					
<u>ltol</u> erance(#)	tolerance for lo	ss criterion; default is ltoler	ance(1e-8)					
<u>iter</u> ate(#)	perform maxim	um # of iterations; default is	iterate(1000)					
<pre>protect(#)</pre>	ect(#) perform # optimizations and report best solution; default is protect(1)							
nolog	suppress the ite	ration log						
trace								
gradient		gradient matrix in iteration lo	g					
<u>sd</u> protect(#)	advanced; see (	Options below						

\* id(varname) is required.

bootstrap, by, jackknife, rolling, statsby, and xi are allowed; see [U] 11.1.10 Prefix commands. The maximum number of observations allowed in mds is the maximum matrix size; see [R] matsize. sdprotect(#) does not appear in the dialog box.

See [U] 20 Estimation and postestimation commands for more capabilities of estimation commands.

#### 2 mds — Multidimensional scaling for two-way data

method	Description
<u>c</u> lassical	<pre>classical MDS; default if neither loss() nor transform() is     specified</pre>
modern	<pre>modern MDS; default if loss() or transform() is specified; except when loss(stress) and transform(monotonic) are specified</pre>
<u>n</u> onmetric	<pre>nonmetric (modern) MDS; default when loss(stress) and transform(monotonic) are specified</pre>
loss	Description
<u>stre</u> ss	stress criterion, normalized by distances; the default
<u>nstr</u> ess	stress criterion, normalized by disparities
<u>sstr</u> ess	squared stress criterion, normalized by distances
<u>nsst</u> ress	squared stress criterion, normalized by disparities
<u>stra</u> in	strain criterion (with transform(identity) is equivalent to classical MDS)
sammon	Sammon mapping
tfunction	Description
<u>i</u> dentity	no transformation; disparity = dissimilarity; the default
power	power $\alpha$ : disparity = dissimilarity <sup><math>\alpha</math></sup>
<u>m</u> onotonic	<pre>weakly monotonic increasing functions (nonmetric scaling); only with loss(stress)</pre>
norm	Description
principal	principal orientation; location $= 0$ ; the default
_ classical	Procrustes rotation toward classical solution
<pre>target(matname)[, copy]</pre>	Procrustes rotation toward <i>matname</i> ; ignore naming conflicts if copy is specified
initopt	Description
<pre>classical random[(#)] from(matname)[, copy]</pre>	start with classical solution; the default start at random configuration, setting seed to # start from <i>matname</i> ; ignore naming conflicts if copy is specified

# Menu

Statistics > Multivariate analysis > Multidimensional scaling (MDS) > MDS of data

# Description

mds performs multidimensional scaling (MDS) for dissimilarities between observations with respect to the variables in *varlist*. A wide selection of similarity and dissimilarity measures is available; see the measure() option. mds performs classical metric MDS (Torgerson 1952) as well as modern metric and nonmetric MDS; see the loss() and transform() options.

mds computes dissimilarities from the observations; mdslong and mdsmat are for use when you already have proximity information. mdslong and mdsmat offer the same statistical features but require different data organizations. mdslong expects the proximity information (and, optionally, weights) in a "long format" (pairwise or dyadic form), whereas mdsmat performs MDS on symmetric proximity and weight matrices; see [MV] mdslong and [MV] mdsmat.

Computing the classical solution is straightforward, but with modern MDS the minimization of the loss criteria over configurations is a high-dimensional problem that is easily beset by convergence to local minimums. mds, mdsmat, and mdslong provide options to control the minimization process 1) by allowing the user to select the starting configuration and 2) by selecting the best solution among multiple minimization runs from random starting configurations.

## Options

Model

id(varname) is required and specifies a variable that identifies observations. A warning message is displayed if varname has duplicate values.

method(method) specifies the method for MDS.

method(classical) specifies classical metric scaling, also known as "principal coordinates analysis" when used with Euclidean proximities. Classical MDS obtains equivalent results to modern MDS with loss(strain) and transform(identity) without weights. The calculations for classical MDS are fast; consequently, classical MDS is generally used to obtain starting values for modern MDS. If the options loss() and transform() are not specified, mds computes the classical solution, likewise if method(classical) is specified loss() and transform() are not allowed.

method(modern) specifies modern scaling. If method(modern) is specified but not loss() or transform(), then loss(stress) and transform(identity) are assumed. All values of loss() and transform() are valid with method(modern).

method(nonmetric) specifies nonmetric scaling, which is a type of modern scaling. If method(nonmetric) is specified, loss(stress) and transform(monotonic) are assumed. Other values of loss() and transform() are not allowed.

loss(loss) specifies the loss criterion.

loss(stress) specifies that the stress loss function be used, normalized by the squared Euclidean distances. This criterion is often called Kruskal's stress-1. Optimal configurations for loss(stress) and for loss(nstress) are equivalent up to a scale factor, but the iteration paths may differ. loss(stress) is the default.

loss(nstress) specifies that the stress loss function be used, normalized by the squared disparities, that is, transformed dissimilarities. Optimal configurations for loss(stress) and for loss(nstress) are equivalent up to a scale factor, but the iteration paths may differ.

loss(sstress) specifies that the squared stress loss function be used, normalized by the fourth power of the Euclidean distances.

loss(nsstress) specifies that the squared stress criterion, normalized by the fourth power of the disparities (transformed dissimilarities) be used.

loss(strain) specifies the strain loss criterion. Classical scaling is equivalent to loss(strain) and transform(identity) but is computed by a faster noniterative algorithm. Specifying loss(strain) still allows transformations.

loss(sammon) specifies the Sammon (1969) loss criterion.

transform(*tfunction*) specifies the class of allowed transformations of the dissimilarities; transformed dissimilarities are called disparities.

transform(identity) specifies that the only allowed transformation is the identity; that is, disparities are equal to dissimilarities. transform(identity) is the default.

transform(power) specifies that disparities are related to the dissimilarities by a power function,

disparity = dissimilarity<sup> $\alpha$ </sup>,  $\alpha > 0$ 

transform(monotonic) specifies that the disparities are a weakly monotonic function of the dissimilarities. This is also known as nonmetric MDS. Tied dissimilarities are handled by the primary method; that is, ties may be broken but are not necessarily broken. transform(monotonic) is valid only with loss(stress).

