Title

svy — The survey prefix command

Description

`svy` fits statistical models for complex survey data by adjusting the results of a command for survey settings identified by `svyset`. Any Stata estimation command listed in [SVY] **svy estimation** may be used with `svy`. User-written programs that meet the requirements in [P] **program properties** may also be used.

Quick start

Data for a two-stage design with sampling weight `wvar1`, strata defined by levels of `svar`, sampling units are identified by `su1`, and second-stage clustering is defined by `su2`

```
svyset su1 [pweight=wvar1], strata(svar) || su2
```

Adjust linear regression for complex survey design settings specified in `svyset`

```
svy: regress ...
```

Same as above, but restrict estimation to the subpopulation where `group` equals 4

```
svy, subpop(if group==4): regress ...
```

Same as above, but use new binary variable `insample` to indicate the subpopulation

```
generate insample = (group==4)
svy, subpop(insample): regress ...
```

Specify that the design degrees of freedom is 135 instead of the difference between the number of unique values of `su1` and the number of levels of `svar`

```
svy, dof(135): regress ...
```

Note: Any estimation command meeting the requirements specified in the *Description* may be substituted for `regress` in the examples above.

Syntax

`svy` [*vcetype*] [, *svy_options eform_option*] : *command*

vcetype	Description
SE	
`linearized`	Taylor-linearized variance estimation
`bootstrap`	bootstrap variance estimation; see [SVY] **svy bootstrap**
`brr`	BRR variance estimation; see [SVY] **svy brr**
`jackknife`	jackknife variance estimation; see [SVY] **svy jackknife**
`sdr`	SDR variance estimation; see [SVY] **svy sdr**

Specifying a *vcetype* overrides the default from `svyset`.

svy_options	Description
if/in	
`subpop(`[*varname*] [*if*]`)`	identify a subpopulation
SE	
`dof(#)`	design degrees of freedom
bootstrap_options	more options allowed with bootstrap variance estimation; see [SVY] ***bootstrap_options***
brr_options	more options allowed with BRR variance estimation; see [SVY] ***brr_options***
jackknife_options	more options allowed with jackknife variance estimation; see [SVY] ***jackknife_options***
sdr_options	more options allowed with SDR variance estimation; see [SVY] ***sdr_options***
Reporting	
`level(#)`	set confidence level; default is `level(95)`
`nocnsreport`	do not display constraints
display_options	control columns and column formats, row spacing, line width, display of omitted variables and base and empty cells, and factor-variable labeling
`noheader`	suppress table header
`nolegend`	suppress table legend
`noadjust`	do not adjust model Wald statistic
`noisily`	display any output from *command*
`trace`	trace *command*
`coeflegend`	display legend instead of statistics

svy requires that the survey design variables be identified using svyset; see [SVY] svyset.

command defines the estimation command to be executed. The by prefix cannot be part of *command*.

collect is allowed; see [U] 11.1.10 Prefix commands. mi estimate may be used with svy linearized if the estimation command allows mi estimate; it may not be used with svy bootstrap, svy brr, svy jackknife, or svy sdr.

noheader, nolegend, noadjust, noisily, trace, and coeflegend are not shown in the dialog boxes for estimation commands.

Warning: Using if or in restrictions will often not produce correct variance estimates for subpopulations. To compute estimates for subpopulations, use the subpop() option.

See [U] 20 Estimation and postestimation commands for more capabilities of estimation commands.

Options

if/in

subpop(*subpop*) specifies that estimates be computed for the single subpopulation identified by *subpop*, which is

[*varname*] [*if*]

Thus the subpopulation is defined by the observations for which *varname* $\neq 0$ that also meet the if conditions. Typically, *varname* $= 1$ defines the subpopulation, and *varname* $= 0$ indicates observations not belonging to the subpopulation. For observations whose subpopulation status is uncertain, *varname* should be set to a missing value; such observations are dropped from the estimation sample.

