

vwls — Variance-weighted least squares

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Description

`vwls` estimates a linear regression using variance-weighted least squares. It differs from ordinary least-squares (OLS) regression in that it does not assume homogeneity of variance, but requires that the conditional variance of *depvar* be estimated prior to the regression. The estimated variance need not be constant across observations. `vwls` treats the estimated variance as if it were the true variance when it computes standard errors of the coefficients.

You must supply an estimate of the conditional standard deviation of *depvar* to `vwls` by using the `sd(varname)` option, or you must have grouped data with the groups defined by the *indepvars* variables. In the latter case, `vwls` treats all *indepvars* as categorical variables, computes the mean and standard deviation of *depvar* separately for each subgroup, and computes the regression of the subgroup means on *indepvars*.

`regress` with analytic weights can be used to produce another kind of “variance-weighted least squares”; see [Remarks and examples](#) for an explanation of the difference.

Quick start

Variance-weighted least-squares regression of *y* on *x1* and *x2*, with the estimated conditional std. dev. of *y* stored in `sd`

```
vwls y1 x1 x2, sd(sd)
```

Add categorical variable *a* using [factor-variable](#) syntax

```
vwls y1 x1 x2 i.a, sd(sd)
```

As above, but restrict the sample to cases where *v* is greater than 1

```
vwls y1 x1 x2 i.a if v>1, sd(sd)
```

Variance-weighted least-squares regression for grouped data with subgroups defined by *a2* and *a3*

```
vwls y2 i.a2 i.a3
```

Menu

Statistics > Linear models and related > Other > Variance-weighted least squares

Syntax

```
vwls depvar indepvars [if] [in] [weight] [, options]
```

<i>options</i>	Description
Model	
noconstant	suppress constant term
sd (<i>varname</i>)	variable containing estimate of conditional standard deviation
Reporting	
level (#)	set confidence level; default is level(95)
display_options	control columns and column formats, row spacing, line width, display of omitted variables and base and empty cells, and factor-variable labeling
coeflegend	display legend instead of statistics
<i>indepvars</i> may contain factor variables; see [U] 11.4.3 Factor variables .	
bootstrap , by , jackknife , rolling , and statsby are allowed; see [U] 11.1.10 Prefix commands .	
Weights are not allowed with the bootstrap prefix; see [R] bootstrap .	
fweights are allowed; see [U] 11.1.6 weight .	
coeflegend does not appear in the dialog box.	
See [U] 20 Estimation and postestimation commands for more capabilities of estimation commands.	

Options

Model

noconstant; see [R] **estimation options**.

sd(*varname*) is an estimate of the conditional standard deviation of *depvar* (that is, it can vary observation by observation). All values of *varname* must be > 0. If you specify **sd**(), you cannot use **fweights**.

If **sd**() is not given, the data will be grouped by *indepvars*. Here *indepvars* are treated as categorical variables, and the means and standard deviations of *depvar* for each subgroup are calculated and used for the regression. Any subgroup for which the standard deviation is zero is dropped.

Reporting

level(#); see [R] **estimation options**.

display_options: **noci**, **nopvalues**, **noomitted**, **vsquish**, **noemptycells**, **baselevels**, **allbaselevels**, **nofvlabel**, **fvwrap**(#), **fvwrapon**(*style*), **cformat**(%*fmt*), **pformat**(%*fmt*), **sformat**(%*fmt*), and **nolstretch**; see [R] **estimation options**.

The following option is available with **vwls** but is not shown in the dialog box:

coeflegend; see [R] **estimation options**.

Remarks and examples

The `vwls` command is intended for use with two special—and different—types of data. The first contains data that consist of measurements from physical science experiments in which all error is due solely to measurement errors and the sizes of the measurement errors are known.

You can also use variance-weighted least-squares linear regression for certain problems in categorical data analysis, such as when all the independent variables are categorical and the outcome variable is either continuous or a quantity that can sensibly be averaged. If each of the subgroups defined by the categorical variables contains a reasonable number of subjects, then the variance of the outcome variable can be estimated independently within each subgroup. For the purposes of estimation, `vwls` treats each subgroup as one observation, with the dependent variable being the subgroup mean of the outcome variable.