normalize (*norm*) specifies a normalization method for the configuration. Recall that the location and orientation of an MDS configuration is not defined ("identified"); an isometric transformation (that is, translation, reflection, or orthonormal rotation) of a configuration preserves interpoint Euclidean distances.

normalize(principal) performs a principal normalization, in which the configuration columns have zero mean and correspond to the principal components, with positive coefficient for the observation with lowest value of id(). normalize(principal) is the default.

normalize(classical) normalizes by a distance-preserving Procrustean transformation of the configuration toward the classical configuration in principal normalization; see [MV] procrustes. normalize(classical) is not valid if method(classical) is specified.

normalize(target(*matname*) [, copy]) normalizes by a distance-preserving Procrustean transformation toward *matname*; see [MV] **procrustes**. *matname* should be an  $n \times p$  matrix, where n is the number of observations and p is the number of dimensions, and the rows of *matname* should be ordered with respect to id(). The rownames of *matname* should be set correctly but will be ignored if copy is also specified.

Note on normalize(classical) and normalize(target()): the Procrustes transformation comprises any combination of translation, reflection, and orthonormal rotation—these transformations preserve distance. Dilation (uniform scaling) would stretch distances and is not applied. However, the output reports the dilation factor, and the reported Procrustes statistic is for the dilated configuration.

- dimension(#) specifies the dimension of the approximating configuration. The default # is 2 and should not exceed the number of observations; typically, # would be much smaller. With method(classical), it should not exceed the number of positive eigenvalues of the centered distance matrix.
- addconstant specifies that if the double-centered distance matrix is not positive semidefinite (psd), a constant should be added to the squared distances to make it psd and, hence, Euclidean. addconstant is allowed with classical MDS only.

Model 2

- unit[(varlist<sub>2</sub>)] specifies variables that are transformed to min = 0 and max = 1 before entering in the computation of similarities or dissimilarities. unit by itself, without an argument, is a shorthand for unit(\_all). Variables in unit() should not be included in std().
- $std[(varlist_3)]$  specifies variables that are transformed to mean = 0 and sd = 1 before entering in the computation of similarities or dissimilarities. std by itself, without an argument, is a shorthand for  $std(_all)$ . Variables in std() should not be included in unit().

measure(measure) specifies the similarity or dissimilarity measure. The default is measure(L2), Euclidean distance. This option is not case sensitive. See [MV] measure\_option for detailed descriptions of the supported measures.

If a similarity measure is selected, the computed similarities will first be transformed into dissimilarities, before proceeding with the scaling; see the s2d() option below.

Classical metric MDS with Euclidean distance is equivalent to principal component analysis (see [MV] **pca**); the MDS configuration coordinates are the principal components.

s2d(standard | oneminus) specifies how similarities are converted into dissimilarities. By default, the command dissimilarity data. Specifying s2d() indicates that your proximity data are similarities.

Dissimilarity data should have zeros on the diagonal (that is, an object is identical to itself) and nonnegative off-diagonal values. Dissimilarities need not satisfy the triangular inequality,  $D(i, j)^2 \leq D(i, h)^2 + D(h, j)^2$ . Similarity data should have ones on the diagonal (that is, an object is identical to itself) and have off-diagonal values between zero and one. In either case, proximities should be symmetric.

The available s2d() options, standard and oneminus, are defined as follows:

standard dissim<sub>ij</sub> =  $\sqrt{\sin_{ii} + \sin_{jj} - 2\sin_{ij}} = \sqrt{2(1 - \sin_{ij})}$ oneminus dissim<sub>ij</sub> =  $1 - \sin_{ij}$ 

s2d(standard) is the default.

s2d() should be specified only with measures in similarity form.

Reporting

neigen(#) specifies the number of eigenvalues to be included in the table. The default is neigen(10).
Specifying neigen(0) suppresses the table. This option is allowed with classical MDS only.

config displays the table with the coordinates of the approximating configuration. This table may also be displayed using the postestimation command estat config; see [MV] mds postestimation.

noplot suppresses the graph of the approximating configuration. The graph can still be produced later via mdsconfig, which also allows the standard graphics options for fine-tuning the plot; see [MV] mds postestimation plots.

Minimization

These options are available only with method(modern) or method(nonmetric):

initialize(initopt) specifies the initial values of the criterion minimization process.

initialize(classical), the default, uses the solution from classical metric scaling as initial values. With protect(), all but the first run start from random perturbations from the classical solution. These random perturbations are independent and normally distributed with standard error equal to the product of sdprotect(#) and the standard deviation of the dissimilarities. initialize(classical) is the default.

initialize(random) starts an optimization process from a random starting configuration. These random configurations are generated from independent normal distributions with standard error equal to the product of sdprotect(#) and the standard deviation of the dissimilarities. The means of the configuration are irrelevant in MDS.

initialize(from(matname) [, copy]) sets the initial value to matname. matname should be an  $n \times p$  matrix, where n is the number of observations and p is the number of dimensions, and the rows of matname should be ordered with respect to id(). The rownames of matname should be set correctly but will be ignored if copy is specified. With protect(), the second-to-last runs start from random perturbations from matname. These random perturbations are independent normal distributed with standard error equal to the product of sdprotect(#) and the standard deviation of the dissimilarities.

- tolerance(#) specifies the tolerance for the configuration matrix. When the relative change in the configuration from one iteration to the next is less than or equal to tolerance(), the tolerance() convergence criterion is satisfied. The default is tolerance(1e-4).
- ltolerance(#) specifies the tolerance for the fit criterion. When the relative change in the fit criterion from one iteration to the next is less than or equal to ltolerance(), the ltolerance() convergence is satisfied. The default is ltolerance(1e-8).

Both the tolerance() and ltolerance() criteria must be satisfied for convergence.

iterate(#) specifies the maximum number of iterations. The default is iterate(1000).

protect(#) requests that # optimizations be performed and that the best of the solutions be reported. The default is protect(1). See option initialize() on starting values of the runs. The output contains a table of the return code, the criterion value reached, and the seed of the random number used to generate the starting value. Specifying a large number, such as protect(50), provides reasonable insight whether the solution found is a global minimum and not just a local minimum.