See [SVY] Subpopulation estimation and [SVY] estat.

SE

dof(#) specifies the design degrees of freedom, overriding the default calculation, df $= N_{\text{psu}} - N_{\text{strata}}$.

bootstrap_options are other options that are allowed with bootstrap variance estimation specified by svy bootstrap or specified as svyset using the vce(bootstrap) option; see [SVY] *bootstrap_options*.

brr_options are other options that are allowed with BRR variance estimation specified by svy brr or specified as svyset using the vce(brr) option; see [SVY] *brr_options*.

jackknife_options are other options that are allowed with jackknife variance estimation specified by svy jackknife or specified as svyset using the vce(jackknife) option; see [SVY] *jackknife_options*.

sdr_options are other options that are allowed with SDR variance estimation specified by svy sdr or specified as svyset using the vce(sdr) option; see [SVY] *sdr_options*.

Reporting

level(#) specifies the confidence level, as a percentage, for confidence intervals. The default is level(95) or as set by set level; see [U] 20.8 Specifying the width of confidence intervals.

nocnsreport; see [R] Estimation options.

display_options: noci, nopvalues, noomitted, vsquish, noemptycells, baselevels, allbaselevels, nofvlabel, fvwrap(#), fvwrapon(*style*), cformat(%*fmt*), pformat(%*fmt*), sformat(%*fmt*), and nolstretch; see [R] Estimation options.

The following options are available with svy but are not shown in the dialog boxes:

noheader prevents the table header from being displayed. This option implies nolegend.

`nolegend` prevents the table legend identifying the subpopulations from being displayed.

`noadjust` specifies that the model Wald test be carried out as $W/k \sim F(k,d)$, where W is the Wald test statistic, k is the number of terms in the model excluding the constant term, d is the total number of sampled PSUs minus the total number of strata, and $F(k,d)$ is an F distribution with k numerator degrees of freedom and d denominator degrees of freedom. By default, an adjusted Wald test is conducted: $(d-k+1)W/(kd) \sim F(k, d-k+1)$.

See Korn and Graubard (1990) for a discussion of the Wald test and the adjustments thereof. Using the `noadjust` option is not recommended.

`noisily` requests that any output from *command* be displayed.

`trace` causes a trace of the execution of *command* to be displayed.

`coeflegend`; see [R] **Estimation options**.

The following option is usually available with `svy` at the time of estimation or on replay but is not shown in all dialog boxes:

eform_option; see [R] ***eform_option***.

Remarks and examples

stata.com

The `svy` prefix is designed for use with complex survey data. Typical survey design characteristics include sampling weights, one or more stages of clustered sampling, and stratification. For a general discussion of various aspects of survey designs, including multistage designs, see [SVY] **svyset**.

Below we present an example of the effects of weights, clustering, and stratification. This is a typical case, but drawing general rules from any one example is still dangerous. You could find particular analyses from other surveys that are counterexamples for each of the trends for standard errors exhibited here.

▷ Example 1: The effects of weights, clustering, and stratification

We use data from the Second National Health and Nutrition Examination Survey (NHANES II) (McDowell et al. 1981) as our example. This is a national survey, and the dataset has sampling weights, strata, and clustering. In this example, we will consider the estimation of the mean serum zinc level of all adults in the United States.

First, consider a proper design-based analysis, which accounts for weighting, clustering, and stratification. Before we issue our `svy` estimation command, we set the weight, strata, and PSU identifier variables:

```
. use https://www.stata-press.com/data/r18/nhanes2f
. svyset psuid [pweight=finalwgt], strata(stratid)
Sampling weights: finalwgt
             VCE: linearized
     Single unit: missing
        Strata 1: stratid
 Sampling unit 1: psuid
           FPC 1: <zero>
```