The `vwls` command fits the model

$$y_i = \mathbf{x}_i\boldsymbol{\beta} + \varepsilon_i$$

where the errors ε_i are independent normal random variables with the distribution $\varepsilon_i \sim N(0, \nu_i)$. The independent variables \mathbf{x}_i are assumed to be known without error.

As described above, `vwls` assumes that you already have estimates s_i^2 for the variances ν_i . The error variance is not estimated in the regression. The estimates s_i^2 are used to compute the standard errors of the coefficients; see [Methods and formulas](#) below.

In contrast, weighted OLS regression assumes that the errors have the distribution $\varepsilon_i \sim N(0, \sigma^2/w_i)$, where the w_i are known weights and σ^2 is an unknown parameter that is estimated in the regression. This is the difference from variance-weighted least squares: in weighted OLS, the magnitude of the error variance is estimated in the regression using all the data.

► Example 1

An artificial, but informative, example illustrates the difference between variance-weighted least squares and weighted OLS.

We measure the quantities x_i and y_i and estimate that the standard deviation of y_i is s_i . We enter the data into Stata:

```
. use http://www.stata-press.com/data/r15/vwlsxmpl
. list
```

	x	y	s
1.	1	1.2	.5
2.	2	1.9	.5
3.	3	3.2	1
4.	4	4.3	1
5.	5	4.9	1
6.	6	6.0	2
7.	7	7.2	2
8.	8	7.9	2

Because we want observations with smaller variance to carry larger weight in the regression, we compute an OLS regression with analytic weights proportional to the inverse of the squared standard deviations:

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```
. regress y x [aweight=s^(-2)]
(sum of wgt is 1.1750e+01)
```

Source	SS	df	MS	Number of obs	=	8
Model	22.6310183	1	22.6310183	F(1, 6)	=	702.26
Residual	.193355117	6	.032225853	Prob > F	=	0.0000
Total	22.8243734	7	3.26062477	R-squared	=	0.9915
				Adj R-squared	=	0.9901
				Root MSE	=	.17952

y	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
x	.9824683	.0370739	26.50	0.000	.8917517 1.073185
_cons	.1138554	.1120078	1.02	0.349	-.1602179 .3879288

If we compute a variance-weighted least-squares regression by using `vwls`, we get the same results for the coefficient estimates but very different standard errors:

```
. vwls y x, sd(s)
```

Variance-weighted least-squares regression	Number of obs	=	8
Goodness-of-fit chi2(6) = 0.28	Model chi2(1)	=	33.24
Prob > chi2 = 0.9996	Prob > chi2	=	0.0000

y	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
x	.9824683	.170409	5.77	0.000	.6484728 1.316464
_cons	.1138554	.51484	0.22	0.825	-.8952124 1.122923

Although the values of y_i were nicely linear with x_i , the `vwls` regression used the large estimates for the standard deviations to compute large standard errors for the coefficients. For weighted OLS regression, however, the scale of the analytic weights has no effect on the standard errors of the coefficients—only the relative proportions of the analytic weights affect the regression.

If we are sure of the sizes of our error estimates for y_i , using `vwls` is valid. However, if we can estimate only the relative proportions of error among the y_i , then `vwls` is not appropriate.

◀

► Example 2

Let's now consider an example of the use of `vwls` with categorical data. Suppose that we have blood pressure data for $n = 400$ subjects, categorized by gender and race (black or white). Here is a description of the data:

```
. use http://www.stata-press.com/data/r15/bp
. table gender race, c(mean bp sd bp freq) row col format(%8.1f)
```

Gender	Race		
	White	Black	Total
Female	117.1	118.5	117.8
	10.3	11.6	10.9
	100.0	100.0	200.0
Male	122.1	125.8	124.0
	10.6	15.5	13.3
	100.0	100.0	200.0
Total	119.6	122.2	120.9
	10.7	14.1	12.6
	200.0	200.0	400.0

Performing a variance-weighted regression using vwls gives

```
. vwls bp gender race
Variance-weighted least-squares regression      Number of obs      =      400
Goodness-of-fit chi2(1)      =      0.88      Model chi2(2)      =      27.11
Prob > chi2      =      0.3486      Prob > chi2      =      0.0000
```

bp	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
gender	5.876522	1.170241	5.02	0.000	3.582892	8.170151
race	2.372818	1.191683	1.99	0.046	.0371631	4.708473
_cons	116.6486	.9296297	125.48	0.000	114.8266	118.4707

By comparison, an OLS regression gives the following result:

```
. regress bp gender race
Source      |      SS      |      df      |      MS      |      Number of obs      =      400
-----+-----|-----+-----|-----+-----|      F(2, 397)      =      15.24
Model      | 4485.66639   |      2      | 2242.83319   |      Prob > F      =      0.0000
Residual   | 58442.7305   |     397     | 147.210908   |      R-squared      =      0.0713
-----+-----|-----+-----|-----+-----|      Adj R-squared   =      0.0666
Total      | 62928.3969   |     399     | 157.71528    |      Root MSE      =      12.133
```

bp	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
gender	6.1775	1.213305	5.09	0.000	3.792194	8.562806
race	2.5875	1.213305	2.13	0.034	.2021938	4.972806
_cons	116.4862	1.050753	110.86	0.000	114.4205	118.552

Note the larger value for the `race` coefficient (and smaller p -value) in the OLS regression. The assumption of homogeneity of variance in OLS means that the mean for black men pulls the regression line higher than in the `vwls` regression, which takes into account the larger variance for black men and reduces its effect on the regression.

Stored results

`vwls` stores the following in `e()`:

Scalars

<code>e(N)</code>	number of observations
<code>e(df_m)</code>	model degrees of freedom
<code>e(chi2)</code>	model χ^2
<code>e(df_gf)</code>	goodness-of-fit degrees of freedom
<code>e(chi2_gf)</code>	goodness-of-fit χ^2
<code>e(rank)</code>	rank of $e(V)$

Macros

<code>e(cmd)</code>	<code>vwls</code>
<code>e(cmdline)</code>	command as typed
<code>e(depvar)</code>	name of dependent variable
<code>e(properties)</code>	<code>b V</code>
<code>e(predict)</code>	program used to implement <code>predict</code>
<code>e(asbalanced)</code>	factor variables <code>fvset</code> as <code>asbalanced</code>
<code>e(asobserved)</code>	factor variables <code>fvset</code> as <code>asobserved</code>

Matrices

<code>e(b)</code>	coefficient vector
<code>e(V)</code>	variance-covariance matrix of the estimators

Functions

<code>e(sample)</code>	marks estimation sample
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Methods and formulas

Let $\mathbf{y} = (y_1, y_2, \dots, y_n)'$ be the vector of observations of the dependent variable, where n is the number of observations. When `sd()` is specified, let s_1, s_2, \dots, s_n be the standard deviations supplied by `sd()`. For categorical data, when `sd()` is not given, the means and standard deviations of y for each subgroup are computed, and n becomes the number of subgroups, \mathbf{y} is the vector of subgroup means, and s_i are the standard deviations for the subgroups.

Let $\mathbf{V} = \text{diag}(s_1^2, s_2^2, \dots, s_n^2)$ denote the estimate of the variance of \mathbf{y} . Then the estimated regression coefficients are

$$\mathbf{b} = (\mathbf{X}'\mathbf{V}^{-1}\mathbf{X})^{-1}\mathbf{X}'\mathbf{V}^{-1}\mathbf{y}$$

and their estimated covariance matrix is

$$\widehat{\text{Cov}}(\mathbf{b}) = (\mathbf{X}'\mathbf{V}^{-1}\mathbf{X})^{-1}$$

A statistic for the goodness of fit of the model is

$$Q = (\mathbf{y} - \mathbf{X}\mathbf{b})' \mathbf{V}^{-1} (\mathbf{y} - \mathbf{X}\mathbf{b})$$

where Q has a χ^2 distribution with $n - k$ degrees of freedom (k is the number of independent variables plus the constant, if any).

References

- Gini, R., and J. Pasquini. 2006. [Automatic generation of documents](#). *Stata Journal* 6: 22–39.
- Grizzle, J. E., C. F. Starmer, and G. G. Koch. 1969. Analysis of categorical data by linear models. *Biometrics* 25: 489–504.
- Press, W. H., S. A. Teukolsky, W. T. Vetterling, and B. P. Flannery. 2007. *Numerical Recipes: The Art of Scientific Computing*. 3rd ed. New York: Cambridge University Press.

Also see

[R] [vpls postestimation](#) — Postestimation tools for vpls

[R] [regress](#) — Linear regression

[U] [11.1.6 weight](#)

[U] [20 Estimation and postestimation commands](#)