If any of the options log, trace, or gradient is also specified, iteration reports will be printed for each optimization run. Beware: this option will produce a lot of output.

- nolog suppresses the iteration log, showing the progress of the minimization process.
- trace displays the configuration matrices in the iteration report. Beware: this option may produce a lot of output.
- gradient displays the gradient matrices of the fit criterion in the iteration report. Beware: this option may produce a lot of output.

The following option is available with mds but is not shown in the dialog box:

sdprotect(#) sets a proportionality constant for the standard deviations of random configurations
(init(random)) or random perturbations of given starting configurations (init(classical) or
init(from())). The default is sdprotect(1).

### **Remarks and examples**

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Remarks are presented under the following headings:

Introduction Euclidean distances Non-Euclidean dissimilarity measures Introduction to modern MDS Protecting from local minimums

### Introduction

Multidimensional scaling (MDS) is a dimension-reduction and visualization technique. Dissimilarities (for instance, Euclidean distances) between observations in a high-dimensional space are represented in a lower-dimensional space (typically two dimensions) so that the Euclidean distance in the lower-dimensional space approximates the dissimilarities in the higher-dimensional space. See Kruskal and Wish (1978) for a brief nontechnical introduction to MDS. Young and Hamer (1987) and Borg and Groenen (2005) offer more advanced textbook-sized treatments.

If you already have the similarities or dissimilarities of the n objects, you should continue by reading [MV] mdsmat.

In many applications of MDS, however, the similarity or dissimilarity of objects is not measured but rather *defined* by the researcher in terms of variables ("attributes")  $x_1, \ldots, x_k$  that are measured on the objects. The pairwise dissimilarity of objects can be expressed using a variety of similarity or dissimilarity measures in the attributes (for example, Mardia, Kent, and Bibby [1979, sec. 13.4]; Cox and Cox [2001, sec. 1.3]). A common measure is the Euclidean distance L2 between the attributes of the objects *i* and *j*:

$$L2_{ij} = \left\{ (x_{i1} - x_{j1})^2 + (x_{i2} - x_{j2})^2 + \dots + (x_{ik} - x_{jk})^2 \right\}^{1/2}$$

A popular alternative is the L1 distance, also known as the cityblock or Manhattan distance. In comparison to L2, L1 gives less influence to larger differences in attributes:

$$L1_{ij} = |x_{i1} - x_{j1}| + |x_{i2} - x_{j2}| + \dots + |x_{ik} - x_{jk}|$$

In contrast, we may also define the extent of dissimilarity between 2 observations as the maximum absolute difference in the attributes and thus give a larger influence to larger differences:

Linfinity<sub>*ij*</sub> = max(
$$|x_{i1} - x_{j1}|, |x_{i2} - x_{j2}|, \dots, |x_{ik} - x_{jk}|$$
)

These three measures are special cases of the Minkowski distance L(q), for q = 2 (L2), q = 1 (L1), and  $q = \infty$  (Linfinity), respectively. Minkowski distances with other values of q may be used as well. Stata supports a wide variety of other similarity and dissimilarity measures, both for continuous variables and for binary variables. See [MV] *measure\_option* for details.

Multidimensional scaling constructs approximations for dissimilarities, not for similarities. Thus, if a similarity measure is specified, mds first transforms the similarities into dissimilarities. Two methods to do this are available. The default standard method,

$$\operatorname{dissim}_{ij} = \sqrt{\operatorname{sim}_{ii} - 2\operatorname{sim}_{ij} + \operatorname{sim}_{jj}}$$

has a useful property: if the similarity matrix is positive semidefinite, a property satisfied by most similarity measures, the standard dissimilarities are Euclidean.

Usually, the number of observations exceeds the number of variables on which the observations are compared, but this is not a requirement for MDS. MDS creates an  $n \times n$  dissimilarity matrix **D** from the *n* observations on *k* variables. It then constructs an approximation of **D** by the Euclidean distances in a matching configuration **Y** of *n* points in *p*-dimensional space:

dissimilarity
$$(x_i, x_j) \approx L2(y_i, y_j)$$
 for all  $i, j$ 

Typically, of course,  $p \ll k$ , and most often p = 1, 2, or 3.

A wide variety of MDS methods have been proposed. mds performs classical and modern scaling. Classical scaling has its roots in Young and Householder (1938) and Torgerson (1952). MDS requires complete and symmetric dissimilarity interval-level data. To explore modern scaling, see Borg and Groenen (2005). Classical scaling results in an eigen decomposition, whereas modern scaling is accomplished by the minimization of a loss function. Consequently, eigenvalues are not available after modern MDS.

### **Euclidean distances**

#### Example 1

The most popular dissimilarity measure is Euclidean distance. We illustrate with data from table 7.1 of Yang and Trewn (2004, 182). This dataset consists of eight variables with nutrition data on 25 breakfast cereals.

```
. use http://www.stata-press.com/data/r13/cerealnut
(Cereal Nutrition)
. describe
Contains data from http://www.stata-press.com/data/r13/cerealnut.dta
  obs:
                   25
                                                 Cereal Nutrition
                                                  24 Feb 2013 17:19
 vars:
                    9
                1,050
                                                  (_dta has notes)
 size:
               storage
                          display
                                      value
variable name
                          format
                                      label
                                                  variable label
                 type
                          %25s
                                                  Cereal Brand
brand
                 str25
calories
                 int
                          %9.0g
                                                  Calories (Cal/oz)
                          %9.0g
                                                 Protein (g)
protein
                 byte
                          %9.0g
                                                 Fat (g)
fat
                 byte
                                                 Na (mg)
Na
                 int
                          %9.0g
                          %9.0g
fiber
                 float
                                                 Fiber (g)
                                                 Carbs (g)
carbs
                 float
                          %9.0g
sugar
                 byte
                         %9.0g
                                                 Sugar (g)
Κ
                 int
                          %9.0g
                                                 K (mg)
```

Sorted by:

. summarize cal	ories-K, sep(	(4)			
Variable	Obs	Mean	Std. Dev.	Min	Max
calories	25	109.6	21.30728	50	160
protein	25	2.68	1.314027	1	6
fat	25	.92	.7593857	0	2
Na	25	195.8	71.32204	0	320
fiber	25	1.7	2.056494	0	9
carbs	25	15.3	4.028544	7	22
sugar	25	7.4	4.609772	0	14
K	25	90.6	77.5043	15	320

. replace brand = subinstr(brand," ","\_",.)