We now estimate the mean by using the proper design-based analysis:

```
. svy: mean zinc
(running mean on estimation sample)

Survey: Mean estimation

Number of strata = 31            Number of obs   =       9,189
Number of PSUs   = 62            Population size = 104,176,071
                                 Design df       =          31

--------------------------------------------------------------
             |             Linearized
             |       Mean   std. err.     [95% conf. interval]
-------------+------------------------------------------------
        zinc |   87.18207   .4944827      86.17356    88.19057
--------------------------------------------------------------
```

If we ignore the survey design and use `mean` to estimate the mean, we get

```
. mean zinc

Mean estimation                          Number of obs = 9,189

--------------------------------------------------------------
             |       Mean   Std. err.     [95% conf. interval]
-------------+------------------------------------------------
        zinc |   86.51518   .1510744      86.21904    86.81132
--------------------------------------------------------------
```

The point estimate from the unweighted analysis is smaller by more than one standard error than the proper design-based estimate. Also, design-based analysis produced a standard error that is 3.27 times larger than the standard error produced by our incorrect analysis.

◁

▷ Example 2: Halfway is not enough—the importance of stratification and clustering

When some people analyze survey data, they say, "I know I have to use my survey weights, but I will just ignore the stratification and clustering information." If we follow this strategy, we will obtain the proper design-based point estimates, but our standard errors, confidence intervals, and test statistics will usually be wrong.

To illustrate this effect, suppose that we used the `svy: mean` procedure with `pweights` only.

```
. svyset [pweight=finalwgt]

Sampling weights: finalwgt
             VCE: linearized
     Single unit: missing
        Strata 1: <one>
 Sampling unit 1: <observations>
           FPC 1: <zero>

. svy: mean zinc
(running mean on estimation sample)

Survey: Mean estimation

Number of strata =     1         Number of obs   =       9,189
Number of PSUs   = 9,189         Population size = 104,176,071
                                 Design df       =       9,188

--------------------------------------------------------------
             |             Linearized
             |       Mean   std. err.     [95% conf. interval]
-------------+------------------------------------------------
        zinc |   87.18207   .1828747      86.82359    87.54054
--------------------------------------------------------------
```

This approach gives us the same point estimate as our design-based analysis, but the reported standard error is less than one-half the design-based standard error. If we accounted only for clustering and weights and ignored stratification in NHANES II, we would obtain the following analysis:

```
. svyset psuid [pweight=finalwgt]

Sampling weights: finalwgt
             VCE: linearized
     Single unit: missing
        Strata 1: <one>
 Sampling unit 1: psuid
           FPC 1: <zero>

. svy: mean zinc
(running mean on estimation sample)

Survey: Mean estimation

Number of strata = 1              Number of obs   =       9,189
Number of PSUs   = 2              Population size = 104,176,071
                                  Design df       =           1

--------------------------------------------------------------
             |             Linearized
             |       Mean   std. err.     [95% conf. interval]
-------------+------------------------------------------------
        zinc |   87.18207   .7426221      77.74616    96.61798
--------------------------------------------------------------
```

Here our standard error is about 50% larger than what we obtained in our proper design-based analysis. ◁

▷ Example 3

Let's look at a regression. We model zinc on the basis of age, weight, sex, race, and rural or urban residence. We compare a proper design-based analysis with an ordinary regression (which assumes independent and identically distributed error).