(20 real changes made)

We replaced spaces in the cereal brand names with underscores to avoid confusing which words in the brand names are associated with which points in the graphs we are about to produce. Removing spaces is not required.

The default dissimilarity measure used by mds is the Euclidean distance L2 computed on the raw data (unstandardized). The summary of the eight nutrition variables shows that K, Na, and calories—having much larger standard deviations—will largely determine the Euclidean distances.

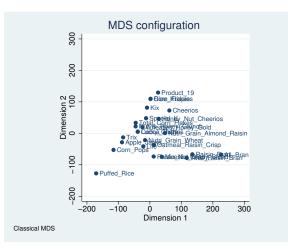
. mds calories-K, id(brand)

Classical metric multidimensional scaling

dissimilarity: L2, computed on 8 variables

			Number	of	obs		=	25
Eigenvalues > 0	=	8	Mardia	fit	measure	1	=	0.9603
Retained dimensions	=	2	Mardia	fit	measure	2	=	0.9970

		abs(eige	nvalue)	(eigenv	alue)^2
Dimension	Eigenvalue	Percent	Cumul.	Percent	Cumul.
1	158437.92	56.95	56.95	67.78	67.78
2	108728.77	39.08	96.03	31.92	99.70
3	10562.645	3.80	99.83	0.30	100.00
4	382.67849	0.14	99.97	0.00	100.00
5	69.761715	0.03	99.99	0.00	100.00
6	12.520822	0.00	100.00	0.00	100.00
7	5.7559984	0.00	100.00	0.00	100.00
8	2.2243244	0.00	100.00	0.00	100.00



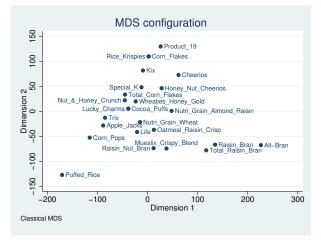
The default MDS configuration graph can be improved upon by using the mdsconfig postestimation command. We will demonstrate this in a moment. But first, we explain the output of mds.

mds has performed classical metric scaling and extracted two dimensions, which is the default action. To assess goodness of fit, the two statistics proposed by Mardia are reported (see Mardia, Kent, and Bibby [1979, sec. 14.4]). The statistics are defined in terms of the eigenvalues of the double-centered distance matrix. If the dissimilarities are truly Euclidean, all eigenvalues are nonnegative. Look at the eigenvalues. We may interpret these as the extent to which the dimensions account for dissimilarity between the cereals. Depending on whether you look at the eigenvalues or squared eigenvalues, it takes two or three dimensions to account for more than 99% of the dissimilarity.

We can produce a prettier configuration plot with the mdsconfig command; see [MV] mds postestimation plots for details.

```
. generate place = 3
. replace place = 9 if inlist(brand,"Rice_Krispies","Nut_&_Honey_Crunch",
> "Special_K","Raisin_Nut_Bran","Lucky_Charms")
(5 real changes made)
. replace place = 12 if inlist(brand,"Mueslix_Crispy_Blend")
(1 real change made)
```

```
. mdsconfig, autoaspect mlabvpos(place)
```



The *marker\_label\_option* mlabvposition() allowed fine control over the placement of the cereal brand names. We created a variable called place giving clock positions where the cereal names were to appear in relation to the plotted point. We set these to minimize overlap of the names. We also requested the autoaspect option to obtain better use of the graphing region while preserving the scale of the x and y axes.

MDS has placed the cereals so that all the brands fall within a triangle defined by Product 19, All-Bran, and Puffed Rice. You can examine the graph to see how close your favorite cereal is to the other cereals.

But, as we saw from the variable summary, three of the eight variables are controlling the distances. If we want to provide for a more equal footing for the eight variables, we can request that mds compute the Euclidean distances on standardized variables. Euclidean distance based on standardized variables is also known as the *Karl Pearson distance* (Pearson 1900). We obtain standardized measures with the option std.

. mds calories-K, id(brand) std noplot

Classical metric multidimensional scaling dissimilarity: L2, computed on 8 variables

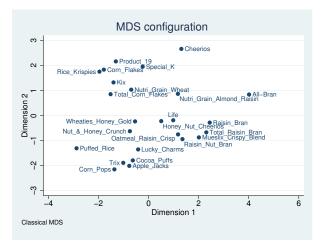
		Number of	obs =	25
0 =	8	Mardia fit	measure 1 =	0.5987
nsions =	2	Mardia fit	measure 2 =	0.7697
r				
	abs(eige	envalue)	(eigenv	alue)^2
Eigenvalue	Percent	Cumul.	Percent	Cumul.
65.645395	34.19	34.19	49.21	49.21
49.311416	25.68	59.87	27.77	76.97
38.826608	20.22	80.10	17.21	94.19
17.727805	9.23	89.33	3.59	97.78
11.230087	5.85	95.18	1.44	99.22
8.2386231	4.29	99.47	0.78	99.99
.77953426	0.41	99.87	0.01	100.00
.24053137	0.13	100.00	0.00	100.00
	Eigenvalue 65.645395 49.311416 38.826608 17.727805 11.230087 8.2386231 .77953426	absions         =         2           abs(eigenvalue         Percent           65.645395         34.19           49.311416         25.68           38.826608         20.22           17.727805         9.23           11.230087         5.85           8.2386231         4.29           .77953426         0.41	0 = 8 Mardia fit hsions = 2 Mardia fit bions = 2 Mardia fit Eigenvalue Percent Cumul. 65.645395 34.19 34.19 49.311416 25.68 59.87 38.826608 20.22 80.10 17.727805 9.23 89.33 11.230087 5.85 95.18 8.2386231 4.29 99.47 .77953426 0.41 99.87	absions         =         2         Mardia fit measure 2 =           abs(eigenvalue)         (eigenvalue)         (eigenvalue)           Eigenvalue         Percent         Cumul.         Percent           65.645395         34.19         34.19         49.21           49.311416         25.68         59.87         27.77           38.826608         20.22         80.10         17.21           17.727805         9.23         89.33         3.59           11.230087         5.85         95.18         1.44           8.2386231         4.29         99.47         0.78           .77953426         0.41         99.87         0.01

In this and the previous example, we did not specify a method() for mds and got classical metric scaling. Classical scaling is the default when method() is omitted and neither the loss() nor transform() option is specified.

Accounting for more than 99% of the underlying distances now takes more MDS-retained dimensions. For this example, we have still retained only two dimensions. We specified the noplot option because we wanted to exercise control over the configuration plot by using the mdsconfig command. We generate a variable named pos that will help minimize cereal brand name overlap.

```
. generate pos = 3
. replace pos = 5 if inlist(brand, "Honey_Nut_Cheerios", "Raisin_Nut_Bran",
> "Nutri_Grain_Almond_Raisin")
(3 real changes made)
. replace pos = 8 if inlist(brand, "Oatmeal_Raisin_Crisp")
(1 real change made)
. replace pos = 9 if inlist(brand, "Corn_Pops", "Trix", "Nut_&_Honey_Crunch",
> "Rice_Krispies", "Wheaties_Honey_Gold")
(5 real changes made)
. replace pos = 12 if inlist(brand, "Life")
(1 real change made)
```

. mdsconfig, autoaspect mlabvpos(pos)



This configuration plot, based on the standardized variables, better incorporates all the nutrition data. If you are familiar with these cereal brands, spotting groups of similar cereals appearing near each other is easy. The bottom-left corner has several of the most sweetened cereals. The brands containing the word "Bran" all appear to the right of center. Rice Krispies and Puffed Rice are the farthest to the left.

Classical multidimensional scaling based on standardized Euclidean distances is actually equivalent to a principal component analysis of the correlation matrix of the variables. See Mardia, Kent, and Bibby (1979, sec. 14.3) for details.

We now demonstrate this property by doing a principal component analysis extracting the leading two principal components. See [MV] pca for details.

. pc	pca calories-K, comp(2)								
Prin	cipal componer	nts/correlation	Number of obs	= 25 = 2					
				Number of comp. Trace	= 2				
	<b>D</b>								
	Rotation: (un)	rotated = princ:	1pal)	Rho	= 0.5987				
	Component	Eigenvalue	Difference	Proportion	Cumulative				
	Comp1	2.73522	.680583	0.3419	0.3419				
	Comp2	2.05464	.436867	0.2568	0.5987				
	Comp3	1.61778	.879117	0.2022	0.8010				
	Comp4	.738659	.270738	0.0923	0.8933				
	Comp5	.46792	.124644	0.0585	0.9518				
	Comp6	.343276	.310795	0.0429	0.9947				
	Comp7	.0324806	.0224585	0.0041	0.9987				
	Comp8	.0100221	•	0.0013	1.0000				

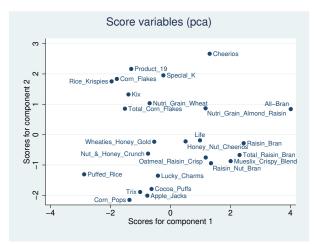
Variable	Comp1	Comp2	Unexplained
calories	0.1992	-0.0632	.8832
protein	0.3376	0.4203	.3253
fat	0.3811	-0.0667	. 5936
Na	0.0962	0.5554	.3408
fiber	0.5146	0.0913	.2586
carbs	-0.2574	0.4492	.4043
sugar	0.2081	-0.5426	.2765
K	0.5635	0.0430	.1278

Principal components (eigenvectors)

The proportion and cumulative proportion of the eigenvalues in the PCA match the percentages from MDS. We will ignore the interpretation of the principal components but move directly to the principal coordinates, also known as the scores of the PCA. We make a plot of the first and second scores, using the scoreplot command; see [MV] scoreplot. We specify the mlabel() option to label the cereals and the mlabvpos() option for fine control over placement of the brand names.

```
. replace pos = 11 if inlist(brand,"All-Bran")
(1 real change made)
```

```
. scoreplot, mlabel(brand) mlabvpos(pos)
```



Compare this PCA score plot with the MDS configuration plot. Apart from some differences in how the graphs were rendered, they are the same.

4

### Non-Euclidean dissimilarity measures

With non-Euclidean dissimilarity measures, the parallel between PCA and MDS no longer holds.

#### Example 2

To illustrate MDS with non-Euclidean distance measures, we will analyze books on multivariate statistics. Gifi (1990) reports on the number of pages devoted to six topics in 20 textbooks on multivariate statistics. We added similar data on five more recent books.

Contains data obs:	from htt <sub>]</sub> 25	p://www.sta	ta-press.c	om/data/r13/mvstatsbooks.dta
vars:	8			15 Mar 2013 16:27
size:	725			(_dta has notes)
	storage	display	value	
variable name	type	format	label	variable label
author	str17	%17s		
math	int	%9.0g		<pre>math other than statistics (e.g., linear algebra)</pre>
corr	int	%9.0g		correlation and regression, including linear structural and functional equations
fact	byte	%9.0g		factor analysis and principal component analysis
cano	byte	%9.0g		canonical correlation analysis
disc	int	%9.0g		discriminant analysis, classification, and cluster analysis
stat	int	%9.0g		<pre>statistics, incl. dist. theory, hypothesis testing &amp; est.; categorical data</pre>
mano	int	%9.0g		manova and the general linear model

. use http://www.stata-press.com/data/r13/mystatsbooks. clear

#### Sorted by:

A brief description of the topics is given in the variable labels. For more details, we refer to Gifi (1990, 15). Here are the data:

. list, noobs

author	math	corr	fact	cano	disc	stat	mano
Roy57	31	0	0	0	0	164	11
Kendal157	0	16	54	18	27	13	14
Kendal175	0	40	32	10	42	60	0
Anderson58	19	0	35	19	28	163	52
CooleyLohnes62	14	7	35	22	17	0	56
-		(0	utput omi	tted)			
GreenCaroll76	290	10	6	0	8	0	2
CailliezPages76	184	48	82	42	134	0	0
Giri77	29	0	0	0	41	211	32
Gnanadesikan77	0	19	56	0	39	75	0
Kshirsagar78	0	22	45	42	60	230	59
Thorndike78	30	128	90	28	48	0	0
MardiaKentBibby79	34	28	68	19	67	131	55
Seber84	16	0	59	13	116	129	101
Stevens96	23	87	67	21	30	43	249
EverittDunn01	0	54	65	0	56	20	30
Rencher02	38	0	71	19	105	135	131

For instance, the 1979 book by Mardia, Kent, and Bibby has 34 pages on mathematics (mostly linear algebra); 28 pages on correlation, regression, and related topics (in this particular case, simultaneous equations); etc. In most of these books, some pages are not classified. Anyway, the number of pages

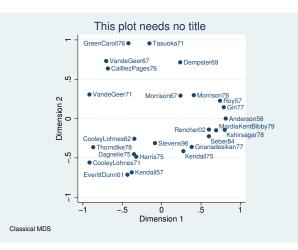
and the amount of information per page vary widely among the books. A Euclidean distance measure is not appropriate here. Standardization does not help us here—the problem is not differences in the scales of the variables but those in the observations. One possibility is to transform the data into *compositional data* by dividing the variables by the total number of classified pages. See Mardia, Kent, and Bibby (1979, 377–380) for a discussion of specialized dissimilarity measures for compositional data. However, we can also use the correlation between observations (not between variables) as the similarity measure. The higher the correlation between the attention given to the various topics, the more similar two textbooks are. We do a classical MDS, suppressing the plot to first assess the quality of a two-dimensional representation.

```
. mds math-mano, id(author) measure(corr) noplot
Classical metric multidimensional scaling
       similarity: correlation, computed on 7 variables
    dissimilarity: sqrt(2(1-similarity))
                                                  Number of obs
                                                                                  25
    Eigenvalues > 0
                                      6
                                                  Mardia fit measure 1 =
                                                                              0.6680
    Retained dimensions
                                      2
                                                  Mardia fit measure 2 =
                                                                              0.8496
                                        abs(eigenvalue)
                                                                   (eigenvalue)<sup>2</sup>
      Dimension
                     Eigenvalue
                                       Percent
                                                   Cumul.
                                                                 Percent
                                                                             Cumul.
               1
                       8.469821
                                         38.92
                                                    38.92
                                                                   56.15
                                                                              56.15
               2
                      6.0665813
                                                                              84.96
                                         27.88
                                                    66.80
                                                                   28.81
               з
                      3.8157101
                                         17.53
                                                    84.33
                                                                   11.40
                                                                              96.35
               4
                      1.6926956
                                          7.78
                                                    92.11
                                                                    2.24
                                                                              98.60
               5
                      1.2576053
                                          5.78
                                                    97.89
                                                                    1.24
                                                                              99.83
               6
                                                                    0.17
                       .45929376
                                          2.11
                                                   100.00
                                                                             100.00
```

Again the quality of a two-dimensional approximation is somewhat unsatisfactory, with 67% and 85% of the variation accounted for according to the two Mardia criteria. Still, let's look at the plot, using a title that refers to the self-referential aspect of the analysis (Smullyan 1986). We reposition some of the author labels to enhance readability by using the mlabvpos() option.

```
. generate spot = 3
. replace spot = 5 if inlist(author,"Seber84","Kshirsagar78","Kendall75")
(3 real changes made)
. replace spot = 2 if author=="MardiaKentBibby79"
(1 real change made)
. replace spot = 9 if inlist(author, "Dagnelie75","Rencher02",
> "GreenCaroll76","EverittDunn01","CooleyLohnes62","Morrison67")
(6 real changes made)
```

. mdsconfig, mlabvpos(spot) title(This plot needs no title)



A striking characteristic of the plot is that the textbooks seem to be located on a circle. This is a phenomenon that is regularly encountered in multidimensional scaling and was labeled the "horseshoe effect" by Kendall (1971, 215–251). This phenomenon seems to occur especially in situations in which a one-dimensional representation of objects needs to be constructed, for example, in *seriation* applications, from data in which small dissimilarities were measured accurately but moderate and larger dissimilarities are "lumped together".

4

#### Technical note

These data could also be analyzed differently. A particularly interesting method is correspondence analysis (CA), which seeks a simultaneous geometric representation of the rows (textbooks) and columns (topics). We used camat to analyze these data. The results for the textbooks were not much different. Textbooks that were mapped as similar using MDS were also mapped this way by CA. The Green and Carroll book that appeared much different from the rest was also displayed away from the rest by CA. In the CA biplot, it was immediately clear that this book was so different because its pages were classified by Gifi (1990) as predominantly mathematical. But CA also located the topics in this space. The pattern was easy to interpret and was expected. The seven topics were mapped in three groups. math and stat appear as two groups by themselves, and the five applied topics were mapped close together. See [MV] ca for information on the ca command.

#### Introduction to modern MDS

We return to the data on breakfast cereals explored above to introduce modern MDS. We repeat some steps taken previously and then perform estimation using options loss(strain) and transform(identity), which we demonstrate are equivalent to classical MDS.

mds is an estimation or eclass command; see program define in [P] program. You can display its stored results using ereturn list. The configuration is stored as e(Y) and we will compare the configuration obtained from classical MDS with the equivalent one from modern MDS.

#### Example 3

```
. use http://www.stata-press.com/data/r13/cerealnut, clear
(Cereal Nutrition)
. replace brand = subinstr(brand, ","_",.)
(20 real changes made)
. quietly mds calories-K, id(brand) noplot
. mat Yclass = e(Y)
. mds calories-K, id(brand) meth(modern) loss(strain) trans(ident) noplot
Iteration 1:
               strain = 594.12657
               strain = 594.12657
Iteration 2:
Modern multidimensional scaling
    dissimilarity: L2, computed on 8 variables
    Loss criterion: strain = loss for classical MDS
    Transformation: identity (no transformation)
                                                  Number of obs
                                                                             25
                                                                   =
                                                  Dimensions
                                                                    =
                                                                              2
    Normalization: principal
                                                  Loss criterion = 594.1266
. mat Ymod = e(Y)
. assert mreldif(Yclass, Ymod) < 1e-6
```

Note the output differences between modern and classical MDS. In modern MDS we have an iteration log from the minimization of the loss function. The method, measure, observations, dimensions, and number of variables are reported as before, but we do not have or display eigenvalues. The normalization is always reported in modern MDS and with normalize(target()) for classical MDS. The loss criterion is simply the value of the loss function at the minimum.