Here is our design-based analysis:

```
. svyset psuid [pweight=finalwgt], strata(stratid)

Sampling weights: finalwgt
             VCE: linearized
     Single unit: missing
        Strata 1: stratid
 Sampling unit 1: psuid
           FPC 1: <zero>

. svy: regress zinc age c.age#c.age weight female black orace rural
(running regress on estimation sample)

Survey: Linear regression

Number of strata = 31                          Number of obs   =       9,189
Number of PSUs   = 62                          Population size = 104,176,071
                                               Design df       =          31
                                               F(7, 25)        =       62.50
                                               Prob > F        =      0.0000
                                               R-squared       =      0.0698

------------------------------------------------------------------------------
             |             Linearized
        zinc | Coefficient  std. err.      t    P>|t|     [95% conf. interval]
-------------+----------------------------------------------------------------
         age |  -.1701161   .0844192    -2.02   0.053    -.3422901     .002058
             |
 c.age#c.age |   .0008744   .0008655     1.01   0.320    -.0008907    .0026396
             |
      weight |   .0535225   .0139115     3.85   0.001     .0251499    .0818951
      female |  -6.134161   .4403625   -13.93   0.000    -7.032286   -5.236035
       black |  -2.881813   1.075958    -2.68   0.012    -5.076244    -.687381
       orace |  -4.118051   1.621121    -2.54   0.016    -7.424349   -.8117528
       rural |  -.5386327   .6171836    -0.87   0.390    -1.797387    .7201216
       _cons |   92.47495   2.228263    41.50   0.000     87.93038    97.01952
------------------------------------------------------------------------------
```

If we had improperly ignored our survey weights, stratification, and clustering (that is, if we had used the usual Stata `regress` command), we would have obtained the following results:

```
. regress zinc age c.age#c.age weight female black orace rural

      Source |       SS           df       MS      Number of obs   =     9,189
-------------+----------------------------------   F(7, 9181)      =     79.72
       Model |  110417.827         7  15773.9753   Prob > F        =    0.0000
    Residual |   1816535.3     9,181   197.85811   R-squared       =    0.0573
-------------+----------------------------------   Adj R-squared   =    0.0566
       Total |  1926953.13     9,188  209.724982   Root MSE        =    14.066

------------------------------------------------------------------------------
        zinc | Coefficient  Std. err.      t    P>|t|     [95% conf. interval]
-------------+----------------------------------------------------------------
         age |   -.090298   .0638452    -1.41   0.157    -.2154488    .0348528
             |
 c.age#c.age |  -.0000324   .0006788    -0.05   0.962    -.0013631    .0012983
             |
      weight |   .0606481   .0105986     5.72   0.000     .0398725    .0814237
      female |  -5.021949   .3194705   -15.72   0.000    -5.648182   -4.395716
       black |  -2.311753   .5073536    -4.56   0.000    -3.306279   -1.317227
       orace |  -3.390879   1.060981    -3.20   0.001    -5.470637   -1.311121
       rural |  -.0966462   .3098948    -0.31   0.755    -.7041089    .5108166
       _cons |   89.49465   1.477528    60.57   0.000     86.59836    92.39093
------------------------------------------------------------------------------
```

The point estimates differ by 3%–100%, and the standard errors for the proper designed-based analysis are 30%–110% larger. The differences are not as dramatic as we saw with the estimation of the mean, but they are still substantial.

◁

Stored results

`svy` stores the following in `e()`:

Scalars

`e(N)`	number of observations
`e(N_sub)`	subpopulation observations
`e(N_strata)`	number of strata
`e(N_strata_omit)`	number of strata omitted
`e(singleton)`	`1` if singleton strata, `0` otherwise
`e(census)`	`1` if census data, `0` otherwise
`e(F)`	model F statistic
`e(df_m)`	model degrees of freedom
`e(df_r)`	variance degrees of freedom
`e(N_pop)`	estimate of population size
`e(N_subpop)`	estimate of subpopulation size
`e(N_psu)`	number of sampled PSUs
`e(stages)`	number of sampling stages
`e(k_eq)`	number of equations in `e(b)`
`e(k_aux)`	number of ancillary parameters
`e(p)`	p-value
`e(rank)`	rank of `e(V)`