4

#### Protecting from local minimums

Modern MDS can sometimes converge to a local rather than a global minimum. To protect against this, multiple runs can be made, giving the best of the runs as the final answer. The option for performing this is protect(#), where # is the number of runs to be performed. The nolog option is of particular use with protect(), because the iteration logs from the runs will create a lot of output. Repeating the minimization can take some time, depending on the number of runs selected and the number of iterations it takes to converge.

#### Example 4

We choose loss(stress), and transform(identity) is assumed with modern MDS. We omit the iteration logs to avoid a large amount of output. The number of iterations is available after estimation in e(ic). We first do a run without the protect() option, and then we use protect(50) and compare our results.

```
. mds calories-K, id(brand) method(modern) loss(stress) nolog noplot
(transform(identity) assumed)
Modern multidimensional scaling
    dissimilarity: L2, computed on 8 variables
    Loss criterion: stress = raw_stress/norm(distances)
    Transformation: identity (no transformation)
                                                     Number of obs
                                                                        =
                                                                                 25
                                                     Dimensions
                                                                                  2
                                                                        =
    Normalization: principal
                                                     Loss criterion
                                                                             0.0263
                                                                        =
. di e(ic)
45
. mat Ystress = e(Y)
. set seed 123456789
. mds calories-K, id(brand) method(modern) loss(stress) nolog protect(50)
(transform(identity) assumed)
          #iter
run
     mrc
                     lossval
                                seed random configuration
  1
       0
              74
                   .02626681
                                Xdb3578617ea24d1bbd210b2e6541937c000400f2
  2
       0
             101
                   .02626681
                                X630dd08906daab6562fdcb6e21f566d100041446
  3
       0
              78
                   .02626681
                                X73b43d6df67516aea87005b9d405c70200040138
  4
       0
              75
                                Xc9913bd7b3258bd7215929fe40ee51c70004111a
                   .02626681
  5
       0
              75
                   .02626681
                                X5fba1fdfb5219b6fd11daa1e739f9a3e000428ab
  6
       0
              57
                   .02626681
                                X3ea2b7553d633a5418c29f0da0b5cbf400044d18
  7
       0
              84
                   .02626681
                                Xeb0a27a9aa7351c262bca0dccbf5dc8e00043683
  8
       0
              75
                   .02626681
                                Xcb99c7f17ce97d3e7fa9d27bafc16ab600041503
  9
       0
              85
                   .02626681
                                X8551c69beb028bbd48c91c1a1b6e6f0e000417b2
 10
       0
              60
                   .02626681
                                X05d89a191a939001044c3710631948c100041017
 11
       0
              63
                   .02626681
                                X2c6eaeaf4dcd2394c628f466db148b3400041552
 12
       0
              45
                   .02626681
                                    <initial nonrandom value>
 13
       0
              55
                   .02626681
                                X167c0bd9c43f462544a474abacbdd93d00043a2f
 14
       0
              57
                   .02626682
                                X51b9b6c5e05aadd50ca4b924a2521240000411dc
 15
       0
              82
                   .02626682
                                Xff9280cf913270440939762f65c2b4d60004179e
 16
       0
              63
                   .02626682
                                X14f4e343d3e32b22014308b4d2407e8900041a03
 17
       0
              63
                   .02626682
                                X1b06e2ef52b30203908f0d1032704417000406e2
 18
       0
              66
                   .02626682
                                X70ceaf639c2f78374fd6a1181468489e00042887
       0
              72
                   .02626682
                                X5fa530eb49912716c3b27b00020b158c00041c8b
 19
 20
       0
              71
                   .02626682
                                Xf2f8a723276c4a1f3c7e5848c07cc43800041369
 21
       0
              52
                   .02626682
                                Xd5f821b18557b512b3ebc04c992e06b40004230d
 22
       0
              66
                   .02626683
                                Xd0de57fd1b0a448b3450528326fab45d0004443d
       0
                                X6d2da541fcdb0d9024a0a92d5d0496230004139e
 23
              61
                   .02626683
 24
       0
              59
                   .02626683
                                Xb8aae22160bc2fd1beaf5a4b98f46a2500044223
 25
       0
              84
                   .02626684
                                Xe2b46cd1199d7ef41b8cd31479b25b27000420e5
 26
       0
             138
                     .026303
                                Xb21a3be5f19dad75f7708eb425730c6e0004471b
 27
       0
             100
                                X22ef41a50a68221b276cd98ee7dfeef50004330e
                     .026303
 28
       0
              74
                     .026303
                                X0efbcec71dbb1b3c7111cb0f62250283000441c9
 29
       0
              55
                     .026303
                                X01c889699f835483fd6182719be301f100041a55
       0
 30
              56
                     .026303
                                X2f66c0cb554aca4ff44e5d6a6cd4b62700043ebd
 31
       0
              67
                     .026303
                                X3a5b1f132e05f86d36e01eb46eff578b000448ae
 32
       0
              67
                     .026303
                                Xac0226dd85d4f2c440745210acea6ceb00041c59
 33
       0
              75
                     .026303
                                X9e59768de92f8d2ab8e9bc0fd0084c7d00040c2d
 34
       0
              58
                     .026303
                                X333e991d0bf2f21be1025e348a68254700043d01
 35
       0
              60
                     .026303
                                Xedef05bfdbdcddd5a2b8abeadcdd5ab700042c81
 36
       0
              59
                     .026303
                                X67e0caf9c38ba588e96cd01d5d908d7f00044a0b
 37
       0
                     .026303
                                X2af205b7aad416610a0dec141b66778a00043995
              53
 38
       0
              52
                     .026303
                                X0b9944753b1c4b3bd3676f624643a915000449b8
 39
       0
              87
                     .026303
                                Xb175975333f6bdee5bc301e7d3055688000426ab
 40
       0
              63
                   .02630301
                                X7e334ce7d25be1deb7b30539d7160266000429d4
 41
       0
              60
                   .02630301
                                Xf2e6bfadef621544c441e8363c8530450004038e
 42
       0
              60
                   .02630301
                                X45c7e0abd63ad668fa94cd4758d974eb00040c37
 43
       0
              58
                   .02630301
                                X60263a35772a812860431439cad14ad900042619
```