Macros

`e(prefix)`	`svy`
`e(cmdname)`	command name from *command*
`e(cmd)`	same as `e(cmdname)` or `e(vce)`
`e(command)`	*command*
`e(cmdline)`	command as typed
`e(wtype)`	weight type
`e(wexp)`	weight expression
`e(weight#)`	variable identifying weight for stage #
`e(wvar)`	weight variable name
`e(singleunit)`	`singleunit()` setting
`e(strata)`	`strata()` variable
`e(strata#)`	variable identifying strata for stage #
`e(psu)`	`psu()` variable
`e(su#)`	variable identifying sampling units for stage #
`e(fpc)`	`fpc()` variable
`e(fpc#)`	FPC for stage #
`e(title)`	title in estimation output
`e(poststrata)`	`poststrata()` variable
`e(postweight)`	`postweight()` variable
`e(vce)`	*vcetype* specified in `vce()`
`e(vcetype)`	title used to label Std. err.
`e(mse)`	`mse`, if specified
`e(subpop)`	*subpop* from `subpop()`
`e(adjust)`	`noadjust`, if specified
`e(properties)`	`b V`
`e(estat_cmd)`	program used to implement `estat`
`e(predict)`	program used to implement `predict`
`e(marginsnotok)`	predictions disallowed by `margins`
`e(marginswtype)`	weight type for `margins`

Matrices

`e(b)`	estimates
`e(V)`	design-based variance
`e(V_srs)`	simple-random-sampling-without-replacement variance, $\widehat{V}_{\text{srswor}}$

`e(V_srssub)`	subpopulation simple-random-sampling-without-replacement variance, $\widehat{V}_{\text{srswor}}$ (created only when `subpop()` is specified)
`e(V_srswr)`	simple-random-sampling-with-replacement variance, $\widehat{V}_{\text{srswr}}$ (created only when `fpc()` option is `svyset`)
`e(V_srssubwr)`	subpopulation simple-random-sampling-with-replacement variance, $\widehat{V}_{\text{srswr}}$ (created only when `subpop()` is specified)
`e(V_modelbased)`	model-based variance
`e(V_msp)`	variance from misspecified model fit, $\widehat{V}_{\text{msp}}$
`e(_N_strata_single)`	number of strata with one sampling unit
`e(_N_strata_certain)`	number of certainty strata
`e(_N_strata)`	number of strata
`e(_N_subp)`	estimate of subpopulation sizes within `over()` groups

Functions

`e(sample)`	marks estimation sample

`svy` also carries forward most of the results already in `e()` from *command*.

In addition to the above, the following is stored in `r()`:

Matrices

`r(table)`	matrix containing the coefficients with their standard errors, test statistics, p-values, and confidence intervals

Note that results stored in `r()` are updated when the command is replayed and will be replaced when any r-class command is run after the estimation command.

Methods and formulas

See [SVY] **Variance estimation** for all the details behind the point estimate and variance calculations made by `svy`.

References

Korn, E. L., and B. I. Graubard. 1990. Simultaneous testing of regression coefficients with complex survey data: Use of Bonferroni t statistics. *American Statistician* 44: 270–276. https://doi.org/10.2307/2684345.

McDowell, A., A. Engel, J. T. Massey, and K. Maurer. 1981. Plan and operation of the Second National Health and Nutrition Examination Survey, 1976–1980. *Vital and Health Statistics* 1(15): 1–144.

Also see

[SVY] **svy estimation** — Estimation commands for survey data

[SVY] **svy postestimation** — Postestimation tools for svy

[SVY] **svy bootstrap** — Bootstrap for survey data

[SVY] **svy brr** — Balanced repeated replication for survey data

[SVY] **svy jackknife** — Jackknife estimation for survey data

[SVY] **svy sdr** — Successive difference replication for survey data

[SVY] **svyset** — Declare survey design for dataset

[SVY] **Calibration** — Calibration for survey data

[SVY] **Poststratification** — Poststratification for survey data

[SVY] **Subpopulation estimation** — Subpopulation estimation for survey data

[SVY] **Variance estimation** — Variance estimation for survey data

[P] **program properties** — Properties of user-defined programs

[P] **_robust** — Robust variance estimates

[U] 20 Estimation and postestimation commands