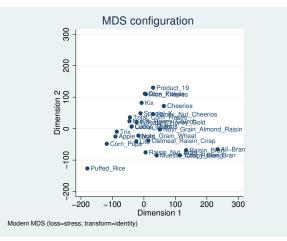
44	0	66	.02630301	X4bf3debb1c7e07f66b533ec5941e1e07000444d7
45	0	63	.02630301	X01f186db4f0db540e749c79e59717c180004351a
46	0	56	.02630302	X66a301f734b575da6762a4edcf9ac6490004313c
47	0	53	.02630302	X5c59c9ffd2e9f2e5bd45f3f9aa22b2f000044969
48	0	131	.19899027	Xe2e15b07d97b0bcb086f194a133dd7b200041db0
49	0	140	.23020403	X065b9333ce65d69bf4d1596e8e8cc72900042a17
50	0	170	.23794378	X075bcd151f123bb5159a55e50022865700043e55

```
Modern multidimensional scaling
dissimilarity: L2, computed on 8 variables
Loss criterion: stress = raw_stress/norm(distances)
Transformation: identity (no transformation)
Number of obs
Dimensions
```

Normalization:	principal

```
. mat YstressP = e(Y)
```

```
. assert mreldif(Ystress, YstressP) < 2e-3
```



The output provided when protect() is specified includes a table with information on each run, sorted by the loss criterion value. The first column simply counts the runs. The second column gives the internal return code from modern MDS. This example only has values of 0, which indicate converged results. The column header mrc is clickable and opens a help file explaining the various MDS return codes. The number of iterations is in the third column. These runs converged in as few as 47 iterations to as many as 190. The loss criterion values are in the fourth column, and the final column contains the seeds used to calculate the starting values for the runs.

In this example, the results from our original run versus the protected run did not differ by much, approximately 1.3e–3. However, looking at runs 46–50 we see loss criterion values that are much higher than the rest. The loss criteria for runs 1–45 vary from .02627 to .02630, but these last runs' loss criteria are all more than .198. These runs clearly converged to local, not global, minimums.

The graph from this protected modern MDS run may be compared with the first one produced. There are obvious similarities, though inspection indicates that the two are not the same.

25

0.0263

2

=

Loss criterion

# Stored results

mds stores the following in e(): Scalars number of observations e(N) number of dimensions in the approximating configuration e(p) number of strictly positive eigenvalues e(np) constant added to squared dissimilarities to force positive semidefiniteness e(addcons) e(mardia1) Mardia measure 1 e(mardia2) Mardia measure 2 e(critval) loss criterion value e(alpha) parameter of transform(power) e(ic) iteration count e(rc) return code e(converged) 1 if converged, 0 otherwise Macros e(cmd) mds e(cmdline) command as typed classical or modern MDS method e(method) e(method2) nonmetric, if method(nonmetric) e(loss) loss criterion e(losstitle) description loss criterion identity, power, or monotonic, transformation function e(tfunction) description of transformation e(transftitle) e(id) ID variable name (mds) e(idtype) int or str; type of id() variable e(duplicates) 1 if duplicates in id(), 0 otherwise e(labels) labels for ID categories e(strfmt) format for category labels e(mxlen) maximum length of category labels e(varlist) variables used in computing similarities or dissimilarities e(dname) similarity or dissimilarity measure name similarity or dissimilarity e(dtype) e(s2d) standard or oneminus (when e(dtype) is similarity) e(unique) 1 if eigenvalues are distinct. 0 otherwise e(init) initialization method seed for init(random) e(iseed) e(seed) seed for solution e(norm) normalization method e(targetmatrix) name of target matrix for normalize(target) e(properties) nob noV for modern or nonmetric MDS; nob noV eigen for classical MDS e(estat\_cmd) program used to implement estat e(predict) program used to implement predict e(marginsnotok) predictions disallowed by margins

Matrices	
e(D)	dissimilarity matrix
e(Disparities)	disparity matrix for nonmetric MDS
e(Y)	approximating configuration coordinates
e(Ev)	eigenvalues
e(idcoding)	coding for integer identifier variable
e(coding)	variable standardization values; first column has value to subtract and second column has divisor
e(norm_stats)	normalization statistics
e(linearf)	two element vector defining the linear transformation; distance equals first element plus second element times dissimilarity
Functions	
e(sample)	marks estimation sample

## Methods and formulas

mds creates a dissimilarity matrix D according to the *measure* specified in option measure(). See [MV] *measure\_option* for descriptions of these measures. Subsequently, mds uses the same subroutines as mdsmat to compute the MDS solution for D. See *Methods and formulas* in [MV] mdsmat for information.

## References

. .

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Also see References in [MV] mdsmat.

Joseph Bernard Kruskal (1928–2010) was born in New York. His brothers were statistician William Henry Kruskal (1919–2005) and mathematician and physicist Martin David Kruskal (1925–2006). He earned degrees in mathematics from Chicago and Princeton and worked at Bell Labs until his retirement in 1993. In statistics, Kruskal made major contributions to multidimensional scaling. In computer science, he devised an algorithm for computing the minimal spanning tree of a weighted graph. His other interests include clustering and statistical linguistics.

## Also see

- [MV] mds postestimation Postestimation tools for mds, mdsmat, and mdslong
- [MV] mds postestimation plots Postestimation plots for mds, mdsmat, and mdslong
- [MV] **biplot** Biplots
- [MV] ca Simple correspondence analysis
- [MV] factor Factor analysis
- [MV] mdslong Multidimensional scaling of proximity data in long format
- [MV] mdsmat Multidimensional scaling of proximity data in a matrix
- [MV] pca Principal component analysis
- [U] 20 Estimation and postestimation commands